MA 519: Homework 4

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September 22, 2016

PROBLEM 4.1 (HANDOUT 5, # 2)

In an urn, there are 12 balls. 4 of these are white. Three players: A, B, and C, take turns drawing a ball from the urn, in the alphabetical order. The first player to draw a white ball is the winner. Find the respective winning probabilities: assume that at each trial, the ball drawn in the trial before is put back into the urn (i.e., selection *with replacement*).

SOLUTION. Denote the events that player A wins, player B wins, and player C wins by A, B, and C respectively.

Note that A wins if some multiple of 3 losses occurs, followed by a win. Also, B wins if some multiple of three losses occurs, followed by a loss, then a win. Last, C wins if some multiple of three losses occurs, followed by two losses then a win.

A win occurs with probability 1/3 each time, and a loss occurs with probability 2/3 each time. Thus, because each draw is independent,

$$P(A) = \sum_{i=0}^{\infty} \frac{1}{3}\left(\frac{2}{3}\right)^{3i}, \qquad P(B) = \sum_{i=0}^{\infty} \frac{1}{3}\left(\frac{2}{3}\right)^{3i+1}, \qquad P(C) = \sum_{i=0}^{\infty} \frac{1}{3}\left(\frac{2}{3}\right)^{3i+2}.$$

That is,

$$P(A) = \sum_{i=0}^{\infty} \frac{9}{19} \approx 0.47, \qquad P(B) = \sum_{i=0}^{\infty} \frac{6}{19} \approx 0.32, \qquad P(C) = \sum_{i=0}^{\infty} \frac{4}{19} \approx 0.21.$$

∎

Problem 4.2 (Handout 5, # 8)

Consider n families with 4 children each. How large must n be to have a 90% probability that at least 3 of the n families are all girl families?

Solution. The probability that a family has all girls is $(0.5)^4$.

In n families, the probability that at least 3 are all-girl is $1 - P(0) - P(1) - P(2)$, where $P(m)$ is the probability that exactly m families are all-girl.

Note that

$$P(0) = (1 - (0.5)^4)^n$$
$$P(1) = \binom{n}{1}(0.5)^4(1 - (0.5)^4)^{n-1}$$
$$P(2) = \binom{n}{2}(0.5)^8(1 - (0.5)^4)^{n-2}$$

(we choose a number of families to be all-girl, and then find the probability that that family is all-girl, while all of the rest are not all-girl).

So that the probability that at least 3 are all girl is

$$1 - \left(1 - (0.5)^4\right)^n - n(0.5)^4\left(1 - (0.5)^4\right)^{n-1} - n(n-1)(0.5)(0.5)^8\left(1 - (0.5)^4\right)^{n-2}.$$

∎

PROBLEM 4.3 (HANDOUT 5, # 10)

(*Yahtzee*). In Yahtzee, five fair dice are rolled. Find the probability of getting a Full House, which is three rolls of one number and two rolls of another, in Yahtzee.

SOLUTION. There are 30 different kinds of full house (6 different three of a kinds, and 5 different kinds of different two of a kind).

The probability of rolling a specific kind of full house is

$$\binom{5}{3}\left(\frac{1}{6}\right)^3\left(\frac{1}{6}\right)^2$$

(Choose 3 dice to be the three of a kind, have them all rolled the specific number, have the other two dice rolled the other specific number.)

So the probability of rolling some kind of full house is

$$30\binom{5}{3}\left(\frac{1}{6}\right)^3\left(\frac{1}{6}\right)^2 = \frac{25}{648} \approx 0.039.$$

∎

PROBLEM 4.4 (HANDOUT 5, # 12)

The probability that a coin will show all heads or all tails when tossed four times is 0.25. What is the probability that it will show two heads and two tails?

SOLUTION. Let H denote the event that, on a given coin toss, that coin is heads, and let T denote the event that, on a given coin toss, that coin is tails. Then, because each coin toss is independent, this says that

$$P(H)^4 + P(T)^4 = \frac{1}{4}.$$

Moreover, because this is a coin,

$$P(H) + P(T) = 1.$$

So,

$$P(H)^4 + (1 - P(H))^4 = \frac{1}{4}$$

$$2P(H)^4 - 4P(H)^3 + 6P(H)^2 - 4P(H) + 1 = \frac{1}{4}$$

We can approximate a solution to the above by $P(H) \approx 0.299$ or by $P(H) \approx 0.701$.

