

Tutorial 3

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- * An n -dimensional random vector, $X : \Omega \rightarrow \mathbb{R}^n$ s.t.

$$X = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix} \quad \text{where } X_i : \Omega \rightarrow \mathbb{R} \quad \forall i \in [n] \text{ are random variables.}$$

Independence of σ -algebras Let \mathcal{F}_i , $i \in I$ are independent σ -algebras ($\mathcal{F}_i \subseteq \mathcal{F}$) if for any collection of events $(A_i \in \mathcal{F}_i : i \in F)$, with finite set $F \subseteq I$

$$P\left(\bigcap_{i \in F} A_i\right) = \prod_{i \in F} P(A_i)$$

Independence of Random Variables A collection of n random variables $X_i : \Omega \rightarrow \mathbb{R}$, $i \in [n]$ defined on (Ω, \mathcal{F}, P) are independent if

$$F_{x_1, \dots, x_n}(x_1, \dots, x_n) = \prod_{i=1}^n F_{X_i}(x_i) \quad \forall x_i \in \mathbb{R}, i \in [n].$$

$$F_X(x) \quad \text{where } x = (x_1, \dots, x_n).$$

σ -algebra generated by a Random Variable: Let Y be an RV on (Ω, \mathcal{F}, P) . $A_Y(y) = \{y^{-1}(-\infty, y] : y \in \mathbb{R}\}$. The σ -algebra generated by Y ,

$$\sigma(Y) = \sigma(\{A_Y(y) : y \in \mathbb{R}\})$$

Let (Ω, \mathcal{F}, P) be the probability space.

Show that $\{X_i : i \in [n]\}$ is a collection of independent RVs if and only if $\{\sigma(X_i) : i \in [n]\}$ is a collection of independent σ -algebras.

Proof (\Rightarrow) Let $\{\sigma(x_i) : i \in [n]\}$ is a collection of independent σ -algebras. Let us choose arbitrary $x = (x_1, \dots, x_n)$ and define

$$A_{x_1, \dots, x_n}(x) = \bigcap_{i \in [n]} A_{x_i}(x_i) \quad \text{where } A_{x_i}(x_i) \in \sigma(x_i) \forall i \in [n]$$

and $x_i \in \mathbb{R}$.

$$\begin{aligned} \therefore F_{x_1, \dots, x_n}(x) &= P(A_{x_1, \dots, x_n}(x)) \\ &= P\left(\bigcap_{i \in [n]} A_{x_i}(x_i)\right) \\ &= \prod_{i \in [n]} P(A_{x_i}(x_i)) \\ &= \prod_{i \in [n]} F_{x_i}(x_i) \end{aligned}$$

(\Leftarrow) Not in the scope of this course.

Joint CDF of two random variables Let X and Y be two RVs defined on the probability space (Ω, \mathcal{F}, P) .

$$F_{X,Y}(x, y) = P(\{\omega | X(\omega) \leq x, Y(\omega) \leq y\})$$

If X, Y are two independent RVs,

$$\begin{aligned} F_{X,Y}(x, y) &= P(\{\omega | X(\omega) \leq x\}) \cdot P(\{\omega | Y(\omega) \leq y\}) \\ &= F_X(x) \cdot F_Y(y) \quad \forall x, y \in \mathbb{R} \end{aligned}$$

Ex.1 Let $F_{X,Y}$ is the joint distribution of RV X, Y on (Ω, \mathcal{F}, P) . Show that

$$\begin{aligned} P(\{\omega : a < X(\omega) \leq b, c < Y(\omega) \leq d\}) \\ = F_{X,Y}(b, d) - F_{X,Y}(a, d) - F_{X,Y}(b, c) + F_{X,Y}(a, c) \end{aligned}$$

$$\underline{\text{Soln}} \quad P\left(\{\omega : a < X(\omega) \leq b\} \cap \{\omega : c < Y(\omega) \leq d\}\right) =: P(C)$$

$$\text{Let } A = \{\omega : X(\omega) \leq b, c < Y(\omega) \leq d\}$$

$$\text{Let } A = \left\{ \omega : x(\omega) \leq b, c < y(\omega) \leq d \right\}$$

$$B = \left\{ \omega : X(\omega) \leq d, a < Y(\omega) \leq b \right\}$$

$$\begin{aligned} P(C) &= P(A \cap B) = P(A) + P(B) - P(A \cup B) \\ &= F_{xy}(b, d) - F_{xy}(b, c) + F_{xy}(b, d) - F_{xy}(a, d) \\ &\quad - F_{xy}(b, d) + F_{xy}(a, c) \\ &= F_{xy}(b, d) - F_{xy}(b, c) - F_{xy}(a, d) + F_{xy}(a, c) \quad \square \end{aligned}$$

Discrete Random Variable : If a random variable takes value in some countable subset of \mathbb{R} .

