

Regularization

- L0 - Count of the non-zero weights – Difficult minimize
- L1 Lasso - Sum of the attribute weights
 - It set the weights to 0 for irrelevant attributes

$$\sum_{i=1}^n (Y_i - \sum_{j=1}^p X_{ij} \beta_j)^2 + \lambda \sum_{j=1}^p |\beta_j|$$

- L2 Ridge (Squared Euclidean norm) - Sum of the squared attribute weights. Used to avoid Overfitting.

$$\sum_{i=1}^n (y_i - \sum_{j=1}^p x_{ij} \beta_j)^2 + \lambda \sum_{j=1}^p \beta_j^2$$

Scaling:

- ◆ Min/Max normalization: $x^{new} = \frac{x - x_{min}}{x_{max} - x_{min}} (x_{max}^{new} - x_{min}^{new}) + x_{min}^{new}$
- ◆ Z-Score normalisierung: $x^{new} = \frac{x - \mu_x}{\sigma_x}$
- ◆ Decimal scaling: $x^{new} = |x| \cdot 10^{-a} \quad a = \max_x \{i \in \mathbb{Z} \mid |x| \cdot 10^i < 1\}$
- ◆ Logarithmic scaling: $x^{new} = \log_a x$

Information Gain

Entropy is the expected information of a message.

$$H(y) = -\sum_{v=1}^k p(y=v) \log_2 p(y=v)$$

Loan	x_1 (Credit report)	x_2 (Employment last 3 months)	x_3 (Collateral > 50% loan)	y (Paid back in full)
1	Positive	Yes	No	Yes
2	Positive	No	Yes	Yes
3	Positive	No	No	No
4	Negative	No	Yes	No
5	Negative	Yes	No	No

$$H_L(y|x_1=n) = -\frac{2}{2} \log_2 \frac{2}{2} - \frac{0}{2} \log_2 \frac{0}{2} = 0 \text{ bit}$$

ID3

ID3(L, X)

1. If all data in L have same class y or X={}, then return leaf node with majority class y.
2. Else
 1. For all attributes $x_j \in X$, calculate split criterion $G_L(x_j)$ or $GR_L(x_j)$.
 2. Choose attribute $x_j \in X$ with highest $G_L(x_j)$ or $GR_L(x_j)$.
 3. Let $L_i = \{(x, y) \in L : x_j = i\}$.
 4. Return test node with attribute x_j and children $ID3(L_1, X \setminus x_j), \dots, ID3(L_k, X \setminus x_j)$.

Information Gain of an Attribute

- Reduction of entropy by splitting the data along an attribute
 - ◆ $G_L(x_j) = H_L(y) - \sum_{v=1}^k p_L(x_j=v) H_L(y|x_j=v)$

Loan	x_1 (Credit report)	x_2 (Employment last 3 months)	x_3 (Collateral > 50% loan)	y (Paid back in full)
1	Positive	Yes	No	Yes
2	Positive	No	Yes	Yes
3	Positive	No	No	No
4	Negative	No	Yes	No
5	Negative	Yes	No	No

- $G_L(x_1) = H_L(y) - p_L(x_1=p) H_L(y|x_1=p) - p_L(x_1=n) H_L(y|x_1=n)$
 $= 0.97 - \frac{3}{5} 0.91 - \frac{2}{5} 0 = 0.42 \text{ bit}$
- Splitting along x_1 reduces the uncertainty regarding the class label y by 0.42 bit.