Now, the probability that two coin flips of four are heads is

$$\binom{4}{2} P(H)^2 P(T)^2 \approx 0.263.$$

∎

PROBLEM 4.5 (HANDOUT 5, # 13)

Let the events $A_1, A_2, \ldots, A_n$ be independent and $P(A_k) = p_k$. Find the probability p that none of the events occurs.

SOLUTION. Let the events $A_1, A_2, \ldots, A_n$ be independent. Then the events $A_1^\sim, A_2^\sim, \ldots, A_n^\sim$ are independent. (Where A denotes the event that A does *not* happen, i.e., $A^\sim = \Omega \setminus A$). The events $A_i^\sim$ each have probability $1 - p_i$ of occurring. Thus, the probability, p, that none of the events occur is

$$p = \prod_{k=1}^{n} (1 - p_k)$$

∎

Problem 4.6 (Handout 6, # 5)

Suppose a fair die is rolled twice and suppose X is the absolute value of the difference of the two rolls. Find the PMF and the CDF of X and plot the CDF. Find a median of X; is the median unique?

Solution. First, we compute the probability mass function. Note that X is an integer between 0 and 5, so computing the probilities that X is each of 0 through 5 suffices to describe the PMF of X.
    We calculate the probabilities, by counting.

$$P(0) = \frac{6}{36} = \frac{1}{6}$$
$$P(1) = \frac{10}{36} = \frac{5}{18}$$
$$P(2) = \frac{8}{36} = \frac{2}{9}$$
$$P(3) = \frac{6}{36} = \frac{1}{6}$$
$$P(4) = \frac{4}{36} = \frac{1}{9}$$
$$P(5) = \frac{2}{36} = \frac{1}{18}$$

We calculate the CDF by summing.

$$\mathrm{CDF}(0) = \frac{6}{36} = \frac{1}{6}$$
$$\mathrm{CDF}(1) = \frac{16}{36} = \frac{4}{9}$$
$$\mathrm{CDF}(2) = \frac{24}{36} = \frac{2}{3}$$
$$\mathrm{CDF}(3) = \frac{30}{36} = \frac{5}{6}$$
$$\mathrm{CDF}(4) = \frac{34}{36} = \frac{17}{18}$$
$$\mathrm{CDF}(5) = \frac{36}{36} = 1$$

The median is 2.   ■

PROBLEM 4.7 (HANDOUT 6, # 7)

Find a discrete random variable X such that $E(X) = E(X^3) = 0$; $E(X^2) = E(X^4) = 1$.

SOLUTION. Set $\Omega = \{0, 1\}$ and define a random variable $X \colon \Omega \to \mathbb{R}$ by $X(0) = -1$, $X(1) = 1$ as well as a probability $P(0) = P(1) = 1/2$. Then

$$E(X) = -1(1/2) + 1(1/2) = 0 = (-1)^3(1/2) + 1^3(1/2) = E(X^3),$$

whereas

$$E(X^2) = (-1)^2(1/2) + 1^2(1/2) = 1 = (-1)^4(1/2) + 1^4(1/2) = E(X^4),$$

as desired.   ■

PROBLEM 4.8 (HANDOUT 6, # 9)

(*Runs*). Suppose a fair die is rolled n times. By using the indicator variable method, find the expected number of times that a six is followed by at least two other sixes. Now compute the value when $n = 100$.

SOLUTION. Let Ω denote the sample space and let A denote the event that in a sequence of n rolls of a die, a subsequence of sixes of length at least three occurs. Define a random variables $X_i \colon \Omega \to \mathbb{R}$, $1 \le i \le n-2$, as the indicator variable that the there are three or more consecutive 6s starting at the ith roll. Then,

$$E(X_1 + \cdots + X_{n-2}) = E(X_1) + \cdots + E(X_{n-2}) = p_1 + \cdots + p_{n-2}.$$

We need to find these p_i. There are 6^n points in the sample space Ω. Starting at the ith place, we can have exactly 5^{n-i-2} ways of choosing the remaining terms in the sequence. Thus,

$$p_i = \frac{5^{n-2}}{6^n}$$

■

Problem 4.9 (Handout 6, # 10)

(*Birthdays*). For a group of n people find the expected number of days of the year which are birthdays of exactly k people. (Assume 365 days and that all arrangements are equally probable.)

