

Tutorial 2

Monday, 8 August 2022 10:18 PM

- (Ω, \mathcal{F}, P) → probability space
 ↙ Sample space ↘ Event space probability measure.

Independence of Events For probability space (Ω, \mathcal{F}, P) , a family of events $(A_i : i \in I)$ is said to be independent if for any finite set $F \subseteq I$, we have

$$P\left(\bigcap_{i \in F} A_i\right) = \prod_{i \in F} P(A_i)$$

* Let A, B, C are 3 events. Their independence is assured by 4 equalities.

$$P(A \cap B) = P(A)P(B), P(B \cap C) = P(B)P(C)$$

$$P(C \cap A) = P(C)P(A), P(A \cap B \cap C) = P(A)P(B)P(C).$$

Examples: 1. Consider two independent tosses of a fair coin and the following events.

$$H_1 = \{1st \text{ toss is head}\} = \{HH, HT\}$$

$$H_2 = \{2nd \text{ toss is head}\} = \{HH, TH\}$$

$$D = \{1st \text{ and 2nd toss have diff. outcome}\} \\ = \{HT, TH\}$$

$$P(H_1 \cap H_2) = P(\{HH\}) = \frac{1}{4} \quad P(H_1) \cdot P(H_2) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}.$$

$$P(H_1 \cap D) = P(\{HT\}) = \frac{1}{4} \quad P(H_1) \cdot P(D) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

$$P(H_2 \cap D) = P(\{TH\}) = \frac{1}{4} \quad P(H_2) \cdot P(D) = \frac{1}{2} \cdot \frac{1}{2} = \frac{1}{4}$$

$$P(H_1 \cap H_2 \cap D) = P(\emptyset) = 0 \quad P(H_1) \cdot P(H_2) \cdot P(D) = \frac{1}{8}$$

2. Consider two independent rolls of a fair 6-sided die, and the following events:

$$A = \{1st \text{ roll is } 1, 2, 3\} = \{(x, y) : x \in [3], y \in [6]\}$$

$$B = \{2nd \text{ roll is } 4, 5, 6\} = \{(x, y) : x \in [6], y \in [3]\}$$

$$\begin{aligned}
 A &= \{(x, y) : x \in [3], y \in [6]\} \\
 B &= \{(x, y) : x \in [6], y \in [3]\} \\
 C &= \{(x, y) : x + y = 9, x \in [6], y \in [6]\}
 \end{aligned}$$

$$P(A) = \frac{1}{2}, \quad P(B) = \frac{1}{2}, \quad P(C) = \frac{1}{9} \Rightarrow P(A)P(B)P(C) = \frac{1}{36}.$$

$$P(A \cap B \cap C) = P(\{(3,6)\}) = \frac{1}{36}$$

$$P(A \cap B) = P(A) \cdot P(B), \quad P(A \cap C) = \frac{1}{36}, \quad P(A)P(C) = \frac{1}{2} \cdot \frac{1}{9} = \frac{1}{18}$$

$$P(B \cap C) = \frac{1}{12}, \quad P(B) \cdot P(C) = \frac{1}{2} \cdot \frac{1}{9} = \frac{1}{18}$$

Independence of σ -algebras Let $\mathcal{F}_i, i \in I$ are independent σ -algebras ($\mathcal{F}_i \subseteq \mathcal{F}$) if for any collection of events $(A_i \in \mathcal{F}_i : i \in F)$, with finite set $F \subseteq I$

$$P\left(\bigcap_{i \in F} A_i\right) = \prod_{i \in F} P(A_i)$$

Random Variable : Consider a probability space (Ω, \mathcal{F}, P) . A random variable $X : \Omega \rightarrow \mathbb{R}$ is a \mathbb{R} -valued function from the sample space to real numbers, s.t for each $x \in \mathbb{R}$ the event

$$X^{-1}(-\infty, x] = A_x(x) \triangleq \{\omega \in \Omega : X(\omega) \leq x\} \in \mathcal{F}$$

That is $X^{-1}(B_x)$ are \mathcal{F} -measurable sets where $B_x = (-\infty, x] \forall x \in \mathbb{R}$.

- X is called \mathcal{F} -measurable random variable.

