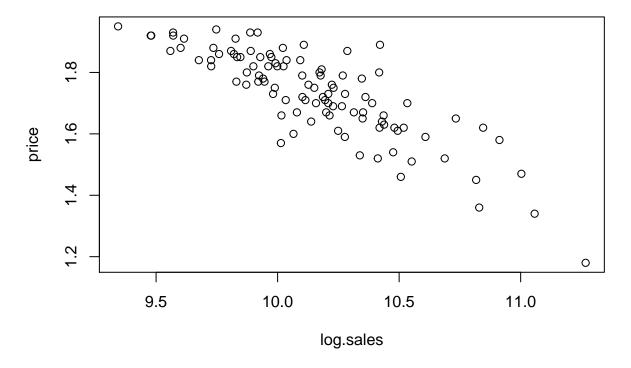
ST565: Time Series HW6

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Question 1

The dataset bluebirdlite contains log sales and prices for the "lite" version of bluebird chips. Quantify the relationship between sales and price.

plot(bluebirdlite)



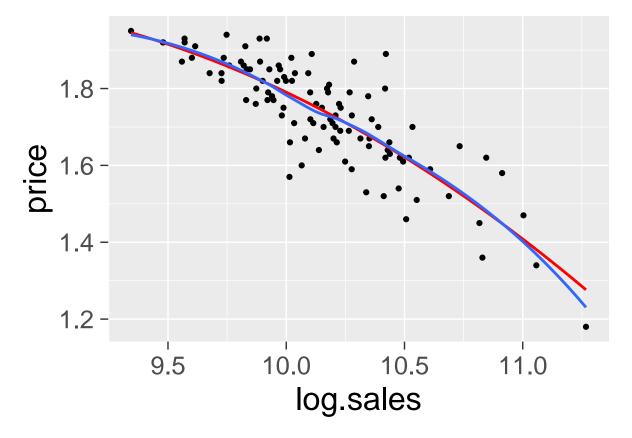
```
big_font <- theme_grey(base_size = 24)

attach(bluebirdlite)

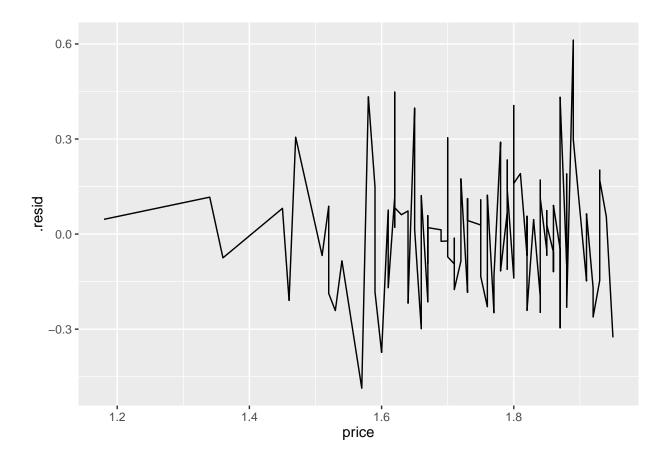
fit_lm <- lm(log.sales ~ price + I(price^2), data = bluebirdlite)
summary(fit_lm)</pre>
```

```
##
## Call:
## lm(formula = log.sales ~ price + I(price^2), data = bluebirdlite)
##
## Residuals:
## Min 1Q Median 3Q Max
## -0.48609 -0.14069 0.01389 0.11249 0.61228
##
```

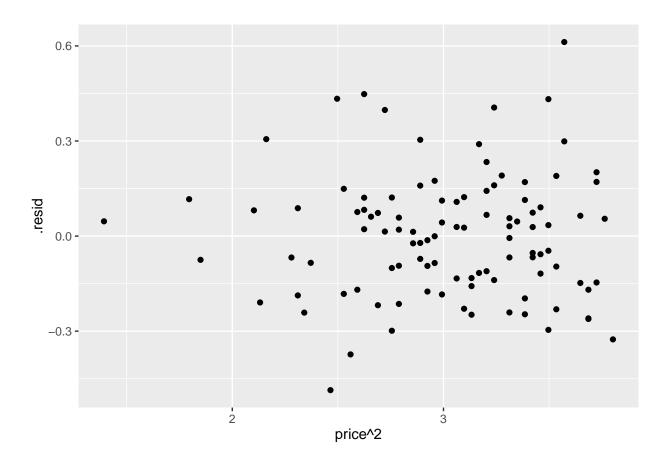
```
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 12.5954
                        1.6938
                                  7.436 3.48e-11 ***
## price
               -0.6507
                           2.0404 -0.319
                                            0.750
                          0.6121 -0.712
## I(price^2)
               -0.4357
                                            0.478
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.1968 on 101 degrees of freedom
## Multiple R-squared: 0.7054, Adjusted R-squared: 0.6996
## F-statistic: 120.9 on 2 and 101 DF, p-value: < 2.2e-16
qplot(log.sales, price, data = bluebirdlite) +
 geom_smooth(method = "lm", formula = y ~ poly(x, 2), se = FALSE, colour = "red") +
 geom_smooth(se = FALSE) +
 big_font
```



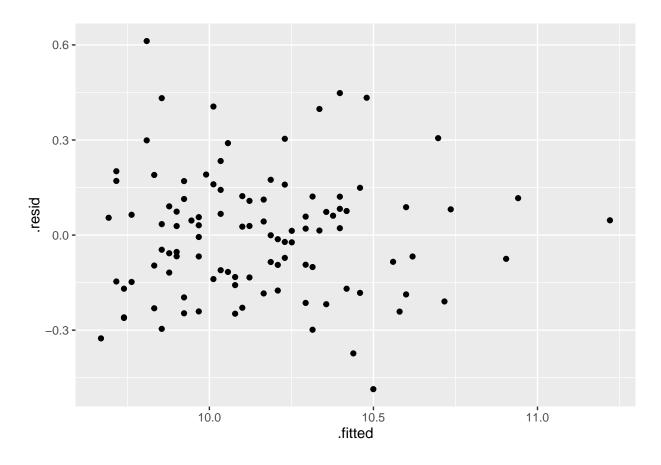
```
# assumptions
# residuals versus covariates
sales_lm <- fortify(fit_lm)
qplot(price, .resid, data = sales_lm, geom= "line")</pre>
```



qplot(price^2, .resid, data = sales_lm)

