MA 519: Homework 6

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Problem 6.1 (Handout 8, # 2)

Identify the parameters n and p for each of the following binomial distributions:

- (a) # boys in a family with 5 children;
- (b) # correct answers in a multiple choice test if each question has a 5 alternatives, there are 25 questions, and the student is making guesses at random.

SOLUTION. For part (a), the distribution is binomial with k being the number of children in a given family and p the probability that a child is born, say, male. In this case, we can reasonably assume that p = 0.5. Thus, the binomial distribution is given by Binom(5,0.5).

For part (b), we use similar reasoning and we have Binom(25, 0.2) where k = 25 is the number of questions and p = 1/5 = 0.2 the probability of guessing a question correctly.

Problem 6.2 (Handout 8, # 10)

A newsboy purchases papers at 20 cents and sells them for 35 cents. He cannot return unsold papers. If the daily demand for papers is modeled as a Binom(50, 0.5) random variable, what is the optimum number of papers the newsboy should purchase?

SOLUTION. Let $X \sim \text{Binom}(50, 0.5)$ denote the daily demand for papers and n the number of copies bought by the newsboy. Then, we must find the ℓ such that the sum

$$35 \sum_{k=\ell}^{50} {50 \choose k} 0.5^{50} - 20 \sum_{k=0}^{\ell-1} {50 \choose k} 0.5^{50} = 35 \left(1 - \sum_{k=0}^{\ell-1} {50 \choose k} 0.5^{50} \right) - 20 \sum_{k=0}^{\ell-1} {50 \choose k} 0.5^{50}$$

$$= 35 - 55 \sum_{k=0}^{\ell-1} {50 \choose k} 0.5^{50}$$

$$= 35 - 55 \cdot 0.5^{51-\ell}$$

equals 0. This can be computed experimetally

Problem 6.3 (Handout 8, # 12)

How many independent bridge dealings are required in order for the probability of a preassigned player having four aces at least once to be 1/2 or better? Solve again for some player instead of a given one.

SOLUTION.

Problem 6.4 (Handout 8, # 13)

A book of 500 pages contains 500 misprints. Estimate the chances that a given page contains at least three misprints.

SOLUTION. Let X be the number of misprints on the given page. The probability that a given misprint is on that page is 1/500. Now, $P(X \ge 3) = 1 - P(X = 0) - P(X = 1) - P(X = 2)$. Also,

$$P(X = 0) = \left(\frac{499}{500}\right)^{500}$$

$$P(X = 1) = 500 \left(\frac{1}{500}\right) \left(\frac{499}{500}\right)^{499}$$

$$P(X = 2) = \frac{500 \cdot 499}{2} \left(\frac{1}{500}\right)^{2} \left(\frac{499}{500}\right)^{498}$$

So that

$$\begin{split} P(X \ge 3) &= 1 - P(X = 0) - P(X = 1) - P(X = 2) \\ &= 1 - \left(\frac{499}{500}\right)^{500} - 500\left(\frac{1}{500}\right)\left(\frac{499}{500}\right)^{499} - \frac{500 \cdot 499}{2}\left(\frac{1}{500}\right)^2\left(\frac{499}{500}\right)^{498} \\ &\approx 0.08 \end{split}$$

that is, the probability that the given page has at least 3 misprints is about 8 percent.

Problem 6.5 (Handout 8, # 14)

Colorblindness appears in 1 per cent of the people in a certain population. How large must a random sample (with replacements) be if the probability of its containing a colorblind person is to be 0.95 or more?

SOLUTION. Let n be the sample size. The probability of the sample containing no colorblind people is 0.99^n . Solving the equation $0.99^n = 0.05$, we see that (for the naturals) taking n = 299 is (minimally) sufficient for 0.99^n to be less than 0.05.

The probability that the sample has some colorblind person is equal to $1 - 0.99^n$. This is at least 95 percent if 0.99^n is less than 0.05. That is, having 299 people in the sample is (minimally) sufficient for there to be a 95 percent chance of having some colorblind person.

Problem 6.6 (Handout 8, #15)

Two people toss a true coin n times each. Find the probability that they will score the same number of heads.

Solution. Let X denote the number of heads that, say, person 1 gets. Then

$$P(X=k) = \frac{\binom{n}{k}}{2^n}.$$

Then, assuming independence, the probability that they score the same number of heads is given by the expression

$$\sum_{k=0}^{n} \left(\frac{\binom{n}{k}}{2^n} \cdot \frac{\binom{n}{k}}{2^n} \right) = \frac{1}{2^{2n}} \sum_{k=0}^{n} \binom{n}{k}^2,$$

which, by the binomial identity of the sum of squares of binomial coefficients, gives us

$$= \frac{1}{2^{2n}} \binom{2n}{n}.$$

MA 519: Homework 6

Problem 6.7 (Handout 8, # 16)

(Binomial approximation to the hypergeometric distribution.) A population of TV elements is divided into red and black elements in the proportion p:q (where p+q=1). A sample of size n is taken without replacement. The probability that it contains exactly k red elements is given by the hypergeometric distribution of II, 6. Show that as $n \to \infty$ this probability approaches Binom(n, p).

