

Exercise 1

There are two bushes. The length (in cm) of the leaves of bush A is a RV whose PDF is:

$$f(x) = \begin{cases} -\frac{k}{256} \cdot x^2 + \frac{k}{32} \cdot x & 0 \leq x \leq 8 \\ 0 & \text{otherwise} \end{cases}$$

The length of the leaves of bush B is a RV whose PDF is:

$$g(x) = \begin{cases} \frac{1}{10} & 0 \leq x \leq 10 \\ 0 & \text{otherwise} \end{cases}$$

- 1) Find k
- 2) Compute the mean and variance of $f(\cdot)$ and $g(\cdot)$

Assume now that we pluck a leaf from one of the bushes at random.

- 3) What is the probability that it is from bush B, given that it is more than 4cm long?

Assume now that we take a sample of 10 leaves from another bush (which is neither A nor B). The length of the leaves is: 5.4, 6.3, 5.8, 7.2, 4.9, 9.2, 7.0, 7.3, 6.9, 10.4

- 4) Compute the 95% confidence interval for the sample mean
- 5) Assume that every leaf is
 - a. 5cm longer than before.
 - b. Double as long as before

What about the sample mean and the confidence interval?

Exercise 2

Consider a system having K gates, through which job requests may arrive. Through each gate, job requests arrive exponentially at a rate λ . Arrivals processes at different gates are independent. When there are $j < K$ requests in the system, $j+1$ gates are open, whereas the others are closed. If there are K or more requests, all the K gates are open. The system has K identical servers, with a serving rate equal to μ .

- 1) Model the above system as a queueing system and draw the transition-rate diagram.
- 2) Compute the steady-state probabilities and state the stability condition
- 3) Compute the state that a random observe is more likely to observe and the mean number of jobs in the system
- 4) Compute the throughput of the system. Compare it to an M/M/1's and discuss the result.
- 5) Compute the probabilities observed by an arriving job request
- 6) Compute the mean response time, the mean waiting time and the mean number of queued jobs.

Exercise 1 - Solution

1) k can be computed based on the normalization condition:

$$\begin{aligned} \int_0^8 \left(-\frac{k}{256} \cdot x^2 + \frac{k}{32} \cdot x \right) dx &= \\ \frac{k}{32} \cdot \int_0^8 \left(-\frac{1}{8} \cdot x^2 + x \right) dx &= \\ \frac{k}{64} \cdot \left[-\frac{1}{12} \cdot x^3 + x^2 \right]_0^8 &= \\ \frac{k}{64} \cdot \left[-\frac{512}{12} + 64 \right] &= \\ \frac{k}{64} \cdot \left[\frac{-128 + 192}{3} \right] &= \\ \frac{k}{3} \end{aligned}$$

Since, by normalization, the integral must be equal to 1, it is $k = 3$

2) $g(\cdot)$ is a uniform RV, hence its mean is $\mu_g = \frac{1}{b-a} = 5$, and its variance is $\sigma^2 = \frac{(b-a)^2}{12} = \frac{25}{3}$.

On the other hand, the mean and variance of $f(\cdot)$ are computed by solving these integrals:

$$\begin{aligned} \mu_f &= \int_0^8 x \cdot 3 \cdot \left(-\frac{1}{256} \cdot x^2 + \frac{1}{32} \cdot x \right) dx = \\ &= \frac{3}{32} \cdot \int_0^8 \left(-\frac{1}{8} \cdot x^3 + x^2 \right) dx \\ &= \frac{3}{32} \cdot \left[-\frac{1}{32} \cdot x^4 + \frac{1}{3} \cdot x^3 \right]_0^8 \\ &= \frac{\cancel{3}}{32} \cdot \left[\frac{-3 \cdot 4096 + 32 \cdot 512}{\cancel{3} \cdot 32} \right] \\ &= \frac{1}{32} \cdot \frac{4096}{32} = 4 \end{aligned}$$

$$\begin{aligned}
\overline{X_f^2} &= \int_0^8 x^2 \cdot 3 \cdot \left(-\frac{1}{256} \cdot x^2 + \frac{1}{32} \cdot x \right) dx = \\
&= \frac{3}{32} \cdot \int_0^8 \left(-\frac{1}{8} \cdot x^4 + x^3 \right) dx \\
&= \frac{3}{32} \cdot \left[-\frac{1}{40} \cdot x^5 + \frac{1}{4} \cdot x^4 \right]_0^8 \\
&= \frac{3}{128} \cdot \left[\frac{-32768 + 40960}{10} \right] \\
&= \frac{3}{128} \cdot \frac{8192}{10} = 3 \cdot \frac{64}{10} = \frac{96}{5} \\
\text{Hence, } \sigma_f^2 &= \frac{96}{5} - (4)^2 = \frac{96 - 16 \cdot 5}{5} = \frac{16}{5}
\end{aligned}$$

3) Call L the event “the leaf is more than 4 cm”. By Bayes’ Theorem, we have:

$$P(B|L) = \frac{P(L|B) \cdot P(B)}{P(L|B) \cdot P(B) + P(L|A) \cdot P(A)}.$$

However, $P(B) = P(A) = 1/2$, since the choice is “at random”. Furthermore, $P(L|B) = (10 - 4)/10 = 0.6$, and:

$$\begin{aligned}
P(L|A) &= \frac{3}{32} \cdot \int_4^8 \left(-\frac{1}{8} \cdot x^2 + x \right) dx \\
&= \frac{3}{64} \cdot \left[-\frac{1}{12} \cdot x^3 + x^2 \right]_4^8 \\
&= \frac{3}{64} \cdot \left[\left(-\frac{512}{12} + 64 \right) - \left(-\frac{64}{12} + 16 \right) \right] \\
&= \frac{3}{64} \cdot \left[48 - \frac{112}{3} \right] \\
&= \frac{3}{64} \cdot \frac{32}{3} = \frac{1}{2}
\end{aligned}$$

