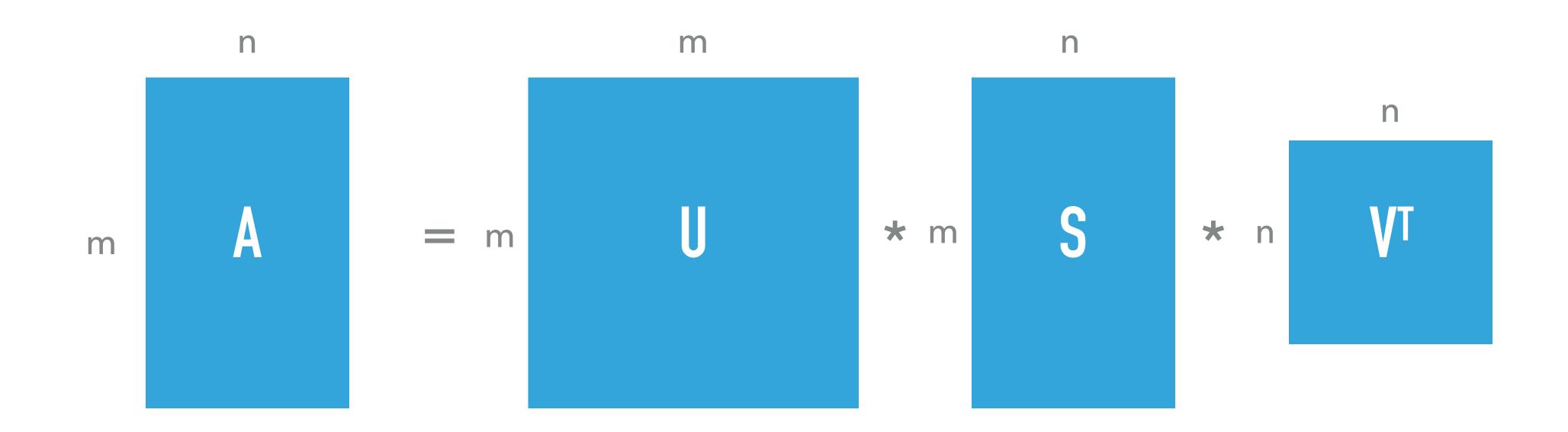
CS57300 PURDUE UNIVERSITY SEPTEMBER 1, 2021

DATA MINING

A rectangular matrix A can be broken down into the product of three matrices: an orthogonal matrix U, a diagonal matrix S, and the transpose of an orthogonal matrix V.



- \triangleright Columns of U are eigenvectors of AA^T
- ► Columns of V are eigenvectors of A^TA
- ▶ Diagonal entries of S are the square roots of the non-zero eigenvalues of AA^T (as well as A^TA)

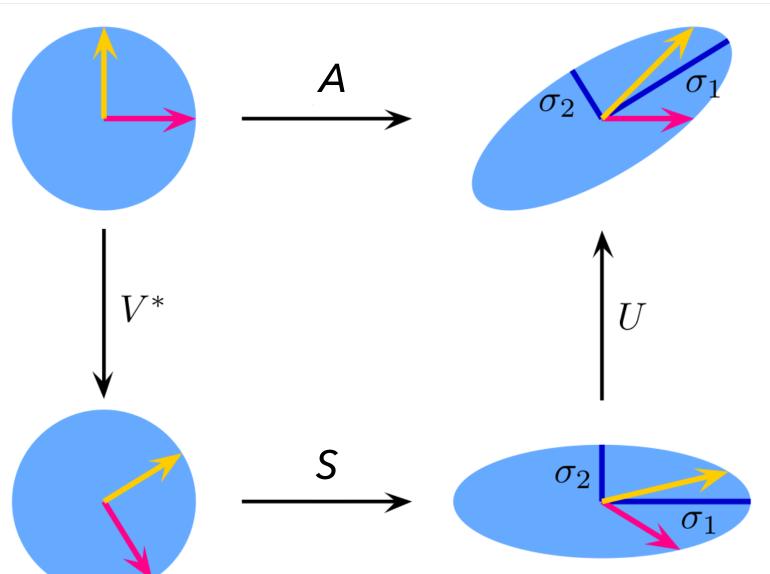
▶ $A = USV^T$; columns of U are eigenvectors of AA^T , columns of V are eigenvectors of A^TA , diagonal entries of S are the square roots of the non-zero eigenvalues of AA^T (as well as A^TA)