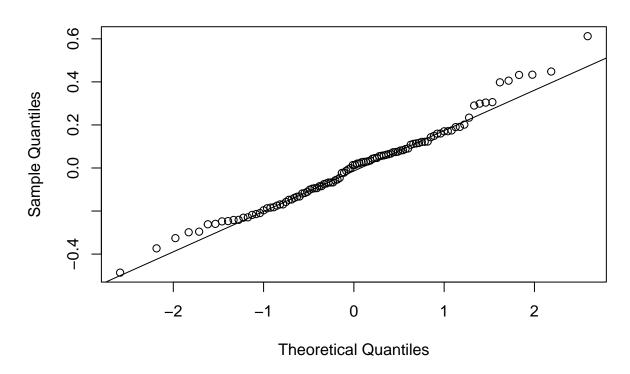


residuals versus fitted
qplot(.fitted, .resid, data = sales_lm)



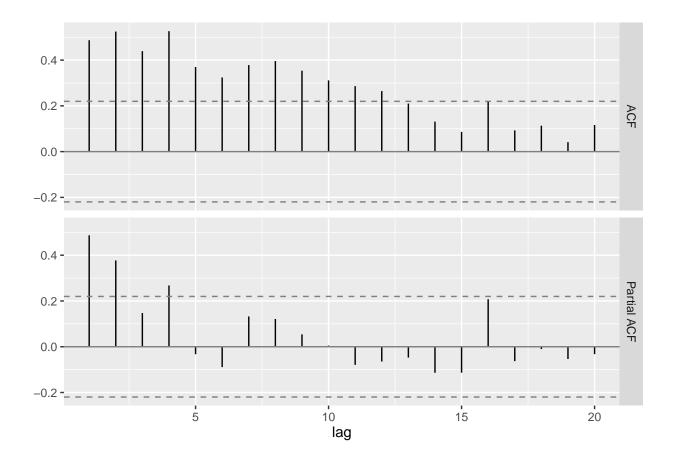
```
# normality of residuals
qqnorm(sales_lm$.resid)
qqline(sales_lm$.resid)
```

Normal Q-Q Plot

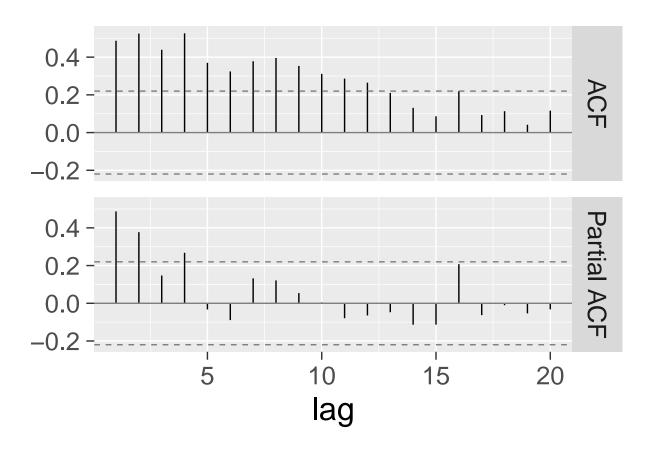


```
# correlation of residuals
source(url("http://stat565.cwick.co.nz/code/get_acf.R")) # my code for examine_corr
examine_corr(residuals(fit_lm))
```

Warning: closing unused connection 5 (http://stat565.cwick.co.nz/code/
get_acf.R)

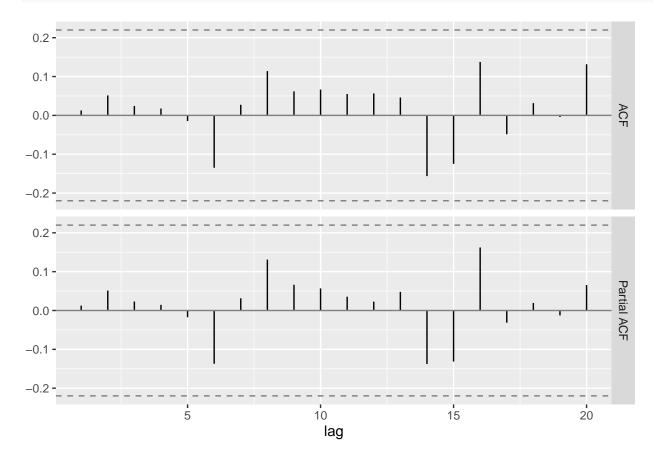


last_plot() + big_font



```
## Generalized least squares fit by maximum likelihood
##
     Model: log.sales ~ price + I(price^2)
     Data: bluebirdlite
##
##
          AIC
                    BIC
                          logLik
##
     -92.6015 -71.44638 54.30075
##
## Correlation Structure: ARMA(4,0)
    Formula: ~1
    Parameter estimate(s):
##
##
        Phi1
                  Phi2
                            Phi3
  0.1851510 0.2132629 0.1345612 0.3039307
##
## Coefficients:
                                    t-value p-value
##
                   Value Std.Error
## (Intercept) 14.899928 1.1651084 12.788448 0.0000
               -3.658602 1.3918337 -2.628620 0.0099
## price
                0.523881 0.4163226 1.258355 0.2112
## I(price^2)
##
##
   Correlation:
```

```
##
              (Intr) price
              -0.994
## price
## I(price^2) 0.985 -0.997
## Standardized residuals:
##
           Min
                         Q1
                                    Med
                                                 QЗ
                                                             Max
## -2.23177526 -0.70553535 0.04652695 0.65792580 2.90668252
##
## Residual standard error: 0.1943017
## Degrees of freedom: 104 total; 101 residual
# arima_fit <- with(bluebirdlite,</pre>
# arima(log.sales, order = c(4, 0, 1), xreg = cbind(price, I(price^2))))
# arima_fit
# diagnostics
fit_lm$residuals <- residuals(gls_fit, type = "normalized")</pre>
bluebirdlite$fitted <- fitted(gls_fit)</pre>
examine_corr(fit_lm$residuals) #looks good
```



```
plot(bluebirdlite$fitted, gls_fit$residuals, data = bluebirdlite)
```

