STATS 225: Bayesian Analysis Supplementary Materials: A brief review of probability ¹

Babak Shahbaba

Department of Statistics, UCI

Winter, 2015

Probability

- We are familiar with statements such as $X \sim Poisson(5)$ distribution. We interpret it as X being a non-negative random variable such that $P(X = k) = 5^k \exp(-5)/k!$.
- We call X a discrete random variable.
- We are also familiar with statements like $Y \sim \text{Normal}(0,1)$. It means, for example, $P(a \le Y \le b) = \int_a^b \frac{1}{\sqrt{2\pi}} \exp(-\frac{y^2}{2}) dy$. We know that P(Y = y) = 0 for any real number y. Y is an example of absolutely continuous random variable.
- Introductory courses on probability group random variables as either discrete or continuous.
- This is not completely correct. There are other types of random variables. For example, consider a random variable Z defined as follows. We toss a coin, if it's head, we set Z=X, otherwise, we set Z=Y.

Probability measure

- A mathematical rigorous probability theory is studied in the context of measure theory.
- A probability measure space (or a probability triple), (Ω, \mathcal{F}, P) , defined as follows:
 - $ightharpoonup \Omega$ is a non-empty set referred to as the sample space (i.e., for example the sample space for the poisson distribution consists of all the non-negative integers).
 - ${\mathcal F}$ is a σ -algebra, which is a collection of measurable (i.e., the probability is defined) subsets of Ω (including Ω itself and the empty set \emptyset), all their complements, and their countable unions. That is, Ω is closed under complement (i.e., if $A \in {\mathcal F}$, then $A^c \in {\mathcal F}$), under countable unions and intersections.
 - P is a probability measure mapping between $\mathcal F$ and a real number between 0 and 1 such that $P(\emptyset)=0,\ P(\Omega)=1,$ and P is countably additive, i.e., if $A_1,A_2,...$ are disjoint subsets included in $\mathcal F$, we have

$$P(A_1 \cup A_2 \cup, ...) = P(A_1) + P(A_2) +, ...$$

Some additional results

- $P(A^c) = 1 P(A)$
- Monotonicity: if $A \subseteq B$, where $A, B \in \mathcal{F}$, then $P(A) \leq P(B)$.
- Countable sub-additivity: if $A_1, A_2, ... \in \mathcal{F}$, which may not be disjoint in general, then $P(\bigcup_n A_n) \leq \sum_n P(A_n)$
- $P(A \cup B) = P(A) + P(B) P(A \cap B)$, where $A, B \in \mathcal{F}$ may not be disjoint.

Shahbaba (UCI) Bayesian Analysis Winter, 2015 4 / 2

Some examples

- Tossing a fair coin:

 - $\mathcal{F} = \{\emptyset, \Omega, \{H\}, \{T\}\}$
 - $P(\emptyset) = 0, P(\Omega) = 1, P(H) = P(T) = 1/2.$
- Tossing *n* fair coins:
 - $\Omega = \{(x_1, x_2, ..., x_n)\}, \text{ where } x_i = 01.$
 - $\mathcal{F} = 2^{\Omega} = \{\text{all subsets}\}\$
 - $P(A) = \frac{|A|}{2^n}$
- Poisson(5) distribution.
 - $\Omega = \{0, 1, 2, ...\}$
 - $\mathcal{F} = 2^{\Omega} = \{\text{all subsets}\}\$
 - $P(A) = \sum_{k \in A} 5^k \exp(-5)/k! \qquad A \in \mathcal{F}$
- What about continuous distributions (where the sample space is not countable) such as Uniform(0, 1)?

Some examples

- The probability triple corresponding to the Uniform(0, 1) distributionin is called the *Lebesgue* measure on [0, 1].
- The sample space is obviously $\Omega = [0, 1]$.
- To construct the corresponding σ -algebra, consider $\mathcal J$ as the set of all intervals (e.g., open, closed, half-open, singleton, etc.) contained in [0, 1].
- Now add all the countable unions of intervals, their complements, their countable intersections, etc. (for original intervals and those created later) to create a σ -algebra.
- The smallest σ -algebra create, $\mathcal{B} = \sigma(\mathcal{J})$ is called the *Borel* σ -algebra, and each of its elements is called a *Borel* set.
- For the Uniform(0, 1) distribution, the probability of each interval is equal to the length of that interval. That is, P([a,b]=P([a,b))=P((a,b])=P((a,b))=b-a for $0 \le a \le b \le 1$ (to define P more precisely, we need to discuss *outer measure* and *extension theorem*).
- Similar procedure is used for other continuous distributions.