- ▶ $A = USV^T$; columns of U are eigenvectors of AA^T , columns of V are eigenvectors of A^TA , diagonal entries of S are the square roots of the non-zero eigenvalues of AA^T (as well as A^TA)
 - $AA^T = USV^TVS^TU^T$
 - ▶ A^TA is a symmetric matrix, so the matrix composed of eigenvectors of A^TA (i.e., V) is orthogonal, thus $V^TV = I$
 - $AA^{T} = USS^{T}U^{T}$ (this is eigendecomposition of matrix AA^{T}).

- \triangleright Columns of U are eigenvectors of AA^T
- \triangleright Columns of V are eigenvectors of A^TA

Diagonal entries of S are the square roots of the non-zero eigenvalues of AA^T (as well as A^TA)

Geometric interpretation:



DISTANCE MEASURES

REPRESENTING DATA IN EUCLIDEAN SPACE

- If data objects have the same fixed set of numeric attributes, then the data objects can be thought of as points in a multi-dimensional space, where each dimension represents a distinct attribute
- Many data mining techniques then use similarity/dissimilarity measures to characterize relationships between the instances

Height	Weight	Heart Rate	BP (Diastolic)	BP (Systolic)
1.79	80	70	73	112
1.60	51	73	69	105

METRIC PROPERTIES

A **metric** d(x,y) (or a distance function) is a function that satisfies the following properties:

► $d(x,y) \ge 0$ for all x,y and d(x,y)=0 iff x=y Positivity

d(x,y) = d(y,x) for all x,y Symmetry

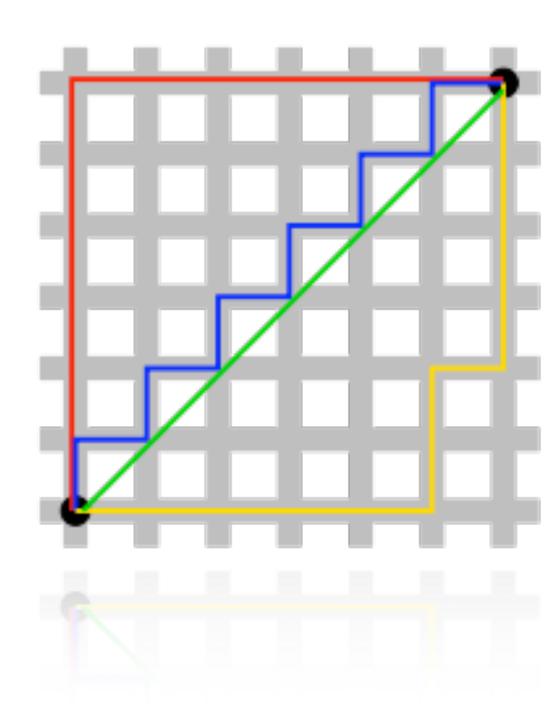
► $d(x,y) \le d(x,k)+d(k,y)$ for all x,y,k Triangle inequality

DIFFERENT TYPES OF METRICS

- Manhattan distance (L1) $d_M(x,y) = \sum_{i=1}^p |x_i y_i|$
- ► Euclidean distance (L2) $d_E(x,y) = \sqrt{\sum_{i=1}^p (x_i y_i)^2}$
 - Most common metric
 - Assumes dimensions are commensurate
- Weighted Euclidean distance

