

**Problem 1:** Using induction on  $i$  we can show that  $Z_i$  are mutually independent and each have exponential distribution with rate parameter of  $\lambda$ .

**Base Case:** Let's show that  $Z_1$  and  $Z_2$  are independent and are distributed exponentially.

Note that  $Z_1 = X_1 - X_0 = X_1$ . So,

$$\begin{aligned} F_{Z_1}(t) &= P(Z_1 \leq t) \\ &= P(X_1 \leq t) = P(N_t \geq 1) \\ &= P(N_t - N_0 \geq 1) \\ &= 1 - P(N_t - N_0 < 1) \\ &= 1 - \text{Pois}_{\lambda t}(0) \\ &= 1 - e^{-\lambda t} \end{aligned}$$

This is the CDF of the exponential distribution.

We know that  $Z_2 = X_2 - X_1$ . Using similar logic (with some handwaviness),

$$\begin{aligned} F_{Z_2}(t) &= P(Z_2 \leq t) \\ &= P(Z_2 \leq t) \\ &= 1 - P(Z_2 > t) \quad (\text{here comes the handwaviness}) \\ &= 1 - P(X_2 > X_1 + t \mid X_1 = s) \\ &= 1 - P(N_{t+s} - N_s = 0) \\ &= 1 - P(N_t = 0) \quad (\text{by poisson}) \\ &= 1 - e^{-\lambda t} \end{aligned}$$

Thus we get that  $Z_2$  is similarly distributed exponentially.

Let's show that  $Z_1$  and  $Z_2$  are independent, so that  $P(Z_1 \leq t \cap Z_2 \leq t) = P(Z_1 \leq t)P(Z_2 \leq t)$ .

We know that by the independent increments of  $N$ ,

$$P(Z_1 \leq t_1 \cap Z_2 \leq t_2) = P(N_{t_1}) \cdot P(N_{t_2}) = P(Z_1 \leq t_1)P(Z_2 \leq t_2)$$

**Induction:** Assume that all  $Z_1, \dots, Z_k$  are mutually independent and distributed exponential with  $\lambda$ .

We want to show that  $Z_{k+1}$  is independent from  $Z_1, \dots, Z_k$  and exponentially distributed.

Using the same idea as before,

$$P(Z_{k+1} \leq t) = 1 - P(Z_{k+1} > t) = 1 - P(N_{X_{k+t}} - N_{X_k} = 0) = 1 - e^{-\lambda t}$$

Using the independent increments of the poisson process,  $Z_{k+1}$  must be independent of the previous time slices.

**Problem 2:** We want to show that  $P(Z > z + w | Z > z) = P(Z > w)$ .

Using the definition of conditional probability

$$\begin{aligned}
 P(Z > z + w | Z > z) &= \frac{P(Z > z + w \cap Z > z)}{P(Z > z)} \\
 &= \frac{P(Z > z + w)}{P(Z > z)} \\
 &= \frac{1 - (1 - e^{-\lambda(z+w)})}{1 - (1 - e^{-\lambda(z)})} \\
 &= \frac{e^{-\lambda(z+w)}}{e^{-\lambda(z)}} \\
 &= e^{-\lambda w} \\
 &= P(Z > w)
 \end{aligned}$$

**Problem 3:** Let  $X$  be a continuous random variable that satisfies the memoryless property.

Then we get the following expression: for all  $z, w \in \mathbb{R}_{\geq 0}$

$$\begin{aligned}
 P(X > w) &= \frac{P(X > z + w)}{P(X > z)} \\
 P(X > z + w) &= P(X > z)P(X > w)
 \end{aligned}$$

Notice that  $0 \leq P \leq 1$ . Define  $G(y) = 1 - f_X(y) = P(X > y)$ .  $G$  must also be continuous and bounded by 0 and 1.

So,  $G(z + w) = G(z)G(w)$ . Taking the log of this, we get  $\ln(G(z + w)) = \ln(G(z)) + \ln(G(w))$ . Notice that  $\ln \circ G$  is a continuous Cauchy function. By a theorem discussed in discussion,  $\ln G(y)$  must be linear. Moreover, the linear coefficient must be negative because  $G$  is bounded by 0 and 1.

Hence we get that  $\ln \circ G(y) = -\lambda y$  for some  $\lambda > 0$ .

Thus we get that  $G(y) = e^{-\lambda y}$ . So  $f_X(y) = 1 - e^{-\lambda y}$ . This is precisely the cdf of an exponential distribution. Hence,  $X$  must be exponentially distributed.

**Problem 4:** We want to show that  $H_t, T_t$  are Poisson Processes with rate  $p\lambda$  and  $(1 - p)\lambda$  respectively.

Obviously,  $H_0 = 0$  and  $T_0 = 0$ , since  $N_0 = 0$  (the underlying poisson process) and we haven't had an event where we could flip a coin.

$H_t - H_s$  must be non-decreasing for some  $t > s$  because we cannot remove a decrement of heads/tails. It is always the sum of the bernoulli variable whose output is either 0 or 1.

Now, we want to show that  $H_t$  has independent increments. Let  $0 \leq s_0 < s_1 < t$ .

$H_t$  is constructed by flipping a coin at every arrival of  $N_t$ . Let  $X$  be a random variable that is 1 when a heads is flipped and 0 when a tails is flipped, at

the rates  $p, (1 - p)$ . If we split the intervals  $[s_0, s_1), [s_1, t]$  into smaller a set of intervals, let's say like a partition  $\{s_0, s_0 + 1, s_0 + 2, \dots, s_1\}$  and  $\{s_1, s_1 + 1, s_1 + 2, \dots, t\}$

We can define  $X_i$  for some time  $i$  to be the same as  $X$  but conditioned on the whether or not an arrival happened during the time period. In other words if  $N_i - N_{i-1} = 1$  and  $X = 1$  then  $X_i = 1$ . Since,  $X$  is independent of previous flips and the poisson process, each  $X_i$  are mutually independent.

Notice that  $H_t$  are constructed by the sum of these mutually independent events. Hence, by the same reasoning as homework 2,  $H_t - H_{s_1}$  and  $H_{s_1} - H_{s_0}$  are mutually independent, and there are independent increments.

Using the ideas of the previous proof, since the coin flip  $X$  and the poisson process are independent and we are summing the  $X_i$  it should follow that the rate of  $H_t$  would be  $p\lambda$ . Similarly can be said about  $T_i$ .