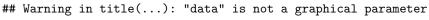
Warning in plot.window(...): "data" is not a graphical parameter

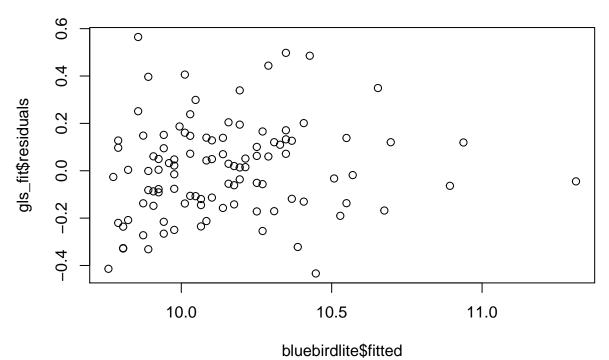
```
## Warning in plot.xy(xy, type, ...): "data" is not a graphical parameter

## Warning in axis(side = side, at = at, labels = labels, ...): "data" is not
## a graphical parameter

## Warning in axis(side = side, at = at, labels = labels, ...): "data" is not
## a graphical parameter

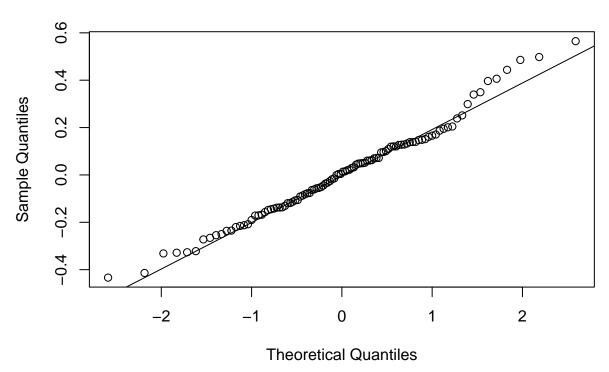
## Warning in box(...): "data" is not a graphical parameter
## Warning in title(___): "data" is not a graphical parameter
```





qqnorm(gls_fit\$residuals)
qqline(gls_fit\$residuals) #Looks okay

Normal Q-Q Plot



```
confint(gls_fit)
                   2.5 %
                              97.5 %
## (Intercept) 12.6163578 17.1834987
## price
              -6.3865463 -0.9306585
## I(price^2) -0.2920958 1.3398587
confint(fit_lm)
##
                   2.5 %
                            97.5 %
## (Intercept) -4.733007 29.923726
## price
              -21.524851 20.223381
## I(price^2)
               -6.697864 5.826414
intervals(gls_fit)
```

```
## Approximate 95% confidence intervals
##
## Coefficients:
## lower est. upper
## (Intercept) 12.5886666 14.8999282 17.2111898
## price -6.4196260 -3.6586024 -0.8975787
## I(price^2) -0.3019906 0.5238814 1.3497534
## attr(,"label")
## [1] "Coefficients:"
##
```

```
Correlation structure:
##
            lower est.
                                   upper
## Phi1 0.10104695 0.1851510 -0.06254519
## Phi2 0.12509406 0.2132629 0.23165667
## Phi3 0.01130536 0.1345612 0.32715590
## Phi4 0.11381837 0.3039307 0.47257119
## attr(,"label")
## [1] "Correlation structure:"
##
##
  Residual standard error:
##
       lower
                  est.
                           upper
## 0.1394571 0.1943017 0.2707152
# backtransform and use % decrease with price increase of 0.1
est_gls <- 100*(1 - exp(0.1 * coef(gls_fit)["price"]))</pre>
ci_gls <- 100*(1 - exp(0.1 * confint(gls_fit)["price", ]))</pre>
```

It is estimated that an increase in price of 10 cents is associated with decrease in median sales of 30.64 (95% CI 47.2 to 8.89).

Question 2

In class we looked at modelling the relationship between mortality, temperature and particulate matter. Repeat the analysis but seasonally difference all three series first. Compare the results.

```
library(ggplot2)
library(dplyr)
# install.packages("tidyr")
library(tidyr)
load(url("http://www.stat.pitt.edu/stoffer/tsa3/tsa3.rda"))
source(url("http://stat565.cwick.co.nz/code/fortify-ts.r"))
source(url("http://stat565.cwick.co.nz/code/get_acf.R")) # my code for examine_corr
big_font <- theme_grey(base_size = 24)</pre>
# mort \leftarrow c(rep(NA, 52), diff(cmort, lag = 52))
# temp <- c(rep(NA, 52), diff(tempr, lag = 52))
# part \leftarrow c(rep(NA, 52), diff(part, lag = 52))
mort <- diff(cmort, lag = 52)</pre>
temp <- diff(tempr, lag = 52)</pre>
part <- diff(part, lag = 52)</pre>
mort <- data.frame(mortality = mort, part = part, temp = temp)</pre>
mort$time <- fortify(cmort)$time[53:508]</pre>
# mort
head(mort)
```

```
## mortality part temp time
## 1 7.54 -28.38 -13.75 1971.000
```

```
## 2 0.35 -8.80 -4.69 1971.019

## 3 8.36 -0.59 4.58 1971.038

## 4 6.62 -16.72 -8.93 1971.058

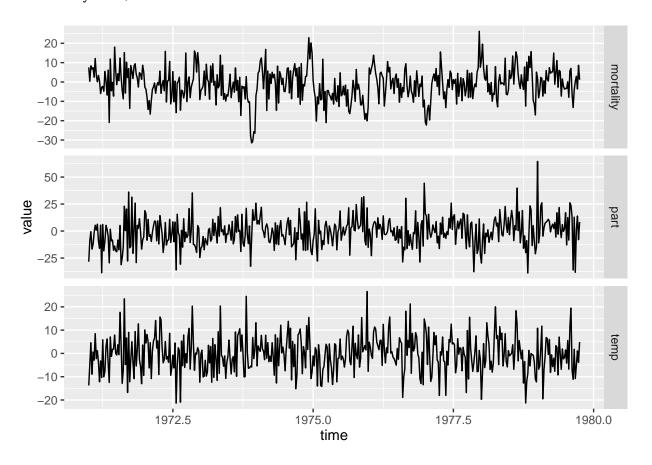
## 5 7.45 -12.48 0.62 1971.077

## 6 2.48 -1.55 -2.90 1971.096
```

