BAYESIAN ESTIMATION OF RISK-NEUTRAL PROBABILITY

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IMPLIED RISK-NEUTRAL PROBABILITY DENSITY

$$S = \sum_{s=1}^{K} \phi(s) q(s)$$

$$\phi^* \stackrel{d}{=} \frac{\phi(s)}{\sum_{s=1}^{K} \phi(s)} = \frac{\phi(s)}{1/(1+r)} = (1+r)\phi(s)$$

$$S = \frac{1}{1+r} \sum_{s=1}^{K} \phi^*(s) q(s) = \frac{1}{1+r} \mathbb{E}^*[q(s)].$$

IMPLIED RISK-NEUTRAL PROBABILITY DENSITY

For TAS utility, the state-prices are

$$\phi(s) = \pi \frac{\beta u'[c(s)]}{u'[c(0)]},$$

$\rightsquigarrow$ implied risk-neutral probability is

$$\phi^*(s) = \pi \frac{\beta u'[c(s)]}{u'[c(0)]}(1+r) = \pi \frac{u'[c(s)]}{u'[c(0)]} \approx \pi,$$

for u' is constant (risk-neutrality) or $c(s) \approx c(0)$.

BREEDEN-LITZENBERGER (1978) FORMULA

For short-term, $r \approx 0$.

$$
\begin{aligned}
C(S,t) &= \mathbb{E}^* \left[[S(T) - \mathcal{K}]^+ \mid S(t) = S \right] \\
&= \int_0^{+\infty} (x - \mathcal{K})^+ \, dP^*(x) \\
&= \int_{\mathcal{K}}^{+\infty} x \, dP^*(x) - \mathcal{K} \int_0^{+\infty} \, dP^*(x) \\
&= \int_{\mathcal{K}}^{+\infty} x \, dP^*(x) - \mathcal{K}(1 - P^*(\mathcal{K}))
\end{aligned}
$$

$$
\begin{aligned}
\frac{\partial C}{\partial \mathcal{K}} &= -\mathcal{K} p^*(\mathcal{K}) - 1 + P^*(\mathcal{K}) + \mathcal{K} p^*(\mathcal{K}) = P^*(\mathcal{K}) - 1 \\
\frac{\partial^2 C}{\partial \mathcal{K}^2} &= p^*(\mathcal{K}).
\end{aligned}
$$

Ait-Sahalia and Duarte (2003)

BAYESIAN FORMULATION

Mark Fisher (2016)

$$y_i = \lambda \sum_{j=1}^{K} X_{ij}\beta_j + \varepsilon_i,$$

where $\varepsilon_i \sim N(0, \sigma^2)$, and $\beta = (\beta_1, \ldots, \beta_K) \in \Delta^{K-1}$, where Δ^{K-1} denotes the simplex of dimension $(K-1)$.

In vector form,

$$y = \lambda X\beta + \varepsilon.$$

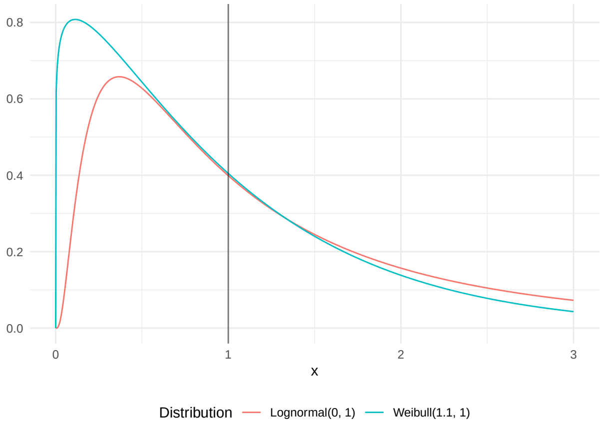
In order to insure $\beta \in \Delta^{K-1}$, we select the (symmetric) Dirichlet prior:

$$p(\beta \mid \alpha) = \mathrm{Dirichlet}(\beta \mid \alpha) \propto \prod_{i=1}^{K} x_i^{\alpha-1}.$$

PRIOR FOR CONCENTRATION COEFFICIENT

α is a vector of concentration parameters for Dirichlet distribution.

It is reasonable to select a distribution that is relatively flat on $[0, 1]$. Mark Fisher (2016) uses lognormal prior.



PRIORS AND POSTERIOR

$$p(\beta \mid \alpha) = \text{Dirichlet}(\beta \mid \alpha) \propto \prod_{i=1}^{K} x_i^{\alpha-1}.$$

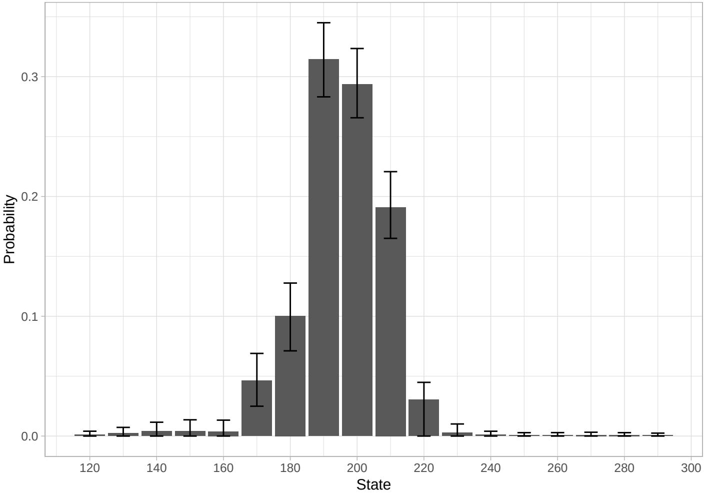
$$p(\alpha) = \text{Weibull}(1.1, 1)$$

$$p(\sigma^2) \propto \frac{1}{\sigma^2}$$

$$p(\alpha, \lambda, \sigma, \beta \mid y) \propto N(y \mid \lambda X\beta, \sigma) \times \text{Dirichlet}(\beta \mid \alpha) \times \frac{1}{\sigma^2}$$

