# Random numerical semigroups and sums of subsets of cyclic groups

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## Contents

1	Expected value proof												2	,				
	1.1	Upper	bound														2	)
		1.1.1	Proof of the	upper bound									•				4	Ŀ
$\mathbf{A}$	Use	ful Boı	ınds														10	)

## Chapter 1

## Expected value proof

### 1.1 Upper bound

Before proving part (b) of the main theorem, we will prove a lemma that shows that a cyclic group of prime order is covered by the sums of a random subset of logarithmic size almost always.

**Lemma 1.1.1.** Let q be a prime number and  $\mathcal{A}$  be a random subset of  $\mathbb{Z}_q$  of size  $4\lfloor 3\log_2 q \rfloor$ . As q tends to infinity,  $2\lfloor 3\log_2 q \rfloor \mathcal{A}$  covers  $\mathbb{Z}_q$  almost always.

**Proof.** Let  $s \in \mathbb{N}$  such that  $s \leq q$ . Let  $\mathcal{A}$  be a uniformly random subset of  $\mathbb{Z}_q$  of size s, that is,

$$\Pr(\mathcal{A}) = \frac{1}{\binom{q}{s}}.$$

For a given  $z \in \mathbb{Z}_q$  and  $k \in \mathbb{N}$  for which  $k \leq s/2$ , let

$$N_z^k := \left\{ K \subseteq \mathbb{Z}_q : |K| = k, \sum_{t \in K} t = z \right\}.$$

Note that  $|N_z^k| = \frac{1}{q} \binom{q}{k}$ , since  $K \in N_0^k$  if and only if  $K + k^{-1}z \in N_z^k$  for every  $z \in \mathbb{Z}_q$ .

For  $K \in \mathbb{N}_z^k$ , let  $E_K$  be the event that  $K \subset \mathcal{A}$ . Let  $X_K$  be the indicator variable of  $E_K$ . We define the random variable

$$X_z = \sum_{K \in N_z^k} X_K.$$

Note that  $X_z$  counts the number of sets of size k which add up to z. We now find  $E[X_z]$ . Since the sum of every subset  $K \subset S$  is in  $\mathbb{Z}_q$ ,

$$\sum_{z \in Z_a} X_z = \binom{s}{k},$$

and so

$$\binom{s}{k} = E\left[\sum_{z \in Z_q} X_z\right] = \sum_{z \in Z_q} E[X_z].$$

As in the argument for finding  $|N_z^k|$ , for every  $z \in \mathbb{Z}_q$ ,

$$E[X_0] = \sum_{K \in N_0^k} E[X_K] = \sum_{K \in N_0^k} E[X_{K+k^{-1}z}] = \sum_{K \in N_z^k} E[X_K] = E[X_z].$$

Therefore, we have that

$$E[X_z] = \frac{1}{q} \binom{s}{k}. \tag{1.1}$$

Now, for  $K, L \in \mathbb{N}_z^k$ , let  $j \in \mathbb{N}$  such that  $j \leq k$  and define

$$\Delta_j := \sum_{|K \cap L| = j} \Pr[E_K \wedge E_L].$$

If  $|K \cap L| = j$ ,

$$\Pr[E_K \wedge E_L] = \frac{\binom{q-2k+j}{s-2k+j}}{\binom{q}{s}}.$$

We can bound the number of events for which  $|K \cap L| = j$ . First we choose K as any set in  $N_z^k$  and then we choose the remaining k - j elements as any subset of  $\mathbb{Z}_q \setminus K$  with size k - j. Thus,

$$\Delta_j \le \frac{1}{q} \binom{q}{k} \binom{q-k}{k-j} \frac{\binom{q-2k+j}{s-2k+j}}{\binom{q}{s}}.$$

This implies that, using 1.1,

$$\begin{split} \frac{\Delta_j}{\mathrm{E}[X_z]^2} &\leq \frac{\binom{q}{k}\binom{q-k}{k-j}\binom{q-2k+j}{s-2k+j}}{\frac{1}{q}\binom{s}{k}\frac{1}{q}\binom{s}{k}q\binom{q}{s}} \\ &= \frac{\frac{q!}{(q-k)!k!}\frac{(p-k)!}{(k-j)!(q-2k+k)!}\frac{(q-2k+j)!}{(s-2k+j)!(q-s)!}}{\frac{1}{q}\binom{s}{k}\frac{s!}{(s-k)!k!}\frac{q!}{(q-s)!s!}} \\ &= \frac{q\binom{s-k}{k-j}}{\binom{s}{k}}. \end{split}$$