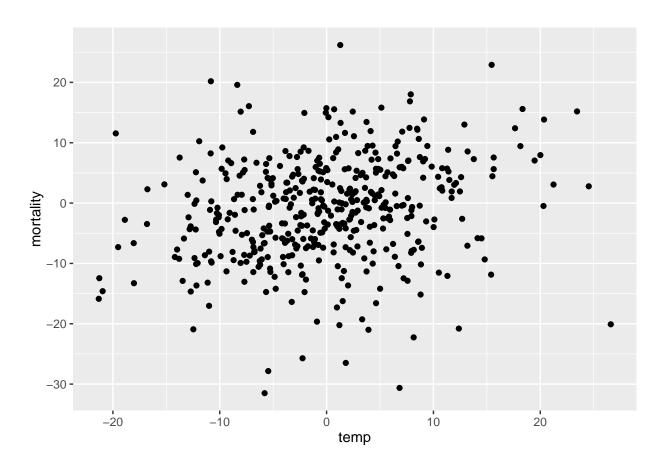
```
qplot(time, value, data = gather(mort, variable, value, -time), geom = "line") +
  facet_grid(variable ~ ., scale = "free")
```

Warning: closing unused connection 7 (http://stat565.cwick.co.nz/code/
get_acf.R)

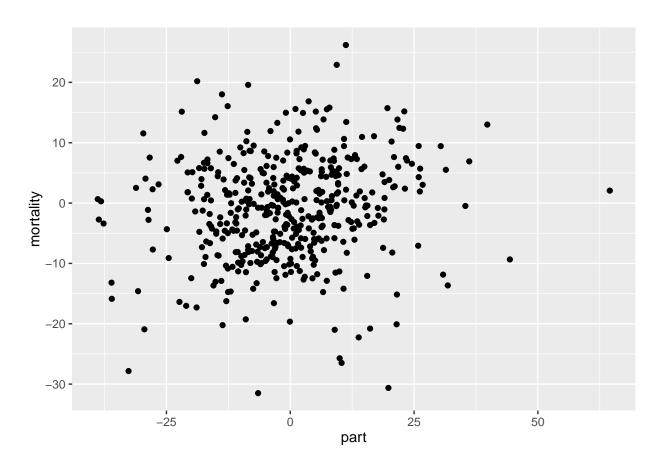
Warning: closing unused connection 6 (http://stat565.cwick.co.nz/code/
fortify-ts.r)



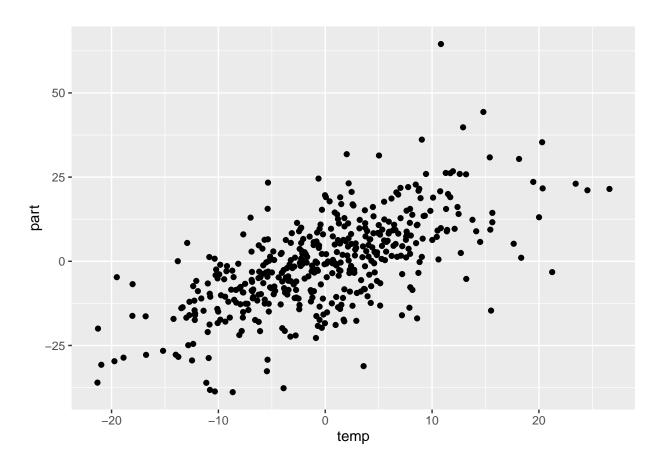
```
qplot(temp, mortality, data = mort)
```



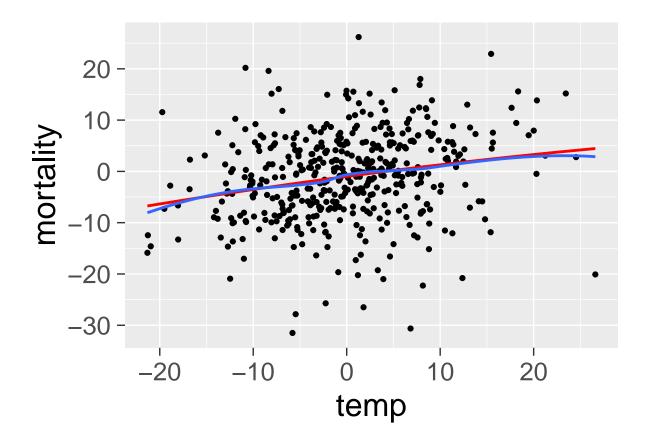
```
qplot(part, mortality, data = mort)
```



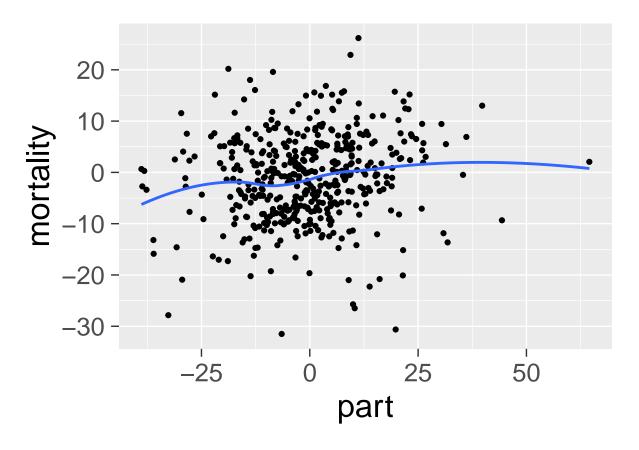
```
qplot(temp, part, data = mort)
```



```
qplot(temp, mortality, data = mort) +
  geom_smooth(method = "lm", formula = y ~ poly(x, 2), se = FALSE, colour = "red") +
  geom_smooth(se = FALSE) +
  big_font
```