Hence:

$$P(B|L) = \frac{P(L|B)}{P(L|B) + P(L|A)} = \frac{0.6}{0.6 + 0.5} = 0.545$$

4) The sample mean is

$$\bar{X} = \frac{1}{n} \cdot \sum_i x_i = 7.04.$$

The sample variance is:

$$S^2 = \frac{1}{n-1} \cdot \sum_i (x_i - \bar{X})^2 = \frac{25.424}{9} = 2.824.$$

From the tabulated Student-s t function, I need $t_{\alpha/2, n-1} = t_{0.025, 9} = 2.262$. The semi-width of the confidence interval centered around \bar{X} is therefore:

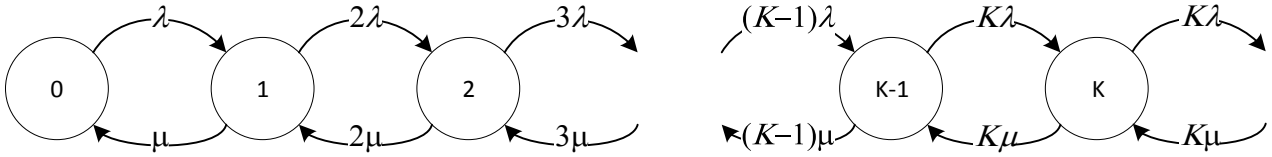
$$w = \frac{S}{\sqrt{n}} \cdot t_{\alpha/2, n-1} = \sqrt{\frac{2.824}{9}} \cdot 2.262 = 1.27.$$

The confidence interval is thus $I = [5.77; 8.31]$

5) if each leaf is 5 cm longer, the mean will be 5cm larger and the variance will stay the same, hence the confidence interval will have the same width, i.e. $I_a = I + 5 = [10.77; 13.31]$. If, on the other hand, the length is doubled, the mean will be doubled as well, and the standard deviation will be double as much. Hence, we will have $I_b = 2I = [11.54; 16.62]$.

Exercise 2 – solution

The TR diagram is the following:



From the above (transitions are nearest-neighbor) one straightforwardly obtains $p_j = \left(\frac{\lambda}{\mu}\right)^j \cdot p_0$. Call $\rho = \lambda/\mu$, then the SS probabilities and the stability condition are the same as an M/M/1's, i.e. $p_j = (1 - \rho) \cdot \rho^j$, as long as $\rho < 1$.

The state that a random observer is more likely to observe is therefore state 0. The mean number of jobs in the system is given by the Kleinrock function $E[N] = \frac{\rho}{1-\rho}$.

The throughput is

$$\begin{aligned}
 \gamma &= \sum_{j=1}^{+\infty} j \cdot \mu_j = \sum_{j=1}^K j \cdot \mu \cdot (1 - \rho) \cdot \rho^j + \sum_{j=K+1}^{+\infty} K \cdot \mu \cdot (1 - \rho) \cdot \rho^j \\
 &= \mu \cdot (1 - \rho) \sum_{j=1}^K j \cdot \rho^j + K \cdot \mu \cdot (1 - \rho) \sum_{j=K+1}^{+\infty} \rho^j \\
 &= \mu \cdot (1 - \rho) \rho \frac{1 - (K+1)\rho^K + K\rho^{K+1}}{(1 - \rho)^2} + K \cdot \mu \cdot (1 - \rho) \frac{\rho^{K+1}}{1 - \rho} \\
 &= \lambda \cdot \frac{1 - (K+1)\rho^K + K\rho^{K+1}}{1 - \rho} + K \cdot \mu \cdot \rho^{K+1} \\
 &= \lambda \cdot \left[\frac{1 - \rho^K - K\rho^K(1 - \rho)}{1 - \rho} + K \cdot \rho^K \right] \\
 &= \lambda \cdot \frac{1 - \rho^K}{1 - \rho}
 \end{aligned}$$

When $K=1$, this system is an M/M/1, and from the above expression we have $\gamma = \lambda$. When $K>1$, the throughput is *larger* than an M/M/1's, since the average arrival rate is larger than λ . Note that we always have $\gamma = \bar{\lambda}$.

The probabilities observed by an arriving job are $r_j = \frac{\lambda_j}{\lambda} \cdot p_j$. Therefore we have:

$$r_j = \begin{cases} (j+1) \cdot \frac{1-\rho}{1-\rho^K} (1-\rho) \cdot \rho^j & j < K \\ K \cdot \frac{1-\rho}{1-\rho^K} (1-\rho) \cdot \rho^j & j \geq K \end{cases}$$

The mean response time, by Little's Law, is:

$$E[R] = \frac{E[N]}{\gamma} = \frac{\rho}{1-\rho} \cdot \frac{1}{\lambda} \cdot \frac{1-\rho}{1-\rho^K} = \frac{1}{\mu} \cdot \frac{1}{1-\rho^K}$$

Moreover, we get:

$$E[W] = E[R] - E[t_s] = \frac{1}{\mu} \cdot \frac{\rho^K}{1-\rho^K}$$

And:

$$E[Nq] = \gamma \cdot E[W] = \frac{1}{\mu} \cdot \frac{\rho^K}{1-\rho^K} \cdot \lambda \cdot \frac{1-\rho^K}{1-\rho} = \frac{\rho^{K+1}}{1-\rho}$$