Chapter 2: Probability and Statistics

Jonathan Roth

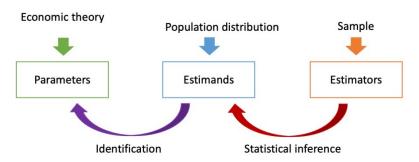
Mathematical Econometrics I Brown University

Course Logistics

- Problem set 1 is posted. It is due on Friday September 19 at 4PM as a GradeScope submission
- TA sessions and OHs start this week. See the aCanvas page for times/tentative locations
- Any logistical questions?

"Big Picture" Recap

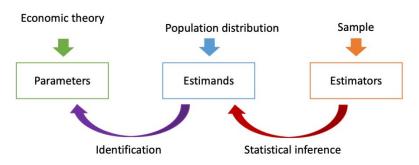
Recall our division of labor:



- **Statistics**: how does the sample data we observe relate to observable features of the population we're interested in?
- **Identification**: how do observable features of population relate to target parameters of interest?

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- **Identification**: how do observable features of population relate to target parameters of interest?
- For both these tasks, we need a mathematical language for talking about how data is generated. Enter probability and statistics



Outline

- 1. Random Variables and Probability Distributions
- 2. Means and Variances
- 3. Identification in Experiments
- 4. Random Sampling and Sample Means
- 5. Hypothesis Testing and Inference

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The realization of a random process is called a random variable.

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Random Variables and CDFs

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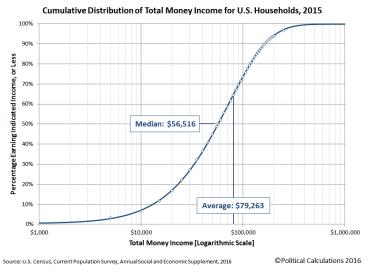
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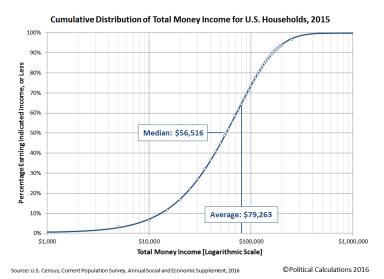
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• Note: we'll typically use lower-case letters like "x" to denote realizations (i.e. non-random numbers) of random variables like "X" ...

7



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• Formally: F(56,516) = 0.5

We All Need Some Support...

- The **support** of a random variable X, denoted X, is the set of values that X can take
 - If X is months in the year you were employed, $X = \{0, 1, ..., 12\}$
 - If X is your income, then $\mathbb{X} = \mathbb{R}_{\geq 0}$ (approximately)

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 - If X is months in the year you were employed, $\mathbb{X} = \{0, 1, ..., 12\}$
 - If X is your income, then $\mathbb{X} = \mathbb{R}_{\geq 0}$ (approximately)
- If the support of X is finite (e.g. $\{0,1\}$), we say X is **discrete**
- If the support of X is a continuum (e.g. \mathbb{R} or [0,1]), we say X is **continuously distributed** (technically, if the CDF is differentiable)

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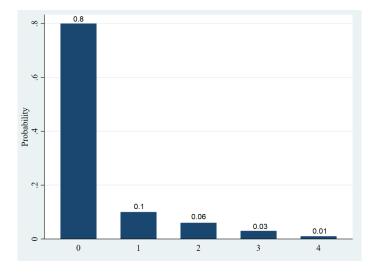
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• Notational note: both p(x) and f(x) are used for PDFs/PMFs

Example of a discrete random variable: number of wifi connection failures

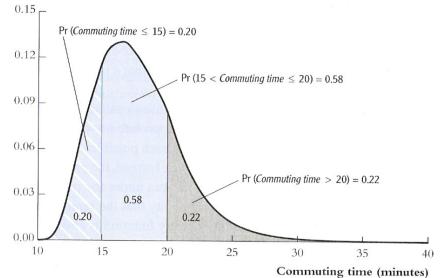
• Here's the PMF; what is the CDF?



Example of a continuous random variable: commuting time

• Here's the PDF; what is the CDF?