Idea: factor the information that is contained in the attribute values into the decision

$$H_L(x_j) = -\sum_{v=1}^k p_L(x_j=v) \log_2 p_L(x_j=v)$$

Information gain ratio:

$$GR_L(x_j) = \frac{G_L(x_j)}{H_L(x_j)}$$

C4.5

Learning Decision Trees with C4.5

C4.5(L)

1. If all data in L have same class y or are identical, then return leaf node with majority class y.
2. Else
 1. For all discrete attributes $x_j \in X$: calculate $G_L(x_j)$.
 2. For all continuous attributes $x_j \in X$ and all values v that occur for x_j in L: calculate $G_L(x_j \leq v)$.
 3. If discrete attribute has highest $G_L(x_j)$:
 1. Let $L_i = \{(x, y) \in L : x_j = i\}$.
 2. Return test node with attribute x_j and children $C4.5(L_1), \dots, C4.5(L_k)$.
 4. If continuous attribute has highest $G_L(x_j \leq v)$:
 1. Let $L_{\leq} = \{(x, y) \in L : x_j \leq v\}$, $L_{>} = \{(x, y) \in L : x_j > v\}$
 2. Return test node with test $x_j \leq v$ and children $C4.5(L_{\leq}), C4.5(L_{>})$.

Regression Trees

- Variance of the target attribute on sample L :

$$\text{Var}(L) = \frac{1}{n} \sum_{i=1}^n (y_i - \bar{y})^2$$
- Variance = MSE of predicting the mean value.
- Splitting criterion: variance reduction of $[x_j \leq v]$:

$$R_L[x_j \leq v] = \text{Var}(L) - \frac{n_{[x_j \leq v]}}{n} \text{Var}(L_{[x_j \leq v]}) - \frac{n_{[x_j > v]}}{n} \text{Var}(L_{[x_j > v]})$$
- Stopping criterion:

$$\text{Do not create a new test node if } n\text{Var}(L) \leq \tau.$$

CART (L)

- If $\sum_{i=1}^n (y_i - \bar{y})^2 < \tau$, then return leaf node with prediction \bar{y} .
- Else
 - For all discrete attributes $x_j \in X$: calculate $R_L(x_j)$.
 - For all continuous attributes $x_j \in X$ and all values v that occur for x_j in L : calculate $R_L(x_j \leq v)$.
 - If discrete attribute has highest $R_L(x_j)$:
 - Let $L_i = \{(x, y) \in L: x_j = i\}$.
 - Return test node with attribute x_j and children $\text{CART}(L_1), \dots, \text{CART}(L_k)$.
 - If continuous attribute has highest $R_L(x_j \leq v)$:
 - Let $L_{\leq} = \{(x, y) \in L: x_j \leq v\}$, $L_{>} = \{(x, y) \in L: x_j > v\}$
 - Return test node with test $x_j \leq v$ and children $\text{CART}(L_{\leq}), \text{CART}(L_{>})$.

Random Forest

- Input: sample L of size n , attributes X .
- For $i=1 \dots k$
 - Draw n instances uniformly with replacement from L into set L_i .
 - Draw m attributes from all attributes X into X_i .
 - Learn model f_i on sample L_i using attributes X_i .
- For classification:
 - Let $f(x)$ be the majority vote among $(f_1(x), \dots, f_k(x))$.
- For regression:
 - Let $f(x) = \frac{1}{k} \sum_{i=1}^k f_i(x)$.

Loss in Classification

- Zero-one loss:

$$\ell_{0/1}(f_{\theta}(x_i), y_i) = \begin{cases} 1 & \text{sign}(f_{\theta}(x_i)) \neq y_i \\ 0 & \text{sign}(f_{\theta}(x_i)) = y_i \end{cases}$$

- Logistic loss:

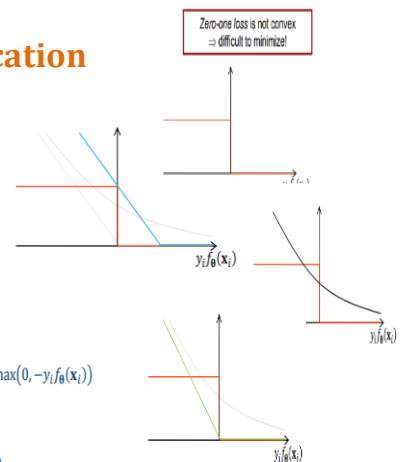
$$\ell_{\log}(f_{\theta}(x_i), y_i) = \log(1 + e^{-y_i f_{\theta}(x_i)})$$

