
Statistical Analysis of Reliability and Survival Data: Rotterdam Dataset

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1 Exploratory Data Analysis

These data sets are used in the paper by Royston and Altman that is referenced below. The Rotterdam data is used to create a fitted model, and the GBSG data for validation of the model. The paper gives references for the data source.

There are 43 subjects who have died without recurrence, but whose death time is greater than the censoring time for recurrence. A common way that this happens is that a death date is updated in the health record sometime after the research study ended, and said value is then picked up when a study data set is created. But it raises serious questions about censoring. For instance subject 40 is censored for recurrence at 4.2 years and died at 6.6 years; when creating the endpoint of recurrence free survival (earlier of recurrence or death), treating them as a death at 6.6 years implicitly assumes that they were recurrence free just before death. For this to be true we would have to assume that if they had progressed in the 2.4 year interval before death (while off study), that this information would also have been noted in their general medical record, and would also be captured in the study data set. However, that may be unlikely. Death information is often in a centralized location in electronic health records, easily accessed by a programmer and merged with the study data, while recurrence may require manual review. How best to address this is an open issue.

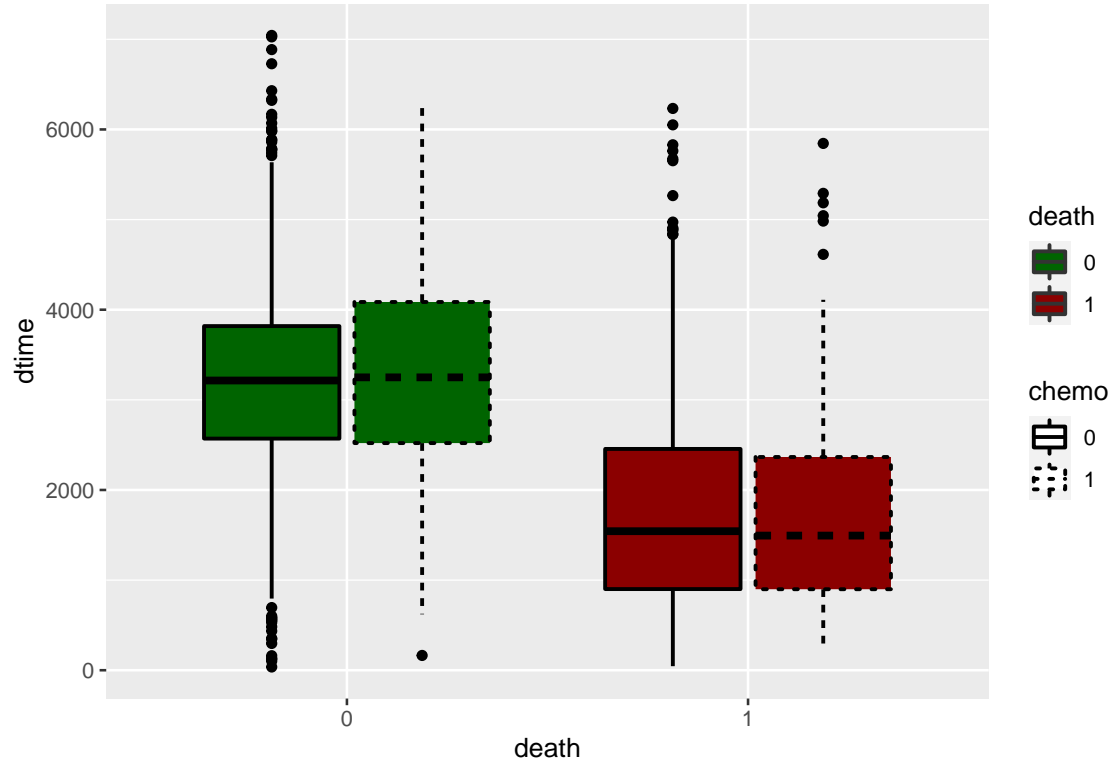
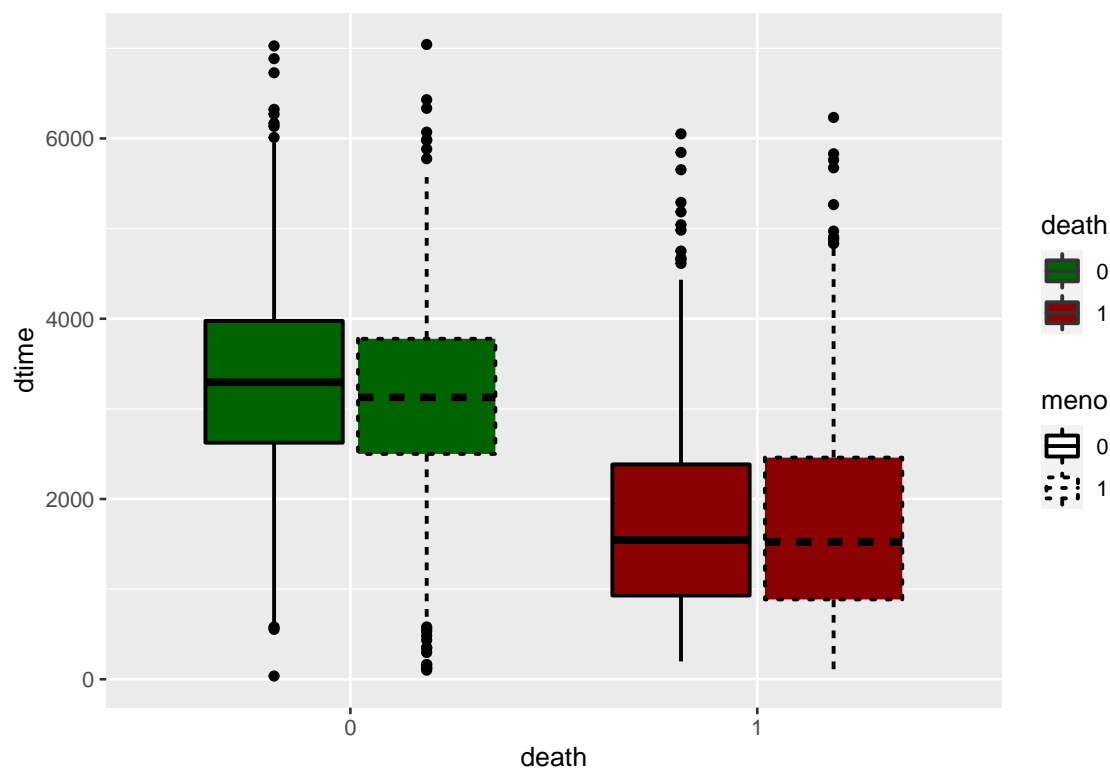
Table 1: Data description