Let  $s=4\lfloor 3\log_2 q\rfloor$  and  $k=2\lfloor 3\log_2 q\rfloor$ . Using that  $\binom{s-k}{k-j}$  is maximized at  $k-j=\lfloor (s-k)/2\rfloor$ ,

$$\frac{\Delta_j}{\mathrm{E}[X_z]^2} \le \frac{q^{\binom{2\lfloor 3\log_2 q\rfloor}{\lfloor 3\log_2 q\rfloor}}}{\binom{4\lfloor 3\log_2 q\rfloor}{2\lfloor 3\log_2 q\rfloor}} \le \frac{q}{\binom{2\lfloor 3\log_2 q\rfloor}{\lfloor 3\log_2 q\rfloor}} \le \frac{q}{2^{\lfloor 3\log_2 q\rfloor}} \sim \frac{1}{q^2},$$

since  $\binom{2\lfloor \log_2 q \rfloor}{\lfloor 3\log_2 q \rfloor}^2 \le \binom{4\lfloor 3\log_2 q \rfloor}{2\lfloor 3\log_2 q \rfloor}$  (Proposition A.0.4).

Hence, by (??) and Theorem ??,

$$\Pr[X_z = 0] \le \frac{E[X_z] + \Delta}{E[X_z]^2} = \frac{1}{E[X_z]} + \sum_{j=0}^k \frac{\Delta_j}{E[X_z]^2}$$

$$\leq \frac{1}{E[X_z]} + \frac{(k+1)}{q^2} = \frac{1}{E[X_z]} + \frac{2\lfloor 3\log_2 q \rfloor + 1}{q^2}.$$

Therefore, by the union bound and since  $q \to \infty$  as  $p \to 0$ ,

$$\Pr\left[\bigvee_{z\in\mathbb{Z}_q} X_z = 0\right] \le \frac{q}{E[X_z]} + \frac{2\lfloor 3\log_2 q\rfloor + 1}{q^2} = o(1).$$

We conclude that  $X_z > 0$  for every  $z \in \mathbb{Z}_q$  almost always. Thus, for every  $z \in \mathbb{Z}_q$ , there exists  $K \in N_z^k$  such that  $K \subset \mathcal{A}$  almost always. This means that  $2\lfloor 3\log_2 q\rfloor \mathcal{A}$  covers  $\mathbb{Z}_q$  almost always.

#### 1.1.1 Proof of the upper bound

**Lemma 1.1.2.** Let  $\psi(x)$  be a function for which  $x(\log x)^2 \in o(\psi(x))$ . Then

$$\lim_{p \to 0} \Pr \left[ F(\mathcal{S}) \le \psi \left( \frac{1}{p} \right) \right] = 1.$$

The proof of this theorem consists of several parts. The strategy is to prove that the Ápery set of a subsemigroup of S is completed before step  $\psi\left(\frac{1}{p}\right)$  with high probability, since F(S) is less than the maximum element of this Ápery set. The proof has the following structure:

- 1. First, we will find a step for which a prime q is chosen with high probability  $(E_1)$ .
- 2. Then, in the spirit of Lemma 1.1.1 we will find a step such that a set  $\mathcal{A}$  of s elements which are different modulo q are chosen with high probability  $(E_2)$ .
- 3. Finally, we will apply Lemma 1.1.1 to  $Ap(\langle A \cup \{q\} \rangle, q)$ .

#### Proof.

#### Part 1

Let h(x) be a function such that  $h(x) \in o(x(\log x)^2)$  and  $x \log x \in o(h(x))$ . Let  $t(x) = 20x \log x$ . Consider the event  $E_1$  that there exists a prime  $q \in \mathcal{S}$ , such that

$$t\left(\frac{1}{p}\right) \le q \le h\left(\frac{1}{p}\right).$$

Let  $q_n$  be the *n*-th prime number. By the prime number theorem [1, Theorem 8],

$$q_n \sim n \log n. \tag{1.2}$$

Let k(x) be the number of primes between t(x) and h(x). Now, for sufficiently large n,  $t(n) \leq q_{20n}$ . Also, for every  $c \in \mathbb{R}^+$ ,  $q_{cn} \in o(h(n))$  since  $cn \log cn \in o(h(n))$ . Thus, for sufficiently large x and every  $c \in \mathbb{R}^+$ , k(x) > cx and we get that

$$\lim_{p \to 0} \Pr[\neg E_1] \le \lim_{p \to 0} (1 - p)^{k(1/p)} \le \lim_{p \to 0} (1 - p)^{\frac{c}{p}} = e^{-c}.$$

Therefore,

$$\lim_{n \to 0} \Pr[E_1] = 1. \tag{1.3}$$

#### Part 2

Now, assume  $E_1$ . Then S contains a prime number q for which

$$t\left(\frac{1}{p}\right) \le q \le h\left(\frac{1}{p}\right).$$

Let q be such a prime. Let  $s = 4\lfloor 3\log_2 q\rfloor$ , as in Lemma 1.1.1. Let  $T = \{1, \ldots, q\}$ . Consider the event  $E_2$  that at least s generators are selected in T. Let  $X_1$  be the number of generators selected in T, then  $X_1 \sim \text{Bin}(q, p)$ . We first show that for sufficiently small p, qp > s in order to use a bound of the left tail of the binomial distribution (Proposition A.0.6).