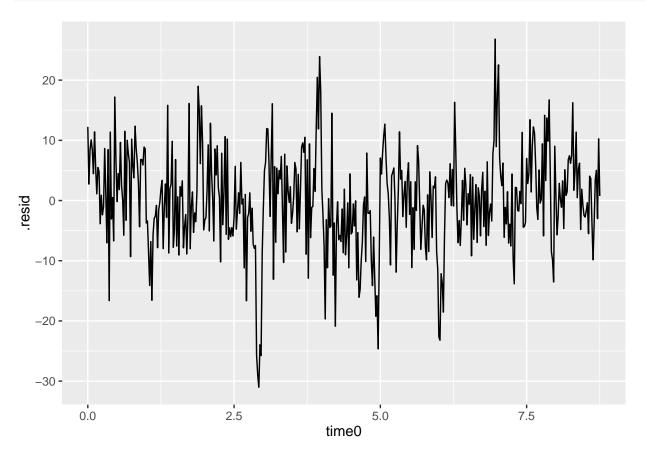
```
qplot(part, mortality, data = mort) + geom_smooth(se = FALSE) +
  big_font
```



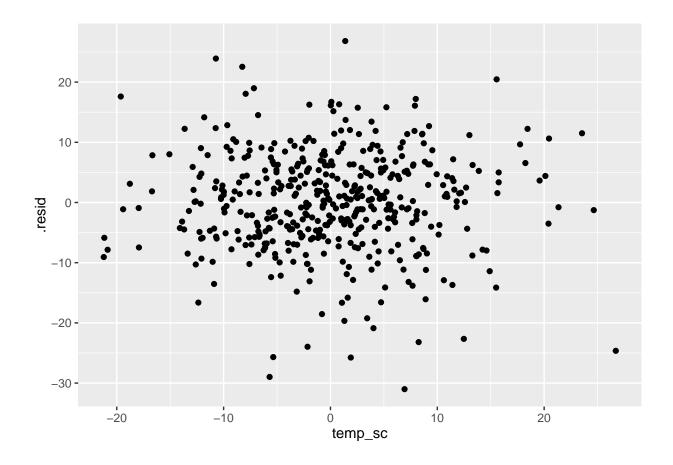
```
mort <- mutate(mort, temp_sc = temp - mean(temp),</pre>
               temp_2 = temp_sc^2,
               time0 = time - min(time))
fit_lm <- lm(mortality ~ time0 + temp_sc + temp_2 + part, data = mort, na.action = na.omit)</pre>
summary(fit_lm)
##
## Call:
## lm(formula = mortality ~ time0 + temp_sc + temp_2 + part, data = mort,
       na.action = na.omit)
##
## Residuals:
       Min
                       Median
                  1Q
                                    ЗQ
                                            Max
## -31.0049 -4.8348
                       0.2759
                                5.3296 26.8181
##
## Coefficients:
                Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -1.281887
                           0.836345 -1.533 0.126045
               0.060907
                           0.157172
                                    0.388 0.698554
## time0
## temp_sc
                                     3.869 0.000125 ***
               0.252980
                           0.065382
## temp_2
               -0.001407
                           0.004095 -0.344 0.731281
               -0.010496
                           0.038038 -0.276 0.782725
## part
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
```

```
## Residual standard error: 8.306 on 451 degrees of freedom
## Multiple R-squared: 0.05112, Adjusted R-squared: 0.0427
## F-statistic: 6.074 on 4 and 451 DF, p-value: 9.107e-05
```

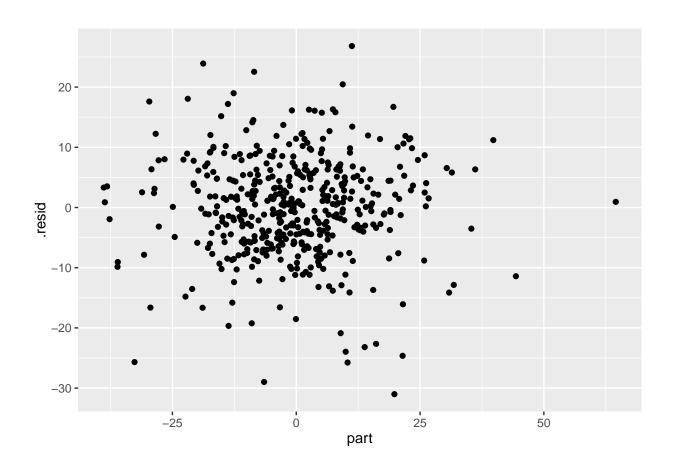
```
# assumptions
# residuals versus covariates
mort_lm <- fortify(fit_lm)
qplot(time0, .resid, data = mort_lm, geom= "line")</pre>
```



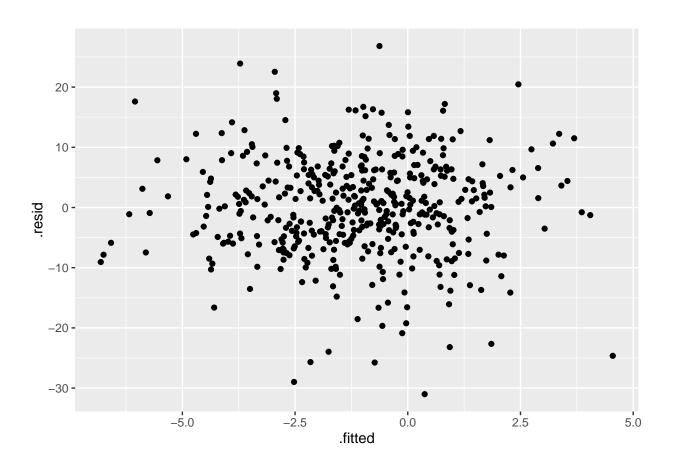
```
qplot(temp_sc, .resid, data = mort_lm)
```



qplot(part, .resid, data = mort_lm)