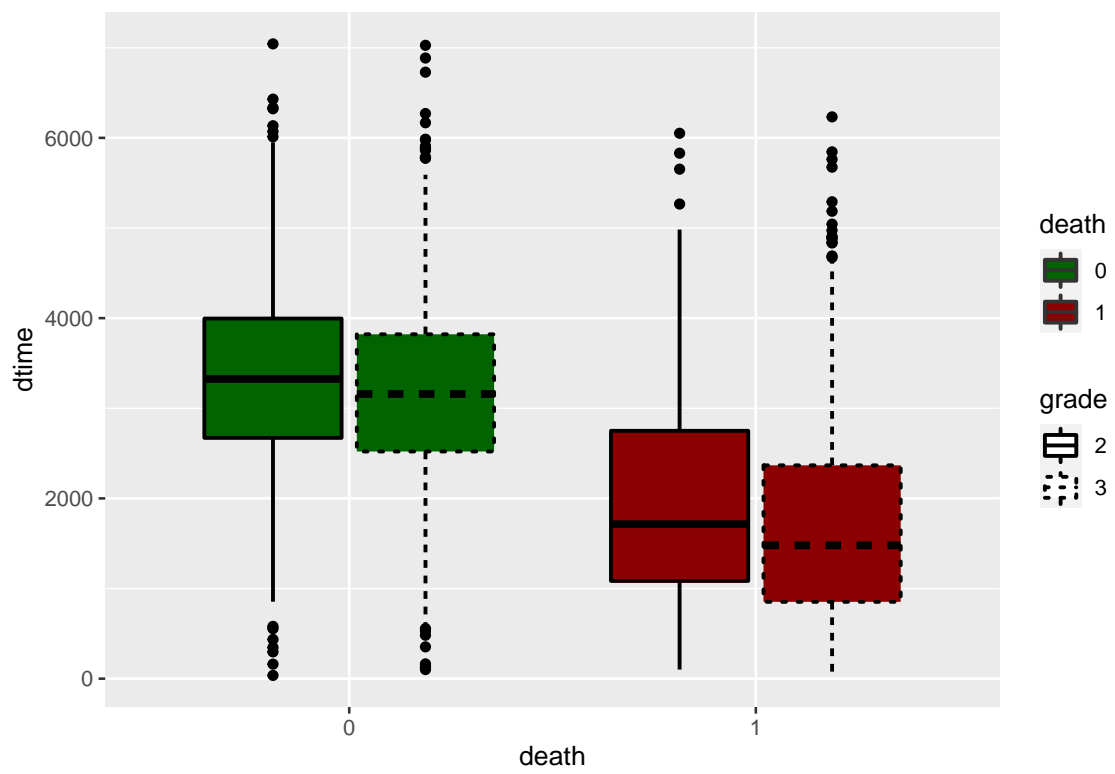
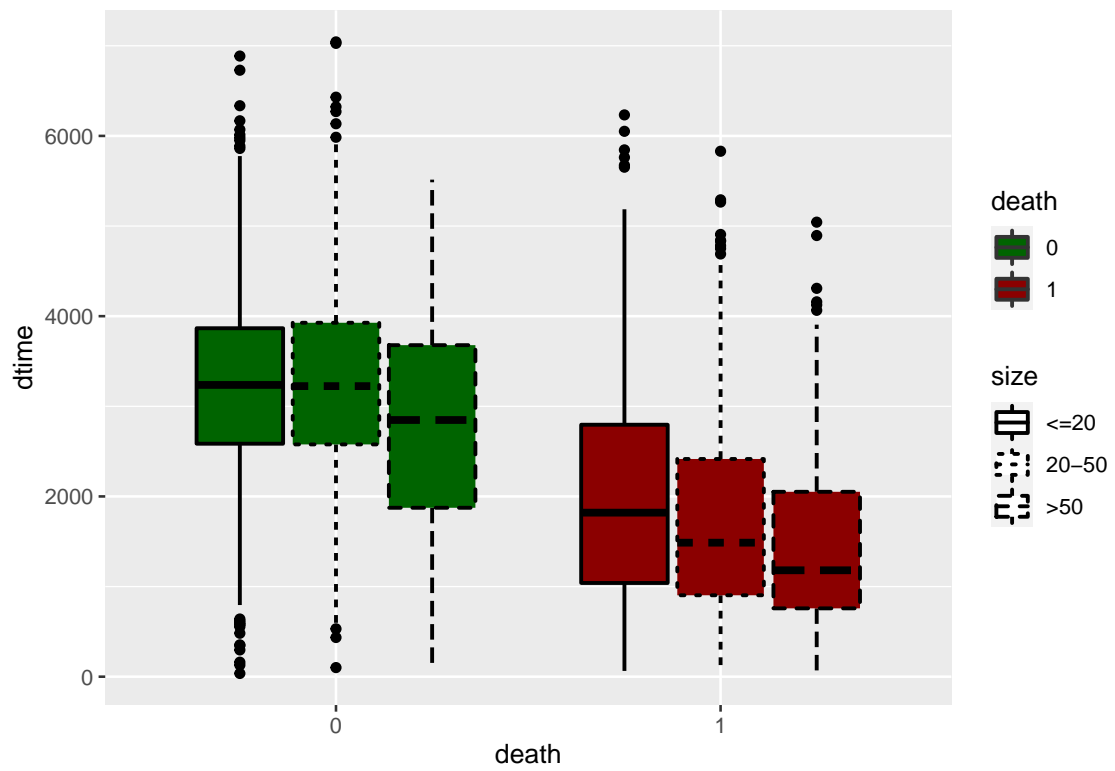
pid	Patient identifier
year	Year of surgery
age	Age at surgery
meno	Menopausal status (0 = premenopausal, 1 = postmenopausal)
size	Tumor size, a factor with levels ≤ 20 , 20-25, > 50
grade	Differentiation grade
nodes	Number of positive lymph nodes
pgr	Progesterone receptors (fmol/l)
er	Estrogen receptors (fmol/l)
hormon	Hormonal treatment (0=no, 1=yes)
chemo	Chemotherapy
rtime	Days to relapse or last follow-up
recur	0 = no relapse, 1 = relapse
dtime	Days to death or last follow-up
death	0 = alive, 1 = dead

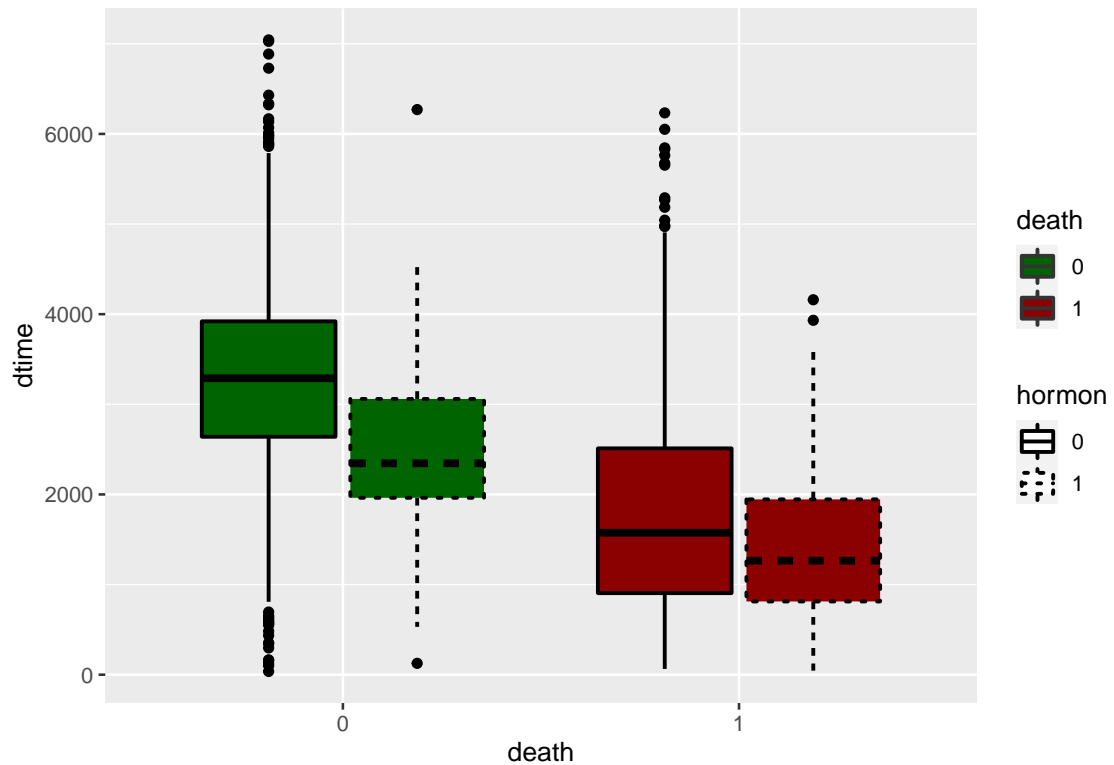
Table 1 explains the covariates in the Rotterdam dataset.