SOLUTION. The hypergeometric distribution has the following PMF

$$\begin{split} P(X=k) &= \binom{r}{k} \binom{n-r}{pn-k} \bigg/ \binom{n}{pn} = \frac{r!(n-r)!(pn)!(qn)!}{k!(r-k)!(pn-k)!(qn-(r-k))!n!} \\ &= \binom{pn}{k} \left(\frac{r!(n-r)!(qn)!}{(r-k)!(qn-(r-k))!n!} \right) \\ &= \binom{pn}{k} \left(\frac{r!/(r-k)!}{n!/(n-k)!} \cdot \frac{(n-r)!/(n-pn-(r-k))!}{(n-pn+(pn-k))!/(qn)!} \right) \\ &= \binom{pn}{k} \prod_{j=1}^k \frac{r-k+j}{n-k+j} \prod_{j=1}^{np-k} \frac{n-r-(np-k)+j}{n-np+j}. \end{split}$$

Now, take the limit as $n \to \infty$

Problem 6.8 (Handout 9, # 3)

Suppose X, Y, Z are mutually independent random variables, and E(X) = 0, E(Y) = -1, E(Z) = 1, $E(X^2) = 4$, $E(Y^2) = 3$, $E(Z^2) = 10$. Find the variance and the second moment of 2Z - Y/2 + eX, where e is the number such that $\ln e = 1$.

SOLUTION. Let W = 2Z - Y/2 + eX. Then

$$\begin{split} E(W^2) &= E((2Z - Y/2 + eX)^2) \\ &= E(4Z^2 - 2ZY + 4eZX + Y^2/4 - eYX + e^2X^2) \\ &= 4E(Z^2) - 2E(Z)E(Y) + 4eE(Z)E(X) + E(Y^2)/4 - eE(Y)E(X) + e^2E(X^2) \\ &= 4 \cdot 10 + 2 + 3/4 + e^2 \cdot 4 \\ &= \frac{171}{4} + 4e^2 \\ &\approx 72.31 \end{split}$$

and

$$Var(W) = E(W^{2}) - E(W)^{2}$$

$$= \frac{171}{4} + 4e^{2} - (2E(Z) + E(Y)/2 + eE(X))^{2}$$

$$= \frac{171}{4} + 4e^{2} - (2 - 1/2)^{2}$$

$$= \frac{171}{4} + 4e^{2} - (3/2)^{2}$$

$$= \frac{81}{2} + 4e^{2}$$

$$\approx 70.06$$

That is, the second moment is "about 72.31" and the variance is "about 70.06"

Problem 6.9 (Handout 9, # 14)

(Variance of Product). Suppose X, Y are independent random variables. Can it ever be true that Var(XY) = Var(X) Var(Y)? If it can, when?

SOLUTION. Yes. In fact, if X and Y are constant random variables $\operatorname{Var} X = \operatorname{Var} Y = 0$ so clearly

$$Var(XY) = Var X Var Y = 0.$$

The problem is finding sufficient conditions for this to be true. Since X and Y are independent, we know that

$$E(XY) = E(X)E(Y)$$

and therefore

$$Var(XY) = E(X^{2}Y^{2}) - E(XY)^{2}$$
$$= E(X^{2})E(Y^{2}) - E(X)^{2}E(Y)^{2}$$

and we want this to be equal to

$$\begin{aligned} \operatorname{Var} X \operatorname{Var} Y &= \left(E(X^2) - E(X)^2 \right) \left(E(Y^2) - E(Y)^2 \right) \\ &= E(X^2) E(Y^2) + E(X)^2 E(Y)^2 - E(X)^2 E(Y^2) - E(X^2) E(Y)^2 \\ &= E(X^2) E(Y^2) - E(X)^2 E(Y^2) \\ &- \left(E(X)^2 E(Y^2) - E(X)^2 E(Y)^2 + E(X^2) E(Y^2) - E(X)^2 E(Y^2) \right) \\ &= \operatorname{Var}(XY) - \operatorname{Var}(Y) E(X^2) - \operatorname{Var}(X) E(Y^2). \end{aligned}$$

This happens precisely when the quantity

$$Var(Y)E(X^{2}) + Var(X)E(Y^{2}) = 0.$$

Since $\operatorname{Var} X, \operatorname{Var} Y, E(X^2), E(Y^2) \geq 0$, this forces the equality

$$Var(Y)E(X^2) = Var(X)E(Y^2) = 0.$$

This happens when both X = Y = 0 or, more generally, if and only if X and Y are constant.