

Accuracy and precision of zero-heat-flux temperature monitoring: a systematic review and meta- analysis

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1 Introduction

Core temperature monitoring is an important intervention within perioperative and intensive care settings. Thermoregulatory dysfunction is commonly associated with the induction of anesthesia and can lead to adverse outcomes including cardiac dysrhythmias, altered hemostasis and increased risk of surgical site infection (Frank et al. 1997; Kurz, Sessler, and Lenhardt 1996; Michelson et al. 1994; Rohrer and Natale 1992). The pulmonary artery catheter is the reference standard for continuous core temperature monitoring(*ref – troublefindingastatementregardingthisintheliterature*), however, the invasive nature of this method renders an increased risk of bloodstream infections and damage to surrounding tissue(Hadian and Pinsky 2006). Although clinically accurate relative to pulmonary artery temperature measurements, surrogate measures of core temperature at esophageal, rectal, and bladder locations remain mildly invasive interventions(*ref*).

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Zero-heat-flux thermometry is an alternative non-invasive method that allows for continuous monitoring of core temperature. Originally developed in the 1970s, zero-heat-flux technology was recently implemented in the 3M SpotOn (3M, St Paul, MN), a single-use, disposable sensor (*ref*). In practice, the zero-heat-flux sensor, such as the 3M SpotOn, is placed on the lateral surface of the forehead and initially warmed to equilibrate the temperature of the skin surface to the underlying core tissues(*ref*). Equipped with a thermal insulator, the zero-heat-flux sensor eliminates heat loss to the environment to allow for changes in core temperature to be directly reflected by a change in skin surface temperature (*ref*).

The agreement between zero-heat-flux thermometers and core as well as other peripheral thermometers has been investigated in several studies conducted within clinical settings. Appraisal of these studies and synthesis of the results would aid clinicians in deciding the appropriate circumstances in which zero-heat-flux thermometers may be used. We aimed to determine if zero-heat-flux thermometers have clinically acceptable accuracy and precision relative to established core and peripheral temperature measurement devices. Accuracy is defined as the average difference between temperature measurements from the zero-heat-flux and comparator device and precision as the variance (standard deviation) in the differences.

2 Methods

A systematic review was conducted in accordance with a predetermined protocol. The protocol was submitted to PROSPERO for registration but due to delays in processing we had started data extraction by the time it was reviewed. Therefore, the protocol did not meet the requirements for registration. A copy of the submitted protocol can be accessed [here](#). The primary comparison for this review was temperature measured from a zero-heat-flux thermometer versus a temperature measured from a core site, which we defined as arterial, esophageal, bladder and rectal sites. Secondary comparisons were made between zero-heat-flux thermometers and temperatures taken at peripheral sites.

2.1 Inclusion criteria

Observational studies that reported temperature measurements from a zero-heat-flux thermometer and comparator thermometer were included. Studies involving a case control design were excluded due to potential for overestimation of the intervention performance. Studies were excluded if conducted on non-human subjects or outside of a clinical healthcare setting. No publication date restrictions were applied. Published conference abstracts were included if there was enough information reported to appraise the quality of the study. There were no language restrictions applied during the search.

2.2 Data sources and searches

Published studies were found by searching Medline and EMBASE from January 2000 to July 2019. The Cochrane-recommended search strategy combining terms for the ‘target condition’ and ‘index test’ was used (Mann and Gilbody 2012). This search strategy is an efficient approach for systematic reviews of diagnostic test accuracy studies. (Preston et al. 2015) We also conducted forward citation searching, by using Google Scholar to search the citations of the first article published on the accuracy of zero-heat-flux thermometers. The search strategies used for each data base can be located in file 1. Selection of studies was undertaken independently by two reviewers using Covidence.

2.3 Data extraction and quality assessment

Information was extracted regarding study characteristics (author, year of publication, country, design, sample size, clinical setting, numbers studied and analyses for each outcome), population characteristics (inclusion and exclusion criteria) and temperature measurement characteristics (placement of sensor, timing and methods of measurements). The outcomes that were extracted included the mean bias (eg, accuracy) and variance (eg, SD, precision) in temperature measurement between the zero-heat-flux and comparator thermometers. We also extracted information about how repeated measurements were handled. In particular we assessed whether studies: (1) analysed each pair of data separately; (2) treated each pair of data as independent; or (3) used either analysis of variance or a random effects model as a way to control for the dependent nature of the repeated measures data (Myles and Cui 2007).