Probability density



Properties of PDFs/CDFs

- Key properties of CDFs $F(x) = Pr(X \le x)$:
 - Non-decreasing: $F(x) \ge F(x')$ if x > x'
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- Corresponding properties of PDFs $f(x) = \frac{\partial}{\partial x} F(x)$:
 - Non-negative: $f(x) \ge 0$ for all $x \in \mathbb{X}$
 - Satisfies $\int_{x \in \mathbb{X}} f(x) dx = 1$
 - For PMFs: $\sum_{x \in \mathbb{X}} p(x) = 1$

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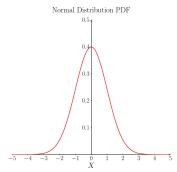
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- Note this is the only distribution of any binary X

An important continuous distribution: Normal $X \in \mathbb{R}$

- Example: the log of annual income (approximately)
- PDF: $f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right)$ for $x \in \mathbb{R}$ and $\sigma > 0$
- Here μ is the mean of X and σ^2 is its variance (also defined soon)
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- Written $X \sim U(a, b)$; also closed under linear transformations

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• Linked to the joint distribution by, e.g., $p(x) = \sum_{y \in \mathbb{Y}} p(x, y)$

Conditional Distributions

Combining joint and marginal distributions gives us the **conditional distribution** of one random variable given another

- Intuitively, the conditional distribution Y|X=x is the distribution of Y among the sub-population with X=x
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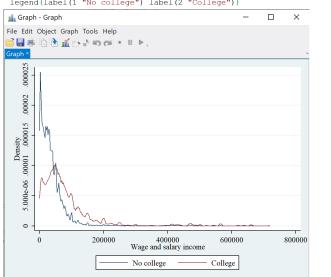
Leads immediately to Bayes' rule:

$$p(y \mid x) = p(x \mid y) \frac{p(y)}{p(x)}$$

Random Variables and Probability Distributions XI

Conditional PDFs of annual income given college completion:

```
1 twoway (kdensity incwage if educ<10) (kdensity incwage if educ>=10), ///
2 xtitle("Wage and salary income") ytitle("Density") ///
3 legend(label(1 "No college") label(2 "College"))
```



- An important concept in this course will be **independence**.
- Intuitively, independence says that knowing the value of X tells us nothing about the value of Y
- Formally, X, Y are independent $(X \perp \!\!\! \perp Y)$ if the conditional PDF/PMF of Y|X=x is the same as the unconditional one:

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- Conditional independence is defined similarly. $Y \perp \!\!\! \perp X \mid W$ if

$$p(y \mid x, w) = p(y \mid w), \forall (y, x, w)$$

• Intuitively, X tells us nothing about Y once we know W.

Multivariate Normals

An important multivariate distribution: joint normal $\mathbf{X} \in \mathbb{R}^K$

- Parameterized by a (mean) vector $\boldsymbol{\mu} \in \mathbb{R}^K$ and a positive-definite (variance-covariance) matrix $\boldsymbol{\Sigma} \in \mathbb{R}^K \times \mathbb{R}^K$
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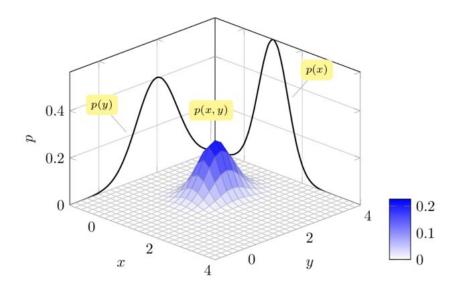
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Many useful facts; here's a few. If (X, Y)' is joint-normally distributed:

- The marginal distributions of X and Y are normal
- ullet The conditional distributions of $X\mid Y$ and $Y\mid X$ are normal
- Any fixed linear combination aX + bY + c is normally distributed

Bivariate Normal PDF



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- **Important fact:** The expectation operator is *linear*: E[a+bX] = a+bE[X] for constants (a,b)
 - Easily proved from the above definitions (make sure you can!)

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- If the die is fair, $Pr(X = 1) = ... = Pr(X = 6) = \frac{1}{6}$
- Plugging this in, we have

$$E[X] = \frac{1}{6}(1 + \dots + 6) = 3.5$$

Variances

Variances measure the squared spread of a distribution:

•
$$Var(X) = E[(X - E[X])^2]$$

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Variance of a linear transformation:

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This implies that $Std(a+bX) = b \cdot Std(X)$.