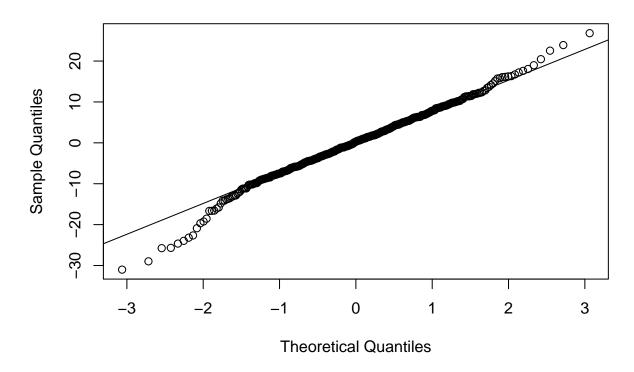


residuals versus fitted
qplot(.fitted, .resid, data = mort_lm)

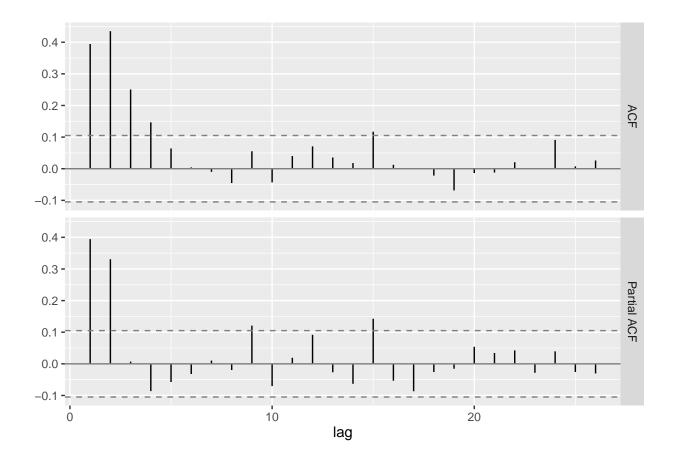


normality of residuals
qqnorm(mort_lm\$.resid)
qqline(mort_lm\$.resid)

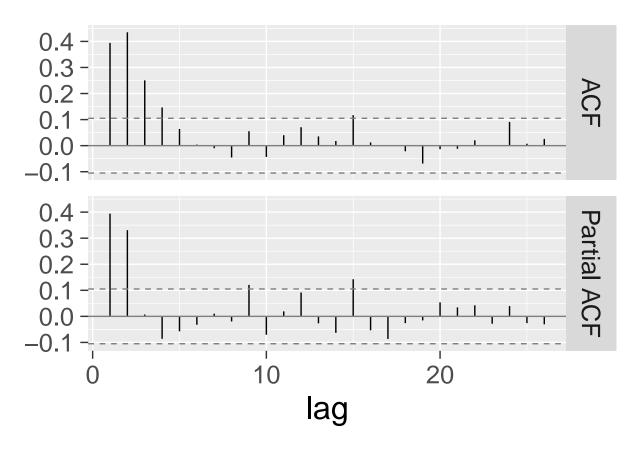
Normal Q-Q Plot



correlation of residuals
examine_corr(residuals(fit_lm))



last_plot() + big_font



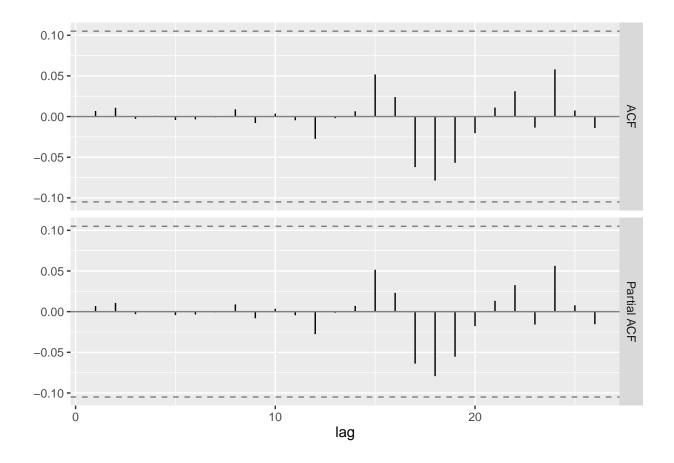
```
# AR (2)? violates regression assumptions
# two ways to fit
library(nlme)
gls_fit <- gls(mortality ~ time0 + temp_sc + temp_2 + part, data = mort,</pre>
    correlation = corARMA(p = 15), method = "ML")
summary(gls_fit)
## Generalized least squares fit by maximum likelihood
##
    Model: mortality ~ time0 + temp_sc + temp_2 + part
##
    Data: mort
##
         AIC
                  BIC
                         logLik
##
    3095.665 3182.237 -1526.832
##
## Correlation Structure: ARMA(15,0)
   Formula: ~1
##
##
   Parameter estimate(s):
##
          Phi1
                                   Phi3
                                                Phi4
                                                            Phi5
                       Phi2
##
   0.281087676  0.376211288  0.022572796  -0.069685677  -0.009436378
##
          Phi6
                       Phi7
                                   Phi8
                                                Phi9
                                                           Phi10
##
   -0.051876224 -0.036587786 -0.008971095
                                         0.135990895 -0.102843892
##
         Phi11
                      Phi12
                                  Phi13
                                               Phi14
                                                           Phi15
   ##
##
## Coefficients:
```