Two reviewers independently assessed the risk of bias for the included studies using the revised Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2) (Whiting et al. 2011). Reviewers rated the risk of bias for patient selection, conduct of the zero-heat-flux measurements, conduct of the comparator thermometer measurements, and timing and flow (eg, timing of ZHF and established core temperature measurements, dropouts) as “high”, “low” or “unclear” risk of bias. We worked to minimize the risk of publication bias by conducting a comprehensive search of multiple databases as well as an international clinical trial registry (Glasziou et al. 2001). Statistical approaches for detection of reporting bias were not conducted due to lack of validated methods (Begg 2005). ASimulations have revealed that tests for detecting funnel plot asymmetry will result in publication bias being incorrectly identified too often (Deeks, Macaskill, and Irwig 2005).

In order to rate the quality of evidence, we applied the Grading Quality of Evidence and Strength of Recommendation methodology (Schünemann et al. 2008). Evidence was downgraded in accordance with study limitations, inconsistency and imprecision. There were no circumstances in which evidence was downgraded for indirectness as this systematic review only included relevant

studies. Although the possibility of publication bias was not excluded, this bias was not formally assessed as it was not considered sufficient enough to reason downgrading the quality of evidence.

2.4 Data synthesis and analysis

The objective for the meta-analysis was to estimate the population limits of agreement between temperature measurements from the ZHF and established comparator thermometers. A framework for meta-analysis of Bland-Altman method comparison studies developed by Tipton and Shuster (2017) based on limits of agreement approach was used. This method was selected because it parallels the approach used in primary Bland-Altman studies, whereby an estimate is generated for the pooled LoAs in the population (not just in the samples studied). The ‘population LoA’ is more broad than the LoA commonly reported in the meta-analyses of Bland-Altman studies (Tipton and Shuster 2017). Pooled limits of agreement were calculated using $\delta \pm 2\sqrt{\sigma^2 + \tau^2}$, where δ is the average bias across studies, σ^2 is the average within-study variation in differences and τ^2 is the variation in bias across studies.

Estimations of δ and σ^2 were made using a weighted least-squares model (similar to a random-effects approach), with the associated standard errors estimated using robust variance estimation. Robust variance estimation was used alternatively to model-based standard errors as some studies included in the systematic review used repeated-measures designs without accommodating for the correlation between measurements (Hedges, Tipton, and Johnson 2010; Tanner-Smith, Tipton, and Polanin 2016; Tipton 2015). We used the method-of-moments estimator from DerSimonian and Laird (1986) for the τ^2 parameter.

Measures of uncertainty were included in our meta-analyses by calculating the outer 95% confidence intervals for pooled limits of agreement. We also accounted for repeated measurements if they were not properly adjusted for in individual studies. This was achieved by using weights proportional to the number of participants, not the total number of measurements. The R statistical program was used to conduct all analyses (Team 2017). All data and R code used in the meta-analyses can be located here.

Prior to conducting the meta-analyses, the results from each study were converted into a standard format, with bias meaning *comparatorthermometer - ZHFthermometer* measured in $^{\circ}C$. In several studies, results were reported for multiple groups of participants, therefore in the meta-analysis each of these groups was treated as a separate ‘comparison’. Other studies reported multiple sets of results, whereby analyses were conducted between ZHF and various comparator devices used on the same participant. These instances were also treated as a separate ‘comparison’ if the comparator devices were apart of separate meta-analyses groups. One study reported intraoperative, postoperative and overall results for the same participants. Only the paired measurements from the overall results were included in the main and low risk bias analyses,