* $(B_x : x \in \mathbb{R})$ is the generating collection for Borel σ -algebra

Ex 1. $B \in \mathcal{B}(\mathbb{R})$. Show that $X^{-1}(B) \in \mathcal{F}$ if X is a \mathcal{F} -measurable r.v.

Solⁿ. We have that the generating sets, $B_x = (-\infty, x] \in \mathcal{B}(\mathbb{R})$ have preimages in \mathcal{F} . We need to show that any set $B \in \mathcal{B}(\mathbb{R})$ has its preimage in \mathcal{F} .

To continue to show that

its preimage in \mathcal{F} .

It suffices to show that

$$\textcircled{1} \quad X^{-1}(B^c) = X^{-1}(B)^c \quad \text{where } B \in \mathcal{B}(\mathbb{R}) \text{ and } X^{-1}(B) \in \mathcal{F}$$

$$\textcircled{2} \quad X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right) = \bigcup_{i=1}^{\infty} X^{-1}(B_i) \quad \text{where } B_i \in \mathcal{B}(\mathbb{R}) \text{ and } X^{-1}(B_i) \in \mathcal{F} \forall i \in \mathbb{N}$$

1. Let $\omega \in X^{-1}(B^c) \Rightarrow X(\omega) \in B^c$

$$\Rightarrow X(\omega) \notin B$$

$$\Rightarrow \omega \notin X^{-1}(B)$$

$$\Rightarrow \omega \in X^{-1}(B^c)^c$$

$$\therefore X^{-1}(B^c) \subseteq X^{-1}(B^c)^c$$

Let $\omega \in X^{-1}(B^c)^c \Rightarrow \omega \notin X^{-1}(B)$

$$\Rightarrow X(\omega) \notin B$$

$$\Rightarrow X(\omega) \in B^c$$

$$\Rightarrow \omega \in X^{-1}(B^c)$$

$$\therefore X^{-1}(B^c)^c \subseteq X^{-1}(B^c)$$

$\rightarrow B$

From A and B, $X^{-1}(B)^c = X^{-1}(B^c)$

2. Let $\omega \in X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right)$

$$\Rightarrow X(\omega) \in \bigcup_{i=1}^{\infty} B_i$$

$$\Rightarrow X(\omega) \in B_i \text{ for some } i \in \mathbb{N}$$

$$\Rightarrow \omega \in X^{-1}(B_i) \text{ for some } i \in \mathbb{N}$$

$$\Rightarrow \omega \in \bigcup_{i=1}^{\infty} X^{-1}(B_i)$$

$$\therefore X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right) \subseteq \bigcup_{i=1}^{\infty} X^{-1}(B_i) \quad \rightarrow C$$

Let $\omega \in \bigcup_{i=1}^{\infty} X^{-1}(B_i)$

$$\Rightarrow \omega \in X^{-1}(B_i) \text{ for some } i \in \mathbb{N}$$

$$\Rightarrow X(\omega) \in B_i \text{ for some } i \in \mathbb{N}$$

$\Rightarrow X(\omega) \in B_i$ for some $i \in \mathbb{N}$

$\Rightarrow X(\omega) \in \bigcup_{i=1}^{\infty} B_i$

$\Rightarrow \omega \in X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right)$

$\therefore \bigcup_{i=1}^{\infty} X^{-1}(B_i) \subseteq X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right)$ - D

From C and D, $\bigcup_{i=1}^{\infty} X^{-1}(B_i) = X^{-1}\left(\bigcup_{i=1}^{\infty} B_i\right)$ \square

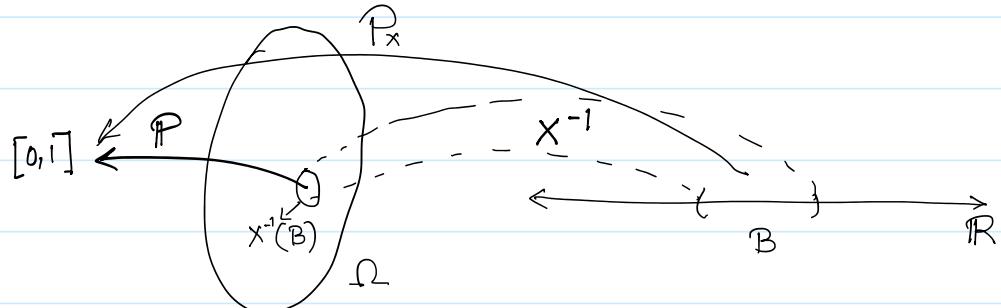
Let (X, \mathcal{X}) and (Y, \mathcal{Y}) be two measurable spaces. Let $f: X \rightarrow Y$. f is called \mathcal{X} -measurable if for every $B \in \mathcal{Y}$, $f^{-1}(B) \in \mathcal{X}$.

