

Spatial Gaussian Process Regression for Bayesian Optical Flow Estimation

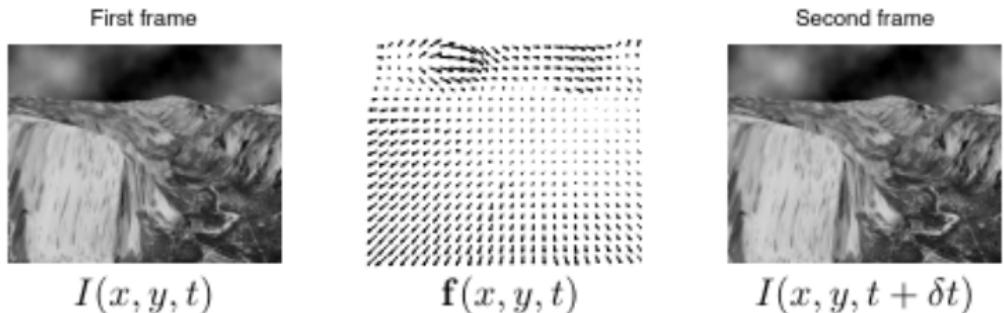
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Motion in Image Sequences



Given two sequential images $I : \mathcal{D} \times [0, T] \rightarrow \mathcal{R}$, i.e.

$$I(x, y, t) \text{ and } I(x, y, t + \delta t)$$

(for fixed t and δt), the optical flow at time t is a vector field

$$\mathbf{f}(x, y, t) = (u(x, y, t), v(x, y, t))$$

that transforms one image into the next:

$$I(x + u(x, y, t), y + v(x, y, t), t + \delta t) \simeq I(x, y, t)$$

Lucas-Kanade Method

Brightness constancy assumption:

$$\frac{d}{dt} I(x(t), y(t), t) = \frac{\partial I}{\partial x} u + \frac{\partial I}{\partial y} v + \frac{\partial I}{\partial t} = 0, \quad u := \frac{dx}{dt}, v := \frac{dy}{dt}$$

Underdetermined, so assume same $(u(x, y), v(x, y)) \forall (x, y) \in W$:

$$\underbrace{\begin{bmatrix} I_x(x_1, y_1) & I_y(x_1, y_1) \\ \vdots & \vdots \\ I_x(x_n, y_n) & I_y(x_n, y_n) \end{bmatrix}}_{\nabla I^T} \underbrace{\begin{bmatrix} u \\ v \end{bmatrix}}_f = \underbrace{\begin{bmatrix} -I_t(x_1, y_1) \\ \vdots \\ -I_t(x_n, y_n) \end{bmatrix}}_{-I_t}$$

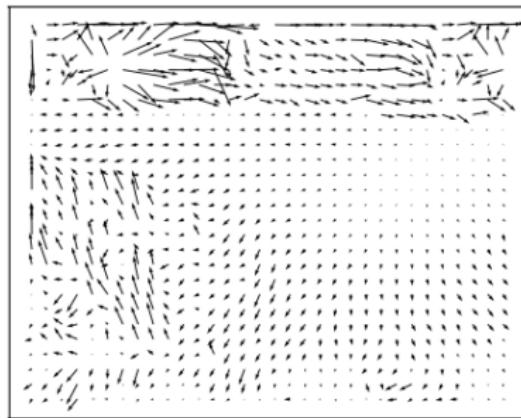
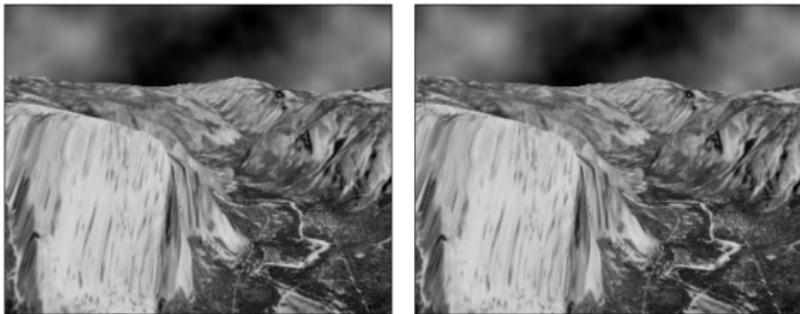
Overdetermined: least squares solution is $v = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T b$:

$$\begin{bmatrix} u \\ v \end{bmatrix} = \left[\begin{array}{cc} \sum I_x(x_i, y_i)^2 & \sum I_x(x_i, y_i) I_y(x_i, y_i) \\ \sum I_x(x_i, y_i) I_y(x_i, y_i) & \sum I_y(x_i, y_i)^2 \end{array} \right]^{-1} \begin{bmatrix} -\sum I_x(x_i, y_i) I_t(x_i, y_i) \\ -\sum I_y(x_i, y_i) I_t(x_i, y_i) \end{bmatrix}$$