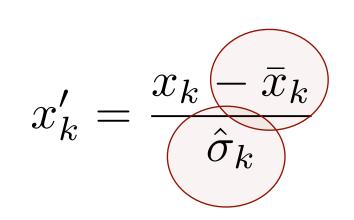
$$d_{WE}(x,y) = \sqrt{\sum_{i=1}^{p} w_i (x_i - y_i)^2}$$

Can weight variables by relative importance



STANDARDIZATION

- Normalization
 - Removes effect of scale



subtract mean divide by stdev

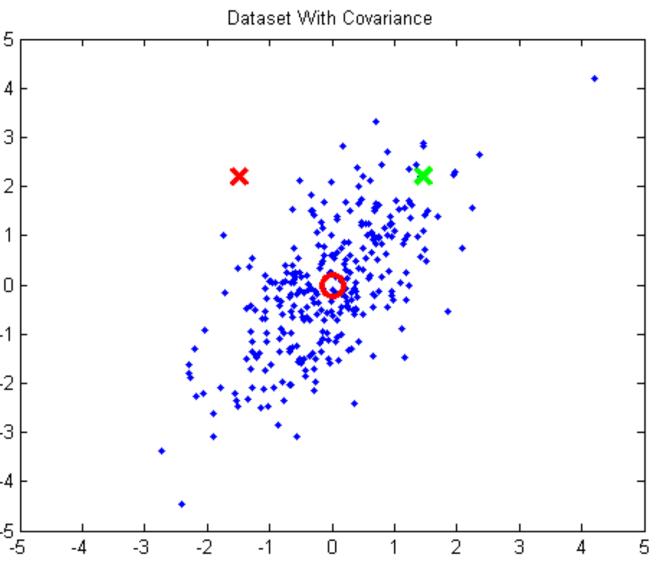
- Divide each variable by its standard deviation
- Weights all variables equally

$$d'_{E}(x,y) = \sqrt{\sum_{i=1}^{p} (x'_{i} - y'_{i})^{2}}$$

CORRELATION AMONG VARIABLES

- Variables contribute independently to additive measure of distance
- May not be appropriate
 if variables are highly correlated
- Can standardize variables in a way that accounts for covariance





MAHALANOBIS DISTANCE

$$d_{MH}(\mathbf{x}, \mathbf{y}) = \sqrt{(\mathbf{x} - \mathbf{y})^T \Sigma^{-1} (\mathbf{x} - \mathbf{y})}$$

pxp covariance matrix

- Automatically accounts for scaling
- Corrects for correlation between attributes
- Tradeoff:
 - Covariance matrix can be hard to estimate accurately
 - Memory and time complexity is quadratic rather than linear

DISTANCE MEASURES FOR BINARY DATA

- d(x,y) when items x and y are p-dimensional binary vectors
- Let n_{11} be the number of attributes where both items have value 1, etc.