Since

$$q \ge t\left(\frac{1}{p}\right) = \frac{20}{p}\log\frac{1}{p},$$

then

$$qp \ge 20 \log \frac{1}{p}$$
.

Also, since

$$q \le h\left(\frac{1}{p}\right) \le \frac{1}{p} \left(\log \frac{1}{p}\right)^2$$
,

then

$$s = 4\lfloor 3\log_2 q \rfloor \leq 4 \left | 3\log_2 \frac{1}{p} \left(\log \frac{1}{p}\right)^2 \right | = 4 \left \lfloor 3\log_2 \frac{1}{p} + 6\log_2 \log \frac{1}{p} \right \rfloor.$$

Thus, for sufficiently small p, qp > s and we can use Proposition A.0.6 with r = s to show that

$$\Pr[\overline{E_2}|E_1] = \Pr[X_1 < s] \le \frac{(q-s)p}{(qp-s)^2}.$$

Thus, bounding by the worst case asymptotically,

$$\lim_{p \to 0} P[\overline{E_2}|E_1] \le \lim_{p \to 0} \frac{\left(h\left(\frac{1}{p}\right) - 4\left\lfloor 3\log_2 t\left(\frac{1}{p}\right)\right\rfloor\right) p}{\left(t\left(\frac{1}{p}\right)p - 4\left\lfloor 3\log_2 h\left(\frac{1}{p}\right)\right\rfloor\right)^2}$$

$$\le \lim_{p \to 0} \frac{\left(h\left(\frac{1}{p}\right) - 4\left\lfloor 3\log_2 \frac{20}{p}\log\frac{1}{p}\right\rfloor\right) p}{\left(20\log\frac{1}{p} - 4\left\lfloor 3\log_2 \frac{1}{p}\left(\log\frac{1}{p}\right)^2\right\rfloor\right)^2}$$

$$= \lim_{p \to 0} \frac{o\left(\frac{1}{p}\left(\log\frac{1}{p}\right)^2\right) p}{\left(20\log\frac{1}{p} - 4\left\lfloor 3\log_2 \frac{1}{p}\left(\log\frac{1}{p}\right)^2\right\rfloor\right)^2}$$

$$= \lim_{p \to 0} \frac{o\left(\left(\log \frac{1}{p}\right)^2\right)}{\left(20\log \frac{1}{p} - 4\left[3\log_2 \frac{1}{p}\left(\log \frac{1}{p}\right)^2\right]\right)^2} = 0.$$

We conclude that

$$\lim_{n\to 0} \Pr[E_2|E_1] = 1,$$

and so, using (1.3),

$$\lim_{p \to 0} \Pr[E_1 \land E_2] = \lim_{p \to 0} \Pr[E_2 | E_1] \Pr[E_1] = 1.$$

#### Part 3

Finally, assume  $E_1$  and  $E_2$ . Let  $\mathcal{A} = \{Y_1, \dots, Y_s\}$  be a randomly selected subset of size s of the generators selected in T. Since the generators are chosen randomly and |T| = q, we can apply Lemma 1.1.1 to  $\mathbb{Z}_q \cong \operatorname{Ap}(\langle \mathcal{A} \cup \{q\} \rangle, q)$  to get that it will be completed before step

$$qs \le h\left(\frac{1}{p}\right) 2\left\lfloor 3\log_2 h\left(\frac{1}{p}\right) \right\rfloor \in O\left(h\left(\frac{1}{p}\right)\log\frac{1}{p}\right),$$

almost always as  $p \to 0$ .

Thus, if

$$\psi(x) = h(x) \, 2 \, \lfloor 3 \log_2 x \rfloor,$$

we have that  $x(\log x)^2 \in o(\psi(x))$  and

$$\lim_{p \to 0} \Pr \left[ F(\langle \mathcal{A} \cup \{q\} \rangle) \le \psi \left( \frac{1}{p} \right) \right] = 1.$$

Since  $F(S) \leq F(\langle A \cup \{q\} \rangle)$ , we conclude that

$$\lim_{p \to 0} \Pr \left[ F(\mathcal{S}) \le \psi \left( \frac{1}{p} \right) \right] = 1.$$

Since the constraints on h(x) are independent of multiplication by constants, the result is true for any function  $\psi$  such that  $x(\log x)^2 \in \psi(g(x))$ .

