

# CSIR-UGC NET-June 2013-Problem(68)

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## Question

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Let  $X_1, \dots, X_n$  be independent and identically distributed random variables with probability density function

$$f(x) = \frac{1}{2}\lambda^3 x^2 e^{-\lambda x}; x > 0; \lambda > 0$$

Then which of the following statements are true?

- ①  $\frac{2}{n} \sum_{i=1}^n \frac{1}{X_i}$  is an unbiased estimator of  $\lambda$
- ②  $\frac{3n}{\sum_{i=1}^n X_i}$  is an unbiased estimator of  $\lambda$
- ③  $\frac{2}{n} \sum_{i=1}^n \frac{1}{X_i}$  is a consistent estimator of  $\lambda$
- ④  $\frac{3n}{\sum_{i=1}^n X_i}$  is a consistent estimator of  $\lambda$

# Definitions

## What is estimator?

An estimator is a statistic that estimates some fact about the population. For example, the sample mean ( $\bar{X}$ ) is an estimator for the population mean,  $\mu$ . The quantity that is being estimated is called the **estimand**.

## Bias of estimator

Let  $\Theta = h(X_1, X_2, \dots, X_n)$  be a point estimator for  $\theta$ . The **bias** of the estimator  $\Theta$  is defined by

$$B(\Theta) = E[\Theta] - \theta \quad (1)$$

where  $E[\Theta]$  is the expectation value of the estimator  $\Theta$  and  $\theta$  is the estimand.

## Definitions Contd.

### Unbiased estimator

Let  $\Theta = h(X_1, X_2, \dots, X_n)$  be a point estimator for a parameter  $\theta$ . We say that  $\Theta$  is an **unbiased estimator** of  $\theta$  if

$$B(\Theta) = 0, \text{ for all possible values of } \theta. \quad (2)$$

### Consistent estimator

Let  $\Theta_1, \Theta_2, \dots, \Theta_n, \dots$ , be a sequence of point estimators of  $\theta$ . We say that  $\Theta_n$  is a **consistent** estimator of  $\theta$ , if

$$\lim_{n \rightarrow \infty} \Pr(|\Theta_n - \theta| \geq \epsilon) = 0, \text{ for all } \epsilon > 0. \quad (3)$$

## Definitions Contd.

### Mean Squared Error (MSE)

The **mean squared error (MSE)** of a point estimator  $\Theta$ , shown by  $MSE(\Theta)$ , is defined as

$$MSE(\Theta) = E[(\Theta - \theta)^2] \quad (4)$$

$$= Var(\Theta) + B(\Theta)^2 \quad (5)$$

where  $B(\Theta)$  is the bias of  $\Theta$ .

### Theorem

Let  $\Theta_1, \Theta_2, \dots$  be a sequence of point estimators of  $\theta$ . If

$$\lim_{n \rightarrow \infty} MSE(\Theta_n) = 0, \quad (6)$$

then  $\Theta_n$  is a consistent estimator of  $\theta$ .

## Solution: Option 1

Now here we have our estimator  $\Theta$  and estimand  $\theta$  as,

$$\Theta = \frac{2}{n} \sum_{i=1}^n \frac{1}{X_i} \text{ and } \theta = \lambda \quad (7)$$

To check if this is an unbiased estimator or not the bias needs to be calculated.

$$B(\Theta) = E[\Theta] - \theta \quad (8)$$

## Option 1 Contd.

The expectation value of the estimator is given by,

$$E[\Theta] = E \left[ \frac{2}{n} \sum_{i=1}^n \frac{1}{X_i} \right] \quad (9)$$

$$= \frac{2}{n} \sum_{i=1}^n E \left[ \frac{1}{X_i} \right] \quad (10)$$

$$= \frac{2}{n} \sum_{i=1}^n \int_{-\infty}^{\infty} \frac{1}{x} f(x) dx \quad (11)$$

$$= \frac{2n}{n} \int_0^{\infty} \frac{1}{x} \frac{1}{2} \lambda^3 x^2 e^{-\lambda x} dx \quad (12)$$

$$= \lambda^3 \int_0^{\infty} x e^{-\lambda x} dx \quad (13)$$

$$= \lambda \quad (14)$$

## Option 1 Contd.

From above calculations we can say,

$$E\left[\frac{1}{X}\right] = \frac{\lambda}{2} \quad (15)$$

So the bias of estimator is given by,

$$B(\Theta) = E[\Theta] - \theta \quad (16)$$

$$= \lambda - \lambda = 0 \quad (17)$$

Therefore  $\frac{2}{n} \sum_{i=1}^n \frac{1}{X_i}$  is an unbiased estimator of  $\lambda$ .