\mathcal{Y} -measurable \mathbb{R} -valued function: A function $f: \Omega \rightarrow \mathbb{R}$ is called \mathcal{Y} -measurable if for any $B \in \mathcal{B}(\mathbb{R})$, $f^{-1}(B) \in \mathcal{Y}$.

Probability law induced by random variable X : Probability law of X , $P_x: \mathcal{B}(\mathbb{R}) \rightarrow [0, 1]$ is defined by

$$P_x(B) = P(X^{-1}(B)) \quad \forall B \in \mathcal{B}(\mathbb{R})$$

$$= P(\{\omega \in \Omega : X(\omega) \in B\})$$



Ex.2. $(\mathbb{R}, \mathcal{B}(\mathbb{R}), P_x)$ is a Probability space for any r.v. X .

Solⁿ Given \mathbb{R} be the sample space $\mathcal{B}(\mathbb{R})$ is a σ -algebra on the Sample space. So $(\mathbb{R}, \mathcal{B}(\mathbb{R}))$ is indeed a measurable space. It remains to check that P_x is a probability measure.

space. It remains to check that P_x is a probability measure.

$$1. P_x(\emptyset) = P(x^{-1}(\emptyset)) = P(\emptyset) = 0$$

$$2. P_x(\Omega) = P(x^{-1}(\Omega)) = P(\Omega) = 1.$$

3. Let $(B_i, i \in \mathbb{N})$ is a collection of disjoint sets in $\mathcal{B}(\mathbb{R})$

Claim: $(x^{-1}(B_i), i \in \mathbb{N})$ are also disjoint.

Let $\omega \in x^{-1}(B_i)$ and $\omega \in x^{-1}(B_j)$ for $i \neq j$

$\Rightarrow x(\omega) \in B_i$ and $x(\omega) \in B_j$ but $B_i \cap B_j = \emptyset$

hence $\nexists \omega \in x^{-1}(B_i)$ and $\omega \in x^{-1}(B_j)$ for $i \neq j$

So $(x^{-1}(B_i), i \in \mathbb{N})$ are disjoint sets.

$$P_x\left(\bigcup_{i \in \mathbb{N}} B_i\right) = P\left(x^{-1}\left(\bigcup_{i \in \mathbb{N}} B_i\right)\right)$$

$$= P\left(\bigcup_{i \in \mathbb{N}} x^{-1}(B_i)\right) \quad [\text{Ex 1}]$$

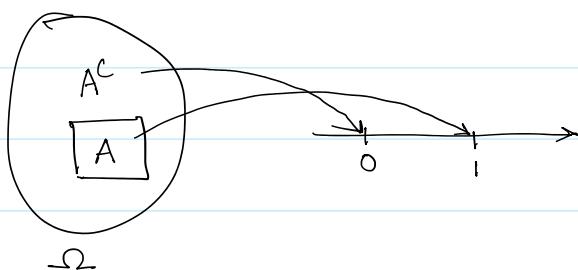
$$= \sum_{i \in \mathbb{N}} P(x^{-1}(B_i)) \quad [\text{as } P \text{ is a prob. meas.}]$$

$$= \sum_{i \in \mathbb{N}} P_x(B_i)$$

□

Indicator Random Variable: $\mathbb{1}_A: \Omega \rightarrow \{0, 1\}$ is call an indicator r.v.

$$\mathbb{1}_A(\omega) = \begin{cases} 1 & \text{if } \omega \in A \\ 0 & \text{if } \omega \notin A \end{cases}$$