Optional weighting of pixels via diagonal matrix \mathbf{W} :

$$v = (\mathbf{A}^T \mathbf{W} \mathbf{A})^{-1} \mathbf{A}^T \mathbf{W} b$$

Lucas-Kanade Results



Barron, J. L., Fleet, D. J., & Beauchemin, S. S. (1994). Performance of Optical Flow Techniques. *International Journal of Computer Vision*, 12(1), 43-77.

Probabilistic Motivation

Classical methods produce deterministic point estimates.

Uncertainty is inherent in optical flow estimation, due to:

- ▶ Image noise
- ▶ Brightness changes
- ▶ Low contrast regions
- ▶ Object occlusion
- ▶ Aperture problem
- ▶ Incompatible motions in localized regions

Quantification of confidence is desirable for safety-critical applications, e.g.:

- ▶ Autonomous navigation
- ▶ Computer-integrated surgery
- ▶ Real-time surveillance

Wang-Orquiza Method

Suppose $\mathbf{y} = \mathbf{Ax} + \boldsymbol{\eta}$, \mathbf{A} fixed, \mathbf{x} fixed unknown, \mathbf{y} observed, $\boldsymbol{\eta} \sim \mathcal{N}(\mathbf{0}, \mathbf{H})$:

$$\hat{\mathbf{x}}_{\text{MLE}}(\mathbf{y}) = \arg \max_{\mathbf{x}} f_{\mathbf{y}}(\mathbf{y} | \mathbf{x}) = \arg \max_{\mathbf{x}} f_{\boldsymbol{\eta}}(\mathbf{y} - \mathbf{Ax}) = \boldsymbol{\Sigma}_{\hat{\mathbf{x}}} \mathbf{A}^T \mathbf{H}^{-1} \mathbf{y}$$

$$\boldsymbol{\Sigma}_{\hat{\mathbf{x}}} = \text{Cov}[\hat{\mathbf{x}}_{\text{MLE}}(\mathbf{y}) - \mathbf{x}] = (\mathbf{A}^T \mathbf{H}^{-1} \mathbf{A})^{-1}$$

Same assumptions as Lucas-Kanade, plus additive Gaussian noise:

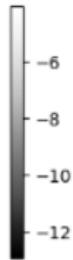
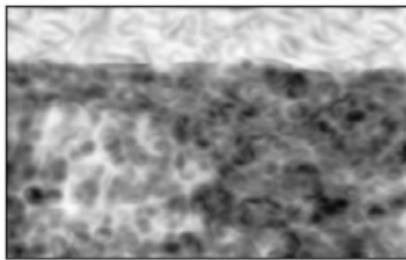
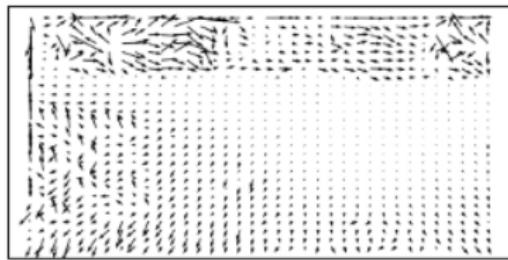
$$\mathbf{I}_t = -\nabla I^T \mathbf{f} + \boldsymbol{\eta} \quad \boldsymbol{\eta} \sim \mathcal{N}(\mathbf{0}, \mathbf{H})$$

Results above give:

$$\hat{\mathbf{f}} = -\boldsymbol{\Sigma}_{\hat{\mathbf{f}}} \nabla I \mathbf{H}^{-1} \mathbf{I}_t \quad \boldsymbol{\Sigma}_{\hat{\mathbf{f}}} = (\nabla I \mathbf{H}^{-1} \nabla I^T)^{-1}$$

Poor texture around $\mathbf{x} \implies$ small $\nabla I \implies$ large $\boldsymbol{\Sigma}_{\hat{\mathbf{f}}}$.