$$n_{11} = \sum_{i}^{p} \mathbb{I}(x_i + y_i = 2)$$

- Matching distance
 - Hamming distance normalized by number of bits

	y=1	y=0
x=1	N 11	n 10
x=0	n ₀₁	n 00

$$d_M(x, y) = 1 - \frac{n_{11} + n_{00}}{n_{11} + n_{00} + n_{10} + n_{01}}$$

- Jaccard distance
 - If we don't care about matches on zeros

$$d_M(x, y) = 1 - \frac{n_{11}}{n_{11} + n_{10} + n_{01}}$$

POPULATIONS AND SAMPLES

ELEMENTARY UNITS, POPULATIONS, AND SAMPLES

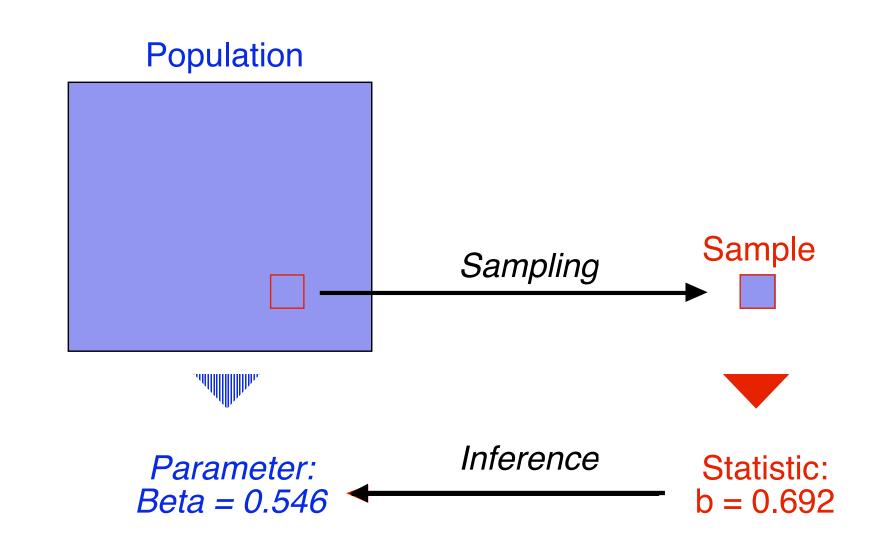
- Elementary units:
 - Entities (e.g., persons, objects, events) that meet a set of specified criteria
 - Example: A person who has purchased something from Walmart in the past month
- Population:
 - Aggregate of elementary units (i.e, all entities of interest)
- Sample:
 - Sub-group of the population

SAMPLING

- Reasons to sample
 - Obtaining the entire set of data of interest is too expensive or time consuming
 - Processing the entire set of data of interest is too expensive or time consuming
- Sampling is the main technique employed for data selection

USE SAMPLES FOR ESTIMATION

- In data mining we often work with a sample of data from the population of interest
- If we had the population we could calculate the properties of interest
- Sample serves as a reference group for estimating characteristics about the population and drawing conclusions



PRINCIPLE FOR EFFECTIVE SAMPLING

- The key principle for effective sampling is the following:
 - Using a sample will work almost as well as using the entire data set, if the sample is representative
 - A sample is representative if it has approximately the same property (of interest) as the original set of data