《□》《圖》《意》《意》。意

Random variables

- Random variables: it assigns numerical values to each possible outcome within a sample space, Ω .
- Therefore, given a probability triple, (Ω, \mathcal{F}, P) , a random variable X is a measurable function from Ω to the real numbers \mathcal{R} .
- For example, we can define X(Taile) = 0 and X(Head) = 1.
- Since X is measurable, we can talk about P(X = 0) by which we mean P(Tail). In general, we can talk about $P(X \in B)$, for any Borel set B.
- Another example: if (Ω, \mathcal{F}, P) is a Lebesgue measure on [0, 1], we can define a random variable $X(\omega) = 3\omega + 4$, for all $\omega \in \Omega$.

Random variables

$$P(X > x) = P(\omega \in \Omega, X(\omega) > x)$$

$$= P\{\omega \in \Omega, 3\omega + 4 > x\}$$

$$= P\{\omega > \frac{x - 4}{3}\}$$

$$P(X > x) = \begin{cases} \frac{1}{7 - x} & x \le 4\\ 0 & x > 7 \end{cases}$$

Independence

 Independence: Two events (or random variables) are independent if they do not affect each other's probability. That is, knowing whether an event A has occurred does not change the probability of event B; we say:

$$P(A \cap B)/P(A) = P(B)$$

or alternatively:

$$P(A \cap B) = P(A)P(B)$$

• We can extend this to any number of events presented as a collection $\{A_{\alpha}\}_{{\alpha}\in I}$:

$$P(A_{\alpha_1} \cap A_{\alpha_2} \cap ... \cap A_{\alpha_j}) = P(A_{\alpha_1})P(A_{\alpha_2})...P(A_{\alpha_j})$$

for any choice of $\alpha_1, \alpha_2, ..., \alpha_j \in I$

Independence

 \bullet SImilarly, we can talk about the independence of random variables. Random variables X and Y are independent if

$$P(X \in S_1, Y \in S_2) = P(X \in S_1)P(Y \in S_2)$$

or alternatively

$$P(X \le x, Y \le y) = P(X \le x)P(Y \le y) \qquad \forall x, y \in \mathcal{R}$$

• For a collection of independent random variables, $\{X_{\alpha}\}_{{\alpha}\in I}$ we have

$$P(X_{\alpha_i} \in S_i, \forall 1 \leq i \leq n) = \prod_{i=1}^n P(X_{\alpha_i} \in S_i)$$

Note that if X and Y are independent, so are f(X) and g(Y).

Expected values

• For simple random variables whose range is finite, we can represent the distinct values as $x_1, x_2, ..., x_n$ and write $X = \sum_{i=1}^{n} x_i \mathbf{1}_{A_i}$, where $\mathbf{1}_A$ is an indicator function such that

$$\mathbf{1}_{A}(\omega) = \left\{ \begin{array}{ll} 1 & \omega \in A \\ 0 & \omega \notin A \end{array} \right.$$

The expected value (mean or expectation) for such variables is defined as:

$$\mu_X = E(X) = E\left(\sum_{i=1}^{n} x_i \mathbf{1}_{A_i}\right) = \sum_{i=1}^{n} x_i P(A_i)$$

where $A_i = \{\omega \in \Omega; X(\omega) = x_i\}$, and $\{A_i\}$ is a finite partition (or in general any collection) of Ω .

Shahbaba (UCI) Bayesian Analysis

Expected values

For example, if we toss a coin and define

$$X(\omega) = \begin{cases} 10 & \text{if Head} \\ 20 & \text{if Tail,} \end{cases}$$

then
$$E(X) = 10 \times 1/2 + 20 \times 1/2 = 15$$
.

ullet Another example: Consider the Lebesgue measure, and let's define X as follows:

$$X(\omega) = \left\{ \begin{array}{ll} 4 & \omega < 0.25 \\ 6 & \omega = 0.25 \\ 8 & \omega > 0.25, \end{array} \right.$$

then
$$E(X) = 4 \times 1/4 + 6 \times 0 + 8 \times 3/4 = 7$$

Some properties of expectation

- $E(\mathbf{1}_A) = P(A)$
- E(c) = c
- E(aX + bY) = aE(X) + bE(Y)
- Expectation is order preserving, i.e., if $X(\omega) \leq Y(\omega)$ for all $\omega \in \Omega$, then $E(X) \leq E(Y)$
- $|E(X)| \le E(|X|)$
- If X and Y are independent, then E(XY) = E(X)E(Y); note that the other direction does not always hold.