Wang-Orquiza Results



Spatial Gaussian Processes

2-dimensional Gaussian process (distribution over functions of 2D inputs):

$$\mathbf{f}(\cdot) = \begin{bmatrix} u(\cdot) \\ v(\cdot) \end{bmatrix} \sim \mathcal{GP}(\mathbf{m}(\cdot), \mathbf{k}(\cdot, \cdot'))$$

Example mean & covariance kernel (affine & RBF):

$$\mathbf{m}(\cdot) = \mathbf{A}(\cdot) + \mathbf{b}, \quad \mathbf{k}(\cdot, \cdot') = \exp\left(-\frac{\|\cdot - \cdot'\|^2}{2\lambda^2}\right) \boldsymbol{\Sigma}$$

Joint distribution of function values at finite subset of points:

$$\begin{bmatrix} \mathbf{f}(\mathbf{x}_1) \\ \vdots \\ \mathbf{f}(\mathbf{x}_n) \end{bmatrix} \sim \mathcal{N} \left(\begin{bmatrix} \mathbf{m}(\mathbf{x}_1) \\ \ddots \\ \mathbf{m}(\mathbf{x}_n) \end{bmatrix}, \begin{bmatrix} \mathbf{k}(\mathbf{x}_1, \mathbf{x}_1) & \cdots & \mathbf{k}(\mathbf{x}_1, \mathbf{x}_n) \\ \vdots & \ddots & \vdots \\ \mathbf{k}(\mathbf{x}_n, \mathbf{x}_1) & \cdots & \mathbf{k}(\mathbf{x}_n, \mathbf{x}_n) \end{bmatrix} \right)$$

Equivalently:

$$\mathbf{F} := \begin{bmatrix} \mathbf{U}(\mathbf{X}) \\ \mathbf{V}(\mathbf{X}) \end{bmatrix} \sim \mathcal{N} \left(\boldsymbol{\mu} := \begin{bmatrix} \boldsymbol{\mu}_u \\ \boldsymbol{\mu}_v \end{bmatrix}, \boldsymbol{\mathcal{K}} := \begin{bmatrix} \mathbf{K}_{uu} & \mathbf{K}_{uv} \\ \mathbf{K}_{vu} & \mathbf{K}_{vv} \end{bmatrix} \right)$$

Gaussian Process Regression

Through Wang-Orquiza method, we obtain noisy observations $\tilde{\mathbf{F}}$:

$$\tilde{\mathbf{f}}(\mathbf{x}) = \mathbf{f}(\mathbf{x}) + \boldsymbol{\eta}(\mathbf{x}), \quad \mathbf{H}(\mathbf{x}) \sim \mathcal{N}(\mathbf{0}, \Sigma_{\tilde{\mathbf{f}}})$$

The noisy distribution is:

$$\tilde{\mathbf{F}} | \mathbf{F} \sim \mathcal{N}(\mathbf{F}, \Sigma_{\tilde{\mathbf{f}}}), \quad \mathbf{F} \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{K}) \implies \tilde{\mathbf{F}} \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{K} + \Sigma_{\tilde{\mathbf{f}}})$$

We want to estimate true optical flows: we form the joint distribution.

$$\begin{bmatrix} \mathbf{F} \\ \tilde{\mathbf{F}} \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix} \boldsymbol{\mu} \\ \boldsymbol{\mu} \end{bmatrix}, \begin{bmatrix} \mathbf{K} & \mathbf{K} \\ \mathbf{K} & \mathbf{K} + \Sigma_{\tilde{\mathbf{f}}} \end{bmatrix}\right)$$

We derive the conditional distribution (posterior):

$$\mathbf{F} | \tilde{\mathbf{F}} \sim \mathcal{N}\left(\boldsymbol{\mu} + \mathbf{K}(\mathbf{K} + \Sigma_{\tilde{\mathbf{f}}})^{-1}(\tilde{\mathbf{F}} - \boldsymbol{\mu}), \mathbf{K} - \mathbf{K}(\mathbf{K} + \Sigma_{\tilde{\mathbf{f}}})^{-1}\mathbf{K}\right)$$

Mean serves as estimate, covariance as uncertainty.