When we investigate the categorical variables, we see that some have a larger difference between time until recurrence than others. When we first investigate menopause, we see that recurring patients are very similar in their recurrence time. However, non-recurring patients have a slightly lower median with a similar distribution. A very similar pattern is observed for chemotherapy but the non-recurring patients had a slightly higher time before leaving the study. While size, having three levels, needs to be interpreted slightly differently. With patients that relapsed into cancer, the larger cancer cells certainly caused it to happen sooner. This makes intuitive sense that more severe cancer would have more long term health risks. Among patients that do not relapse, only the extremely large cancer patients had a much lower recurrence time. While small and medium

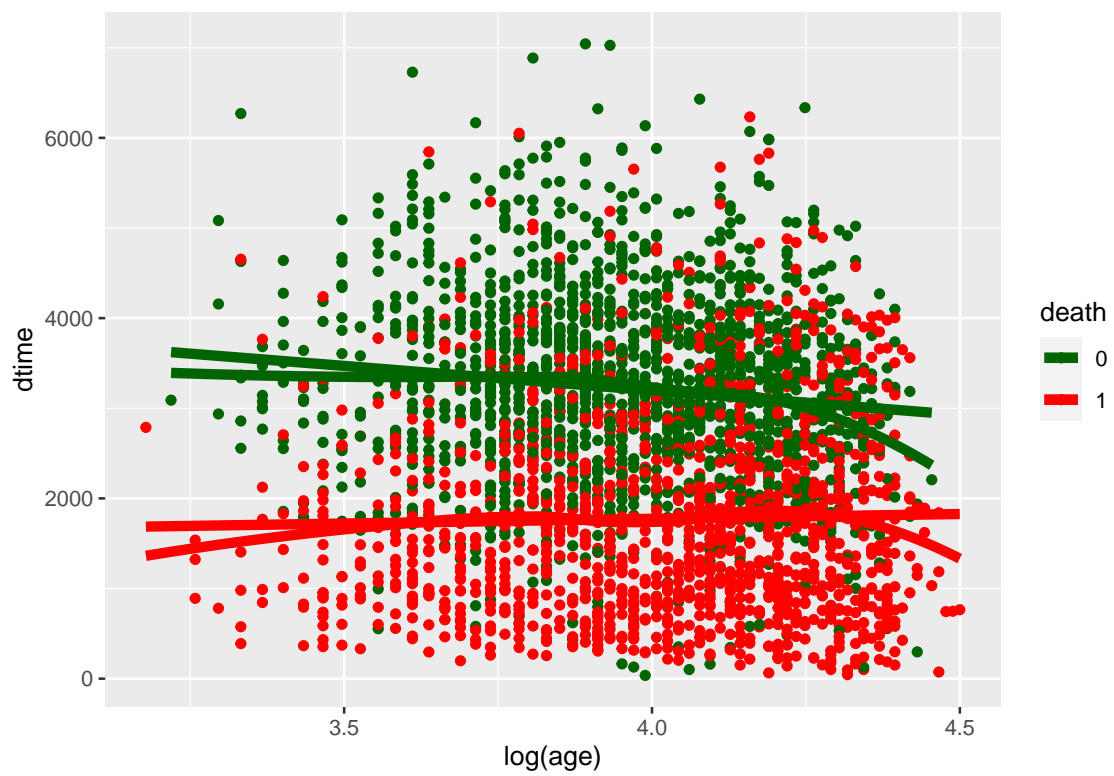
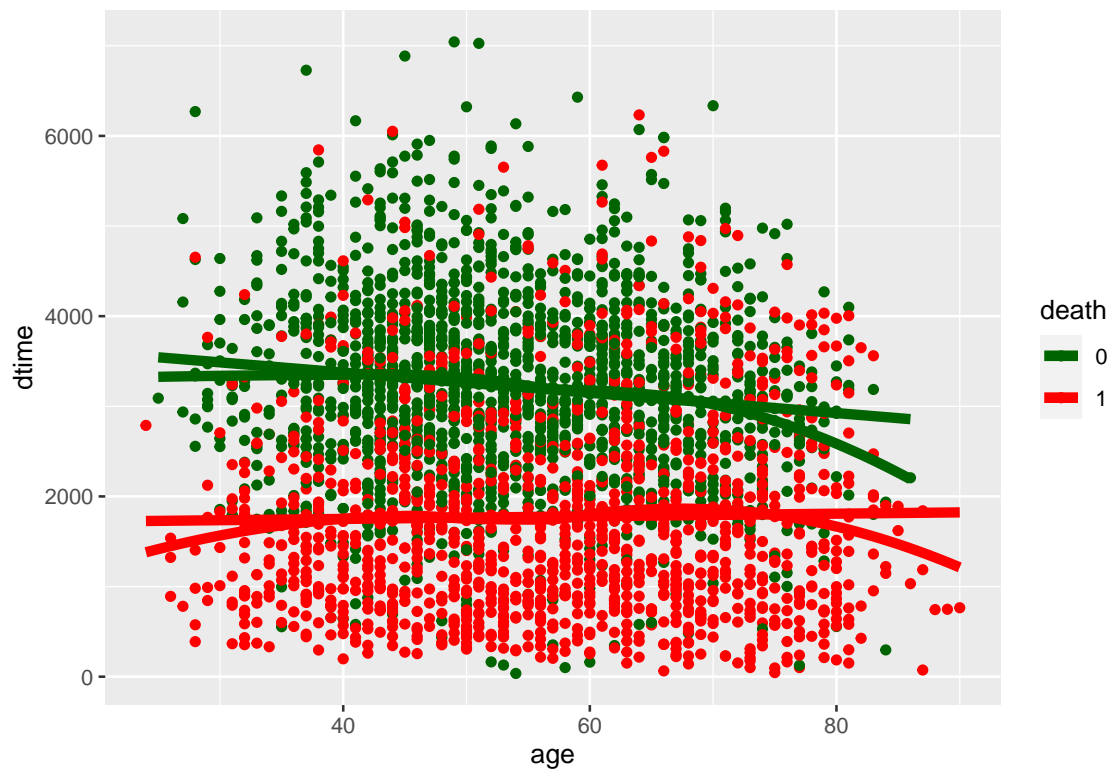
patients had a very similar response. Grade had a very similar response among relapsing and non-relapsing patients. Patients given hormonal therapy seem to have a smaller tail among relapsing patients, meaning those who relapse are more likely to do it soon after treatment. Non-relapsing patients that had hormonal therapy also appear to be censored more quickly than those who do not receive the treatment.





