

Homework 1

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1. In a simple symmetric random walk, show that S_k and $S_n, k \neq n$, are dependent random variables.

Proof. WLOG, $n > k$. Consider the covariance between S_k and S_n :

$$\begin{aligned} \text{Cov}(S_k, S_n) &= \text{Cov} \left(\sum_{i=1}^k X_i, \sum_{j=1}^n X_j \right) \\ &= \sum_{i=1}^k \sum_{j=1}^n \text{Cov}(X_i, X_j) \end{aligned}$$

Note that $\text{Cov}(X_i, X_j) = 0$ for $i \neq j$, otherwise it is

$$\text{Cov}(X_i, X_i) = \text{Var}(X_i) = E[X_i^2] - (E[X_i])^2 = 1 - 0^2 = 1$$

Thus, the value of the double summation is exactly $k \neq 0$, so S_k and S_n are dependent, as desired. \square

2. Consider a gambler who on each gamble is equally likely to either win or lose 1 unit. Starting with i units, show that the expected time until the gambler's fortune reaches either 0 or k is $i(k-i), i = 0, 1, 2, \dots, k$.

Proof. Let M_i denote this expected time, and let $X_i = \pm 1$ be the result of the i -th gamble, with probability $1/2$ of either result. Then we have the recurrence

$$M_i = 1 + \frac{1}{2}M_{i+1} + \frac{1}{2}M_{i-1}$$

since we will have taken a step already. Rearranging, we have

$$\begin{aligned} -2 &= M_{i+1} - 2M_i + M_{i-1} \\ -2 &= M_{i+2} - 2M_{i+1} + M_i \\ \implies 0 &= M_{i+2} - 3M_{i+1} + 3M_i - M_{i-1} \end{aligned}$$

This is a homogeneous linear recurrence with characteristic equation

$$x^3 - 3x^2 + 3x - 3 = (x-1)^3$$

so the general form for M_i is given by

$$M_i = ai^2 + bi + c, \quad a, b, c \in \mathbb{R}$$

We know that $M_0 = 0$ since we are already at 0, so

$$M_0 = c = 0 \implies M_i = ai^2 + bi$$

is the general form. Substituting into the recurrence relation, we have

$$\begin{aligned} -2 &= [a(i+1)^2 + b(i+1)] - 2(ai^2 + bi) + [a(i-1)^2 + b(i-1)] = 2a \\ \implies -1 &= a \end{aligned}$$

Now, we also know that $M_k = 0$ since we are already at k , so

$$M_k = -k^2 + bk = 0 \implies b = k$$

Thus, the closed form is given by $M_i = -i^2 + ki = i(k-i)$, as desired. \square

3. A particle moves on the set of integers 0 through n so that at each step it is equally likely to move to any of its neighbors. If the particle starts at 0, show that the expected number of steps it takes to reach n is n^2 .

Proof. Let T_i represent the number of steps it takes to get from position $i - 1$ to i . On the first step, we can either get to i with probability $1/2$, or get to $i - 2$ with probability $1/2$. If we are at $i - 2$, we have already used a step, and we will have to get back to $i - 1$, with an expected time of $E[T_{i-1}]$ steps, then get back to i again with an expected time of $E[T_i]$ steps. Combining everything, we have

$$\begin{aligned} E[T_i] &= \frac{1}{2}(1) + \frac{1}{2}(1 + E[T_{i-1}] + E[T_i]) \\ \implies E[T_i] &= 2 + E[T_{i-1}] \end{aligned}$$

If we are at position 0, we are guaranteed to get to position 1, so $E[T_1] = 1$. Then $E[T_2] = 3$, and in general, $E[T_k] = 2k - 1$. The expected number of steps it takes to reach n from 0 is the sum

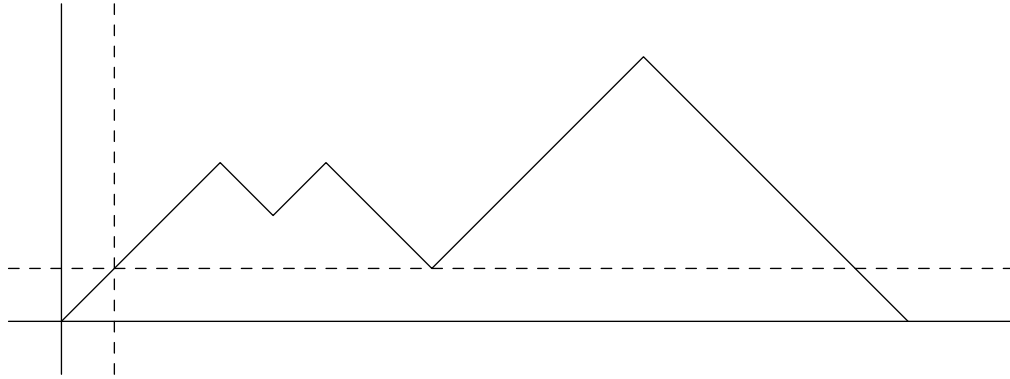
$$\sum_{i=1}^n E[T_i] = \sum_{i=1}^n (2i - 1) = n^2$$

since we are summing the first n odd integers. □

4. For a simple symmetric random walk, give a geometric proof that

$$P[S_1 \geq 0, S_2 \geq 0, \dots, S_{2n-1} \geq 0, S_{2n} = 0] = 2f_{2n+2}$$

Proof. Consider a path that only touches the origin at 0 and $2n + 2$. The probability of this happening is f_{2n+2} , and it can either be strictly positive or strictly negative. The positive case is shown:



Since it is strictly positive, we may restrict the portion of the path we are considering to the points S_1, \dots, S_{2n+1} , which when translated becomes a path such that $S_0 = 0, S_1 \geq 0, S_2 \geq 0, \dots, S_{2n} = 0$. Note that the case where the path is negative is nearly identical, we just have to reflect about the t axis in addition to translating. Thus, for every 2 paths whose first return is $2n + 2$, there is 1 path that is nonnegative with $S_{2n} = 0$, so

$$P[S_1 \geq 0, S_2 \geq 0, \dots, S_{2n-1} \geq 0, S_{2n} = 0] = 2f_{2n+2}$$

as desired. □

5. If X is a nonnegative integer-value random variable, show that

$$E[X] = \sum_{n=1}^{\infty} P[X \geq n] = \sum_{n=0}^{\infty} P[X > n]$$

Proof. Let I_n be an indicator variable, where

$$I_n = \begin{cases} 1, & X \geq n \\ 0, & X < n \end{cases}$$

Then we can express X as

$$X = \sum_{j=1}^{\infty} I_j$$

since X will be incremented for every $j \leq n$, for a total of n times, and never again after that. Thus,

$$E[X] = E\left[\sum_{j=1}^{\infty} I_j\right] = \sum_{j=1}^{\infty} E[I_j] = \sum_{j=1}^{\infty} P[X \geq j]$$

We also have

$$P[X \geq n] = P[X > n-1]$$

since X takes on integer values, so the rightmost sum is the same. □