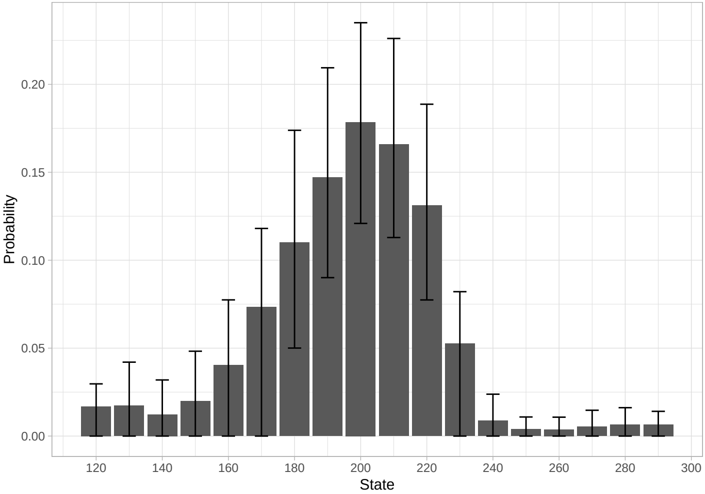
EXAMPLE. AAPL

Calculated in 2025-05-23 with expiration date 2025-06-13

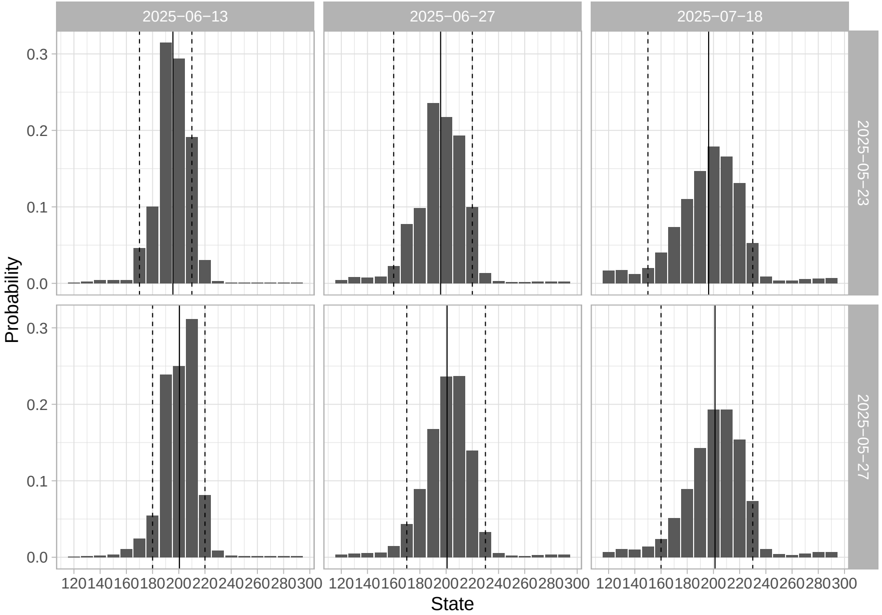


EXAMPLE. AAPL

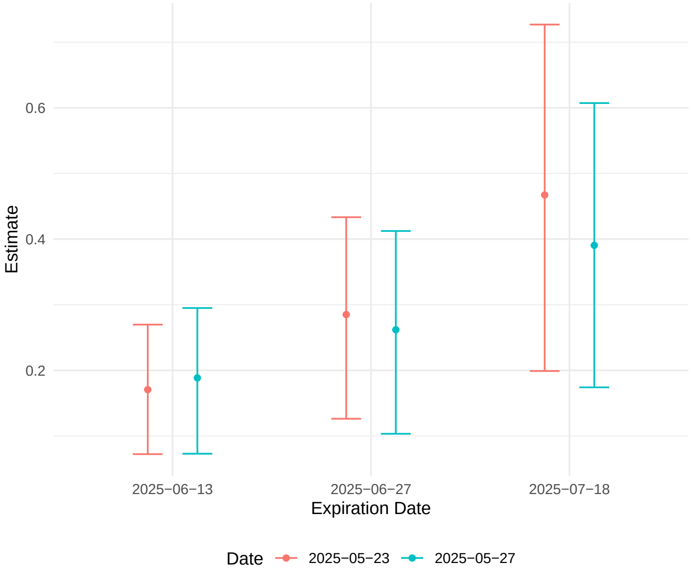
Calculated in 2025-05-23 (same) with expiration date 2025-07-18 (+35 days).



BETA ESTIMATES



ALPHA ESTIMATES



USE CASES

REFERENCES [1]

Ait-Sahalia, Yacine and Jefferson Duarte (Sept. 2003). "Nonparametric Option Pricing under Shape Restrictions". In: *Journal of Econometrics* 116(1), pp. 9–47. ISSN: 0304-4076. DOI: 10.1016/S0304-4076(03)00102-7. (Visited on 05/31/2025).

Breeden, Douglas T. and Robert H. Litzenberger (1978). "Prices of State-Contingent Claims Implicit in Option Prices". In: *The Journal of Business* 51(4), pp. 621–651. ISSN: 0021-9398. JSTOR: 2352653. (Visited on 05/31/2025).

Mark Fisher (May 2016). *Simplex Regression*.