 Intuitively, if I measure income in cents, the standard deviation should be 100 times if I measure it in dollars

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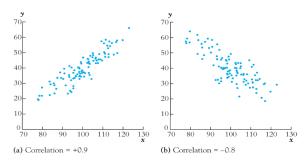
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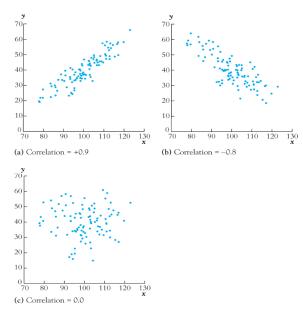
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- But not vice-versa! Independence is a stronger notion of association

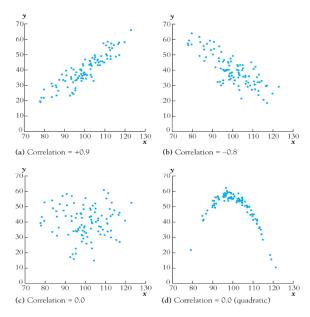
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- Covariances are linear: Cov(aX + c, bY + d) = abCov(X, Y)and Cov(X + Z, Y) = Cov(X, Y) + Cov(Z, Y)

In economics we are especially interested in conditional expectations

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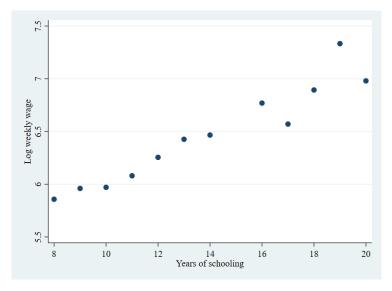
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- Conditioning on X makes functions of it constant: e.g. E[f(X)+g(X)Y|X=x]=f(x)+g(x)E[Y|X=x] for any $f(\cdot)$, $g(\cdot)$

Conditional Expectation Example

CEF of (log) annual income given years of schooling



A very important result for us: the **Law of Iterated Expectations** (LIE)

Let's start with an example. Suppose I want to calculate the average height of people in the United States. The LIE says I can:

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= $E[E[height|gender]]$

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Note that the expectation on the LHS uses p(y), while the outer expectation on the RHS uses p(x) and the inner expectation uses $p(y \mid x)$

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ullet Also, of course, independent \Longrightarrow mean independent (but not \Longleftarrow)

Quick Aside on Vector/Matrix Notation

Often it will be useful to work with random vectors $\mathbf{X} = [X_1, \dots, X_N]^T$

- These will always be "columns" in this course, and denoted in bold
- I will also use bold for matrices, with K columns and N rows

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Expectations are elementwise: e.g.
$$E\begin{bmatrix} X_{11} & X_{12} \\ X_{21} & X_{22} \end{bmatrix} = \begin{bmatrix} E[X_{11}] & E[X_{12}] \\ E[X_{21}] & E[X_{22}] \end{bmatrix}$$

• Define
$$Var\begin{pmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix} \end{pmatrix} = \begin{bmatrix} Var(X_1) & Cov(X_1, X_2) \\ Cov(X_1, X_2) & Var(X_2) \end{bmatrix}$$
 and
$$Cov\begin{pmatrix} \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}, \begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix} \end{pmatrix} = \begin{bmatrix} Cov(X_1, Y_1) & Cov(X_1, Y_2) \\ Cov(X_2, Y_1) & Cov(X_2, Y_2) \end{bmatrix}$$
, etc

Outline

- 1. Random Variables and Probability Distributions ✓
- 2. Means and Variances ✓
- 3. Identification in Experiments
- 4. Random Sampling and Sample Means
- 5. Hypothesis Testing and Inference

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- Recall the *potential outcomes* framework:
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• Suppose that for each person we assign D_i by flipping a coin. This implies that $D_i \perp \!\!\! \perp (Y_i(1), Y_i(0))$. Why?

• By virtue of the experiment, $D_i \perp \!\!\! \perp (Y_i(1), Y_i(0))$.

