STATS 230: Computational Statistics Posterior Approximation

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Outline

- When it is difficult to sample from a distribution, we can approximate it with one of the standard distributions such as normal
- In this lecture, we discuss such approximation methods in general; however, in statistics these methods are usually used for posterior approximation
- In what follows, we start with a simple method called Laplace's approximation
- We then discuss variational methods and illustrate it with a simple application

- To approximate a distribution, we could focus on the high probability regions, more specifically, where the probability is the highest (at least locally), i.e., the mode (\hat{x}) of the distribution.
- ullet For this, we can either obtain \hat{x} analytically or use optimization algorithms such as Newton's method
- We first discuss the univariate case, and then generalize it to multivariate problems

- Assume that we know the distribution (e.g., posterior distribution) up a constant: $f(x) = f^*(x)/Z$.
- Denote the log of unnormalized density as $\ell(x) = \log(f^*(x))$.
- Find the first derivative $\ell'(x)$ and the second derivative $\ell''(x)$.
- Start with an initial value x₀.
- At each iteration n, use the Taylor series expansion (up to the quadratic term) around the current value of x_n

$$\ell(x) \simeq \ell(x_n) + \ell'(x_n)(x - x_n) + \frac{1}{2}\ell''(x_n)(x - x_n)^2$$

• Taking the derivative of $\ell(x)$, setting it to zero, and solving for x gives the next guess x_{n+1}

$$x_{n+1} = x_n - \frac{\ell'(x_n)}{\ell''(x_n)}$$

- Now assume that we have found the mode \hat{x} .
- We write down the Taylor series expansion (up to the quadratic term) around \hat{x} (note that $\ell'(\hat{x}) = 0$)

$$\ell(x) \simeq \ell(\hat{x}) + \frac{1}{2}\ell''(\hat{x})(x-\hat{x})^2$$

Rewrite the above approximation as follows

$$\ell(x) \simeq \ell(\hat{x}) - \frac{1}{2c}(x - \hat{x})^2$$

$$c = [-\ell''(\hat{x})]^{-1}$$

which means

$$f^*(x) \simeq f^*(\hat{x}) \exp[-\frac{1}{2c}(x-\hat{x})^2]$$

- The right hand side is the density of a normal up to a constant.
- Therefore, we can approximate the distribution with

$$N(\hat{x}, [-\ell''(\hat{x})]^{-1})$$

For multivariate distributions, we have

$$N(\hat{\mathbf{x}}, [-\nabla^2 \ell(\hat{\mathbf{x}})]^{-1})$$

Normalizing constant

• We can use the above approach to approximate an integral

$$Z=\int f^*(x)dx$$

for example, to find the normalizing constant of the posterior distribution

• As before, assume that the function $f^*(x)$ peaks at a \hat{x} . We have

$$\int f^*(x)dx \simeq \int f^*(\hat{x}) \exp[-\frac{1}{2c}(x-\hat{x})^2]dx$$

• Then, we can approximate $Z = \int f^*(x) dx$ as follows:

$$Z = f^*(\hat{x})\sqrt{2\pi c}$$

• When $x = (x_1, ..., x_k)$,

$$Z = f^*(\hat{x})\sqrt{(2\pi)^k|\Sigma|}$$

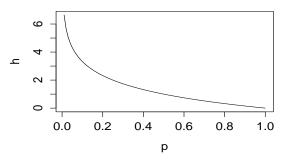
Variational methods

- Normal approximation does not always work.
- In what follows, we discuss a more general approach based on variational methods.
- This approach is inspired by information theory, so we first review some fundamental concepts in this field.

Information theory

- Information theory deals with communication problems.
- Information content (in bits) of an outcome x is

$$h(X = x) = \log_2 \frac{1}{P(X = x)}$$

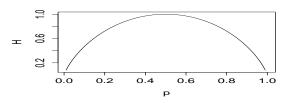


Entropy

- Shannon information content is highest for outcomes with low probability.
- For a set of outcomes, entropy is defined as the average Shannon information:

$$H(X) = \sum_{x} P(x) \log_{2} \frac{1}{P(x)}$$
$$= -\sum_{x} P(x) \log_{2} P(x)$$

