

# CSYS 300 Assignment 5

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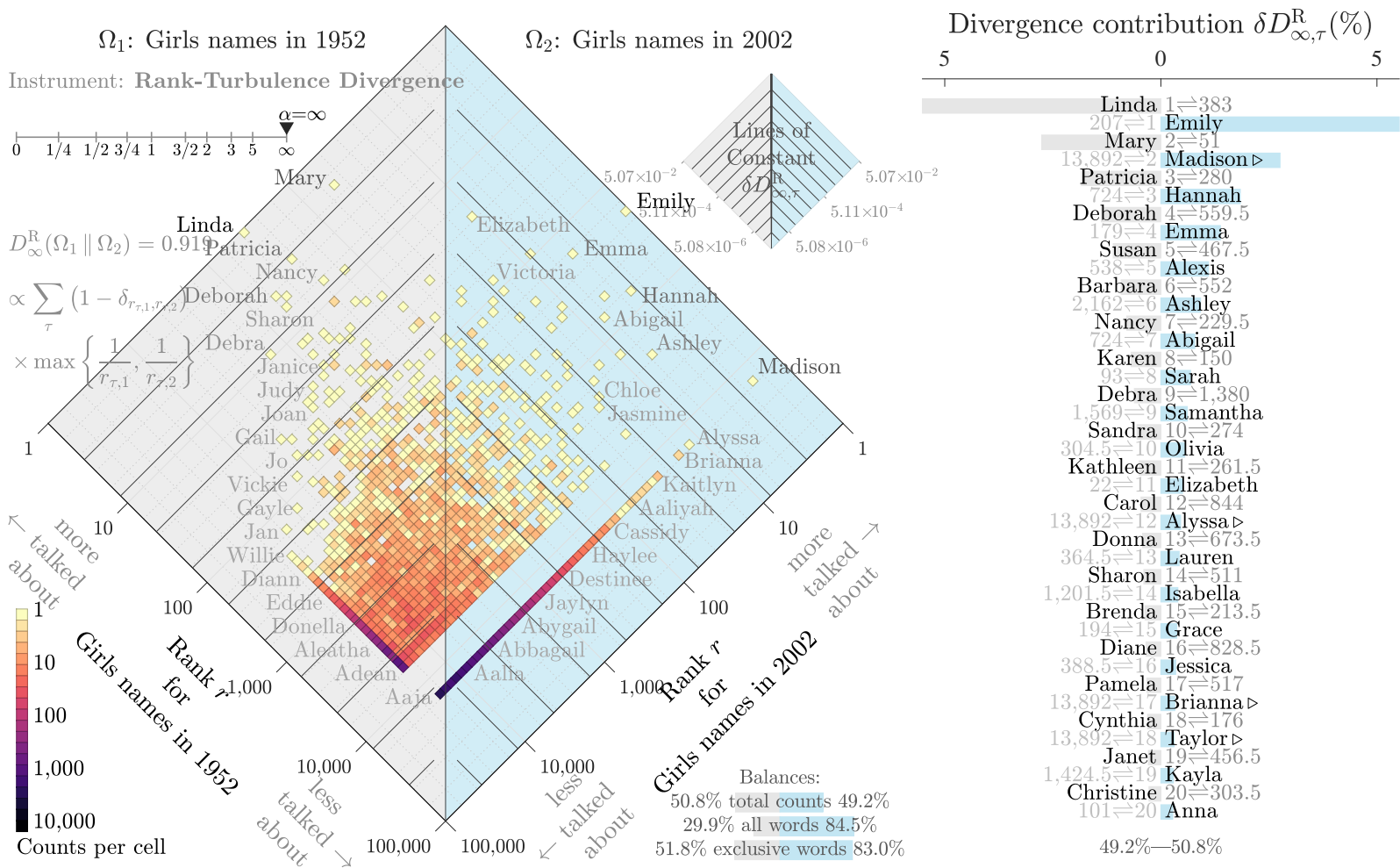
October 1, 2021