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- Similarly, $E[Y_i|D_i = 0] = E[Y_i(0)|D_i = 0] = E[Y_i(0)].$
- Combining these results, we can see that

$$\underbrace{E[Y_i|D_i=1]}_{\text{Pop mean for treated}} - \underbrace{E[Y_i|D_i=0]}_{\text{Pop mean for control}} = \underbrace{E[Y_i(1)-Y_i(0)]}_{\text{Avg treatment effect}} = \tau$$

Thus, the difference in treated/control population means in an experiment identifies the ATE!

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 - "Quasi-experiment"/"natural experiment": we think D_i is (effectively) as good as random among people with same value of \mathbf{X}_i

- Park et al (2021) study the impact of hot days (D_i) during the school year on test scores (Y_i)
 - Note Their D_i is not binary, although we could imagine a binarized treatment, e.g. $D_i = 1[Hotdays > 10]$
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 - Heat in a given year is effectively random conditional on historical weather patterns.
- Does this seem reasonable to you?

Park et al. (2021) Abstract

Human capital generally, and cognitive skills specifically, play a crucial role in determining economic mobility and macroeconomic growth. While elevated temperatures have been shown to impair short-run cognitive performance, much less is known about whether heat exposure affects the rate of skill formation. We combine standardized achievement data for 58 countries and 12,000 US school districts with detailed weather and academic calendar information to show that the rate of learning decreases with an increase in the number of hot school days. These results provide evidence that climatic differences may contribute to differences in educational achievement both across countries and within countries by socioeconomic status and that may have important implications for the magnitude and functional form of climate damages in coupled human-natural systems.

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- Essentially, they argue that once we know what colleges you applied/ were admitted to, where you choose to go is effectively random
- Do you believe this? Why might this assumption go wrong?

- <u>Dale and Krueger (2002)</u> studied a Q similar to our ongoing example: What is the effect on earnings of attending a selective college?
- Clearly, $D_i \not\perp \!\!\! \perp (Y_i(1), Y_i(0))$ because students who attend selective college will tend to have different academic ability.
- Dale and Krueger argue that $D_i \perp \!\!\! \perp (Y_i(1), Y_i(0)) | \mathbf{X}_i$, where \mathbf{X}_i is the set of colleges to which someone applied and got admitted.
- Essentially, they argue that once we know what colleges you applied/ were admitted to, where you choose to go is effectively random
- Do you believe this? Why might this assumption go wrong?
 - Students who choose to go to selective college may still differ in family background, motivation, career plans, etc.

Dale and Krueger (2002)

Abstract

Estimates of the effect of college selectivity on earnings may be biased because elite colleges admit students, in part, based on characteristics that are related to future earnings. We matched students who applied to, and were accepted by, similar colleges to try to eliminate this bias. Using the College and Beyond data set and National Longitudinal Survey of the High School Class of 1972, we find that students who attended more selective colleges earned about the same as students of seemingly comparable ability who attended less selective schools. Children from low-income families, however, earned more if they attended selective colleges.

Using Conditional Unconfoundedness

- Suppose that $D_i \perp \!\!\! \perp (Y_i(1), Y_i(0)) | \mathbf{X}_i$, where \mathbf{X}_i is a vector of observable characteristics.
- Similar to in the experiment, we have for all x:

$$E[Y_i|D_i = 1, X_i = x] = E[Y_i(1)|X_i = x].$$

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$$E[Y_i|D_i=0,\mathbf{X}_i=\mathbf{x}]=E[Y_i(0)|\mathbf{X}_i=\mathbf{x}].$$

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This implies that

$$\underbrace{E[Y_i|D_i=1,\mathbf{X}_i=\mathbf{x}]}_{\text{Pop avg treated w/}} - \underbrace{E[Y_i|D_i=0,\mathbf{X}_i=\mathbf{x}]}_{\text{Pop avg control w/}} = \underbrace{E[Y_i(1)-Y_i(0)|\mathbf{X}_i=\mathbf{x}]}_{\text{ATE with }\mathbf{X}_i=\mathbf{x}}$$

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• $E[Y_i(1) - Y_i(0) | \mathbf{X}_i = \mathbf{x}]$ is often called the *conditional average* treatment effect, written $CATE(\mathbf{x})$.