Take a Borel set $B \in \mathcal{B}(\mathbb{R})$

$$\mathbb{1}_A^{-1}(B) = \begin{cases} \emptyset & \text{if } 0 \notin B, 1 \notin B \\ A & \text{if } 0 \notin B, 1 \in B \\ A^c & \text{if } 0 \in B, 1 \notin B \\ \Omega & \text{if } 0 \in B, 1 \in B \end{cases}$$

or,

$$\mathbb{1}_A^{-1}(-\infty, x] = \begin{cases} \emptyset & \text{if } x < 0 \\ A^c & \text{if } x \in [0, 1] \\ \Omega & \text{if } x \geq 1 \end{cases}$$

$$F_{\mathbb{1}_A}(x) = P_{\mathbb{1}_A}((-\infty, x]) = P \circ \mathbb{1}_A^{-1}((-\infty, x])$$

$$= \begin{cases} P(\emptyset) & \text{if } x < 0 \\ P(A^c) & \text{if } x \in [0, 1] \end{cases} = \begin{cases} 0 & \text{if } x < 0 \\ 1 - P(A) & \text{if } x \in [0, 1] \\ \dots & \end{cases}$$

$$= \begin{cases} 1 & \text{if } x \in [0, 1) \\ 0 & \text{if } x \geq 1 \end{cases} = \begin{cases} 1 - \mathbb{P}(A) & \text{if } x \in [0, 1) \\ 1 & \text{if } x \geq 1 \end{cases}$$

So, $\mathbb{1}_A$ is $\{\emptyset, A, A^c, \Omega\}$ -measurable.

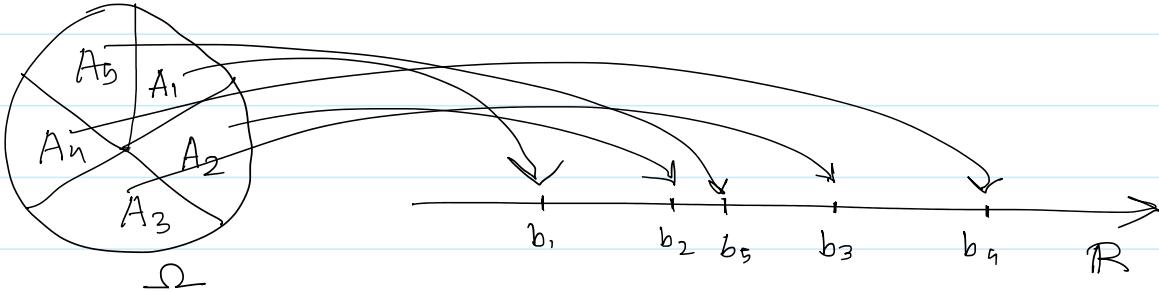
Let $\mathcal{F} \subseteq \mathcal{G}$. $\mathbb{1}_A$ is \mathcal{G} -measurable too.

Simple Random Variable : $X: \Omega \rightarrow \mathcal{X} \subseteq \mathbb{R}$ is called simple r.v. if

$|x| = n$ for some $n \in \mathbb{N}$ Let $\mathcal{X} = \{b_1, \dots, b_n\}$

$$X(\omega) = \sum_{i=1}^n b_i \cdot \mathbb{1}_{A_i}(\omega)$$

$\mathbb{R}, \mathcal{B}(\mathbb{R}), P_x$.



σ -algebra generated by a random variable : Let $X: \Omega \rightarrow \mathbb{R}$ be an \mathcal{F} -measurable random variable defined on $(\Omega, \mathcal{F}, \mathbb{P})$. The smallest event space generated by $A_x(x) = X^{-1}(B_x) = X^{-1}(-\infty, x]$ for $x \in \mathbb{R}$ is called σ -algebra generated by X .

$$\sigma(X) = \sigma(\{A_x(x) : x \in \mathbb{R}\})$$

• Prove that $\sigma(X) := \{X^{-1}(B) : B \in \mathcal{B}(\mathbb{R})\}$

Ex 3. Show that if a simple random variable $X = \sum_{i=1}^n b_i \mathbb{1}_{A_i}$ then $\sigma(X) = \sigma(\{A_i : i \in [n]\})$.

Solⁿ : Let $A \in \sigma(X) \Rightarrow X(A) = B \in \mathcal{B}(\mathbb{R})$.