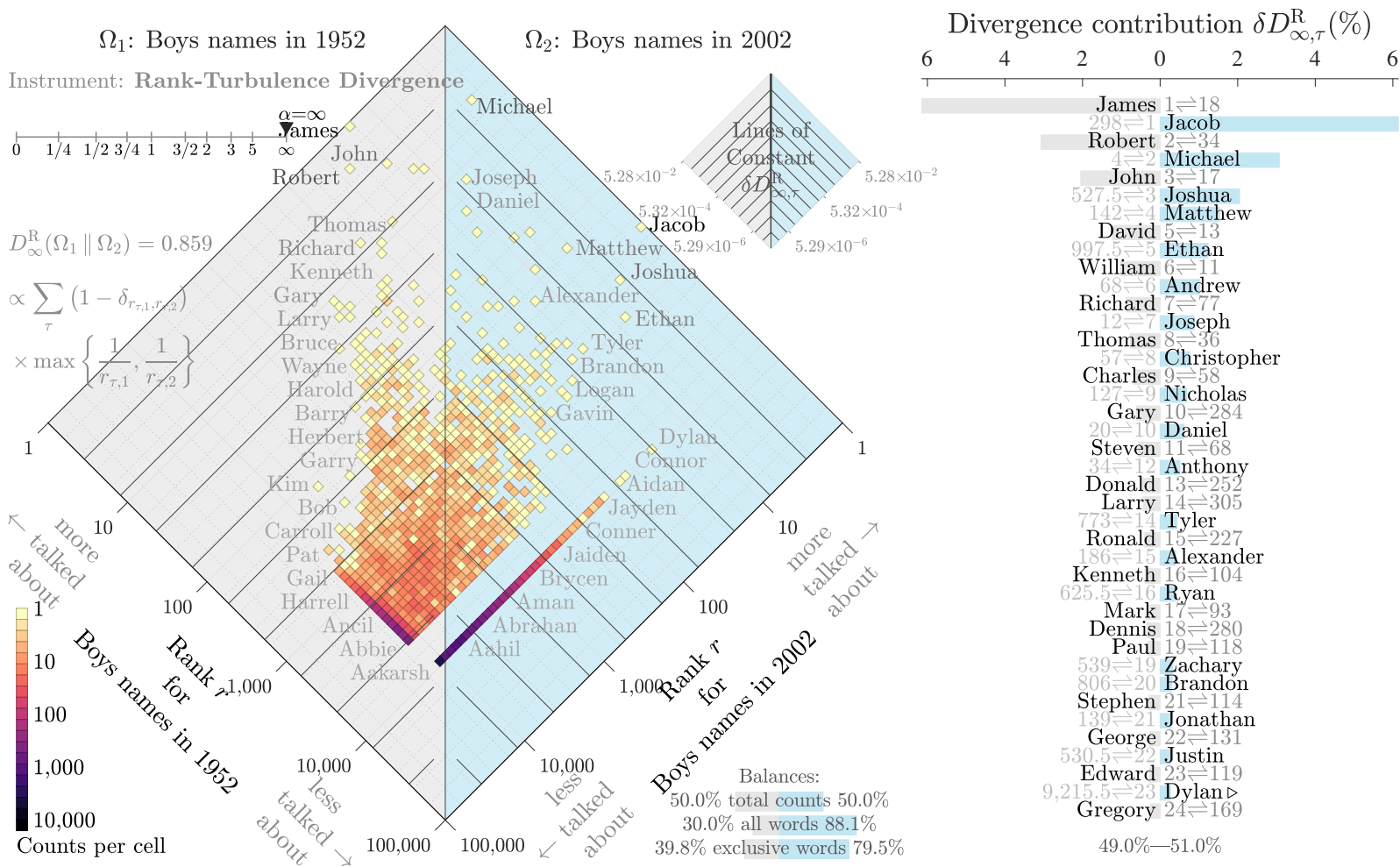
**Link to Github for Assignment Code:** YOURLINKTOYOURGITHUB