- Perceptron loss:

$$\ell_p(f_{\theta}(x_i), y_i) = \begin{cases} -y_i f_{\theta}(x_i) & -y_i f_{\theta}(x_i) > 0 \\ 0 & -y_i f_{\theta}(x_i) \leq 0 \end{cases} = \max(0, -y_i f_{\theta}(x_i))$$

- Hinge loss:

$$\ell_h(f_{\theta}(x_i), y_i) = \begin{cases} 1 - y_i f_{\theta}(x_i) & 1 - y_i f_{\theta}(x_i) > 0 \\ 0 & 1 - y_i f_{\theta}(x_i) \leq 0 \end{cases} = \max(0, 1 - y_i f_{\theta}(x_i))$$



Gradient descent

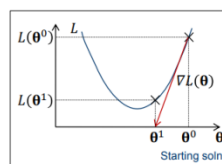
Linear classification model: minimize

$$L(\theta) = \sum_{i=1}^n \ell(x_i^T \theta, y_i) + \lambda \Omega(\theta)$$

Gradient descent method:

```

RegERM(Data: (x1, y1), ..., (xn, yn))
Set θ0 = 0 and t = 0
DO
  Compute gradient ∇L(θt)
  Compute step size αt
  Set θt+1 = θt - αt ∇L(θt)
  Set t = t + 1
WHILE ||θt - θt+1|| > ε
RETURN θt
  
```



with Step size (line search)

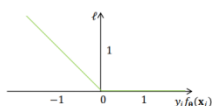
Determine step size through line search:

```

RegERM-LineSearch(Data: (x1, y1), ..., (xn, yn))
Set θ0 = 0 and t = 0
DO
  Compute gradient ∇L(θt)
  Choose step size αt:
    αt = argmin_{α>0} L(θt - α ∇L(θt))
  Set θt+1 = θt - αt ∇L(θt)
  Set t = t + 1
WHILE ||θt - θt+1|| > ε
RETURN θt
  
```

Loss function:

$$\ell_p(f_{\theta}(x_i), y_i) = \begin{cases} -y_i f_{\theta}(x_i) & -y_i f_{\theta}(x_i) > 0 \\ 0 & -y_i f_{\theta}(x_i) \leq 0 \end{cases} = \max(0, -y_i f_{\theta}(x_i))$$



Stochastic Gradient descent

RegERM-Stoch(Data: (x1, y1), ..., (xn, yn))

```

Set θ0 = 0 and t = 0
DO
  Shuffle data randomly
  FOR i = 1, ..., n
    Compute subset gradient ∇xi L(θt)
    Compute step size αt
    Set θt+1 = θt - αt ∇xi L(θt)
    Set t = t + 1
  END
WHILE ||θt - θt+1|| > ε
RETURN θt
  
```

Perceptron

Perceptron(Instances {(xi, yi)})

```

Set θ = 0
DO
  FOR i = 1, ..., n
    IF yi f_{\theta}(xi) ≤ 0
      THEN θ = θ + yi xi
    END
  WHILE θ changes
  RETURN θ
  
```

Stochastic gradient method:

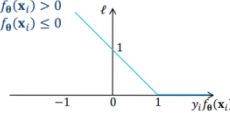
$$\nabla L_{x_i}(\theta) = \begin{cases} -y_i x_i & -y_i f_{\theta}(x_i) > 0 \\ 0 & -y_i f_{\theta}(x_i) \leq 0 \end{cases}$$

Loss function:

$$\ell_h(f_\theta(\mathbf{x}_i), y_i) = \begin{cases} 1 - y_i f_\theta(\mathbf{x}_i) & \text{if } 1 - y_i f_\theta(\mathbf{x}_i) > 0 \\ 0 & \text{if } 1 - y_i f_\theta(\mathbf{x}_i) \leq 0 \end{cases} = \max(0, 1 - y_i f_\theta(\mathbf{x}_i))$$

Regularizer:

$$\Omega_2(\theta) = \theta^T \theta = \sum_{j=1}^m |\theta_j|^2 = \|\theta\|_2^2$$



SVM

Loss function is 0, if...