Shahbaba (UCI) Bayesian Analysis Winter, 2015 13 /

Some properties of expectation

- If f(X) is a function of X, $f: \mathcal{R} \to \mathcal{R}$, then f itself is a simple random variable and can be written as $f(X) = \sum_{i=1}^{n} f(x_i) \mathbf{1}_{A_i}$ and $E(f(x)) = \sum_{i=1}^{n} f(x_i) P(A_i)$.
- Especially, if $f(X) = (x \mu_X)^2$, the expectation of f is the *variance* of X: $Var(X) = E((x \mu_X)^2)$, which leads to the well known conclusion that $Var(X) = E(X^2) E(X)^2$ and also $Var(X) \le E(X^2)$.

Some other properties of variance

- $Var(aX + b) = a^2 Var(X)$
- Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y), where $Cov(X, Y) = E((x \mu_X)(y \mu_y)) = E(XY) E(X)E(Y)$
- If X and Y are independent, then Cov(X, Y) = 0 and Var(X + Y) = Var(X) + Var(Y).
- Variance of X is in fact its second central moment.
- In general, the k^{th} moment of a random variable is defined as $E(X^k)$.
- With some mathematical precautions, the above properties can be extended to non-simple random variables.

The integration connection

- Similar to the integral, expectation has some nice properties such as linearity, oerder-preserving and so forth.
- In fact, it can be shown that given a probability triple (Ω, \mathcal{F}, P) , and a measurable function X,

$$E(X) = \int_{\Omega} X dP = \int_{\Omega} X(\omega) P(d\omega)$$

which is the integral of X with respect to the probability measure.

- If (Ω, \mathcal{F}, P) is the Lebesgue measure on [0, 1], and X is Riemann integrable, then the above integral is the common calculus-style integral: $E(X) = \int_0^1 X(t) dt$.
- Even if X is not Riemann integrable (but nevertheless a measurable function with respect to the Lebesgue measure), we can still get the expectation which in this case called the *Lebesgue integral*. That is, the Lebesgue integral is the generalization of the Riemann integral.

Shahbaba (UCI) Bayesian Analysis

Distributions

• Given a random variable X on a probability triple (Ω, \mathcal{F}, P) , its distribution μ is a probability measure on the sample space \mathcal{R} (with the Borel σ -algebra) defined as

$$\mu(B) = P(X \in B)$$
 B Borel $X \sim \mu$

- Moreover, the *cumulative distribution function* of a random variable X is defined as $F_X(x) = P(X \le x)$ for $x \in \mathcal{R}$.
- Note that $\lim_{x\to -\infty} F_X(x) = 0$ and $\lim_{x\to \infty} F_X(x) = 1$.
- Recall that we defined the expected value of a measurable function f(X) as $E(f(X)) = \int_{\Omega} f(X(\omega))P(d\omega)$ with respect to the probability measure P.
- Alternatively, we can define $E(f(X)) = \int_{-\infty}^{\infty} f(t)\mu(dt) = \int_{-\infty}^{\infty} f(t)d\mu$. This is known as the *change of variable theorem*.

Some simple distributions

- One simple distribution is the point mass δ_c , which is the distribution of random variable X where P(X=c)=1.
- Another simple distribution is the Poisson(θ) distribution, where $\mu(X) = \sum_{i=0}^{\infty} (\theta^{j} \exp(-\theta)/j!) \delta_{j}$
- Normal(0, 1) distribution is defined as $\mu_N(B) = \int_{-\infty}^{\infty} f(t) \mathbf{1}_B(t) \lambda(dt)$. where $f(x) = \frac{1}{\sqrt{2\pi}} \exp(-x^2/2)$ is called the *density function*, and λ is the Lebesgue measure on R.
- If a distribution has, μ , has a density, f, instead of taking the integral of a function g(t) with respect to μ , we can take the integral g(t)f(t) with respect to λ .

$$\int_{-\infty}^{\infty} g(t)\mu(dt) = \int_{-\infty}^{\infty} g(t)f(t)\lambda(dt)$$

 That is, for such cases we mainly take a calculus-style integral using the density function.

