Model Selection for Gaussian Process Regression

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Preliminary

In a regression setting, we are interested in finding an optimal map f(x) between data points and labels/ function values.

#Gaussian-Process can be used to represent a prior distribution over a space of functions. Gaussian Process can be written as,

$$y = f(x) \sim GPigg(m(x), \; k(x,x')igg)$$

Note ∨

This Note contains the following:

- (1) Compute the marginal likelihood for Gaussian Process Model
- (2) Compute the gradients of (1) w.r.t the hyper-parameters of kernel
- (3) Check (2) using Standard Finite Difference
- (4) Evaluate (1) and (2) to scipy's 'SLSQP' to maximize log marginal likelihood

Model Selection for GP Regression

Although Gaussian Process is a non-parametric model, the so called 'hyper-parameters' in the kernel heavily influence the inference process. It is therefore essential to select the best possible parameters for the kernel. For example in the R.B.F-kernel; the *length* (l) and *scale* (σ) are the hyper-parameters.

$$k(x,x') = \sigma^2 \exp\left(rac{|x-x'|^2}{2l^2}
ight)$$

where, l controls the 'reach of influence on neighbors' and σ dictates the average amplitude from the mean of the function.

Marginal Likelihood

$$\label{eq:posterior} \text{Posterior} = \frac{\text{Likelihood} \times \text{Prior}}{\text{Marginal Likelihood}}$$

$$\mathbb{P}(heta|y,X) = rac{\mathbb{P}(y|X, heta) imes \mathbb{P}(heta)}{\mathbb{P}(y|X)}$$

A #Marginal-Likelihood is a likelihood function that has been integrated over the parameter space. It represents the probability of generating the observed sample from a #prior and is often referred to as the #model-evidence or #evidence.

Let θ represent the parameters of the model. We now formulate a "prior" over the output of the function as a "Gaussian-Process".

$$\mathbb{P}(f|X, heta) = \mathcal{N}\Big(0,k(x,x')\Big)$$

We can always transform the data to have zero mean and (1) can be viewed as a general case. Assume that the (#likelihood) takes the following form

$$\mathbb{P}(Y|f) \sim \mathcal{N}(f, \sigma_n^2 I) \tag{2}$$

(2) tells that the observations y are subject to additive Gaussian noise. Now, the joint distribution is given by;

$$\mathbb{P}(Y, f|X, \theta) = \mathbb{P}(Y|f) \, \mathbb{P}(f|X, \theta) \tag{3}$$

It is worth noting that we would eventually like to optimize the hyper-parameters θ for the kernel function. However, the "prior here is over the mapping f and not any parameters directly. In the **evidence-based** framework, which approximates Bayesian averaging by optimizing the "Marginal-Likelihood" we can make use of the denominator part in the **Bayes' Rule** as an objective function for optimization. For this we take the joint distribution (3) and marginalize over f since we are not directly interested in optimizing it. This can be done in the following way:

$$\mathbb{P}(Y|X, heta) = \int \mathbb{P}(Y,f|X, heta) \ df = \int \mathbb{P}(Y|f) \ \mathbb{P}(f|X, heta) \ df$$

(4) is an integration performed all possible spaces of f and it aims to remove f from the distribution of Y. After marginalization Y is no longer dependent on f but it depends on the hyper-parameters θ .

As per Rasmussen & Williams, the log marginal likelihood is given by;

$$\log \mathbb{P}(y|X,\theta) = -\frac{1}{2}y^T K_y^{-1} y - \frac{1}{2} \log |K_y| - \frac{n}{2} \log (2\pi)$$
 (5)

where $K_y = K_f + \sigma_n^2 I$ is the covariance matrix for the noisy targets y and K_f is the covariance matrix for the noise-free latent f. The first term penalizes wrong prediction, second penalizes model complexity and the third monomial is normalization term.

Gradients of Marginal Likelihood

```
egin{aligned} & egin{aligned} rac{\partial}{\partial 	heta} K^{-1} &= -K^{-1} rac{\partial K}{\partial 	heta} K^{-1} \ & rac{\partial}{\partial 	heta} \log |K| &= \operatorname{trace} \left( K^{-1} rac{\partial K}{\partial 	heta} 
ight) \end{aligned}
```

Now, the partial derivatives w.r.t. the hyper-parameters is given by;

$$\frac{\partial}{\partial \theta_{j}} \log \mathbb{P}(y|X,\theta) = \frac{1}{2} y^{T} K^{-1} \frac{\partial K}{\partial \theta_{j}} K^{-1} y - \frac{1}{2} \operatorname{trace} \left(K^{-1} \frac{\partial K}{\partial \theta_{j}} \right)$$

$$= \frac{1}{2} \operatorname{trace} \left((\alpha \alpha^{T} - K^{-1}) \frac{\partial K}{\partial \theta_{j}} \right) \tag{6}$$

where, $lpha=K^{-1}y$

▲ Warning ∨

In general, computing the inverse of a matrix directly (e.g. np.linalg.inv()) is not stable and there is a loss of precision. In the case when the matrix is positive definite, Cholesky decompostion can be used to compute inverse.