Continuous Random Variable : X is called continuous RV if we can write

$$F_x(x) = \int_{-\infty}^x f_x(u) du \quad \text{for } f_x : \mathbb{R} \rightarrow [0, \infty)$$

Ex 2 X and Y have joint distribution function

$$F_{xy}(x, y) = \begin{cases} 0 & \text{if } x < 0 \\ (1-e^{-x}) \left(\frac{1}{2} + \frac{1}{\pi} \tan^{-1}(y) \right) & \text{if } x \geq 0 \end{cases}$$

$$\text{Soln} \quad \frac{\partial}{\partial x} F_{xy}(x, y) = \begin{cases} 0 & \text{if } x < 0 \\ e^{-x} \left(\frac{1}{2} + \frac{1}{\pi} \tan^{-1}(y) \right) & \text{if } x \geq 0 \end{cases}$$

$$\frac{\partial^2}{\partial x \partial y} F_{xy}(x, y) = \begin{cases} 0 & \text{if } x < 0 \\ e^{-x} / \pi (1+y^2) & \text{if } x \geq 0. \end{cases}$$

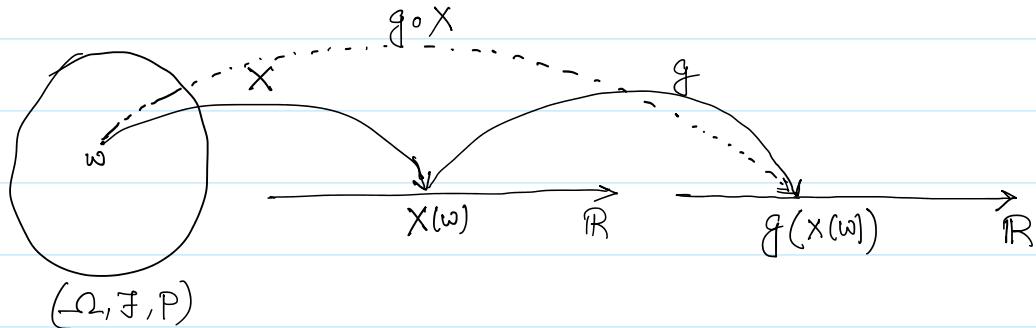
$$\left(= \frac{\partial^2}{\partial y \partial x} F_{xy}(x, y) \right)$$

$$\Rightarrow F_{xy}(u, v) = \int_{-\infty}^u du \int_{-\infty}^v dv \left(\frac{\partial}{\partial u} \left(\frac{\partial}{\partial v} F_{xy}(u, v) \right) \right)$$

$$\therefore f_{xy}(x, y) = \begin{cases} 0 & \text{if } x < 0 \\ e^{-x} / \pi (1+y^2) & \text{if } x \geq 0 \end{cases}$$

$$\therefore f_{XY}(x, y) = \begin{cases} e^x / \pi(1+y^2) & \text{if } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

Transformation of Random Variables



Q1. Is $g \circ X$ a RV? If yes, when?

Q2. How to find the probability law of $g \circ X$?

1. One sufficient condition is if g is Borel measurable function.
2. If g is monotonically increasing.

$$\begin{aligned} F_{g \circ X}(y) &= P(\{w : g(X(w)) \leq y\}) \\ &= P(\{w : X(w) \leq g^{-1}(y)\}) \\ &= F_X(g^{-1}(y)) \\ &= \int_{-\infty}^{g^{-1}(y)} f_X(x) dx \end{aligned}$$

As g is monotonically increasing function, $g(x) \leq y \Rightarrow x \leq g^{-1}(y)$

Let $x = g^{-1}(t)$ so $g'(x) dx = dt$

$$\therefore F_{g \circ X}(y) = \int_{-\infty}^y f_X(g^{-1}(t)) \cdot \frac{dt}{g'(g^{-1}(t))}$$

$$\Rightarrow f_{g \circ X}(y) = f_X(g^{-1}(y)) \cdot \frac{1}{g'(g^{-1}(y))}$$

If g is monotonically decreasing, $f_{g \circ X}(y) = f_X(g^{-1}(y)) \cdot \frac{-1}{g'(g^{-1}(y))}$

For any monotonic function, g

$$f_{g \circ X}(y) = \frac{f_X(g^{-1}(y))}{|g'(g^{-1}(y))|}$$

Ex 3. Let X be a Gaussian RV of mean 0 and variance 1. i.e,
 $X \sim N(0, 1)$. Find the distribution of $Y = e^X$.