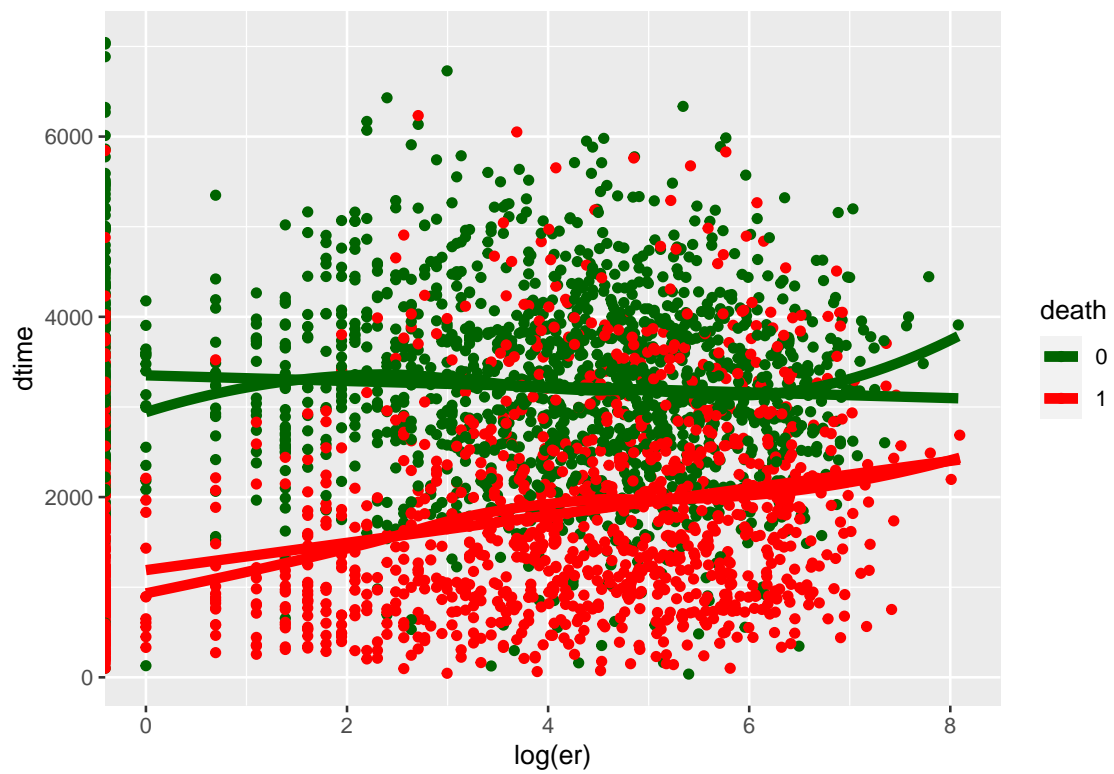


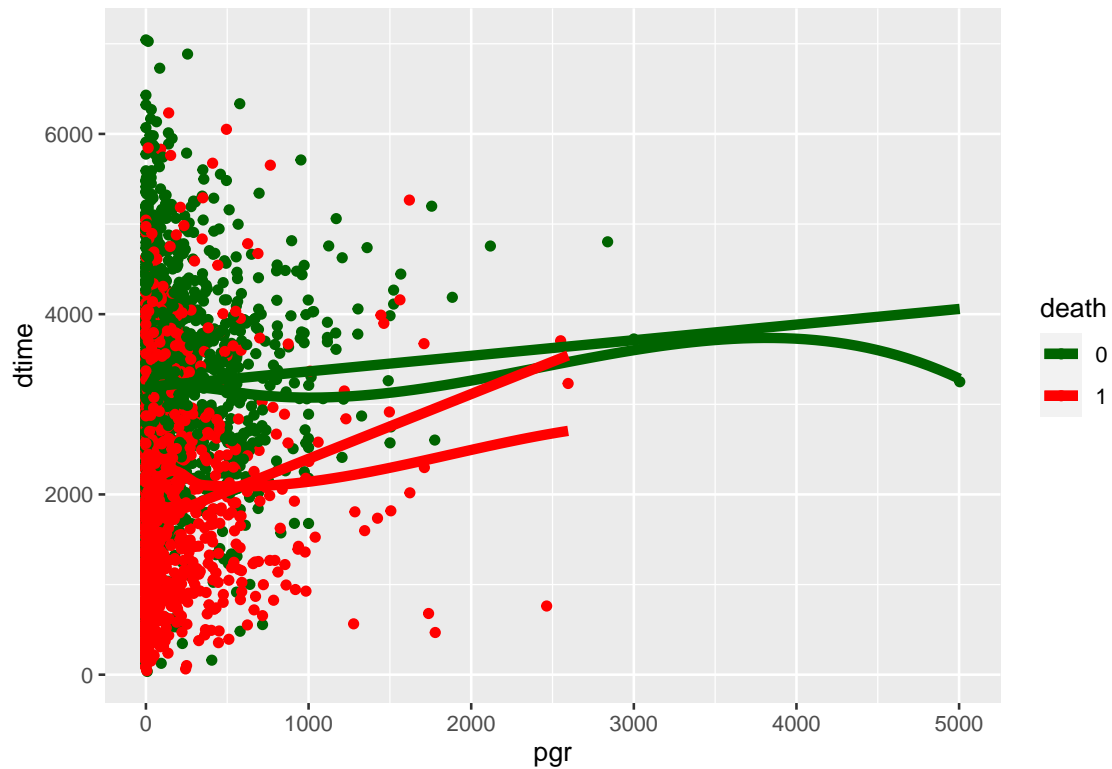
Among continuous variables, we first investigated age. We plotted a both the lowess curve and the linear model to see if they match. If so, they most likely have data with a limited skew and do not require transformation for modeling. We first evaluated age and there was a difference between the linear and lowess curves. To adjust for this, a log transform was performed and the fit generally improved for most of the data. Next, we investigated the relationship between er and relapse time. Here we see a large deviation which was corrected for again with a log transform. Finally, we evaluated pgr and it had a good fit.