Convergence

- Convergence with probability 1 (almost surely): $P(\lim_{n\to\infty} X_n = X) = 1$.
- Convergence in probability: $\lim_{n\to\infty} P(|X_n-X|\geq \varepsilon)=0$.
- Convergence in distribution: $\lim_{n\to\infty} P(X_n \le x) = P(X \le x)$.
- Convergence with probability $1 \Rightarrow$ convergence in probability \Rightarrow convergence in distribution.
- Weak law of large numbers: Let $X_1, X_2, ...$ be a sequence of independent random variables with the same mean m and finite variance, then their partial average, $\frac{1}{n}(X_1 + X_2 + ... + X_n)$ converges in probability to m.
- Strong law of large numbers: If besides the above conditions the forth momen, $E((X_i m)^4)$ is also finite, the partial average converges to m with probability 1 (i.e., almost surely).
- Central limit theorem: Let $X_1, X_2, ...$ be *iid* with finite mean m and finite variance v. Set $S_n = X_1 + X_2 + ... + X_n$. Then as $n \to \infty$, $\frac{S_n nm}{\sqrt{nv}}$ convergence in distribution to $Z \sim N(0, 1)$.

Conditional probability and expectation

- Conditional probability is simply defined as $P(A|B) = \frac{P(A \cap B)}{P(B)}$, where P(B) > 0. This is the proportion of the event B that also includes the event A. In other words, you investigate the event A in a smaller subset of the sample space where the event B occurs.
- Similarly, for random variables, we can define the conditional distribution as $P(Y \in S|B) = \frac{P(Y \in S,B)}{P(R)}$.
- Using this new constructed (conditional) distribution, ν, we can define conditional expectations in the usual way: $E(Y|B) = \int y \nu d(y)$, $E(f(Y)|B) = \int f(y)\nu(dy).$
- This seems quite straightforward as long as P(B) > 0. But what happens when P(B) = 0; for example, can we discuss P(A|X = 0.5) where X is a random variable with Uniform(0, 1) distribution.

Conditional probability and expectation

• To resolve this issue, we regard the conditional probability P(A|X) and expectation E(Y|X) as themselves being random variables that are functions of X. These new values should have the correct expected values:

$$E(P(A|X)) = P(A)$$
 $E(E(Y|X)) = E(Y)$

• However, having the above correct expected values is not enough to specify the distribution of P(A|X) and E(E(Y|X)). More specifically, we need these for any Borel $S \subseteq \mathcal{T}$

$$E(P(A|X)\mathbf{1}_{X\in S}) = P(A\cap \{X\in S\})$$

$$E(E(Y|X)\mathbf{1}_{X\in S}) = E(Y\mathbf{1}_{X\in S})$$

• Since the above expectations would not be affected by changes on a set of measure 0, the above definitions are only unique up to a set of measure 0. We can in fact change P(A|X) without restriction whenever P(X=x)=0.

Conditional probability and expectation

• Also note that when $S = \mathcal{R}$, we again obtain

$$E(P(A|X)) = P(A)$$
 $E(E(Y|X)) = E(Y)$

- Some useful properties:
 - $P(A \cap B) = P(A)P(B|A) = P(B)P(A|B)$
 - ▶ If A is independent of B, then P(A|B) = P(A) and $P(A \cap B) = P(A)P(B)$
 - ► The total probability rule: if $B_1, B_2, ..., B_n$ partition the sample space (i.e., their are mutually exclusive and $\bigcup_{i=1}^n B_i = \Omega$), then

$$P(A) = P(A|B_1)P(B_1) + P(A|B_2)P(B_2) + ... + P(A|B_n)P(B_n)$$

- The multiplication rule: If $A_1, A_2, ..., A_n$ is a sequence of events, then $P(A_1 \cap A_2 \cap ... \cap A_n) = P(A_1)P(A_2|A_1)P(A_3|A_1 \cap A_2)...P(A_n|A_1 \cap ... \cap A_{n-1})$
- Conditional probabilities play a very important role in Bayesian statistics, and we will discuss them more in future.