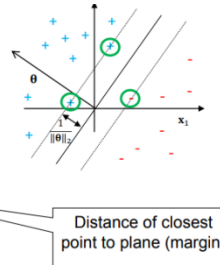
$$\sum_{i=1}^n \max(0, 1 - y_i f_\theta(\mathbf{x}_i)) = 0$$

$$\Leftrightarrow \forall_{i=1}^n: y_i f_\theta(\mathbf{x}_i) \geq 1$$

$$\Leftrightarrow \forall_{i=1}^n: y_i \mathbf{x}_i^T \theta \geq 1$$

$$\Leftrightarrow \forall_{i=1}^n: y_i \mathbf{x}_i^T \frac{\theta}{\|\theta\|_2} \geq \frac{1}{\|\theta\|_2}$$

$$\Leftrightarrow \forall_{i=1}^n: \mathbf{x}_i^T \frac{\theta}{\|\theta\|_2} \begin{cases} \geq \frac{1}{\|\theta\|_2} & \text{if } y_i = +1 \\ \leq -\frac{1}{\|\theta\|_2} & \text{if } y_i = -1 \end{cases}$$



Linear classification model: minimize

$$L(\theta) = \sum_{i=1}^n \left[\max(0, 1 - y_i \mathbf{x}_i^T \theta) + \frac{\lambda}{n} \theta^T \theta \right]$$

Gradient:

$$\nabla L(\theta) = \sum_{i=1}^n \nabla_{\mathbf{x}_i} L(\theta)$$

Stochastic gradient for \mathbf{x}_i :

$$\nabla_{\mathbf{x}_i} L(\theta) = \begin{cases} \frac{2\lambda}{n} \theta & \text{if } y_i \mathbf{x}_i^T \theta > 1 \\ \frac{2\lambda}{n} \theta - y_i \mathbf{x}_i & \text{if } y_i \mathbf{x}_i^T \theta < 1 \end{cases}$$

Linear Regression

Loss Functions for Regression

Absolute loss:

$$\ell_{\text{abs}}(f_\theta(\mathbf{x}_i), y_i) = |f_\theta(\mathbf{x}_i) - y_i|$$

Squared loss:

$$\ell_2(f_\theta(\mathbf{x}_i), y_i) = (f_\theta(\mathbf{x}_i) - y_i)^2$$

ϵ -insensitive loss:

$$\ell_\epsilon(f_\theta(\mathbf{x}_i), y_i) = \begin{cases} |f_\theta(\mathbf{x}_i) - y_i| - \epsilon & |f_\theta(\mathbf{x}_i) - y_i| - \epsilon > 0 \\ 0 & |f_\theta(\mathbf{x}_i) - y_i| - \epsilon \leq 0 \end{cases}$$

Regularizer for Regression

L1 regularization:

$$\Omega_1(\theta) \propto \|\theta\|_1 = \sum_{j=1}^m |\theta_j|$$

L2 regularization:

$$\Omega_2(\theta) \propto \|\theta\|_2^2 = \sum_{j=1}^m \theta_j^2$$

Special Cases

Lasso: squared loss + L1 regularization

$$L(\theta) = \sum_{i=1}^n \ell_2(f_\theta(\mathbf{x}_i), y_i) + \lambda \|\theta\|_1$$

Ridge regression: squared loss + L2 regularization

$$L(\theta) = \sum_{i=1}^n \ell_2(f_\theta(\mathbf{x}_i), y_i) + \lambda \|\theta\|_2^2$$