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2 Further Analysis

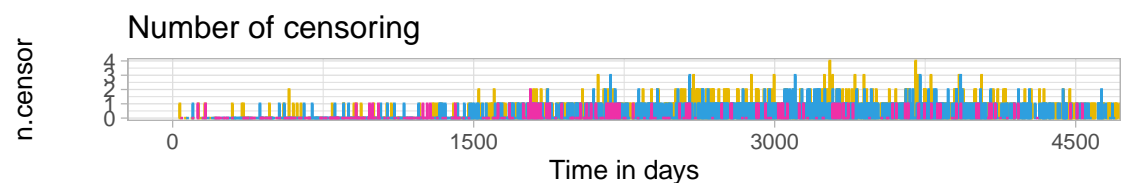
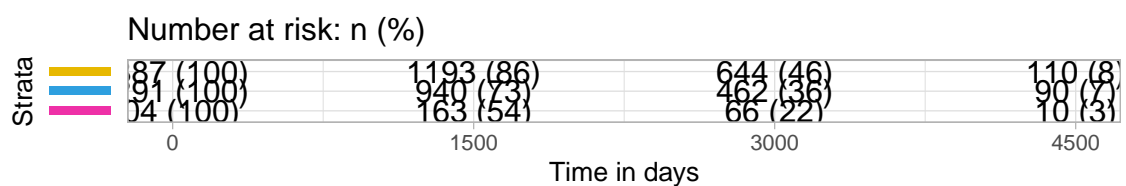
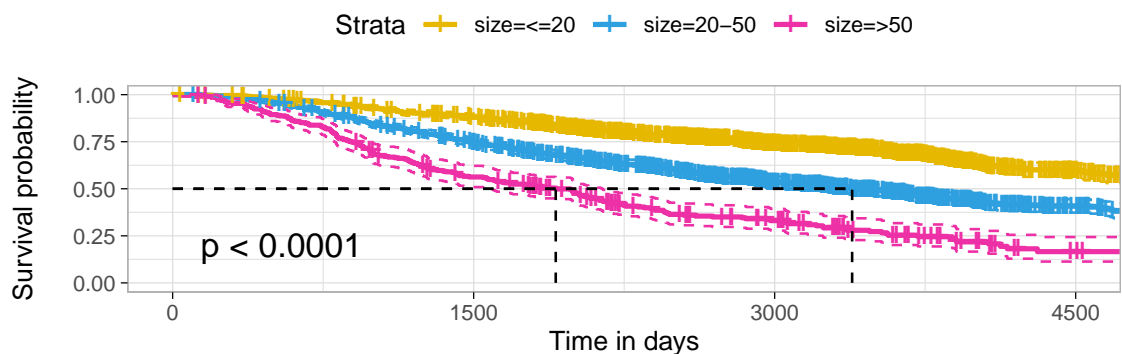
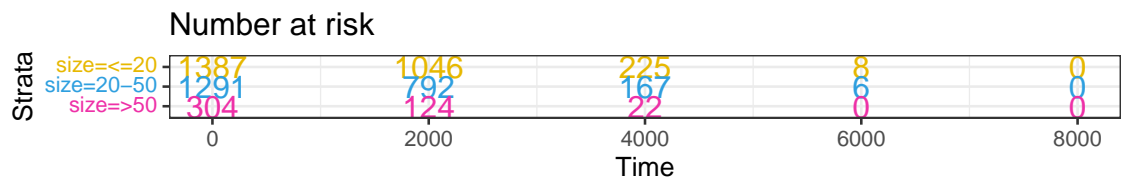
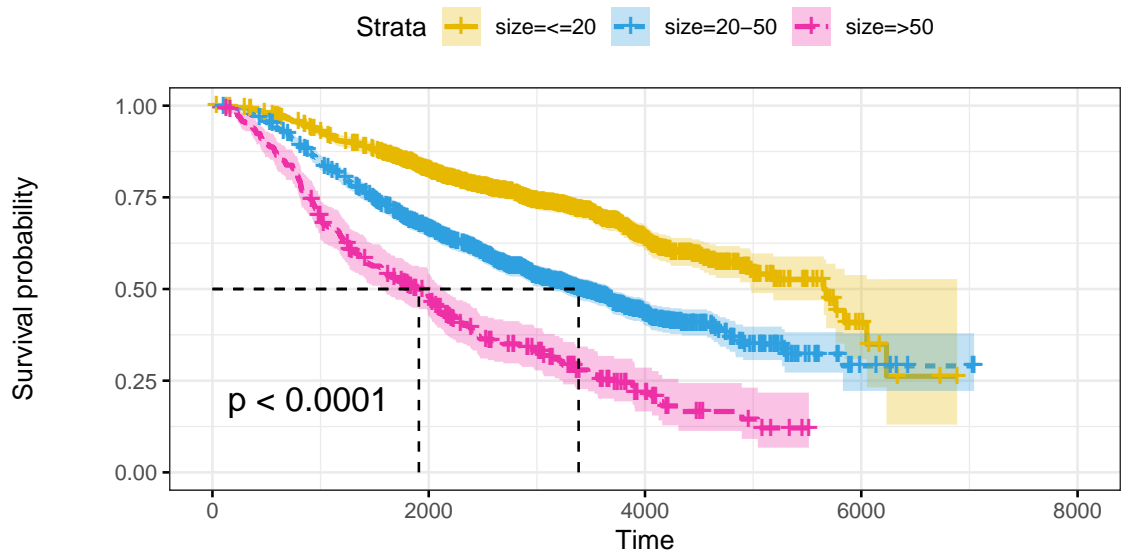
Now the focus will be on the response variable, the censoring indicator, and the categorical variable.

2.1 Survival Distrubution by Levels of ...

- (a) For each of the levels of the categorical variable, compute the survival distribution. Plot them on the same graph. What do the graphs suggest?

2.1.1 Size

##	time	n.risk	n.event	n.censor	surv	upper	lower
## 1	36	1387	0	1	1.0000000	1.0000000	1.0000000
## 2	64	1386	1	0	0.9992785	1.0000000	0.9978659
## 3	97	1385	1	0	0.9985570	1.0000000	0.9965606
## 4	101	1384	1	0	0.9978355	1.0000000	0.9953918
## 5	129	1383	0	1	0.9978355	1.0000000	0.9953918
## 6	141	1382	1	0	0.9971135	0.9999421	0.9942928



The horizontal axis represents time in days, and the vertical axis shows the probability of surviving, or the proportion of people surviving. The lines represent survival curves of the three groups. A vertical drop in the curves indicates an event. The vertical tick mark on the curves means that a patient was censored at this time.

At time zero, the survival probability is 1.0 (or 100% of the participants are alive). At time 2250, the probability of survival is approximately 0.625 for size= \geq 50, and 0.85 for size= \leq 50.

Table 2: Summary of the model.

	records	n.max	n.start	events	rmean	se(rmean)	median	0.95LCL	0.95UCL
Size (≤ 20)	1387	1387	1387	414	4721.199	119.40159	5653	4983	
Size (20–50)	1291	1291	1291	646	3807.025	95.52799	3386	3084	3690
Size (> 50)	304	304	304	212	2537.178	148.10071	1909	1566	2141

From Table 2 can be seen that the median survival for Size ≤ 20 is 5653, for Size 20-50 is 3386, and for Size > 50 is 1909. This suggests slightly worse survival for patients with tumor of larger size. However, to evaluate whether this difference is statistically significant requires a formal statistical test, a subject that is discussed in the next sections.

2.1.2 Menopause

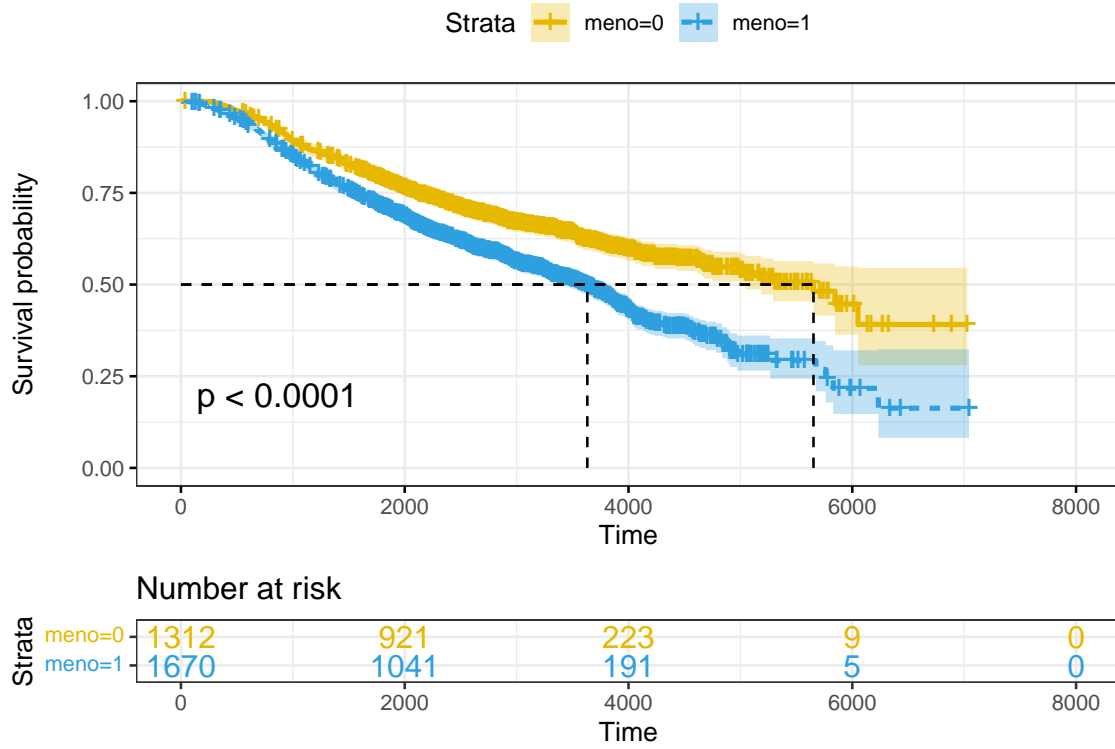


Table 3: Median survival times for each group.