Using Conditional Unconfoundedness (cont.)

- We showed that under conditional unconfoundedness, $CATE(\mathbf{x}) = E[Y_i(1) Y_i(0)|\mathbf{X}_i = \mathbf{x}]$ is identified.
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- Requires $0 < Pr(D_i = 1 | \mathbf{X}_i = \mathbf{x}]) < 1$: called an **overlap** condition
- Intuitively, we need there to be some treated and some control units for each value of X_i , in order to learn about the overall ATE

Learning about Population Means

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- We need to learn about these estimands from the observed sample
- Enter statistical inference...

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- 1. Random Variables and Probability Distributions ✓
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 - In the Dewey v. Truman example, not representative!

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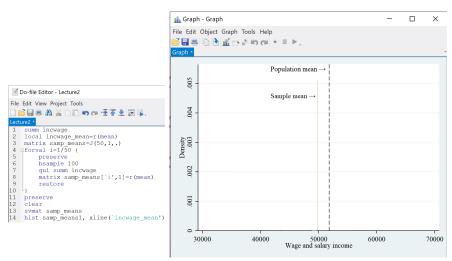
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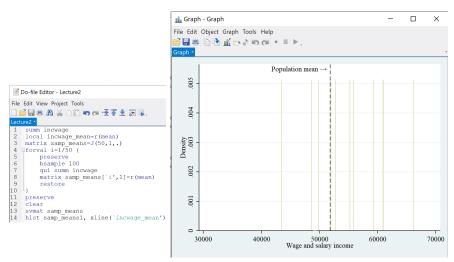
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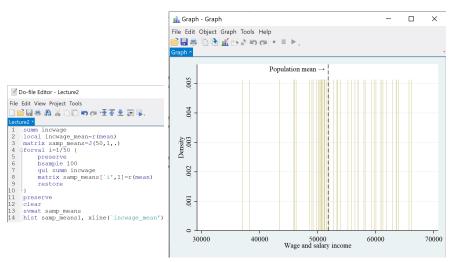
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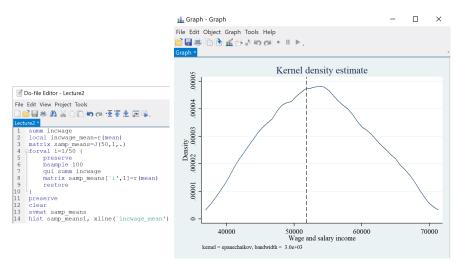
- Equation (1) says that $\hat{\mu}$ is *unbiased*: its average value is μ
- Equation (2) says that the standard deviation of $\hat{\mu}$ from its mean (i.e. μ) shrinks with the sample size N (\approx consistency)

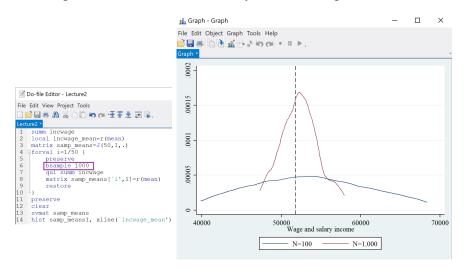


Simulations of Random Sampling









So: two reasons why $\hat{\mu} = \frac{1}{N} \sum_i Y_i$ is a good estimator of $\mu = E[Y_i]$:

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In the next chapter we'll see another nice property of $\hat{\mu}$: when N is large, its distribution is approximately normal

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• What if X_i is not discrete, or $N_x \nrightarrow \infty$? Coming soon...

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Hypothesis Testing – an Introduction

- \bullet We've shown that when N gets large, the sample mean $\hat{\mu}$ gets close to the population mean μ
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- It tells us whether it is likely to see a sample mean of \$50,000 if the truth is \$55,000, \$70,000, etc.