Elastic net: squared loss, L1 + L2 regularization

$$L(\theta) = \sum_{i=1}^n \ell_2(f_\theta(\mathbf{x}_i), y_i) + \lambda \|\theta\|_2^2 + \lambda' \|\theta\|_1$$

Ridge

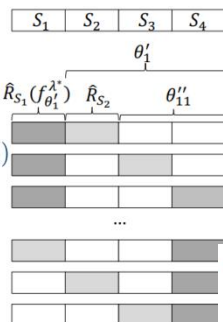
$$\theta = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}$$

Lasso

$$L(\theta) = (\mathbf{X}\theta - \mathbf{y})^T (\mathbf{X}\theta - \mathbf{y}) + \lambda \|\theta\|_1$$

Nested Cross Validation

- For $i = 1 \dots k$
 - Iterate over values λ
 - For $j = 1 \dots k \setminus i$
 - Train $f_{\theta_{ij}}^\lambda$ on $S \setminus S_i \setminus S_j$
 - Determine $\hat{R}_{S_j}(f_{\theta_{ij}}^\lambda)$
 - Average \hat{R}_{S_j} to determine $\hat{R}_{S \setminus S_i}(f_{\theta_i}^\lambda)$
 - Choose λ_i^* that minimizes $\hat{R}_{S \setminus S_i}(f_{\theta_i}^\lambda)$
 - Train $f_{\theta_i}^{\lambda_i^*}$ on $S \setminus S_i$
 - Determine $\hat{R}_{S_i}(f_{\theta_i}^{\lambda_i^*})$
- Average $\hat{R}_{S_i}(f_{\theta_i}^{\lambda_i^*})$ to determine $\hat{R}_S(f_{\theta}^{\lambda^*})$
- Determine λ^* by averaging λ_i^*
- Train $f_{\theta}^{\lambda^*}$ on S
- Return $f_{\theta}^{\lambda^*}$ and $\hat{R}_S(f_{\theta}^{\lambda^*})$



F_α measures combine precision and recall values into single value:

$$F_\alpha = \frac{n_{TP}}{\alpha(n_{TP} + n_{FP}) + (1 - \alpha)(n_{TP} + n_{FN})}$$

- $\alpha = 1$: Precision
- $\alpha = 0$: Recall
- $\alpha = 0.5$: "F-measure", harmonic mean of precision and recall.
- Alternative definition: F_β measures.

$$\text{Relationship: } \alpha = \frac{1}{1 + \beta}$$

Precision & Recall

- TP (True Positive) - Predicted & Actuals are versicolor
- TN (True Negative) - Predicted & Actuals are not versicolor
- FP (False Positive) - Predicted is versicolor but actual is not versicolor
- FN (False Negative) - Predicted is not versicolor but actual is versicolor

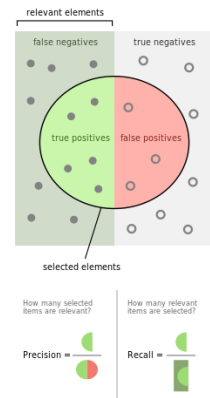
$$\text{Accuracy} = (TP + TN) / (TP + TN + FP + FN)$$

$$\text{Precision} = TP / (TP + FP)$$

$$\text{Predicted \& Actuals are versicolor} / (\text{Predicted \& Actuals are versicolor}) + (\text{Predicted is versicolor but actual is not versicolor})$$

$$\text{Recall} = TP / (TP + FN)$$

$$\text{Predicted \& Actuals are versicolor} / (\text{Predicted \& Actuals are versicolor}) + (\text{Predicted is not versicolor but actual is versicolor})$$

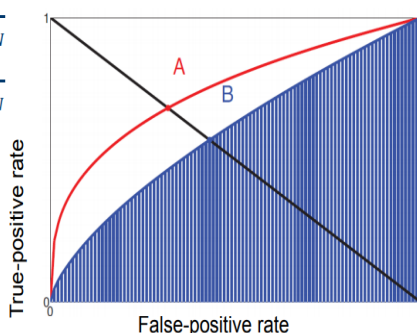


Area under the ROC curve (AUC):