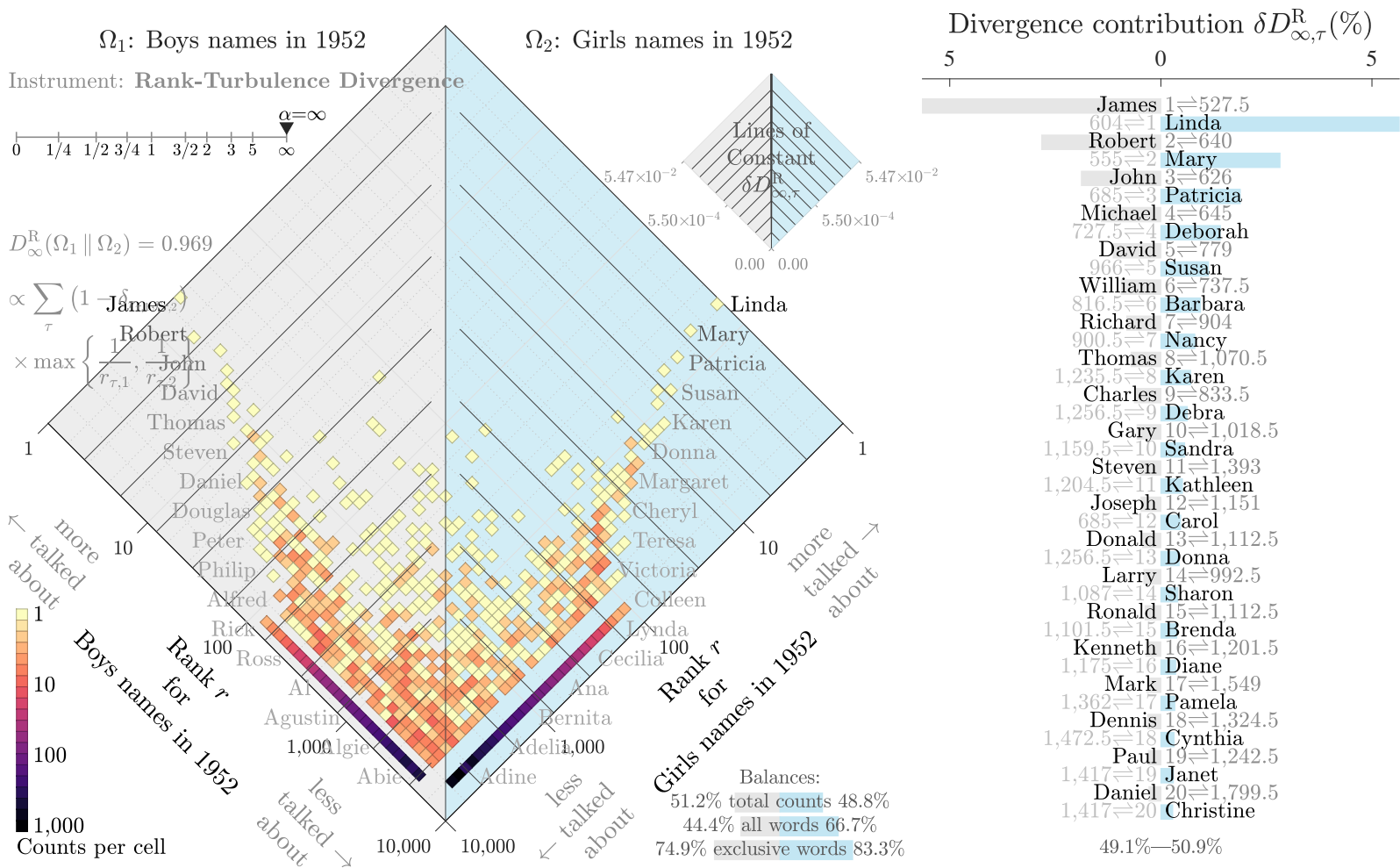
## **Question 1.**

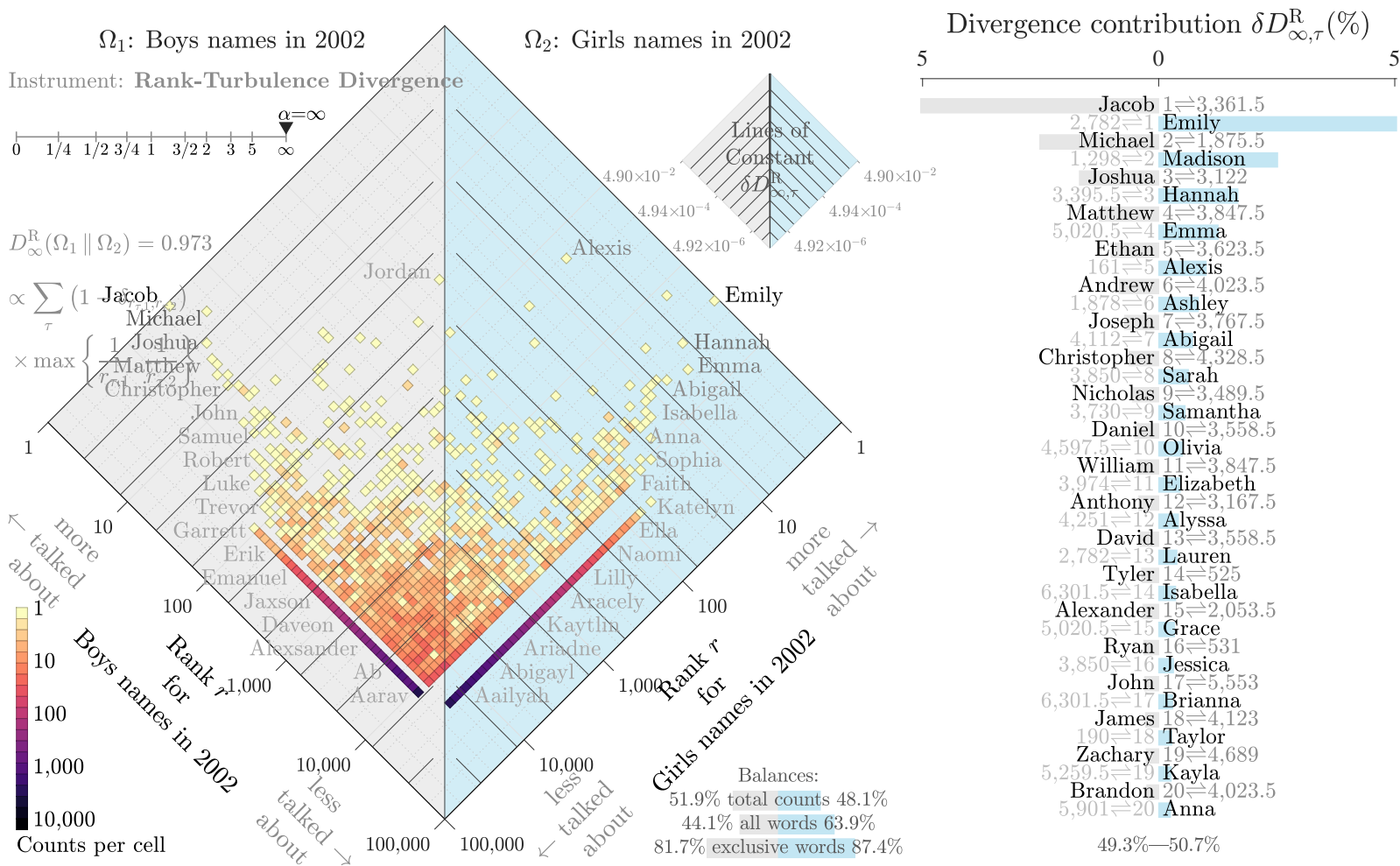
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Generate allotaxonographs comparing the following four pairs: **Responses:**









**Question 2.**

Everyday random walks and the Central Limit Theorem:

Show that the observation that the number of discrete random walks of duration  $t = 2n$  starting at  $x_0 = 0$  and ending at displacement  $x_2n = 2k$  where  $k \in (0, \pm 1, \pm 2, \dots, \pm n)$  is

$$N(0, 2k, 2n) = \binom{2n}{n+k} = \binom{2n}{n-k}$$

leads to a Gaussian distribution for large  $t = 2n$ :

$$\Pr(x_t \equiv x) \simeq \frac{1}{\sqrt{2\pi t}} e^{-\frac{x^2}{2t}}$$

Please note that  $k \ll n$ .

Stirling's sterling approximation will prove most helpful.

**Hint:** You should be able to reach this form:

$$\frac{\text{Some stuff not involving penguins}}{\text{Some other penguin-free stuff} \times (1 - k^2/n^2)^{n+1/2} (1 + k/n)^k (1 - k/n)^{-k}}$$

Lots of sneakiness here. You'll want to examine the natural log of the piece shown above, and see how it behaves for large  $n$ .

You may very well need to use the Taylor expansion  $\ln(1 + z) \equiv z$ .

Exponentiate and carry on.