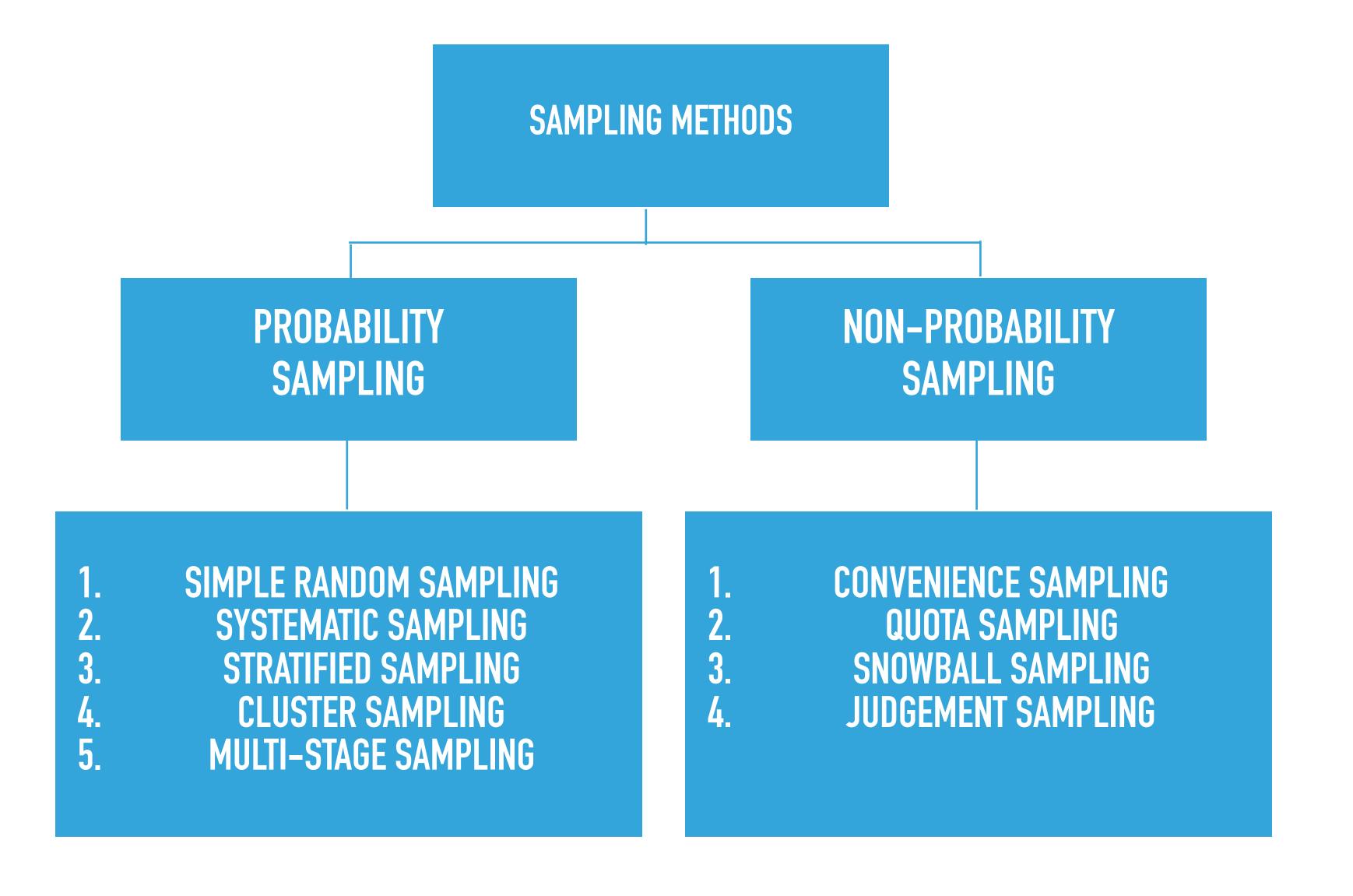
BACKGROUND & BASICS 20

SAMPLE SIZE

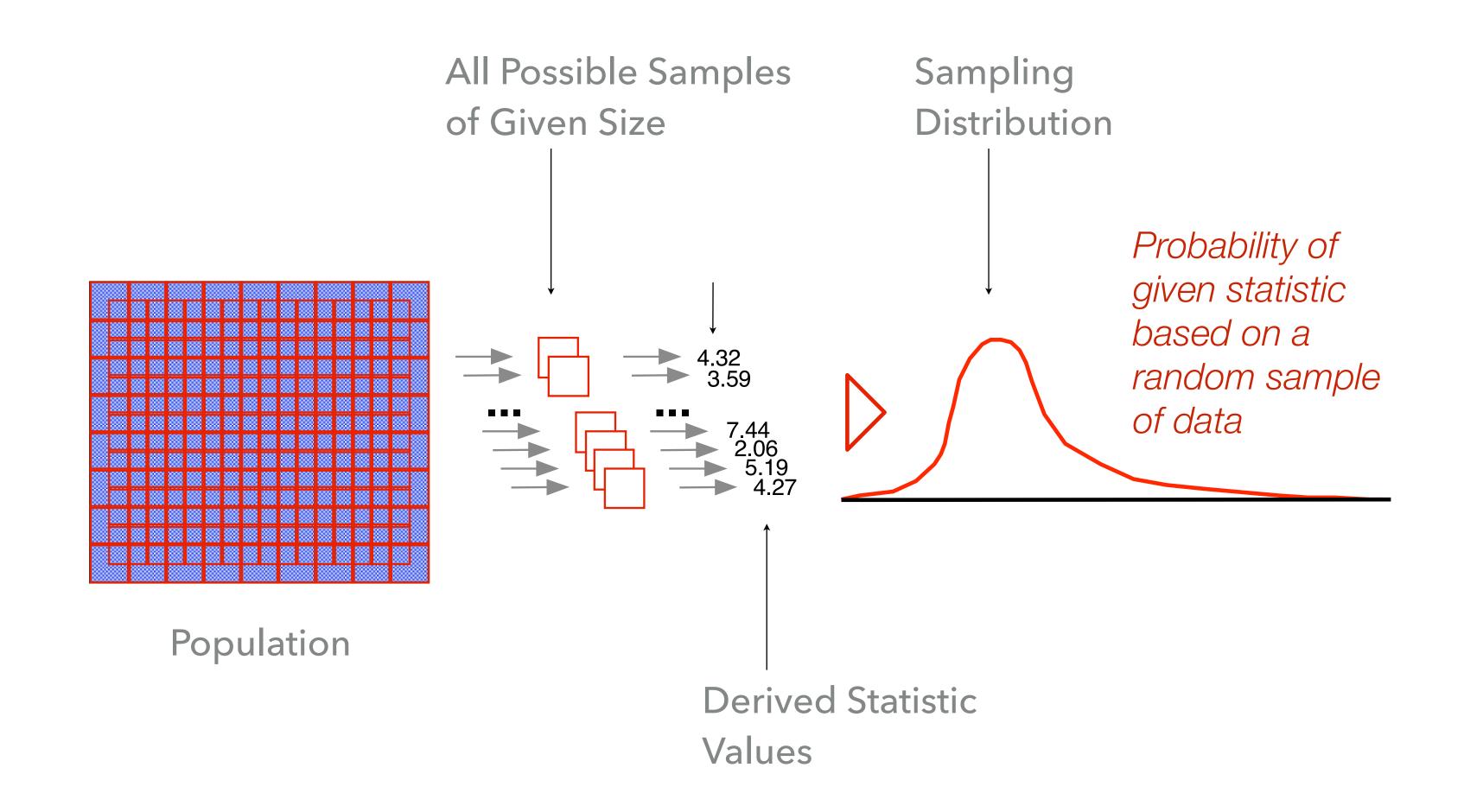


TYPES OF PROBABILITY SAMPLING

- Simple random sampling
 - There is an equal probability of selecting any particular item
 - Sampling without replacement
 - As each item is selected, it is removed from the population
 - Sampling with replacement
 - Items are not removed from the population as they are selected for the sample; the same item can be picked up more than once
- Stratified sampling
 - > Split the data into several partitions; then draw random samples from each partition



SAMPLING DISTRIBUTIONS



STATISTICAL INFERENCE

STATISTICAL INFERENCE

- Infer properties of an unknown distribution with sample data generated from that distribution
- Parameter estimation
 - Infer the value of a population parameter based on a sample statistic (e.g., estimate the mean)
- Hypothesis testing
 - Infer the answer to a question about a population parameter based on a sample statistic (e.g., is the mean non-zero?)

PARAMETER ESTIMATION

- \blacktriangleright Infer the value of population parameters (heta) from data
- $m{ heta}$ can take values in the parameter space $m{\Theta}$
- Frequentist approach
 - Population parameters are fixed but unknown
 - Data is a random sample drawn from population
 - Use maximum likelihood estimation (MLE)
- Bayesian approach
 - Parameters are random variables with a distribution of possible values
 - Data is fixed and known, provides evidence for different parameter values
 - Use maximum aposteriori estimation (MAP)

MAXIMUM LIKELIHOOD ESTIMATION (MLE)

- Suppose we have a set of data $X = \{x_i\}_{i=1}^N$ independently drawn from the population
- The maximum likelihood estimation finds the parameter values that maximize the likelihood of observing the data

$$egin{aligned} heta_{MLE} &= rg\max_{ heta} P(X| heta) \ &= rg\max_{ heta} \prod_{i} P(x_i| heta) \ &= rg\max_{ heta} \log \prod_{i} P(x_i| heta) \ &= rg\max_{ heta} \sum_{i} \log P(x_i| heta) \end{aligned}$$

MAXIMUM A-POSTERIORI ESTIMATION (MAP)