	records	n.max	n.start	events	rmean	se(rmean)	median	0.95LCL	0.95UCL
Meno = 0	1312	1312	1312	468	4622.427	108.80722	5653	4983	
Meno = 1	1670	1670	1670	804	3672.887	92.46545	3632	3368	3813

The horizontal axis represents time in days, and the vertical axis shows the probability of surviving, or the proportion of people surviving. The lines represent survival curves of the three groups. A vertical drop in the curves indicates an event. The vertical tick mark on the curves means that a patient was censored at this time.

At time zero, the survival probability is 1.0 (or 100% of the participants are alive). At time 2200, the probability of survival is approximately 0.75 for premenopausal patients, and 0.825 for postmenopausal patients.

From Table 3 can be seen that the median survival is 5653 for premenopausal patients, and 3632 for postmenopausal, suggesting slightly worse survival for patients that have gone through menopause. However, to evaluate whether this difference is statistically significant requires a formal statistical test, a subject that is discussed in the next sections.

2.1.3 Hormonal Treatment

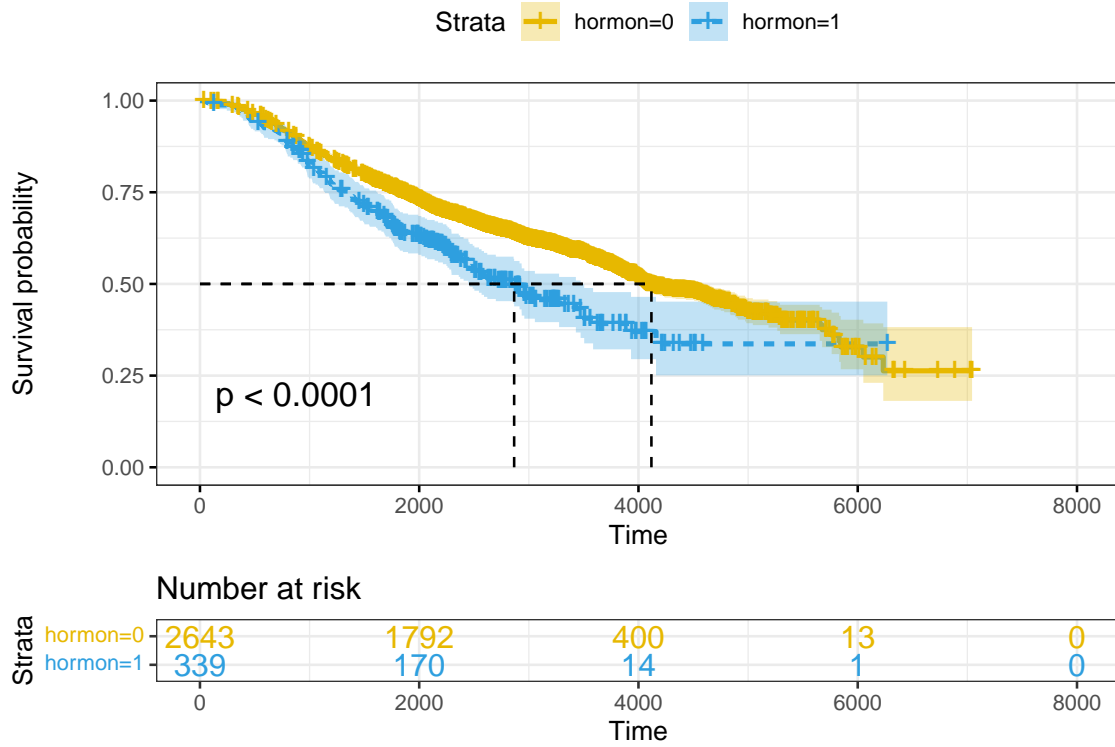


Table 4: Median survival times for each group.

	records	n.max	n.start	events	rmean	se(rmean)	median	0.95LCL	0.95UCL
hormon=0	2643	2643	2643	1113	4159.588	76.57099	4118	3988	4614
hormon=1	339	339	339	159	3659.665	203.38456	2866	2450	3472

The horizontal axis represents time in days, and the vertical axis shows the probability of surviving, or the proportion of people surviving. The lines represent survival curves of the three groups. A vertical drop in the curves indicates an event. The vertical tick mark on the curves means that a patient was censored at this time.

At time zero, the survival probability is 1.0 (or 100% of the participants are alive). At time 2200, the probability of survival is approximately 0.75 for premenopausal patients, and 0.825 for postmenopausal patients.

From Table 4 can be seen that the median survival is 4118 for patients that went through hormonal therapy, and 2866 for those who did not, suggesting slightly worse survival for patients that have gone through the hormonal therapy. However, to evaluate whether this difference is statistically significant requires a formal statistical test, a subject that is discussed in the next sections.

2.1.4 Chemotherapy

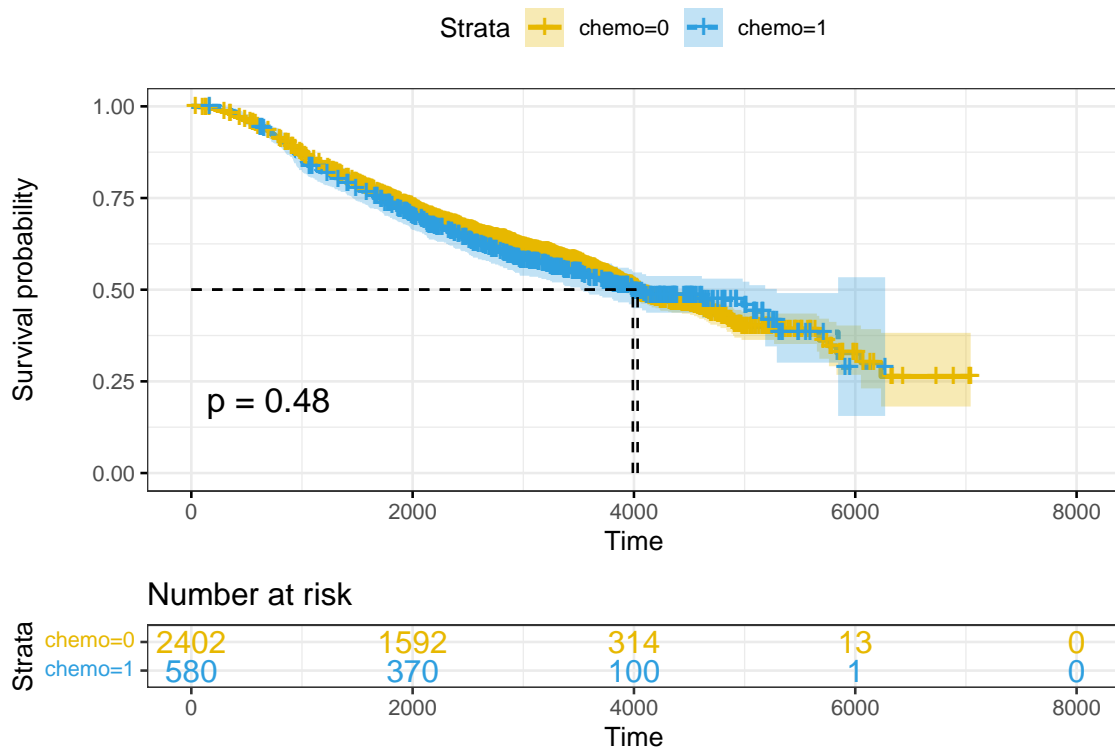


Table 5: Median survival times for each group.