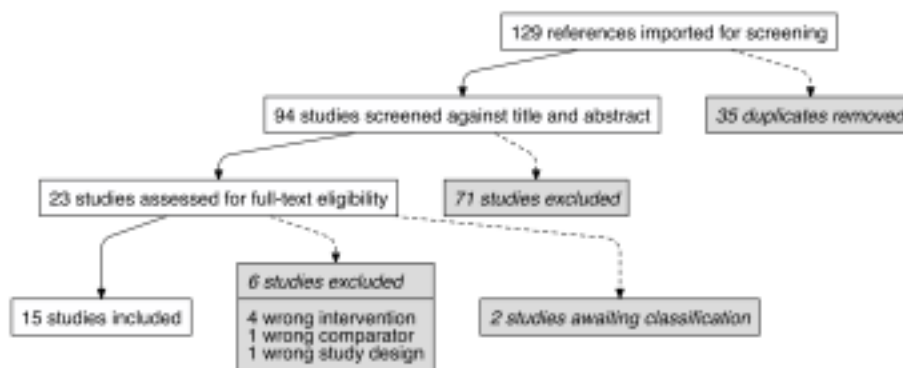


Fig. 1 PRISMA Flow Diagram

leaving paired measurements exclusively from the intraoperative and postoperative timepoints to be included in respective meta-analyses subgroups.

The conventionally cited clinically acceptable agreement between ZHF and comparator devices is 0.5°C (*ref—most of the studies sit in this range from Eshragi...*). It was deemed that outer confidence bounds for 95% LoA between zero-heat-flux and core temperature measurements (termed as ‘population limits of agreement’) outside of these bounds would be clinically unacceptable.

A sensitivity analysis for the primary comparison (ZHF versus temperature measurement at core site) was performed based on risk of bias, whereby ‘unclear risk of bias’ was treated as ‘high risk’ and ‘high risk of bias’ studies were excluded from analyses. As clinicians would be interested in the accuracy of ZHF relative to the thermometer devices they use, and within the clinical setting in which they use it, we conducted subgroup analyses according to the comparator device used (either core, sublingual or nasopharyngeal) and clinical setting (either intraoperative or intensive care unit).

3 Results

Fifteen studies were included in this systematic review (Figure 1). Two studies reported only in abstract form were not included and assigned as ‘studies awaiting classification’ because there was insufficient information provided.

The characteristics of each study are in online supplementary file 1.

The primary comparison of zero-heat-flux versus core temperature measurements (eg, arterial, bladder, esophageal or rectal) consisted of 19 comparisons from 13 individual studies. In total, data from 576 participants with 179821 paired measurements were included in this comparison (two studies did not report the total number of measurements included in their analysis). The sensitivity analysis for the primary comparison with only studies that were judged as low risk of bias across all domains included 10 comparisons

from 5 studies that enrolled 273 participants with 104294 paired measurements (two studies did not report the total number of measurements included in their analysis). The third group compared zero-heat-flux to core temperature measurements in ICU patients and included 7 (23%) comparisons from 5 (31%) studies, and enrolled 246 participants with 155,598 paired measurements. The fourth group compared ZHF to core temperature measurement in intraoperative patients and included 16 (53%) comparisons from 10 (63%) studies, and enrolled 464 participants with 24,223 paired measurements. The fifth group compared ZHF to nasopharyngeal temperature measurements and included 4 (13%) comparisons from 4 (25%) studies, and enrolled 297 participants with 109,819 paired measurements. The sixth group compared ZHF to sublingual temperature measurement and included 2 (7%) comparisons from 2 (13%) studies, and enrolled 107 participants with 22731 paired measurements. The studies included in this systematic review evaluated a ZHF temperature monitoring system manufactured by 3M. Previously known as the SpotOn Temperature Monitoring System, the 3M ZHF device is now referred to as the Bair Hugger Temperature Monitoring System. All comparisons reported adherence to the ZHF device manufacturer instructions and placed the sensor on the forehead of the participants. High risk of bias was associated with patient selection for 13 (43%) comparisons from 9 (56%) studies, conduct of ZHF and comparator measurements in 10(33%) comparisons from 8 (50%) studies and 12 (40%) comparisons 9 (30%) studies, respectively (mostly due to ZHF measurements being taken with knowledge of the comparator measurements and vice versa) and participant flow for 13 (43%) studies. In 19 (50%) comparisons from 8 (50%) studies, the authors had declared a conflict of interest or receipt of funding or equipment from the manufacturer of the ZHF device under evaluation.

3.1 Study selection and description

3.2 Agreement between ZHF and core temperature measurements

4 Discussion

4.1 Limitations

5 Conclusion

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