- Suppose we have a set of data $X = \{x_i\}_{i=1}^N$ independently drawn from the population, and the prior distribution for the parameter is $P(\theta)$
- The maximum a-posteriori estimation finds the mode of the posterior distribution of the parameters

$$egin{aligned} heta_{MAP} &= rg \max_{ heta} P(X| heta)P(heta) \ &= rg \max_{ heta} \log P(X| heta)P(heta) \ &= rg \max_{ heta} \log \prod_{i} P(x_i| heta)P(heta) \ &= rg \max_{ heta} \sum_{i} \log P(x_i| heta)P(heta) \end{aligned}$$

MLE VS. MAP EXAMPLE

- ▶ Flip a coin for *N* times and observe *n* heads; what's the probability of seeing the head if tossing the coin once?
- Likelihood of observing the data: $P(D \mid \theta) = \binom{N}{n} \theta^n (1 \theta)^{N-n}$
 - The number of heads observed follows a binomial distribution
- Maximum likelihood estimation:

$$\theta_{MLE} = \operatorname{argmax}_{\theta} P(D \mid \theta) = \frac{n}{N}$$

MLE VS. MAP EXAMPLE

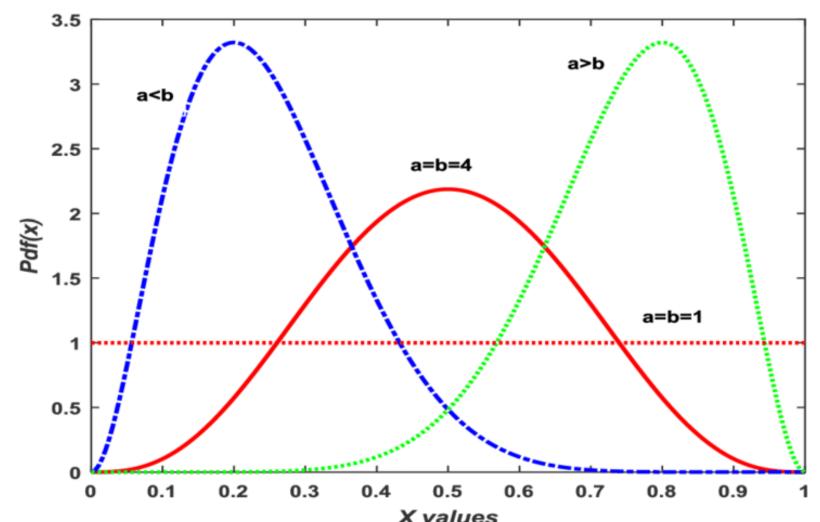
- Maximum a-posteriori estimation:
 - Suppose the prior is a Beta distribution