	records	n.max	n.start	events	rmean	se(rmean)	median	0.95LCL	0.95UCL
chemo=0	2402	2402	2402	1014	4103.663	80.06765	4033	3885	4239
chemo=1	580	580	580	258	4080.311	162.59049	3990	3522	5291

The horizontal axis represents time in days, and the vertical axis shows the probability of surviving, or the proportion of people surviving. The lines represent survival curves of the three groups. A vertical drop in the curves indicates an event. The vertical tick mark on the curves means that a patient was censored at this time.

At time zero, the survival probability is 1.0 (or 100% of the participants are alive). At time 2200, the probability of survival is approximately 0.75 for premenopausal patients, and 0.825 for postmenopausal patients.

From Table 5 can be seen that the median survival is 4118 for patients that went through hormonal therapy, and 2866 for those who did not, suggesting slightly better survival for patients that have gone through chemotherapy. However, to evaluate whether this difference is statistically significant requires a formal statistical test, a subject that is discussed in the next sections.

2.2 Confidence Intervals and Estimators by Levels of ...

- (b) For each level obtain an appropriate estimator and confidence interval for the 3 quartiles of the survival curves. Interpret the results.

We first evaluate size and see similar results to what was observed in the EDA barcharts ???. We see that at the 25th percentile, size >50 has a very significant difference from the other two groups.

Table 6: Estimates and confidence intervals for 3 quartiles for each level of covariates *size*, *chemo*, *meno*, *hormon*.

	\hat{q}_1	5%	95%	\hat{q}_2	5%	95%	\hat{q}_3	5%	95%
Size (≤ 50)	2880	2590	3315	5653	4983	NA	NA	6051	NA
Size (20–50)	1476	1361	1623	3386	3084	3690	NA	5830	NA
Size (> 50)	890	809	999	1909	1566	2141	3714	3240	4309
Chemo = 0	1812	1677	1957	4033	3885	4239	NA	6051	NA
Chemo = 1	1699	1455	1954	3990	3522	5291	NA	5845	NA
Meno = 0	2115	1944	2371	5653	4983	NA	NA	NA	NA
Meno = 1	1571	1424	1723	3632	3368	3813	5762	5266	NA
Hormon = 0	1882	1742	1994	4118	3988	4614	NA	6051	NA
Hormon = 1	1361	1140	1618	2866	2450	3472	NA	NA	NA

With this being significantly smaller, their *dttime* is significantly smaller. This intuitively makes sense that a person who has larger cancer would have a worse life expectancy. However, this phenomenon has faded by the 50th percentile and persists through the 75th. This means that if a person survives the initial hurdle after surgery, they are not expected to be significantly different from the rest of the population.

When we investigate the effect of menopause on survival, we see that generally people that are menopausal have lower survival times. This is supported by that at the 25th, 50th, and 75th percentiles, there is no overlap in the quantiles' confidence intervals. This means that at all stages, a menopausal person would have a shorter life. However, this could be due to other covariates such as the fact that all menopausal people are women or that they tend to be older.

We next evaluate the effect of hormonal treatment. We see here that those who received hormonal treatment generally have a lower survival time. We also see that there is no overlap between the confidence intervals indicating that not only did they have a large difference, but it sustained itself throughout the entire distribution of patients.

Finally, we evaluate the effect of chemotherapy. Here we see that generally chemotherapy patients live longer. However, this only becomes significant at the 75th percentile. This means that generally patients that receive chemotherapy are not significantly affected until they live for a long time.

2.3 Test of Differences Between the Survival Curves

- (c) Conduct a single test of differences between the survival curves. Justify your choice of test.

Table 7: Stratified log-rank test for differences in menopause.

	χ^2	df	p-value
Stratified by Age	12.05	1	0.0005177
Stratified by Size	42.53	1	0.0000000
Stratified by Hormon	44.06	1	0.0000000

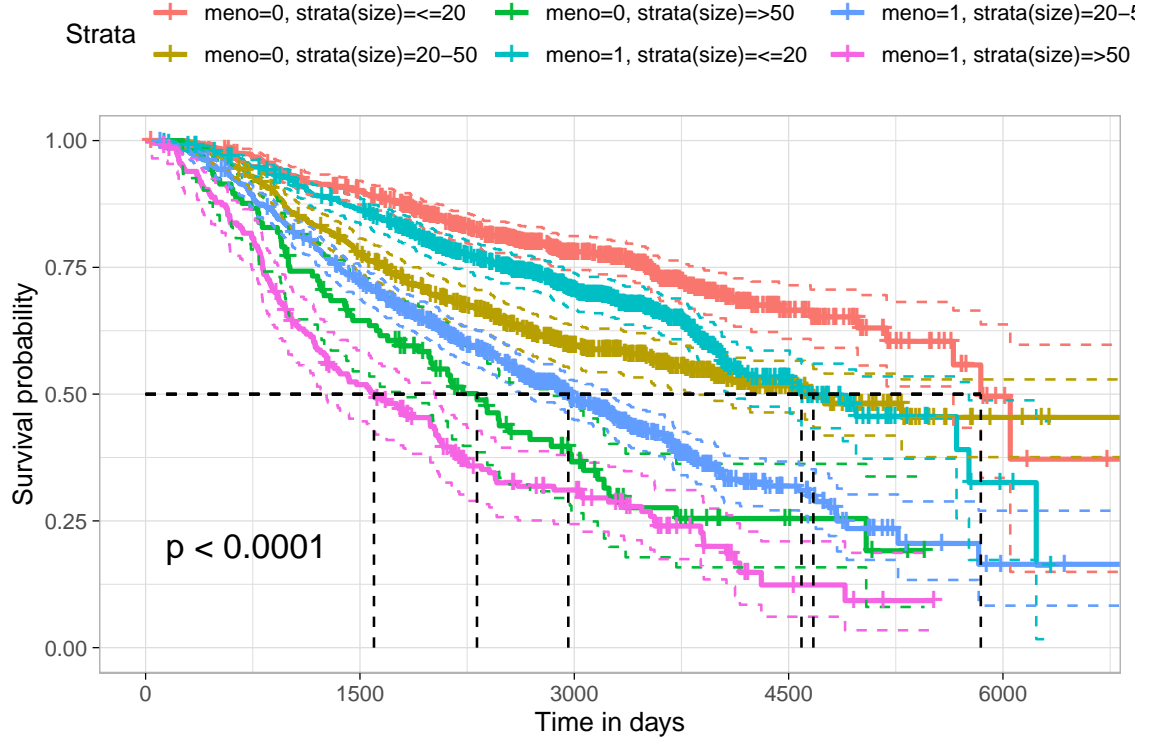


Figure 1: Independent variable meno stratified by size of the tumor.