Solⁿ . $g(x) = e^x$, $g'(x) = e^x$, $g^{-1}(y) = \ln(y)$, $g'(\ln(y)) = y$

$$f_Y(y) = \frac{f_X(\ln(y))}{y} = \frac{1}{\sqrt{2\pi}} e^{-\frac{(\ln y)^2}{2}} \text{ for } y > 0$$

Ex 4. Let X_1, \dots, X_n be random variables on (Ω, \mathcal{F}, P) with joint CDF $F_{X_1 \dots X_n}$. Define

$$Y = \min(X_1, \dots, X_n) \text{ and}$$

$$Z = \max(X_1, \dots, X_n)$$

- a. Show that Y and Z are Random variables
- b. Find distributions of Y & Z .

Solⁿ. a. Consider the event

$$\begin{aligned} \{\omega \in \Omega : Z(\omega) \leq z\} &= \{\omega \in \Omega : X_1(\omega) \leq z, \dots, X_n(\omega) \leq z\} \\ &= \bigcap_{i=1}^n \{\omega \in \Omega : X_i(\omega) \leq z\} \in \mathcal{F} \end{aligned}$$

$\Rightarrow Z$ is an RV.

$$\begin{aligned} \{\omega \in \Omega : Y(\omega) \leq y\} &= \{\omega \in \Omega : Y(\omega) > y\}^c \\ &= \{\omega \in \Omega : X_1(\omega) > y, \dots, X_n(\omega) > y\}^c \\ &= \left(\bigcap_{i=1}^n \{\omega \in \Omega : X_i(\omega) > y\} \right)^c \\ &= \bigcup_{i=1}^n \{\omega \in \Omega : X_i(\omega) > y\}^c \in \mathcal{F}. \end{aligned}$$

$\Rightarrow Y$ is an RV

$$\begin{aligned} b. \quad F_z(z) &= P(\{\omega \in \Omega : Z(\omega) \leq z\}) \\ &= P\left(\bigcap_{i=1}^n \{\omega \in \Omega : X_i(\omega) \leq z\}\right) \\ &= F_{X_1 \dots X_n}(z) \end{aligned}$$

$$\begin{aligned} \bar{F}_y(y) &= P(\{\omega \in \Omega : Y(\omega) > y\}) \\ &= P\left(\bigcap_{i=1}^n \{\omega \in \Omega : X_i(\omega) > y\}\right) \\ &= \bar{F}_{X_1 \dots X_n}(z). \end{aligned}$$

• What if X_1, \dots, X_n iid random variables?

$$\begin{aligned} 1. \quad F_z(z) &= F_{X_1 \dots X_n}(z) = F_{X_1}(z) \dots F_{X_n}(z) \\ &= F_{X_1}(z)^n \end{aligned}$$

$$\begin{aligned} 2. \quad \bar{F}_y(y) &= \bar{F}_{X_1 \dots X_n}(y) = \bar{F}_{X_1}(y) \dots \bar{F}_{X_n}(y) \\ &= (\bar{F}_{X_1}(y))^n \end{aligned}$$

$$\begin{aligned} F_y(y) &= P(\{\omega \in \Omega : Y(\omega) \leq y\}) = P(\{\omega \in \Omega : Y(\omega) > y\}^c) \\ &\equiv 1 - P(\{\omega \in \Omega : Y(\omega) > y\}) \\ &= 1 - \bar{F}_y(y) \end{aligned}$$

$$\therefore 1 - F_y(y) = (1 - \bar{F}_{X_1}(y))^n$$

$$\Rightarrow F_y(y) = 1 - (1 - \bar{F}_{X_1}(y))^n$$

Ex 5. Let X_1, \dots, X_n be independent exponentially distributed RVs with parameters $\lambda_1, \dots, \lambda_n$. $F_{X_i}(x) = 1 - e^{-\lambda_i x}$ for $x > 0$. Let $Y = \min(X_1, \dots, X_n)$. Show that Y is exponentially distributed with parameter $\lambda_1 + \dots + \lambda_n$.

Solⁿ $F_y(y) = 1 - \prod_{i=1}^n (1 - F_{X_i}(y))$

$$\begin{aligned}
 \underline{\text{Sol}}^n \quad F_Y(y) &= 1 - \prod_{i=1}^n (1 - F_{X_i}(x)) \\
 &= 1 - \prod_{i=1}^n e^{-\lambda_i x} \\
 &= 1 - e^{-(\sum_{i=1}^n \lambda_i)x}
 \end{aligned}$$

$$\Rightarrow Y \sim \text{Exp}\left(\sum_{i=1}^n \lambda_i\right).$$

□