

## Solution : Homework 10

$$1. \quad E(X) = \frac{1}{\beta} \int_0^{\infty} x e^{-\frac{\beta}{x}} dx = \beta$$

~~$\Rightarrow E\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$~~

$$\text{Hence, } E(\bar{X}_n) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \sum_{i=1}^n \beta = \beta.$$

2. Done in class.

3. Consider

$$T_n(X_1, \dots, X_n) = \begin{cases} 1 & \text{if } X_1 = 0, X_2 = 1 \\ 0 & \text{otherwise} \end{cases}$$

$$\begin{aligned} \text{Now, } E(T_n) &= 1 \cdot P[X_1 = 0, X_2 = 1] \\ &= P[X_1 = 0] P[X_2 = 1] \\ &= e^{-\theta} \times \frac{e^{-\theta} \theta^1}{1!} = \theta \cdot e^{-2\theta}. \end{aligned}$$

$\Rightarrow T_n$  is an unbiased estimator for  $\theta \cdot e^{-2\theta}$ .

4. Consider

$$T_n(x_1, \dots, x_n) = \begin{cases} 1 & \text{if } x_1 = 1, x_2 = 1, x_3 = 0 \\ 0 & \text{otherwise} \end{cases}$$

Now,

$$E(T_n) = P[X_1 = 1, X_2 = 1, X_3 = 0]$$

$$= P[X_1 = 1] P[X_2 = 1] P[X_3 = 0]$$

$$= \theta^2(1-\theta).$$

$\Rightarrow T_n$  is an unbiased estimator of  $\theta^2(1-\theta)$ .

5. Joint p.d.f. of  $X_1$  and  $X_2$  is

$$\begin{aligned} f_{X_1, X_2}(x_1, x_2) &= (\theta \cdot e^{-\theta x_1}) \times (\theta^2 e^{-2\theta x_2}) \\ &= 2\theta^2 e^{-\theta(x_1 + 2x_2)} \\ &= \left\{ \theta^2 e^{-\theta(x_1 + 2x_2)} \right\} \times (2). \end{aligned}$$

Hence, by FT,  $T(X_1, X_2) = X_1 + 2X_2$  is a sufficient statistic for  $\theta$ .

6. Done in class.

7.  $f(x_1, \dots, x_n | \theta) = \begin{cases} 1 & \text{if } \theta - \frac{1}{2} \leq x_{(1)} \leq \dots \leq x_{(n)} \leq \theta + \frac{1}{2} \\ 0 & \text{otherwise} \end{cases}$

ie,

i.e.,

$$f(x|\theta) = 1 \times 1_{\{\theta - \frac{1}{2}, x\}} \times 1_{\{x, \theta + \frac{1}{2}\}}$$

[Here  $1_{\{a, b\}} = 1$  if  $a < b$   
= 0 otherwise]

$$= g\left(0, \underset{\parallel}{(x_{(1)}, x_{(n)})}\right) \underset{\perp}{h(x)}$$

$$1_{\{0 - \frac{1}{2}, x_{(1)}\}} \times 1_{\{x_{(1)}, 0 + \frac{1}{2}\}}.$$

Hence, by FT,  $T_n(x_1, \dots, x_n) = (x_{(1)}, x_{(n)})$

is jointly sufficient for  $\theta$ .