The bound on the Frobenius number also implies bounds on the genus and the embedding dimension.

Corollary 1.1.1. Let  $\psi(x)$  be a function for which  $x(\log x)^2 \in o(\psi(x))$ . Then

$$\lim_{p\to 0} \Pr\left[g(\mathcal{S}) \le \psi\left(\frac{1}{p}\right)\right] = 1.$$

**Proof.** Use Proposition ??.

**Corollary 1.1.2.** Let  $\varphi(x)$  be a function for which  $(\log x)^2 \in o(\varphi(x))$ . Then

$$\lim_{p \to 0} \Pr\left[e(\mathcal{S}) \le \varphi\left(\frac{1}{p}\right)\right] = 1.$$

**Proof.** Since

$$\lim_{p\to 0} \Pr\left[F(\mathcal{S}) \leq \psi\left(\frac{1}{p}\right)\right] = 1,$$

and the maximal element of the minimal generating set is at most  $2F(\mathcal{S})$ , the elements of the minimal generating set are chosen before step  $2\psi\left(\frac{1}{p}\right)$  with high probability. Since

$$\left| \mathcal{A} \cap \left\{ 1, \dots, \left| 2\psi\left(\frac{1}{p}\right) \right| \right\} \right| \sim \operatorname{Bin}\left( \left| 2\psi\left(\frac{1}{p}\right) \right|, p \right),$$

by the bound on the right tail of the binomial distribution (Proposition A.0.5), we have that

$$\lim_{p \to 0} \Pr \left[ e(\mathcal{S}) \le (3p)\psi \left(\frac{1}{p}\right) \right] = 1.$$

Thus, if  $\varphi(x) = \frac{3}{x}\psi(x)$ , then  $(\log x)^2 \in \varphi(x)$  and

$$\lim_{p \to 0} \Pr \left[ e(\mathcal{S}) \le \varphi \left( \frac{1}{p} \right) \right] = 1. \quad \Box$$

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### Appendix A

### Useful Bounds

We include some bounds that are useful in the proofs of the main results. By Stirling's Formula, we have that

$$k! \sim \sqrt{2\pi k} \left(\frac{k}{e}\right)^k$$
 (A.1)

**Proposition A.0.1.**  $\binom{n}{k} \le \left(\frac{en}{k}\right)^k$  for  $1 \le k \le n$ .

**Proof.** Using (A.1), we have that, for  $k \ge 1$ ,

$$k! \ge \left(\frac{k}{e}\right)^k.$$

Then

$$\binom{n}{k} \le \frac{n^k}{\left(\frac{k}{e}\right)^k} = \left(\frac{en}{k}\right)^k. \quad \Box$$

**Proposition A.0.2.**  $\left(\frac{n}{k}\right)^k \geq \binom{n}{k}$  for  $1 \leq k \leq n$ .

Proof.

$$\binom{n}{k} = \prod_{i=0}^{k-1} \frac{n-i}{k-i} \ge \left(\frac{n}{k}\right)^k. \quad \Box$$

**Proposition A.0.3.**  $(1 - p) \le e^{-p} \text{ for } 0 \le p \le 1.$ 

**Proof.** The Taylor series of  $e^{-p}$  is alternating with a decreasing sequence, so

$$e^{-p} = 1 - p + \frac{p^2}{2!} - \frac{p^3}{3!} + \dots \ge 1 - p.$$

We also give a combinatorial proof of the following result.

**Proposition A.0.4.**  $\binom{2n}{k}^2 \leq \binom{4n}{2k}$  for  $n \geq 1$ .

**Proof.** The number of subsets of size 2k of a set of size 4n is  $\binom{4n}{2k}$ . This is greater than the number of subsets that can be expressed as the product of two subsets of size k of a set of size 2n, which is  $\binom{2n}{k}^2$ .

The proof of the following bound can be found in [2, Section 6.3].

**Proposition A.0.5.** Let  $X \sim \text{Bin}(n, p)$ . If r > np,

$$\Pr[X \ge r] \le \frac{r(1-p)}{(r-np)^2}.$$

Since the binomial distribution is symmetric, we also have the following.

**Proposition A.0.6.** Let  $X \sim \text{Bin}(n, p)$ . If r < np,

$$\Pr[X \le r] \le \frac{(n-r)p}{(np-r)^2}.$$