

3.7 (Derivation of the one-way ANOVA F -test using the extra SS method):

- a. The ANOVA identity follows immediately from (3.13) by noting that $\hat{y}_{ij} = \bar{y}_i$ and $e_{ij} = y_{ij} - \hat{y}_{ij} = y_{ij} - \bar{y}_i$ for $j = 1, \dots, n_i$ and $i = 1, \dots, k$. The squared norms of the vectors are then given by the corresponding sums of squares.
- b. Under H_0 , the one-way ANOVA model becomes $y_{ij} = \mu + \varepsilon_{ij}$, where μ is the common mean of all groups under H_0 . It is easy to show that the LS estimate of μ is $\hat{\mu} = \bar{y}$. So $SSE_0 = \sum_{i=1}^k \sum_{j=1}^{n_i} (y_{ij} - \bar{y})^2 = SST$. Therefore $SSH_0 = SSE_0 - SSE = SST - SSE = SSG$ from the ANOVA identity. The hypothesis d.f. is $k - 1$ since without H_0 there are k independent parameters, μ_1, \dots, μ_k , while under H_0 , there is only one, μ ; so H_0 imposes $k - 1$ linearly independent restrictions.. The error d.f. is, of course, $N - k$. This explains the extra SS F -statistic, which equals the ANOVA F -statistic.

6.2 (Extra sum of squares test in terms of R^2): The extra SS test statistic for comparing the two models is given by

$$F = \frac{(SSE_q - SSE_p)/(p - q)}{SSE_p/[n - (p + 1)]}.$$

We have $SSE_p = SST(1 - R_p^2)$ and $SSE_q = SST(1 - R_q^2)$. Substituting in the extra SS F -statistic we get

$$\begin{aligned} F &= \frac{SST[(1 - R_q^2) - (1 - R_p^2)]}{SST(1 - R_p^2)/[n - (p + 1)]} \\ &= \frac{(R_p^2 - R_q^2)/(p - q)}{(1 - R_p^2)/[n - (p + 1)]}. \end{aligned}$$

For the given data

$$F = \frac{(0.90 - 0.80)/(5 - 3)}{(1 - 0.90)/(26 - 6)} = 10 > f_{2,20,.01} = 5.85.$$

So the increase R^2 is significant at $\alpha = .01$.

3.9 (Alternate coding of categorical variables): With the new coding: $x_1 = \pm 1$ and $x_2 = \pm 1$, we get the following equations:

$$\beta_0 - \beta_1 - \beta_2 + \beta_3 = 40$$

$$\beta_0 + \beta_1 - \beta_2 - \beta_3 = 45$$

$$\beta_0 - \beta_1 + \beta_2 - \beta_3 = 50$$

$$\beta_0 + \beta_1 + \beta_2 + \beta_3 = 65.$$

The solution to this square system of equations $\mathbf{X}\boldsymbol{\beta} = \mathbf{y}$ is the LS estimate $\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$ where

$$\mathbf{X} = \begin{bmatrix} 1 & -1 & -1 & 1 \\ 1 & 1 & -1 & -1 \\ 1 & -1 & 1 & -1 \\ 1 & 1 & 1 & 1 \end{bmatrix}, \mathbf{y} = \begin{bmatrix} 40 \\ 45 \\ 50 \\ 65 \end{bmatrix} \quad \text{and} \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \end{bmatrix}.$$

Note that \mathbf{X} is an orthogonal matrix and $\mathbf{X}'\mathbf{X} = 4\mathbf{I}$. Hence

$$\begin{aligned} \hat{\boldsymbol{\beta}} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} \\ &= \frac{1}{4} \begin{bmatrix} 1 & 1 & 1 & 1 \\ -1 & 1 & -1 & 1 \\ -1 & -1 & 1 & 1 \\ 1 & -1 & -1 & 1 \end{bmatrix} \begin{bmatrix} 40 \\ 45 \\ 50 \\ 65 \end{bmatrix} \\ &= \frac{1}{4} \begin{bmatrix} 40 + 45 + 50 + 65 \\ -40 + 45 - 50 + 65 \\ -40 - 45 + 50 + 65 \\ 40 - 45 - 50 + 65 \end{bmatrix} = \begin{bmatrix} 50 \\ 5 \\ 7.5 \\ 2.5 \end{bmatrix}. \end{aligned}$$

The interpretations of these coefficients are as follows.

1. $\hat{\beta}_0$ is the overall mean of all the y 's.
2. $\hat{\beta}_1$ is (1/4)th times the change in y when x_1 is changed from -1 to $+1$ summed over the two levels of x_2 (called the main effect of x_1).
3. $\hat{\beta}_2$ is (1/4)th times the change in y when x_2 is changed from -1 to $+1$ summed over the two levels of x_1 (called the main effect of x_2).
4. $\hat{\beta}_3$ is (1/4)th times the difference in the changes in y when x_1 is changed from -1 to $+1$ between the two levels of x_2 (called the interaction between x_1 and x_2). This is also equal to (1/4)th times the difference in the changes in y when x_2 is changed from -1 to $+1$ between the two levels of x_1 .