**Option 1 is correct.**



## Option 2

Here in this option, we have our estimator  $\Theta$  and estimand  $\theta$  as,

$$\Theta = \frac{3n}{\sum_{i=1}^n X_i} \text{ and } \theta = \lambda \quad (18)$$

The expectation value of the estimator is given by,

$$E[\Theta] = E \left[ \frac{3n}{\sum_{i=1}^n X_i} \right] \quad (19)$$

$$= \frac{3n}{\sum_{i=1}^n} E \left[ \frac{1}{X_i} \right] \quad (20)$$

## Option 2 Contd.

The value of  $E\left[\frac{1}{X_i}\right]$  can be obtained from (15) as so we have,

$$E[\Theta] = \frac{3n}{\sum_{i=1}^n} \frac{\lambda}{2} \quad (21)$$

$$= \frac{3n}{n} \frac{\lambda}{2} = \frac{3\lambda}{2} \quad (22)$$

So the bias of estimator is given by,

$$B(\Theta) = E[\Theta] - \theta \quad (23)$$

$$= \frac{3\lambda}{2} - \lambda \neq 0 \quad (24)$$

Therefore  $\frac{3n}{\sum_{i=1}^n X_i}$  is not an unbiased estimator of  $\lambda$

**Option 2 is not correct.**

## Option 3

Now here we have our estimator  $\Theta$  and estimand  $\theta$  as,

$$\Theta = \frac{2}{n} \sum_{i=1}^n \frac{1}{X_i} \text{ and } \theta = \lambda \quad (25)$$

Now the variance of  $\Theta$  is calculated as

$$\text{Var}(\Theta) = \text{Var} \left( \frac{2}{n} \sum_{i=1}^n \frac{1}{X_i} \right) \quad (26)$$

$$= \frac{4}{n^2} \sum_{i=1}^n \text{Var} \left( \frac{1}{X_i} \right) \quad (27)$$

$$= \frac{4n}{n^2} \left( E \left[ \frac{1}{X_i}^2 \right] - E \left[ \frac{1}{X_i} \right]^2 \right) \quad (28)$$

$$= \frac{4}{n} \left( \int_{-\infty}^{\infty} \frac{1}{x^2} f(x) dx - \left( \frac{\lambda}{2} \right)^2 \right) \quad (29)$$

## Option 3 Contd.

$$\text{Var}(\Theta) = \frac{4}{n} \left( \int_0^\infty \frac{1}{x^2} \frac{1}{2} \lambda^3 x^2 e^{-\lambda x} dx - \frac{\lambda^2}{4} \right) \quad (30)$$

$$= \frac{\lambda^2}{n} \quad (31)$$

The bias of  $\Theta$  from option 1 is given as  $B(\Theta) = 0$  So we have,

$$\text{MSE}(\Theta_n) = \text{Var}(\Theta) + B(\Theta)^2 \quad (32)$$

$$= \frac{\lambda^2}{n} \quad (33)$$

$$\lim_{n \rightarrow \infty} \text{MSE}(\Theta_n) = \lim_{n \rightarrow \infty} \frac{\lambda^2}{n} \quad (34)$$

$$= 0 \quad (35)$$

Therefore,  $\frac{2}{n} \sum_{i=1}^n \frac{1}{X_i}$  is a consistent estimator of  $\lambda$ .

**Option 3 is correct.**

## Option 4

Now here we have our estimator  $\Theta$  and estimand  $\theta$  as,

$$\Theta = \frac{3n}{\sum_{i=1}^n X_i} \text{ and } \theta = \lambda \quad (36)$$

Now the variance of  $\Theta$  is calculated as

$$\text{Var}(\Theta) = \text{Var}\left(\frac{3n}{\sum_{i=1}^n X_i}\right) \quad (37)$$

$$= \frac{9n^2}{\sum_{i=1}^n} \text{Var}\left(\frac{1}{X_i}\right) \quad (38)$$

$$(39)$$

Now the value of  $\text{Var}\left(\frac{1}{X_i}\right)$  from (28) is substituted, we have

## Option 4 Contd.

$$Var(\Theta) = \frac{9n^2}{\sum_{i=1}^n} \frac{\lambda^2}{4} \quad (40)$$

$$= \frac{9n^2 \lambda^2}{4n} \quad (41)$$

$$= \frac{9n \lambda^2}{4} \quad (42)$$

The bias of  $\Theta$  from option 2 is given as  $B(\Theta) = \frac{\lambda}{2}$  So we have,

$$MSE(\Theta_n) = Var(\Theta) + B(\Theta)^2 \quad (43)$$

$$= \frac{9n \lambda^2}{4} + \left(\frac{\lambda}{2}\right)^2 \quad (44)$$

$$= \frac{\lambda^2}{4}(9n + 1) \quad (45)$$

## Option 4 Contd.

Now,

$$\lim_{n \rightarrow \infty} MSE(\Theta_n) = \lim_{n \rightarrow \infty} \frac{\lambda^2}{4} (9n + 1) \quad (46)$$

(47)

Clearly as  $n$  grows larger  $9n + 1$  also grows larger, so

$$\lim_{n \rightarrow \infty} MSE(\Theta_n) \neq 0 \quad (48)$$

Therefore,  $\frac{3n}{\sum_{i=1}^n X_i}$  is not a consistent estimator of  $\lambda$ .

Option 4 is not correct.

**Therefore option 1 and option 3 are correct.**