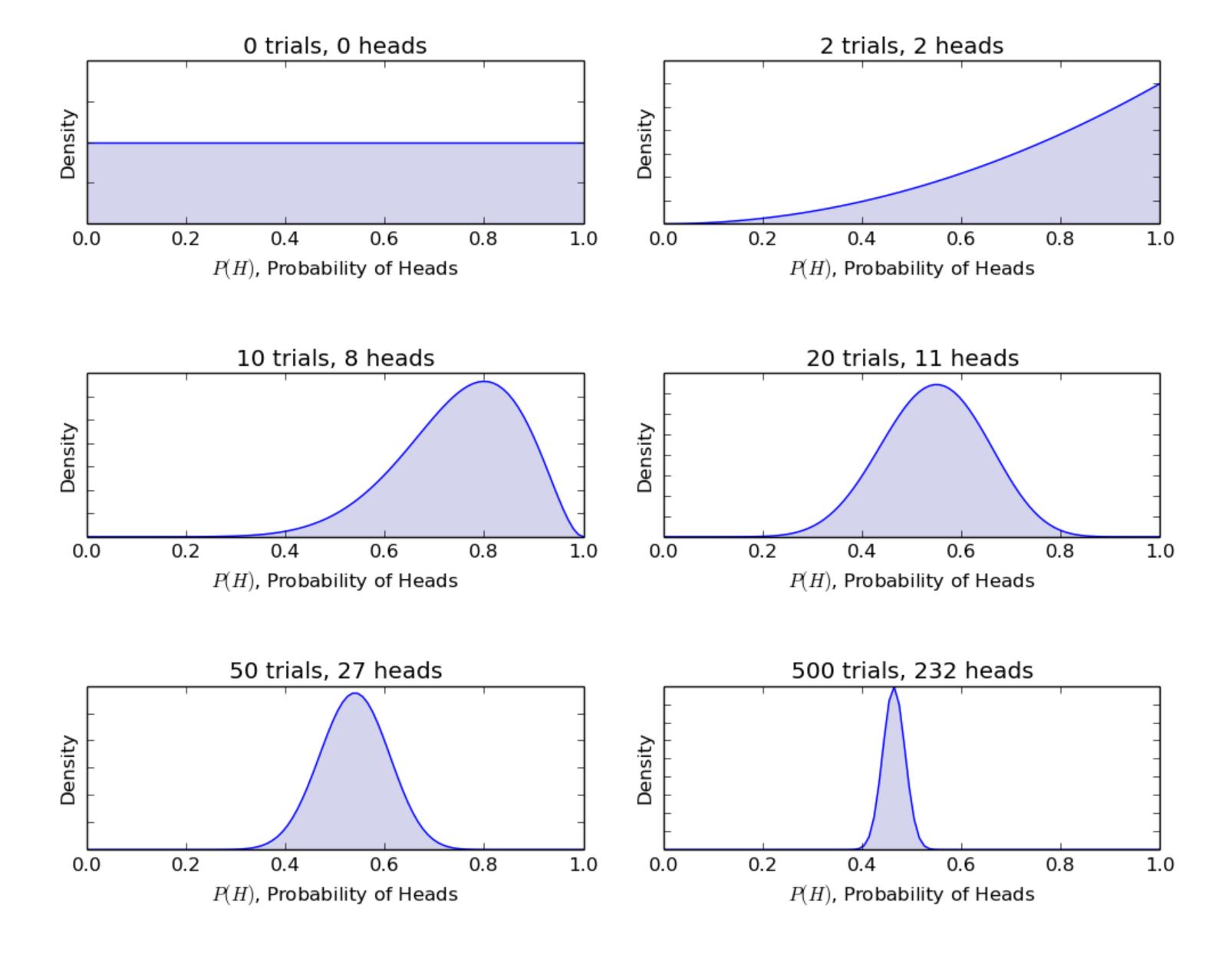
$$P(\theta) = \frac{\theta^{a-1}(1-\theta)^{b-1}}{B(a,b)} \sim Beta(a,b), \text{ where } B(a,b) = \int_0^1 \theta^{a-1}(1-\theta)^{b-1}d\theta = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}, \Gamma(z) = \int_0^\infty x^{z-1}e^{-x}dx$$

Then, the posterior is:

$$P(\theta \mid D) = \frac{P(D \mid \theta)P(\theta)}{P(D)} = \frac{\binom{N}{n}(\theta^{a+n-1}(1-\theta)^{b+N-n-1}/B(a,b))}{\int_{0}^{1} \binom{N}{n}(\theta^{a+n-1}(1-\theta)^{b+N-n-1}/B(a,b))d\theta}$$

$$\sim Beta(a+n,b+N-n)$$





MLE VS. MAP EXAMPLE

Maximum a-posteriori estimation:

$$\theta_{MAP} = \operatorname{argmax}_{\theta} P(\theta | D)$$

$$= \operatorname{argmax}_{\theta} Beta(a + n, b + N - n)$$

$$= \frac{a + n - 1}{a + b + N - 2}$$

- Notice that in this example, the posterior distribution is in the same probability distribution family as the prior distribution.
 - The prior and posterior are called **conjugate distributions**, and the prior is called a **conjugate prior** for the likelihood function.
 - The beta distribution is a conjugate prior to the binomial likelihood.