**Tip:** If at any point penguins appear in your expression, you're in real trouble. Get some fresh air and start again.

**Responses:**

$$\begin{aligned}
& \frac{(2n)!}{(n+k)!(n-k)!} \\
&= \frac{\sqrt{2\pi}}{\sqrt{2\pi}^2} * \frac{(2n)^{2n+1/2} e^{-2n}}{(n+k)^{n+k+1/2} (n-k)^{n-k+1/2} e^{-(n+k) - (n-k)}} \\
&= \frac{1}{\sqrt{2\pi}} * \frac{(2n)^{2n+1/2}}{(n+k)^{n+k+1/2} (n-k)^{n-k+1/2}} \\
&= \frac{1}{\sqrt{2\pi}} * \frac{(2n)^{2n+1/2}}{(n(1+\frac{k}{n}))^{n+k+1/2} (n(1-\frac{k}{n}))^{n-k+1/2}} \\
&= \frac{1}{\sqrt{2\pi n}} * \frac{(2)^{2n+1/2}}{(1+\frac{k}{n})^{n+k+1/2} (1-\frac{k}{n})^{n-k+1/2}} \\
&= \frac{1}{\sqrt{2\pi n}} * \frac{(2)^{2n+1/2}}{(1-k^2/n^2)^{n+1/2} (1+\frac{k}{n})^k (1-\frac{k}{n})^{-k}} \\
&= (2n) \ln(2^{1/2}) - \ln(\sqrt{2\pi n}) - (n+1/2) \ln(1 - \frac{k^2}{n^2}) + k \ln(1 + \frac{k}{n}) - k \ln(1 - \frac{k}{n}) \quad \text{then use Taylor expansion} \\
&= (2n) \ln(2^{1/2}) - \ln(\sqrt{2\pi n}) + (n+1/2) (\frac{k^2}{n^2}) - k \frac{k}{n} - k (\frac{k}{n}) \\
&= (2n) \ln(2^{1/2}) - \ln(\sqrt{2\pi n}) - \frac{k^2}{n} - \frac{k^2}{2n^2} + k \frac{k}{n} - k \frac{k}{n} \\
&= (2n) \ln(2^{1/2}) - \ln(\sqrt{2\pi n}) - \frac{k^2}{n} \\
&= \exp 2n \ln(2^{1/2}) - \ln(\sqrt{2\pi n}) - \frac{k^2}{n^2} \\
&= \frac{2^{1/2}}{\sqrt{2\pi n}} e^{-\frac{k^2}{n}} = \dots \text{I know I'm somewhat close, just running out of time}
\end{aligned}$$

### Question 3.

From the lectures, show that the number of distinct 1-d random walk that start at  $x = i$  and end at  $x = j$  after  $t$  time steps is

$$N(i, j, t) = \binom{t}{(t+j-i)/2}$$

Assume that  $j$  is reachable from  $i$  after  $t$  time steps.

#### Responses:

We know that the number of random walks is  $N_{distinct} = \binom{t}{N}$

$$\begin{aligned}
P + N &= t \\
P - N &= j - i \\
\implies P &= t - N \\
\implies t - N - N &= j - i \\
N &= \frac{t - j + i}{2} \\
\implies \binom{t}{N} &= \binom{t}{\frac{t-j+i}{2}}
\end{aligned}$$

**Question 4.**

*Discrete random walks:*

In class, we argued that the number of random walks returning to the origin for the first time after  $2n$  time steps is given by  $N_{\text{first return}}(2n) = N_{\text{fr}}(2n) = N(1, 1, 2n-2) - N(-1, 1, 2n-2)$  where  $N(i, j, t) = \binom{t}{(t+j-i)/2}$ . Find the leading order term for  $N_{\text{fr}}(2n)$  as  $n \rightarrow \infty$ .

Two-step approach:

- (a) Combine the terms to form a single fraction,
- (b) and then again use Stirling's bonza approximation

If you enjoy this sort of thing, you may like to explore the same problem for random walks in higher dimensions. Seek out George Pólya.

And we are connecting to much other good stuff in combinatorics; more to come in the solutions.