t-value p-value

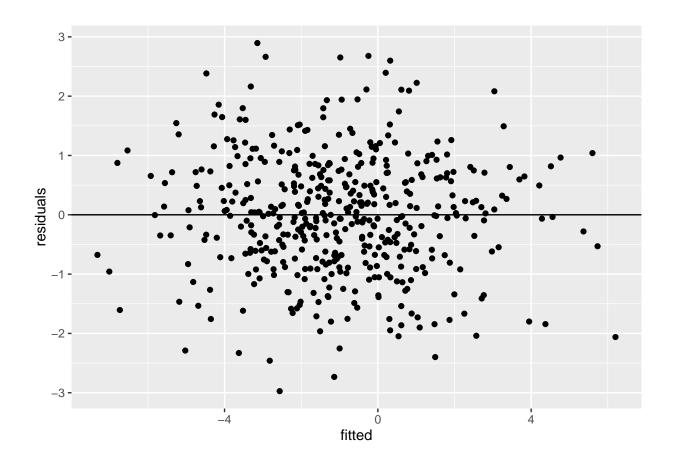
Value Std.Error

##

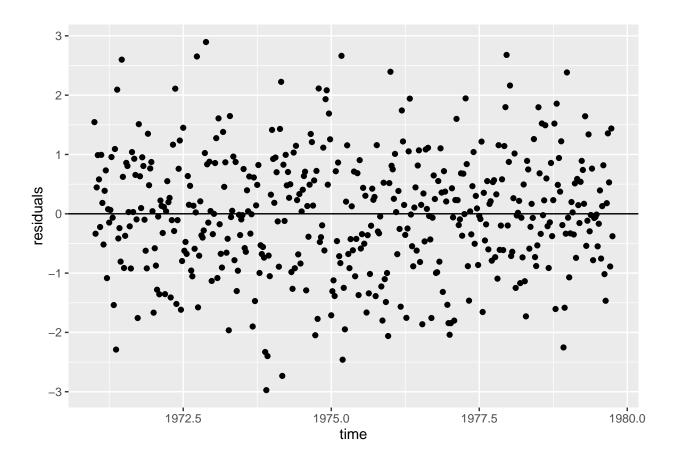
```
## (Intercept) -0.9142608 1.7486759 -0.522830 0.6013
## time0
             -0.0203654 0.3421046 -0.059530 0.9526
## temp sc
              0.2300512 0.0501430 4.587903 0.0000
## temp_2
              0.0001867 0.0029156 0.064029 0.9490
## part
              0.0437174 0.0282271 1.548771 0.1221
##
## Correlation:
##
          (Intr) time0 tmp_sc temp_2
## time0
          -0.857
## temp_sc -0.030 0.042
## temp_2 -0.109 0.003 -0.120
          0.068 -0.070 -0.683 0.040
## part
##
## Standardized residuals:
##
          Min
                      Q1
                                 Med
                                             QЗ
                                                        Max
## -3.87947076 -0.60426277 0.01300854 0.65797250 3.19226482
##
## Residual standard error: 8.281303
## Degrees of freedom: 456 total; 451 residual
# or
# auto.arima(mort$mortality)
arima fit <- with(mort,
 arima(mortality, order = c(15, 0, 0), xreg = cbind(time0, temp_sc, temp_2, part)))
arima_fit
##
## arima(x = mortality, order = c(15, 0, 0), xreg = cbind(time0, temp_sc, temp_2,
##
      part))
## Coefficients:
##
                  ar2
                          ar3
                                   ar4
                                           ar5
                                                    ar6
                                                            ar7
                                                                     ar8
##
        ## s.e. 0.0470 0.0487 0.0514
                               0.0521
                                       0.0526
                                                0.0507
                                                        0.0513
                                                                  0.0513
##
           ar9
                  ar10
                          ar11
                                  ar12
                                          ar13
                                                           ar15 intercept
                                                   ar14
        0.1360 - 0.1029 \ 0.0066 \ 0.1173 - 0.0702 - 0.0979 \ 0.1440
##
                                                                  -0.9144
## s.e. 0.0507
                0.0513 0.0516 0.0511
                                        0.0524
                                               0.0487 0.0465
                                                                   1.7469
##
          time0 temp_sc temp_2
                                   part
##
        -0.0213
                0.2300
                         2e-04 0.0437
                          3e-03 0.0292
## s.e. 0.3412
                0.0541
## sigma^2 estimated as 47.29: log likelihood = -1526.83, aic = 3093.66
# diagnostics
mort$residuals <- residuals(gls_fit, type = "normalized")</pre>
mort$fitted <- fitted(gls_fit)</pre>
examine_corr(mort$residuals)
```



qplot(fitted, residuals, data = mort) + geom_hline(yintercept = 0)

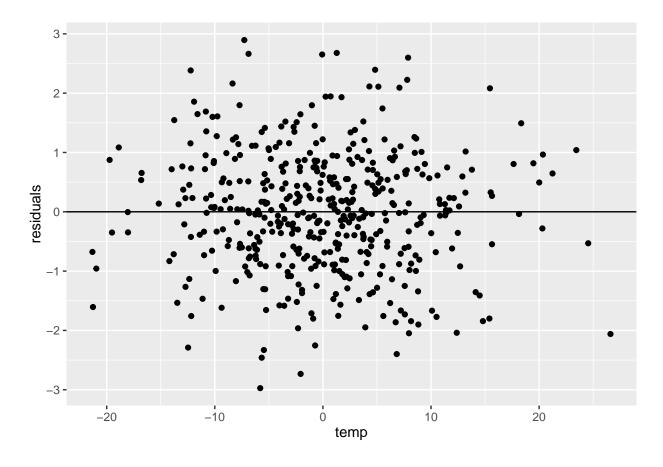


qplot(time, residuals, data = mort) + geom_hline(yintercept = 0)



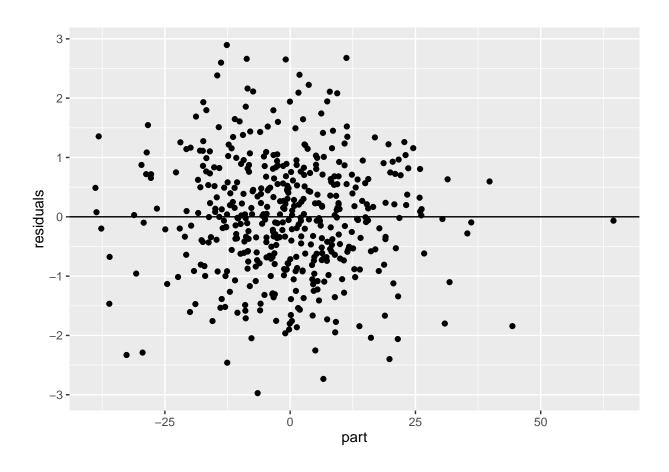
qplot(temp, residuals, data = mort) + geom_hline(yintercept = 0)

Don't know how to automatically pick scale for object of type ts. Defaulting to continuous



