Assignment 2: CS215

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Solution 1

Task A

When $X \sim \text{Ber}(p)$, PGF of X is

$$G_{\text{Ber}}(z) = \mathbb{E}(z^X) \tag{1}$$

$$=\sum_{n=0}^{\infty}P[X=n]z^n\tag{2}$$

Since P[X = 0] = (1 - p), P[X = 1] = p, P[X = n] = 0 when n > 1,

$$G_{\text{Ber}}(z) = P[X=0]z^0 + P[X=1]z^1$$
 (3)

$$= (1 - p) + pz \tag{4}$$

Task B

When $X \sim Bin(n, p)$, PMF of X is

$$P[X=k] = \binom{n}{k} p^k (1-p)^{n-k} \text{ for } k \le n.$$
 (5)

and P[X = k] = 0 for k > n.

$$G_{\text{Bin}}(z) = \sum_{k=0}^{\infty} P[X=k] z^k \tag{6}$$

$$= \sum_{k=0}^{n} \binom{n}{k} p^{k} (1-p)^{n-k} z^{k}$$
 (7)

$$= \sum_{k=0}^{n} \binom{n}{k} (pz)^k (1-p)^{n-k}$$
 (8)

$$= (1 - p + pz)^n. (9)$$

By equation 4, $G_{Bin}(z) = (1 - p + pz)^n = (G_{Ber}(z))^n$. Hence proved.

Task D

When $X \sim \text{Geo}(p)$, PMF of X,

$$P[X = k] = (1 - p)^{k-1}p (10)$$

for k > 0. P[X = 0] = 0. Now, PGF of X,

$$G_{\text{Geo}}(z) = \sum_{k=0}^{\infty} P[X=k] z^k$$
(11)

$$=\sum_{k=1}^{\infty}P[X=k]z^k\tag{12}$$

$$= \sum_{k=1}^{\infty} p(1-p)^{k-1} z^k \tag{13}$$

$$= \sum_{k=1}^{\infty} pz(z - zp)^{k-1}$$
 (14)

$$=pz\sum_{k=0}^{\infty}(z-zp)^k\tag{15}$$

$$=\frac{pz}{1-z+pz}\tag{16}$$

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Task E

By equation 9, $G_{Bin}(z) = (1 - p + pz)^n = G_X^{(n,p)}(z)$. For $Y \sim \text{NegBin}(n,p)$

$$P[Y = k] = {\binom{k-1}{n-1}} p^n (1-p)^{k-n} \text{ for } k \ge n$$
(17)

Otherwise, P[Y = k] = 0. PGF of Y is

$$G_Y^{(n,p)}(z) = \sum_{k=0}^{\infty} P[Y=k] z^k$$
 (18)

$$= \sum_{k=n}^{\infty} {k-1 \choose n-1} p^n (1-p)^{k-n} z^k$$
 (19)

$$= \sum_{k=0}^{\infty} {k+n-1 \choose n-1} p^n (1-p)^k z^{n+k}$$
 (20)

$$= (pz)^n \sum_{k=0}^{\infty} {k+n-1 \choose n-1} (z-pz)^k$$
 (21)

We know $\sum_{k=0}^{\infty} {k+n-1 \choose n-1} x^k = (1-x)^{-n}$. Thus

$$G_Y^{(n,p)}(z) = (pz)^n (1 - z + pz)^{-n}$$
(22)

$$= \left((1 - p^{-1} + p^{-1}z^{-1})^n \right)^{-1} \tag{23}$$

$$= (G_X^{(n,p^{-1})}(z^{-1}))^{-1}. (24)$$

Hence Proved.

Solution 2

Task A

To prove:

Let X be a continuous real-valued random variable with CDF: $\mathbb{R} \to [0,1]$. Assume that F_X is invertible. Then the random variable $Y := F_X(X) \in [0,1]$ is uniformly distributed in [0,1]

Proof:

 F_X by definition can also be written as

$$F_X(x) = P(X \le x)$$

Define a new random variable Y,

$$Y = F_X(X)$$

Y is the result of applying CDF F_X to the random variable X. To prove the theorem, assume $y \in [0,1]$. So, the probability that $Y \leq y$ is:

$$P(Y \le y) = P(F_X(X) \le y)$$

It is assumed that $F_X(x)$ is invertible, so,

$$P(Y \le y) = P(X \le F_X^{-1}(y))$$

which is basically, probablity that X is less that or equal to $F_X^{-1}(y)$. This can be written in the CDF form, which is $F_X(F_X^{-1}(y))$. So,

$$P(Y \le y) = P(X \le F_X^{-1}(y)) = F_X(F_X^{-1}(y)) = y$$

So,

$$P(Y \le y) = y$$

where $y \in [0,1]$, which is the CDF of uniform distribution in [0,1]. So, Y is a uniform distribution in [0,1] regardless of X.

Task B

According to the theorem proved above, CDF of any random variable X mapped with itself gives a uniform random variable Y in [0,1]. So, let $Y \sim \text{Uniform}(0,1)$. Then for any random variable X,

$$F_{\mathbf{X}}(\mathbf{X}) = \mathbf{Y} \tag{25}$$

$$X = F_X^{-1}(Y) \tag{26}$$

Algorithm A:

- 1. Input: A sample y from the uniform distribution on [0,1].
- 2. Transformation:
 - Apply the inverse CDF to *y* to compute a sample *u*.
 - Define $A(u) = u = F_X^{-1}(y)$
- 3. Output: The random variable $U = F_X^{-1}(Y)$

This gives us the correct required random variables as, CDF of U is $F_U(u)$,

$$P(U \le u) = P(F_X(Y) \le u) \tag{27}$$

$$F_U(u) = P(F_X(F_X^{-1}(X) \le u))$$
 (28)

$$F_{U}(u) = P(X \le u) \tag{29}$$

$$F_{U}(u) = F_{X}(u) \tag{30}$$

(31)

U and *X* have the same CDF, which was initially required.

Solution 3

Solution 4

Solution 5