• Suppose there are only two possible outcomes with probabilities p and 1-p,



Relative entropy

• The relative entropy between two probability distributions P(x) and Q(x) is defined as

$$D_{KL}(P||Q) = \sum_{x} P(x) \log \frac{P(x)}{Q(x)}$$

which satisfies Gibbs' inequality

$$D_{KL}(P||Q) \geq 0$$

ullet Note that this is not symmetric in general so $D_{\mathit{KL}}(P||Q)$ is not the same as

$$D_{KL}(Q||P) = \sum_{x} Q(x) \log \frac{Q(x)}{P(x)}$$

This is known as Kullback-Leibler divergence

Variational methods

• Now suppose we have a complex (e.g., high dimensional) probability distribution P(x),

$$P(x) = \frac{1}{z}e^{-E(x)}, \quad \text{where } x = (x_1, \dots, x_d)$$

- We want to approximate P(x) by $Q(x, \theta)$ through adjusting θ to get the best approximation.
- ullet θ is called "variational parameters."
- We define "best" in terms of minimum Kullback-Leibler divergence between the two distributions.

• As an example, consider the following model:

$$x|\mu, \sigma^2 \sim N(\mu, \sigma^2)$$
 $\mu \sim N(\mu_0, \sigma_0^2)$
 $\sigma^2 \sim \text{Inv-Gamma}(\alpha, \beta)$

Posterior distribution of model parameters is

$$P(\mu, \sigma^{2}|x) = \frac{P(\mu, \sigma^{2}, x)}{P(x)}$$

$$= \frac{P(x|\mu, \sigma^{2})P(\mu)P(\sigma^{2})}{\int P(x|\mu, \sigma^{2})P(\mu)P(\sigma^{2})d\mu d\sigma^{2}}$$

- This is not a tractable distribution in general.
- We want to approximate $P(\mu, \sigma^2|x)$ with a tractable distribution, $Q(\theta)$ that depends on variational parameters θ .

• We find a member of $Q(\theta)$ family (i.e., find optimum θ) as an approximation to the posterior distribution by minimizing the KL divergence,

$$D = \int_{Q} \log \frac{Q(\mu, \sigma^{2}|\theta)}{P(\mu, \sigma^{2}|x)} Q(\mu, \sigma^{2}|\theta) d\mu d\sigma^{2}$$

$$= E_{Q}[\log Q(\mu, \sigma^{2}|\theta)] - E_{Q}[\log P(\mu, \sigma^{2}, x)] + \log P(x)$$

$$= -\mathcal{L}(Q) + \text{constant}$$

• To minimize D, we need to minimize $-\mathcal{L}(Q)$, or alternatively, maximize $\mathcal{L}(Q)$,

$$\mathcal{L}(Q) = E_Q[\log P(\mu, \sigma^2, x)] - E_Q[\log Q(\mu, \sigma^2 | \theta)]$$

 In the above example, using Jensen's inequality and concavity of the logarithm function, we have

$$\begin{split} \log P(x) &= \log \int P(\mu, \sigma^2, x) d\mu d\sigma^2 \\ &= \log \int P(\mu, \sigma^2, x) \frac{Q(\mu, \sigma^2 | \theta)}{Q(\mu, \sigma^2 | \theta)} d\mu d\sigma^2 \\ &= \log E_Q \left[\frac{P(\mu, \sigma^2, x)}{Q(\mu, \sigma^2 | \theta)} \right] \\ &\geq E_Q \left[\log \frac{P(\mu, \sigma^2, x)}{Q(\mu, \sigma^2 | \theta)} \right] \\ &= E_Q \left[\log P(\mu, \sigma^2, x) \right] - E_Q \left[\log Q(\mu, \sigma^2 | \theta) \right] = \mathcal{L}(Q) \end{split}$$

• Therefore, $\mathcal{L}(Q)$ is the lower bound of log P(x), i.e., the logarithm of the marginal probability of the observed data

- Minimizing $-\mathcal{L}(Q)$ in general would not be easy.
- ullet To make this simple, we usually assume that $Q(\mu, \sigma^2 | \theta)$ factorizes.
- For the above example, we assume

$$Q(\mu, \sigma^2 | \theta) = Q(\mu | m, v) Q(\sigma^2 | a, b)$$

More specifically,

$$\mu | m, v \sim N(\mu | m, v)$$

 $\sigma^2 | a, b \sim \text{Inv-Gamma}(\sigma^2 | a, b)$

• We can minimize $-\mathcal{L}(Q)$ using gradient descent (or coordinate descent, when possible) methods,

$$\theta^{t+1} = \theta^t - \eta \nabla_{\theta} \mathcal{L}(Q)$$



Approximate posterior distribution using variational Bayes. Left panel: True posterior distribution; Right panel: Variational Bayes approximation.