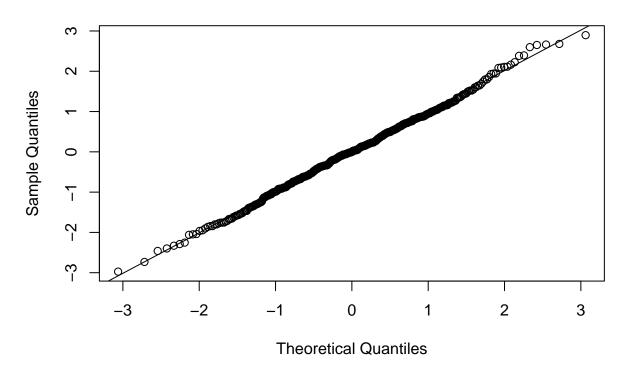
```
qplot(part, residuals, data = mort) + geom_hline(yintercept = 0)
```

Don't know how to automatically pick scale for object of type ts. Defaulting to continuous



qqnorm(mort\$residuals)
qqline(mort\$residuals)

Normal Q-Q Plot



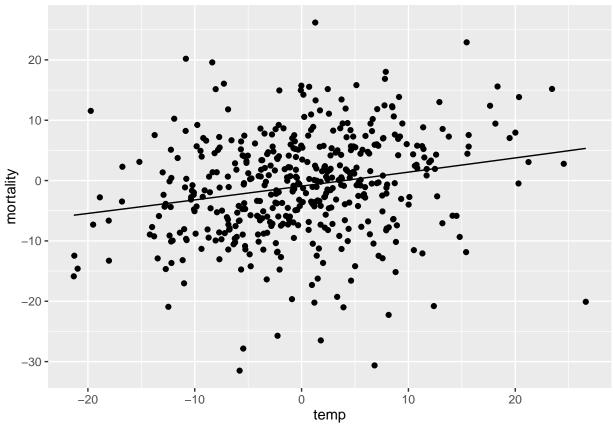
confint(arima_fit)

confint(fit_lm)

```
2.5 %
##
                               97.5 %
## ar1
              0.18887839
                          0.373302306
## ar2
              0.28067323
                          0.471767161
## ar3
             -0.07819816
                          0.123256413
## ar4
             -0.17179392
                          0.032500759
## ar5
             -0.11245853
                          0.093541568
             -0.15128301 0.047550479
## ar6
## ar7
             -0.13722378 0.063954810
## ar8
             -0.10943930
                          0.091526134
## ar9
              0.03666492 0.235335205
## ar10
             -0.20337533 -0.002367212
## ar11
             -0.09447303 0.107653745
## ar12
              0.01720230
                          0.217465367
## ar13
             -0.17296648 0.032542975
## ar14
             -0.19331659 -0.002469489
## ar15
              0.05277746
                          0.235243173
## intercept -4.33830438
                          2.509522744
## time0
             -0.69001009
                          0.647365473
## temp_sc
              0.12391961
                          0.336159556
## temp_2
             -0.00566624
                          0.006041411
## part
             -0.01360277
                          0.101041880
```

2.5 % 97.5 %

```
## (Intercept) -2.925503168 0.361729983
## time0
              -0.247972794 0.369787229
              0.124489216 0.381471429
## temp sc
              -0.009455327 0.006640811
## temp_2
## part
              -0.085249027 0.064257196
intervals(gls_fit)
## Approximate 95% confidence intervals
##
##
   Coefficients:
##
                     lower
                                    est.
                                               upper
## (Intercept) -4.350825057 -0.9142608019 2.522303453
              -0.692682374 -0.0203654481 0.651951478
              0.131508264 0.2300511561 0.328594048
## temp_sc
## temp_2
              -0.011755686 0.0437173682 0.099190422
## part
## attr(,"label")
## [1] "Coefficients:"
##
##
  Correlation structure:
##
              lower
                            est.
                                        upper
        0.13296541 0.281087676 0.166031008
## Phi1
## Phi2 0.25778977 0.376211288 0.310109711
## Phi3 -0.06733034 0.022572796 0.019586118
## Phi4 -0.16190861 -0.069685677 -0.057467380
## Phi5 -0.10097754 -0.009436378 -0.004552784
## Phi6 -0.13706313 -0.051876224 -0.039901390
## Phi7 -0.12968346 -0.036587786 -0.008844102
## Phi8 -0.09685403 -0.008971095 0.026611710
## Phi9
         0.04799215 0.135990895 0.162242420
## Phi10 -0.17031658 -0.102843892 -0.052828465
## Phi11 -0.05241476  0.006614401  0.050032658
## Phi12 0.05774212 0.117305504 0.148565429
## Phi13 -0.11844013 -0.070173766 -0.018916003
## Phi14 -0.15913453 -0.097902762 -0.004405277
## Phi15 0.05133356 0.144008956 0.234225287
## attr(,"label")
## [1] "Correlation structure:"
##
  Residual standard error:
##
     lower
               est.
                       upper
## 7.546556 8.281303 9.087585
# if you refit the model with white noise errors using gls, you can actually
# test to see if the correlation structure improved the fit
gls_wn <- gls(mortality ~ time0 + temp_sc + temp_2 + part, data = mort, method = "ML")</pre>
anova(gls_wn, gls_fit)
          Model df
                                 BIC
                                        logLik
                        AIC
                                                 Test L.Ratio p-value
## gls_wn
              1 6 3231.778 3256.513 -1609.889
## gls_fit
              2 21 3095.665 3182.237 -1526.832 1 vs 2 166.1131 <.0001
```



My analysis shows that removing the seasonality by differencing with lag 52, results in a better-fitting model, in comparison with when seasonality is present in data and is not dealt with properly. This is because when there exists a clear seasonal trend in the data, the data is non-stationary and we should remove non-stationarity for our analysis and model to be meaningful. In our case, the annual changes in temperature and other seasonal trends are a result of one confounding factor "the time of the year". Differencing gets rid of this, and gives us a clean stationary time series. Our model shows different significant explanatory variables, and the AIC for our model is better than Charlotte's provided code without removing seasonality.