

## Homework 5

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1. Let  $X_1, X_2, X_3, \dots$  be iid random variables. Let  $M(t) = E[e^{tX_1}]$  be the MGF of  $X_1$  (and thus of each  $X_i$ ). Fix  $t$  and assume that  $M(t) < \infty$ . Define the partial sum process by letting  $S_0 = 0$  and for  $n > 0$ ,

$$S_n = X_1 + \dots + X_n.$$

Let

$$Z_n = \frac{e^{tS_n}}{M(t)^n}$$

Show that  $\{Z_n\}_{n=0}^\infty$  is a martingale with respect to  $\{X_n\}_{n=0}^\infty$ .

2. Consider a Markov chain  $\{X_n, n \geq 0\}$  with state space consisting of  $N + 1$  states which are real numbers  $x_0 < x_1 < x_2 < \dots < x_N$ , and with transition matrix  $P(i, j) = P[X_{n+1} = x_j \mid X_n = x_i]$  for  $0 \leq i, j \leq N$ . Suppose that  $\{X_n, n \geq 0\}$  is also a martingale. Show that the states  $x_0$  and  $x_N$  are absorbing states.
3. Calculate the PGF for a random variable  $X$  which has
- (a) a Geometric( $\frac{1}{2}$ ) distribution.
  - (b) a Poisson( $\lambda$ ) distribution.
4. Let  $\{X_1, X_2, X_3, \dots\}$  be a sequence of iid random variables with mean  $\mu$  and variance  $\sigma^2$ . Let  $S_n = X_1 + X_2 + \dots + X_n$  for each integer  $n \geq 1$ . Let  $N$  be a positive integer random variable which is independent of the  $\{X_i\}_{i \geq 1}$ , and has mean  $\nu$  and variance  $\tau^2$ . Calculate the variance of  $S_N$ .

*Solution.* We have

$$\text{Var}(S_N) = E[S_N^2] - (E[S_N])^2$$

Using the law of total probability, we have

$$\begin{aligned} E[S_N] &= E[E[S_N \mid N]] = \sum_{n=1}^{\infty} E[S_N \mid N = n] P[N = n] \\ &= \sum_{n=1}^{\infty} E\left[\sum_{i=1}^N X_i \mid N = n\right] P[N = n] = \sum_{n=1}^{\infty} E\left[\sum_{i=1}^n X_i \mid N = n\right] P[N = n] \\ &= \sum_{n=1}^{\infty} \left(\sum_{i=1}^n E[X_i]\right) P[N = n] = \sum_{n=1}^{\infty} (n\mu) P[N = n] \\ &= \mu \sum_{n=1}^{\infty} n P[N = n] = \mu\nu \end{aligned}$$

and

$$\begin{aligned}
E[S_N^2] &= E[E[S_N^2 \mid N]] = \sum_{n=1}^{\infty} E[S_N^2 \mid N=n]P[N=n] \\
&= \sum_{n=1}^{\infty} E\left[\left(\sum_{i=1}^N X_i\right)^2 \mid N=n\right]P[N=n] = \sum_{n=1}^{\infty} E\left[\left(\sum_{i=1}^n X_i\right)^2 \mid N=n\right]P[N=n] \\
&= \sum_{n=1}^{\infty} E\left[\left(\sum_{i=1}^n X_i\right)^2\right]P[N=n] \\
&= \sum_{n=1}^{\infty} E\left[\left(\sum_{i=1}^n X_i^2\right) + \left(\sum_{j \neq k} X_j X_k\right)\right]P[N=n] \\
&= \sum_{n=1}^{\infty} \left(\sum_{i=1}^n E[X_i^2] + \sum_{j \neq k} E[X_j X_k]\right)P[N=n] \\
&= \sum_{n=1}^{\infty} \left[\sum_{i=1}^n (E[X_i^2] - (E[X_i])^2 + (E[X_i])^2) + \sum_{j \neq k} E[X_j]E[X_k]\right]P[N=n] \\
&= \sum_{n=1}^{\infty} \left[\sum_{i=1}^n (\sigma^2 + \mu^2) + \sum_{j \neq k} \mu^2\right]P[N=n] \\
&= \sum_{n=1}^{\infty} [n\sigma^2 + n\mu^2 + (n^2 - n)\mu^2]P[N=n] \\
&= \sum_{n=1}^{\infty} (n^2\mu^2 + n\sigma^2)P[N=n] = \mu^2 \sum_{n=1}^{\infty} n^2 P[N=n] + \sigma^2 \sum_{n=1}^{\infty} n P[N=n] \\
&= \mu^2 E[N^2] + \sigma^2 \nu = \mu^2 (E[N^2] - (E[N])^2 + (E[N])^2) + \sigma^2 \nu \\
&= \mu^2 (\tau^2 + \nu^2) + \sigma^2 \nu = \mu^2 \tau^2 + \mu^2 \nu^2 + \sigma^2 \nu
\end{aligned}$$