Parameter Fitting

Mean/covariance parameters (e.g. lengthscale λ^2) are optimized by maximizing the marginal likelihood of observations $\tilde{\mathbf{F}}$. This likelihood is:

$$p(\tilde{\mathbf{F}}) \propto \frac{1}{\sqrt{\det(\mathbf{K} + \boldsymbol{\Sigma}_{\tilde{f}})}} \exp\left(-\frac{1}{2} (\tilde{\mathbf{F}} - \boldsymbol{\mu})^T (\mathbf{K} + \boldsymbol{\Sigma}_{\tilde{f}})^{-1} (\tilde{\mathbf{F}} - \boldsymbol{\mu})\right)$$

In practice we minimize negative log-likelihood via gradient descent:

$$-\log p(\tilde{\mathbf{F}}) \propto \frac{1}{2} (\tilde{\mathbf{F}} - \boldsymbol{\mu})^T (\mathbf{K} + \boldsymbol{\Sigma}_{\tilde{f}})^{-1} (\tilde{\mathbf{F}} - \boldsymbol{\mu}) + \frac{1}{2} \det(\mathbf{K} + \boldsymbol{\Sigma}_{\tilde{f}})$$

We minimize this numerically via gradient descent.

Parameter Fitting (cont.)

Let φ be a parameter of \mathbf{m} , $\mathbf{y} = \tilde{\mathbf{F}} - \mu$, $\mathbf{K}_y = \mathbf{K} + \sigma_\eta^2 \mathbf{I}$:

$$\frac{\partial \mathcal{L}}{\partial \varphi} = \frac{1}{2} \left[\left(\frac{\partial \mathbf{y}^T}{\partial \varphi} \right) \mathbf{K}_y^{-1} \mathbf{y} + \mathbf{y}^T \mathbf{K}_y \left(\frac{\partial \mathbf{y}}{\partial \varphi} \right) \right] = \frac{1}{2} \left(2 \mathbf{y}^T \mathbf{K}_y \left(\frac{\partial \mathbf{y}}{\partial \varphi} \right) \right) = -\mathbf{y}^T \mathbf{K}_y \left(\frac{\partial \mathbf{m}}{\partial \varphi} \right)$$

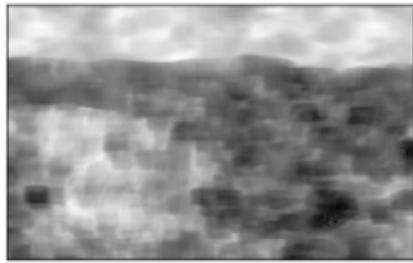
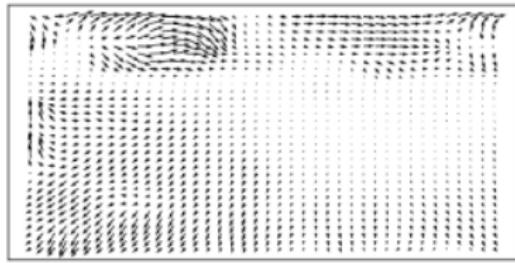
Let θ be a parameter of \mathbf{k} :

$$\frac{\partial \mathcal{L}}{\partial \theta} = \mathbf{y}^T \left(\frac{\partial \mathbf{K}_y^{-1}}{\partial \theta} \right) \mathbf{y} + \frac{\partial |\mathbf{K}_y|}{\partial \theta} = -\frac{1}{2} \mathbf{y}^T \mathbf{K}^{-1} \left(\frac{\partial \mathbf{K}_y}{\partial \theta} \right) \mathbf{K}^{-1} \mathbf{y} + \frac{1}{2} \text{tr} \left(\mathbf{K}_y^{-1} \left(\frac{\partial \mathbf{K}_y}{\partial \theta} \right) \right)$$

RBF parameter λ^2 :

$$\frac{\partial \mathbf{K}_y}{\partial \lambda} = \mathbf{K} \odot \frac{\|\mathbf{x} - \mathbf{x}'\|}{\lambda^3}$$

GP Regression Results



Discussion

Our maximum-likelihood method produced an initial estimate and uncertainty.

Our Gaussian-process Bayesian method provided updated these (posterior), after enforcing a global smoothness constraint (Gaussian process prior).

Possible next steps:

- ▶ Obtain better observations: FlowNet, RAFT, etc.
- ▶ Experiment with more complex covariance kernels.
- ▶ Perform segmentation, smooth each segment independently.
- ▶ Extend to spatiotemporal regression (3D, multiple frames).