**Responses:**

$$\begin{aligned}
 \binom{2n-2}{(2n-2+1)/2} - \binom{2n-2}{(2n-2+1+1)/2} &= \binom{2n-2}{n-1} - \binom{2n-2}{n} \\
 &= \frac{(2n-2)!}{(n-1)!(n-1)!} - \frac{(2n-2)!}{n!(n-2)!} \\
 &= \frac{(2n-2)!}{(n-1)!(n-1)!} - \frac{(2n-2)!(n-1)}{n(n-1)!(n-1)!} \\
 &= \frac{n(2n-2)!}{n(n-1)!(n-1)!} - \frac{(2n-2)!(n-1)}{n(n-1)!(n-1)!} \\
 &= \frac{(2n-2)!(n-(n-1))}{n(n-1)!(n-1)!} \\
 &= \frac{(2n-2)!}{n(n-1)!(n-1)!}
 \end{aligned}$$

Now apply Stirling's formula

$$\begin{aligned}
 &= \frac{1}{\sqrt{2\pi}} \frac{(2n-2)^{2n-3/2} e^{2-2n}}{n(n-1)^{2(n-1/2)} e^{-2(n-1)}} \\
 &= \frac{1}{\sqrt{2\pi}} \frac{(2n-2)^{2n-3/2}}{n(n-1)^{2n-1}} \\
 &= \frac{1}{\sqrt{2\pi}} \frac{2^{2n-3/2}}{n(n-1)^{1/2}}
 \end{aligned}$$

The leading term in this as  $n \rightarrow \infty$  is  $2^{2n-3/2}$

**Question 5.**

More on the peculiar nature of distributions of power law tails:

Consider a set of  $N$  samples, randomly chosen according to the probability distribution  $P_k = ck^{-\gamma}$  where  $k = 1, 2, 3, \dots$

- (a) Estimate  $\min k_{\max}$ , the approximate minimum of the largest sample in the system, finding how it depends on  $N$ .

(Hint: we expect on the order of 1 of the  $N$  samples to have a value of  $\min k_{\max}$  or greater.)

**Hint - Some visual help on setting this problem up:**



<http://www.youtube.com/watch?v=4tqlEuXA7QQ>

(b) Determine the average value of samples with value  $k \geq \min k_{max}$  to find how the expected value of  $k_{max}$  (i.e.,  $\langle k_{max} \rangle$ ) scales with  $N$ .

Notes:

- For language, this scaling is known as Heap's law.
- In a later assignment, we will test this scaling by (thoughtfully) sampling from power-law size distributions

### Responses:

(a) Since at any given point, the maximum in a given sample is given by the probability that 1 in  $N$  samples is greater than the rest, with value  $k^*$ , we have that

$$\begin{aligned}\frac{1}{N} &= \int_{k^*}^{\infty} cx^{-\gamma} dx \\ \frac{1}{N} &= \frac{c}{-\gamma+1} x^{-\gamma+1} \Big|_{k^*}^{\infty} \\ \frac{1}{N} &= 0 - \frac{c}{-\gamma+1} k^{*- \gamma+1} \\ N^{-1} &= -c^* k^* \\ k^* &\propto N^{\frac{1}{\gamma-1}}\end{aligned}$$

(b) To find the average, we find the first moment of the distribution evaluated from  $k^*$  to infinity.

$$\begin{aligned}\langle x \rangle &= N \int_{k^*}^{\infty} x cx^{-\gamma} dx \\ \langle x \rangle &= N \int_{k^*}^{\infty} cx^{-\gamma+1} dx \\ \langle x \rangle &= N \frac{c}{-\gamma+2} x^{-\gamma+2} \Big|_{k^*}^{\infty}\end{aligned}$$

- For  $1 < \gamma \leq 2$ , we have infinite mean
- For  $\gamma > 2$ , we have:

$$\begin{aligned}\langle k^* \rangle &= 0 - \frac{cN}{-\gamma+2} k^{*- \gamma+2} \\ \langle k^* \rangle &= -\frac{cN}{-\gamma+2} k^{*- \gamma+1} k^* \\ \langle k^* \rangle &= c^* \frac{1}{N} k^* \\ \langle k^* \rangle &= c_3^* N^{\frac{1}{\gamma-1}} \\ N^{\frac{1}{\gamma-1}} &\propto \langle k^* \rangle\end{aligned}$$

[illegible]

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settings.imageformat.open = 'no';

%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%
%% make some flip books
%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%

%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%
%% example flipbook series for rank divergence for 1 and 2
%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%%

settings.plotkind = 'rank';
%% settings.plotkind = 'probability';
%% settings.plotkind = 'count';

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%% settings.instrument = 'probability divergence';
%% settings.instrument = 'alpha divergence type 2';

%% move the shift (adds to 0.60)
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