Combining these two, we have

$$\text{Var}(S_N) = (\mu^2 \tau^2 + \mu^2 \nu^2 + \sigma^2 \nu) - (\mu \nu)^2 = \mu^2 \tau^2 + \sigma^2 \nu$$

□

5. Consider a branching process with offspring distribution given by the frequency function  $f$ , where  $f(2) = a$ ,  $f(1) = b$ , and  $f(0) = c$ , with  $a + b + c = 1$ . Assume that the probability of extinction is  $d$ ,  $0 < d < 1$ . Express  $d$  in terms of  $a, b, c$ .

*Solution.* The generating function for the offspring distribution is

$$G(s) = as^2 + bs + c$$

and the extinction probability satisfies  $d = G(d)$ , so we have

$$\begin{aligned}
d &= G(d) = ad^2 + bd + c \\
0 &= ad^2 + (b-1)d + c
\end{aligned}$$

and solving for  $d$  we have

$$d = \frac{-(b-1) \pm \sqrt{(b-1)^2 - 4ac}}{2a}$$

Since  $0 < d < 1$  but 1 is a root of the quadratic, we must have the greater root be 1, so thus

$$d = \frac{1 - b - \sqrt{(b-1)^2 - 4ac}}{2a}$$

□

6. Verify that if  $\{Z_n\}$  is a branching process, then  $\left\{\frac{Z_n}{\mu^n}\right\}$  is a martingale, where  $\mu$  denotes the mean of the offspring distribution.

*Proof.* We have

$$E\left[\left|\frac{Z_n}{\mu^n}\right|\right] = \frac{1}{\mu^n} E[Z_n] = \frac{1}{\mu^n} (\mu^n E[Z_0]) = E[Z_0] < \infty$$

and

$$\begin{aligned} E\left[\frac{Z_{n+1}}{\mu^{n+1}} \mid \frac{Z_0}{\mu^0}, \frac{Z_1}{\mu^1}, \dots, \frac{Z_n}{\mu^n}\right] &= \frac{1}{\mu^{n+1}} E[Z_{n+1} \mid Z_0, Z_1, \dots, Z_n] \\ &= \frac{1}{\mu^{n+1}} E[Z_{n+1} \mid Z_n] \\ &= \frac{1}{\mu^{n+1}} (\mu Z_n) = \frac{Z_n}{\mu^n} \end{aligned}$$

so this is indeed a martingale. □

7. A particle moves according to a Markov chain on  $\{1, 2, \dots, c+d\}$  where  $c$  and  $d$  are positive integers. Starting from any one of the first  $c$  states, the particle jumps in one transition to a state chosen uniformly from the last  $d$  states. Starting from any of the last  $d$  states, the particle jumps in one transition to a state chosen uniformly from the first  $c$  states.

- (a) Show that the chain is irreducible.

*Proof.* Let  $C$  and  $D$  are sets of the first  $c$  and last  $d$  states, respectively. Then if  $i \in C$  and  $j \in D$ , then  $i$  and  $j$  communicate because they can directly transition between themselves. If  $i, j \in C$ , then if  $n \in D$ , we can have  $i \rightarrow n \rightarrow j$ , so  $i$  and  $j$  communicate. By a similar argument, if  $i, j \in D$ , then  $i$  and  $j$  communicate. Thus, all states communicate, so the chain is irreducible. □

- (b) Find the invariant distribution.

*Solution.* The chain is periodic, half the time we are in  $C$  and half the time we are in  $D$ . Since the transition probabilities of the

□

$i \rightarrow j$