Solution. Let Ω denote the sample space and let A denote the event that exactly k people share a birthday. Define a random variable $X_i \colon \Omega \to \mathbb{R}$, $1 \le i \le 365$, as the indicator variable that the ith day of the year is the birthday of exactly k people. Then,

$$E(X_1 + \cdots + X_{365}) = E(X_1) + \cdots + E(X_{365}) = 365p,$$

where p is the probability that a exactly k people have a given day of the year as their birthday.

The latter probability is not difficult to compute. There are exactly $\binom{n}{k}$ ways to chose k people having the 1st of January as their birthday and $(365 - 1)^{n-k} = 364^{n-k}$ choices for we can assign to the remaining $n - k$ people and of course, the sample space has cardinality $\#\Omega = 365^n$. Thus, the probability that exactly k people have the 1st of January as their birthday is

$$P = \frac{\binom{n}{k}364^{n-k}}{365^n}.$$

Thus,

$$E[X] = 365\left(\frac{\binom{n}{k}364^{n-k}}{365^n}\right) = \frac{\binom{n}{k}364^{n-k}}{365^{n-1}}.$$

∎

PROBLEM 4.10 (HANDOUT 6, # 11)

(*Continuation*). Find the expected number of multiple birthdays. How large should n be to make this expectation exceed 1?

SOLUTION. For any given person, the probability that the other $n-1$ people do not share a birthday with them is

$$\frac{\binom{365}{1}(365-1)^{n-1}}{365^n} = \frac{365 \cdot 364^{n-1}}{365^n} = \left(\frac{364}{365}\right)^{n-1}.$$

Thus, the probability that any given person shares a birthday with somebody else is

$$p = 1 - \left(\frac{364}{365}\right)^{n-1}.$$

Now, define a random variable $X_i \colon \Omega \to \mathbb{R}$, $1 \leq i \leq n$, the indicator random variable that the ith person shares a birthday. Then

$$E(X_1 + \cdots + X_n) = E(X_1) + \ldots + E(X_1)$$
$$= np$$
$$= n\left(1 - \left(\frac{364}{365}\right)^{n-1}\right).$$

For the expectation to exceed 1, we must have

$$1 < n\left(1 - \left(\frac{364}{365}\right)^{n-1}\right)$$
$$\frac{1}{n} < 1 - \left(\frac{364}{365}\right)^{n-1}$$
$$\frac{365^n}{n} < 365^n - 364^n.$$

Experimentally, this number seems to be 20.   ∎

PROBLEM 4.11 (HANDOUT 6, # 12)

(*The blood-testing problem*). A large number, N, of people are subject to a blood test. This can be administered in two ways, (i) Each person can be tested separately. In this case N tests are required, (ii) The blood samples of k people can be pooled and analyzed together. If the test is negative, this one test suffices for the k people. If the test is positive, each of the k persons must be tested separately, and in all $k+1$ tests are required for the k people. Assume the probability p that the test is positive is the same for all people and that people are stochastically independent.

(b) What is the expected value of the number, X, of tests necessary under plan (ii)?
(c) Find an equation for the value of k which will minimize the expected number of tests under the second plan. (Do not try numerical solutions.)

SOLUTION.   ∎

PROBLEM 4.12 (HANDOUT 6, # 13)

(*Sample structure*). A population consists of r (classes whose sizes are in the proportion $p_1 : p_2 : \cdots : p_r$. A random sample of size n is taken with replacement. Find the expected number of classes not represented in the sample.

SOLUTION. Let A_i be the event that the ith class is not represented in the sample of size n. Let A be the number of classes not represented in the sample of size n. Let A_i also represent the number of times that the ith class is not represented. Then

$$
\begin{aligned}
E(A) &= \sum_{i=1}^{r} E(A_i) \\
&= \sum_{i=1}^{r} P(A_i) \\
&= \sum_{i=1}^{r} (1 - p_i)^n
\end{aligned}
$$

∎