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- **3** Specify a **null hypothesis** that the population mean is a particular value, $H_0: \mu = \mu_0$.
 - E.g. A population mean is \$55,000 would be H_0 : $\mu = 55,000$
- ② Calculate how likely it would be to observe $\hat{\mu}$ at least this far from μ_0 if the null is true. This is called a **p-value**
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 - The CI, by construction, contains the true value μ in 95% of the realizations of the data when $\alpha=0.05$

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$$\begin{split} \rho(\hat{t}) &= 1 - (\Phi(|\hat{t}|) - \Phi(-|\hat{t}|)) = 1 - \left(\Phi\left(\frac{|\hat{\mu} - \mu_0|}{\sigma/\sqrt{N}}\right) - \Phi\left(\frac{-|\hat{\mu} - \mu_0|}{\sigma/\sqrt{N}}\right)\right) \\ &= 2\left(1 - \Phi\left(\frac{|\hat{\mu} - \mu_0|}{\sigma/\sqrt{N}}\right)\right) \end{split}$$

• Intuitively, p is the probability we would see a $|\hat{t}|$ at least this big if the null is true.

Illustration of P-Value Construction

Standard Normal PDF (mean zero, unit std. dev.)

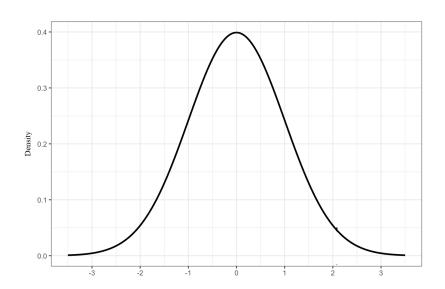


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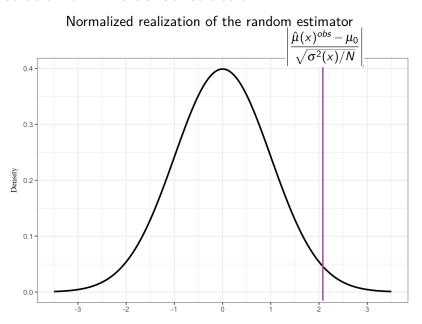
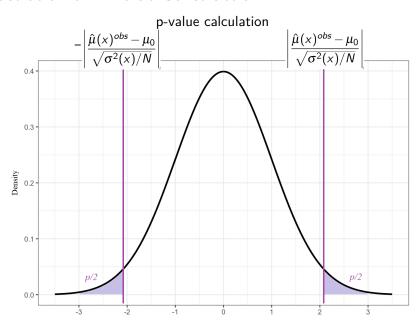


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Recall that our p-value takes the form

$$1 - \left(\Phi(|\hat{t}|) - \Phi(-|\hat{t}|)\right)$$

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- ullet The interval $\hat{\mu}\pm 1.96\sigma/\sqrt{N}$ is thus the 95% confidence interval (CI)
 - It has the property that $Pr(\mu_0 \in CI) = 0.95$ when $H_0: \mu = \mu_0$ is true

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- The *significance level* (or *size*) of a test is the pre-specified probability of incorrectly rejecting the null when it is true (type-I error rate)
 - E.g. a 5% level test rejects when p < 0.05.
- The power of a test is the probability of correctly rejecting the null when it is false (1 - type-II error rate)
 - The power is a function of the *alternative* hypothesis. I.e., the probability that we reject $H_0: \mu = \mu_0$ when in fact $\mu = \mu_A$

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 - By Bayes' rule, $P(H_0|Data) = P(Data|H_0) * P(H_0)/P(Data)$. But to formalize this, we need to take a stand on our *prior* belief that H_0 is true, $P(H_0)$.

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- People often interpret a p < 0.05 as strong evidence of an effect and $p \ge 0.05$ as evidence of no effect.
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 - It's better to view the *p*-value as a spectrum indicating how likely is the observed data given the null.
 - Moreover, *p*-values can be large even if the null is false (low power!)

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09 March 2015 This story originally asserted that "The closer to zero the P value gets, the greater the chance the null hypothesis is false." P values do not give the probability that a null hypothesis is false, they give the probability of obtaining data at least as extreme as those observed, if the null hypothesis was true. It is by convention that smaller P values are interpreted as stronger evidence that the null hypothesis is false. The text has been changed to reflect this.

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 - Have I just been wasting your time!?! No.
- We will next review powerful **asymptotic** results. These will allow us to apply similar inference tools if the sample is "large" even when Y_i is not normally distributed.