Example:

Let K be a symmetric positive definite matrix. Now, if we want to calculate $\alpha = K^{-1}y$, we can do the following:

$$K = ext{Cholesky}
ightarrow LL^T$$
 $K^{-1} = (L^T)^{-1}L^{-1}$

 $\alpha = \text{np.linalg.solve}(L.T, \text{np.linalg.solve}(L, y))$

Implementation

```
def kernel(x, xp, σ, l):
    '''k(x,x') = sigma^2 exp(-0.5*length^2*|x-x'|^2)'''
    length = l
    sq_norm = (scipy.spatial.distance.cdist(x, xp))**2
    return σ**2 * np.exp(-0.5*sq_norm/(length**2))
```

```
computes partial derivative of K w.r.t length (l)
                arg: x1 = (N1, D), x2 = (N2, D)
                return: (N1, N2)
                1.1.1
                sq_norm = (scipy.spatial.distance.cdist(x1, x2))**2
                return (\sigma**2) * np.exp(-sq_norm/(2*l**2)) * (sq_norm) / (l**3)
def dKd\sigma(x1, x2, \sigma, l):
                computes partal derivatice of K w.r.t sigma (std not variance)
                arg: x1 = (N1, D), x2 = (N2, D)
                return: (N1, N2)
                1.1.1
                sq_norm = (scipy.spatial.distance.cdist(x1, x2))**2
                return 2*\sigma*np.exp(-sq_norm/(2*l**2))
def dLdt(a, iKxx, dKdt):
                1.1.1
                computes gradient of log marginal likelihood w.r.t. a hyper-parameter
                i.e. either sigma or length
                1.1.1
                return 0.5**np.trace(np.dot(a @ a.T - iKxx), dKdt)
def f_{opt}(kernel, X, y, \sigma, l):
                Evalaute Negative-Log Marginal Likelihood
                \sigma_n = 0.1 # std of noise hard-coded for now
                K = kernel(X, X, \sigma = \sigma, l = l) + (\sigma_n **2) *np.eye(X.shape[0])
                L = np.linalg.cholesky(K) + 1e-12 # Cholesky decomposition
                a = np.linalg.solve(L.T, np.linalg.solve(L, y)) # compute alpha
                \#\log_{i,k}(0) = -0.5 * y.T @ a - 0.5 * np.trace(np.log(L)) - 0.5 * X.shape[0] * np.log(2*np.pi)
                \log_{1} ikelihood = -0.5 * y.T @ a - 0.5 * np.\log_{1} np.\log_{2} 
                return -log_likelihood
def grad_f(kernel, X, y, l, \sigma):
                Compute gradient of objective function w.r.t. two parameters
                l, \sigma = params
                \sigma_n = 0.1 # std of noise hard-coded for now
                K = kernel(X, X, \sigma = \sigma, l = l) + (\sigma_n**2)*np.eye(X.shape[0])
                L = np.linalg.cholesky(K) # Cholesky decomposition
                a = np.linalg.solve(L.T, np.linalg.solve(L, y)) # compute alpha
                inv_k = np.linalg.inv(K)
                grad = np.empty([2,])
                grad[0] = dLdt(a = a, iKxx = inv_k, dKdt = dKd\sigma(X, X, \sigma, l)) # gradient w.r.t sigma
                grad[1] = dLdt(a = a, iKxx = inv_k, dKdt = dKdL(X, X, \sigma, l)) # gradient w.r.t length
                return grad
def marginal(params, X, y):
                1.1.1
                Evalaute Negative-Log Marginal Likelihood -- for scipy optimization
                #print (params)
                l, \sigma = params
                \sigma_n = 0.1 \text{ # std of noise hard-coded for now}
                K = kernel(X, X, \sigma = \sigma, l = l) + (\sigma_n **2)*np.eye(X.shape[0])
                L = np.linalg.cholesky(K) + 1e-12 # Cholesky decomposition
                a = np.linalg.solve(L.T, np.linalg.solve(L, y)) # compute alpha
                \#\log_{\text{likelihood}} = -0.5 * y.T @ a - 0.5 * np.trace(np.log(L)) - 0.5 * X.shape[0] * np.log(2*np.pi)
                \log_{\text{likelihood}} = -0.5 * \text{y.T } @ \text{a} - 0.5 * \text{np.log(np.linalg.det(K))} - 0.5 * \text{X.shape[0]} * \text{np.log(2*np.pi)}
                return -log_likelihood
```

def dKdL(x1, x2, σ , l):

```
True function f(x) = sin(x) & X ~ Unif(-4,4) with 10 samples
'''
f_sin = lambda x: (np.sin(x)).flatten()
X = np.random.uniform(-4, 4, size = (10, 1))
y = f_sin(X)

# Scipy-optimization via SLSQP

lim = [10**-3, 10**3]
bound = [lim, lim]
start = [0.3, 0.1] # initial hyper-parameters
result = scipy.optimize.minimize(fun = marginal, x0 = start, args = (X, y), method = 'SLSQP', options = {'disp':True}, bounds = bound, tol = 0.0001)
```

```
Contour Plot
 1.1.1
L = np.linspace(10**-3, 10**2, 1000)
S = np.linspace(10**-3, 10**3, 1000)
\sigma, l = np.meshgrid(L, S)
func_val = np.zeros_like(\sigma)
for i in range(\sigma.shape[0]):
                                          for j in range(l.shape[0]):
                                                                                    func_val[i, j] = f_opt(kernel = kernel, X = X.reshape(-1,1), y = y.reshape(-1,1), \sigma = S[i], l = f_opt(kernel = kernel, X = X.reshape(-1,1), y = y.reshape(-1,1), reshape(-1,1), reshape(
L[j])
plt.contourf(\sigma, l, func_val, cmap = 'plasma')
plt.xscale('log')
plt.yscale('log')
plt.scatter(result.x[0], result.x[1], color = 'black', marker = 'x')
plt.colorbar()
plt.show()
```