In this setting we chose to use the log-rank test. This is because we do not want to weight any observations more heavily than others. Thus, we $W(Y(j)) = 1$. The log-rank test is the most powerful so it is the best choice. However, this test is only the most powerful when the hazard rates are proportional to each other within populations. For this reason, we stratify the model to identify its impacts and make sure no assumptions are violated.

3 Modeling

3.1 Semi Parametric PH models

For Semi Parametric Proportional Hazards Model, We proceed as follows:

The variable size is a categorical variable with 3 levels so we recode the levels 20-50 and >50 as binary variables and take the level ≤ 20 as reference level. We then estimate the full model only excluding rtime and recur variables since they are not the chosen event of interest.

For model selection with all covariates, we first fit a full model to apply backward selection procedure. By excluding the covariate with the largest p-value in each iteration, we ended up with the model “Surv(dtime,as.numeric(death))~meno+hormon+size+ age+grade+nodes+pgr” where all covariates are significant. Then we also run stepwise AIC selection in both direction and it resulted the model with size, age, grade, nodes, pgr as variables. To decide which criteria should be used to choose the model, we approach from 2 different points. From a logical point of view, the model including hormonal treatment is reasonable since it might have a direct influence on survival time. If we compare AIC of both models, we see that there is no huge difference (18543 and 18546). So we decide to go with the model selected based on p-value since it includes hormon in exchange of acceptable increase in AIC.

Table 8: Relative Risks and CI for each covariate

	HR	95% CI Lower Bound	95% CI Upper Bound
meno1	1.0555649	0.8701282	1.2805208
hormon1	0.9348205	0.7863100	1.1113801
as.factor(size)20-50	1.5573452	1.3701840	1.7700718
as.factor(size)>50	2.2862035	1.9120724	2.7335401
age	1.0135491	1.0061192	1.0210340
as.factor(grade)3	1.3705019	1.1929473	1.5744831
nodes	1.0762694	1.0662294	1.0864039
pgr	0.9996186	0.9993875	0.9998497
size>50 vs size20-50	1.4680133	1.2521226	1.7211279

Table 9: Estimators and confidence intervals for relative risks under xxxxx model.

	$exp(\hat{\beta})$	5%	95%
Chemo	1.052	0.896	1.235
Meno	1.073	0.881	1.307
Hormon	0.937	0.788	1.114
Size (20-50)	1.557	1.369	1.769
Size (>50)	2.276	1.902	2.722
Age	1.014	1.007	1.022
Grade = 3	1.371	1.193	1.575
Nodes	1.076	1.065	1.086
Pgr	1.000	0.999	1.000
Er	1.000	1.000	1.000

Estimation and CI of relative risks for every pair of levels of the categorical variables are as below: HR of chemo vs non-chemo: 1.0423 with 95%CI [0.91353 1.18914] HR of meno vs non-meno: 0.8160 with 95%CI [0.69415 0.95916] HR of hormon treatment vs no treatment: 1.7815 with 95%CI [0.69415 0.95916] HR of size20-50 vs size<=20: 0.8650 with 95%CI [0.78503 0.95308] HR of size>50 vs size<=20: 0.9678 with 95%CI [0.79736 1.17457] HR of size>50 vs size20-50: $\exp(\text{Beta}(\text{size}>50)-\text{Beta}(\text{size}20-50))=\{\exp(\text{fit_ph4coefficients}['\text{as.factor}(\text{size}) > 50'])/\exp(\text{fit_ph4coefficients}['\text{as.factor}(\text{size})20-50'])\}=1.118818$ with 95%CI [??] HR of age i vs age i-1: 1.0075 with 95%CI [1.00110 1.01387] HR of grade 3 vs grade 2: 1.0575 with 95%CI [0.95655 1.16910] HR of recur vs non-recur: 0.0131 with 95%CI [0.01043 0.01646] HR of rtime i vs rtime i-1: 0.9985 with 95%CI [0.99841 0.99854]

```
##               chisq df      p
## meno          4.33  1 0.03745
## hormon         0.73  1 0.39284
## as.factor(size) 5.22  2 0.07357
## age          13.96  1 0.00019
## as.factor(grade) 3.16  1 0.07544
## nodes         3.96  1 0.04655
## pgr          41.79  1 1.0e-10
## GLOBAL       60.91  8 3.1e-10
```

We run a statistical test based on Schoenfeld residuals for proportional hazard assumption for each covariate included in the cox fit. From the output, one can see that test is statistically

significant for some covariates like size, grade, rtime and recur and so is the global test. So we can state that proportional hazard assumption is violated since there is significant dependency between Schoenfeld residuals and time.

##next step Q: should we also check outliers and nonlinearity between log hazard and covariates??

3.2 Parametric Regression Models

	log(normal)	weibull	exponential	log(logistic)
AIC	24031.88	24120.92	24259.09	24045.35

We first evaluate parametric models based on variables already selected in the previous section. Here we are comparing log (normal), weibull, exponential, and log (logistic) distributions. This is done by using AIC and identifying if the additional parameters are worth it. What we see is that the log(normal) distribution has the best AIC at 2.4031876×10^4 . It should be noted that these AIC values can not be directly compared to the semi-parametric values. We have calculated the linear coefficient predictions to be the negative of the log(normal) coefficients.

	term	estimate	std.error
(Intercept)	(Intercept)	9.3600956671	1.420210e-01
meno1	meno1	-0.0686644484	7.914145e-02
hormon1	hormon1	0.1555160089	7.493306e-02
as.factor(size)20-50	as.factor(size)20-50	-0.3431421186	5.075681e-02
as.factor(size)>50	as.factor(size)>50	-0.6096778052	8.020362e-02
age	age	-0.0087558536	2.942332e-03
as.factor(grade)3	as.factor(grade)3	-0.2730317406	5.496964e-02
nodes	nodes	-0.0763493950	5.307671e-03
pgr	pgr	0.0004211039	8.749922e-05

	statistic	p.value	2.5 %	97.5 %
(Intercept)	65.9063998	0.000000e+00	9.0817395370	9.6384517971
meno1	-0.8676167	3.856042e-01	-0.2237788481	0.0864499514
hormon1	2.0753992	3.794955e-02	0.0086499123	0.3023821055
as.factor(size)20-50	-6.7605135	1.375035e-11	-0.4426236443	-0.2436605930
as.factor(size)>50	-7.6016247	2.924359e-14	-0.7668740091	-0.4524816012
age	-2.9758214	2.922050e-03	-0.0145227177	-0.0029889894
as.factor(grade)3	-4.9669556	6.801214e-07	-0.3807702462	-0.1652932350
nodes	-14.3847261	6.453304e-47	-0.0867522391	-0.0659465508
pgr	4.8126594	1.489350e-06	0.0002496086	0.0005925992

	linCoeffs
(Intercept)	-9.3600956671
meno1	0.0686644484
hormon1	-0.1555160089
as.factor(size)20-50	0.3431421186
as.factor(size)>50	0.6096778052
age	0.0087558536
as.factor(grade)3	0.2730317406
nodes	0.0763493950
pgr	-0.0004211039

When we look at our categorical variable, chemo, it is highly significant and the confidence interval confirms this as it does not contain 0.

4 References