If $A \neq \emptyset$, $B \in \{b_i : i \in [n]\}$

$$\Rightarrow A = \bigcup_{b \in B} X^{-1}(b) = \bigcup_{i \in [n]} A_i \in \sigma(\{A_i : i \in [n]\})$$

s.t. $b_i \in B$

If $A = \emptyset$, $A \in \sigma(\{A_i : i \in [n]\}) \Rightarrow \sigma(A) \subseteq \sigma(\{A_i : i \in [n]\})$

$\{b_i\}$'s are singleton Borel sets. ← A

$$\therefore X^{-1}\{b_i\} = A_i \in \sigma(X) \quad \forall i \in [n]$$

$$\Rightarrow \{A_i : i \in [n]\} \subseteq \sigma(X)$$

$$\Rightarrow \sigma\{A_i : i \in [n]\} \subseteq \sigma(X) = B.$$

From A & B. result follows □

* σ -algebra generated by random variables are independent

\Leftrightarrow random variables are independent. $F_{XY}(x, y) = F_X(x) F_Y(y)$

$$\text{Ex 4} \quad P(X=x) = F(x) - \lim_{y \uparrow x} F(y)$$

$$\begin{aligned} \text{Soln.} \quad \{\omega : X(\omega) = x\} &= \bigcap_{n \in \mathbb{N}} \{\omega : x - \frac{1}{n} < X(\omega) \leq x\} \\ &= \lim_{n \rightarrow \infty} \{\omega : x - \frac{1}{n} < X(\omega) \leq x\} \end{aligned}$$

$$\therefore P(X=x) = P(\{\omega : X(\omega) = x\})$$

$$= P\left(\lim_{n \rightarrow \infty} \{\omega : x - \frac{1}{n} < X(\omega) \leq x\}\right)$$

$$= \lim_{n \rightarrow \infty} P\left(\{\omega : x - \frac{1}{n} < X(\omega) \leq x\}\right)$$

$$= \lim_{n \rightarrow \infty} P\left(\{\omega : X(\omega) \leq x\} \setminus \{\omega : X(\omega) \leq x - \frac{1}{n}\}\right)$$

$$= F_X(x) - \lim_{n \rightarrow \infty} F_X(x - \frac{1}{n})$$

$$= F_X(x) - \lim_{y \uparrow x} F_X(y)$$

□

Law of total probability: If $(B_n : n \in \mathbb{N})$ is partition of Ω , i.e., $B_n \cap B_m = \emptyset \quad \forall n \neq m, n, m \in \mathbb{N}$ and $\bigcup_{n \in \mathbb{N}} B_n = \Omega$, then for any $A \in \mathcal{F}$

$$P(A) = \sum_{n \in \mathbb{N}} P(A \cap B_n)$$

Conditional Probability: Conditional probability of an event A given another event B s.t. $P(B) > 0$ is

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

From the def'n of conditional probability -

$$\cdot P(A \cap B) = P(A|B) \cdot P(A)$$

From the above equation and Law of total probability -

$$\cdot P(A) = \sum_{n \in I} P(A|B_n) \cdot P(B_n) \quad \text{if } P(B_n) > 0, \quad \forall n$$

From Conditional probability and the above equation -

$$\cdot P(B_i|A) = \frac{P(B_i \cap A)}{P(A)} = \frac{P(A|B_i) \cdot P(B_i)}{\sum_{n \in I} P(A|B_n) \cdot P(B_n)} \quad \text{if } P(A) > 0.$$

• For events A_1, A_2, \dots, A_n satisfying $P(A_1 \cap A_2 \cap \dots \cap A_{n-1}) > 0$, prove that

$$P(A_1 \cap A_2 \cap \dots \cap A_n) = P(A_1) \cdot P(A_2|A_1) \cdot \dots \cdot P(A_n|A_1 \cap A_2 \cap \dots \cap A_{n-1})$$

(multiplication rule)

Ex-5: There are n urns of which the n^{th} urn contains $n-1$ red balls and $n-r$ magenta balls. You pick an urn at random and remove two balls at random w/o replacement. Find the probability that

a. The 2nd ball is magenta.

b. The 2nd ball is magenta given the first is magenta.

c. Given the first ball is magenta, what is the 1:1:1 ratio?

b. The 2nd ball is magenta given the first is magenta.

c. Given the first ball is magenta what is the probability that ith urn was chosen?

Solⁿ. Let the n^{th} urn is picked \equiv event B_n

Chosen an urn, the first draw could be red or magenta

let $D_n^1 \equiv$ The first draw is magenta.

$D_n^1 \equiv$ The first draw is red.

let $D_m^2 \equiv$ The 2nd draw is magenta.

We need to find $P(D_m^2)$.

$\bigcup_{n=1}^n B_n = \Omega$. Using law of total probability

$$P(D_m^2) = \sum_{i=1}^n P(D_m^2 \cap B_i) = \sum_{i=1}^n P(D_m^2 | B_i) \cdot P(B_i)$$

Also, $D_m^1 \cup D_n^1 = \Omega$, So, for fixed n , using law of total probability we get

$$\begin{aligned} P(D_m^2 | B_i) &= P(D_m^1 \cap D_m^2 | B_i) + P(D_n^1 \cap D_m^2 | B_i) \\ &= P(D_m^1 | B_i) P(D_m^2 | B_i \cap D_m^1) \\ &\quad + P(D_n^1 | B_i) P(D_m^2 | B_i \cap D_n^1) \\ &= \frac{n-i}{n-1} \cdot \frac{n-i-1}{n-2} + \frac{i-1}{n-1} \cdot \frac{n-i}{n-2} \\ &= \frac{n-i}{n-1} \end{aligned}$$

$$\begin{aligned} P(D_m^2) &= \sum_{i=1}^n P(B_i) P(D_m^2 | B_i) \\ &= \sum_{i=1}^n \frac{1}{n} \cdot \frac{n-i}{n-1} = \frac{n}{n-1} - \frac{n \cdot (n+1)}{2 \cdot n \cdot (n-1)} \\ &= \frac{2n - n - 1}{2 \cdot (n-1)} \\ &= \frac{n-1}{2(n-1)} = \frac{1}{2} \text{ Ans.} \end{aligned}$$

$$\begin{aligned}
 P(D_m^2 | D_m^1) &= \frac{\sum_{i=1}^n P(D_m^1 \cap D_m^2 \cap B_i)}{\sum_{i=1}^n P(D_m^1 \cap B_i)} \\
 &= \frac{\sum_{i=1}^{n-2} P(D_m^1 \cap D_m^2 \cap B_i)}{\sum_{i=1}^{n-1} P(D_m^1 \cap B_i)} \\
 &= \frac{\sum_{i=1}^{n-2} \cdot P(D_m^2 | D_m^1 \cap B_i) \cdot P(D_m^1 | B_i) \cdot P(B_i)}{\sum_{i=1}^{n-1} P(D_m^1 | B_i) \cdot P(B_i)}
 \end{aligned}$$

$$\begin{aligned}
 &= \frac{\sum_{i=1}^{n-2} \frac{n-i-1}{n-2} \cdot \frac{n-i}{n-1} \cdot \frac{1}{n}}{\sum_{i=1}^{n-1} \frac{n-i}{n-1} \cdot \frac{1}{n}} \\
 &= \frac{n(n-1)}{n(n-1)(n-2)} \sum_{i=1}^{n-2} \frac{n^2 - (2n-1)i + i^2 - n}{\sum_{i=1}^{n-1} n - i} \\
 &= \frac{1}{n-2} \cdot \frac{n(n-1)(n-2) - \frac{(2n-1)(n-1)(n-2)}{2} + \frac{(n-2)(n-1)(2n-3)}{6}}{n^2 - \frac{(n-1)n}{2}} \\
 &= \frac{2}{n(n-1)(n-2)} \cdot \left(n(n-1)(n-2) - \frac{(2n-1)(n-1)(n-2)}{2} + \frac{(n-2)(n-1)(2n-3)}{6} \right) \\
 &= 2 - \frac{2n-1}{n} + \frac{2n-3}{3n} = 2 - 2 + \frac{1}{n} + \frac{2}{3} - \frac{1}{n} \\
 &= \frac{2}{3}
 \end{aligned}$$

$$\begin{aligned}
 P(B_i | D_m^1) &= \frac{P(D_m^1 | B_i) \cdot P(B_i)}{\sum_{i=1}^n P(D_m^1 | B_i) \cdot P(B_i)} \\
 &= \frac{\frac{n-i}{n-1} \cdot \frac{1}{n}}{\sum_{i=1}^n \frac{n-i}{n-1} \cdot \frac{1}{n}} = \frac{n-i}{\sum_{i=1}^n n-i} = \frac{n-i}{n^2 - \frac{n(n+1)}{2}}
 \end{aligned}$$

$$\Leftarrow \frac{2(n-i)}{n(n-1)}$$