- Let \mathbf{x}_+ be a randomly drawn positive instance.
- Let \mathbf{x}_- be a randomly drawn negative instance.
- $AUC(\theta) = P(f_\theta(\mathbf{x}_+) > f_\theta(\mathbf{x}_-))$.

$$\hat{r}_{TP} = \frac{n_{TP}}{n_{TP} + n_{FN}}$$

$$\hat{r}_{FP} = \frac{n_{FP}}{n_{FP} + n_{TN}}$$



Neural Network

Softmax Activation

One output unit per class:

- $x_k^d = \sigma_{sm}(h_k^d) = \frac{e^{h_k^d}}{\sum_{k'} e^{h_{k'}^d}}$
- x_k^d : predicted probability for class k .

Softmax activation function:

- $x_k^d = \sigma_{sm}(h_k^d) = \frac{e^{h_k^d}}{\sum_{k'} e^{h_{k'}^d}}$
- $\frac{\partial \sigma_{sm}(h_k^d)}{\partial h_k^d} = \sigma_{sm}(h_k^d)(1 - \sigma_{sm}(h_k^d))$

Cost function:

- $\ell(y, x^d) = \sum_k y_k \log x_k^d$
- $\frac{\partial \ell(y, x^d)}{\partial h_k^d} = x_k^d - y_k$

Linear Activation

Linear:

- $x^d = h^d$.
- Output unbounded.

Linear activation function:

- $x_k^d = \sigma_s(h_k^d) = h_k^d$
- $\frac{\partial \sigma_s(h_k^d)}{\partial h_k^d} = 1$

Cost function:

- $\ell(y, x^d) = \frac{1}{2} \sum_k (x_k^d - y_k)^2$
- $\frac{\partial \ell(y, x^d)}{\partial x_k^d} = x_k^d - y_k$

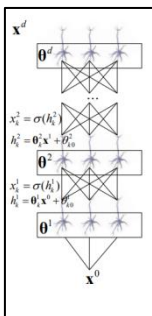
Rectified Linear Units

$$x_k^i = \sigma_{ReLU}(h_k^i) = \max(0, h_k^i)$$

Rectified linear activation function:

- $x_k^i = \sigma_{ReLU}(h_k^i) = \max(0, h_k^i)$
- $\frac{\partial \sigma_{ReLU}(h_k^i)}{\partial h_k^i} = \begin{cases} 1 & \text{if } h_k^i > 0 \\ 0 & \text{otherwise} \end{cases}$

Back Propagation



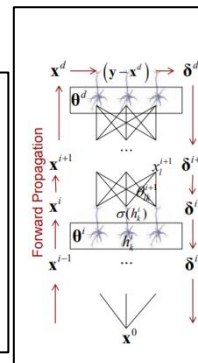
- Loss function $\hat{R}(\theta) = \frac{1}{2m} \sum_{j=1}^m \ell(y_j, x^d)$
- Gradient descent:
 - $\theta' = \theta - \alpha \nabla_{\theta} \hat{R}(\theta) = \theta - \alpha \frac{\partial \hat{R}(\theta)}{\partial \theta}$
 - $= \theta - \frac{\alpha}{2m} \sum_{j=1}^m \frac{\partial \ell(y_j, x^d)}{\partial \theta}$
- Stochastic gradient for instance x_j :
 - $\theta' = \theta - \alpha \nabla_{\theta} \ell(y_j, x^d)$
 - $= \theta - \alpha \frac{\partial \ell(y_j, x^d)}{\partial \theta}$

- Stochastic gradient for output units for instance x_j :

$$\frac{\partial \ell(y_j, x^d)}{\partial \theta_k^d} = \frac{\partial \ell(y_j, x^d)}{\partial x^d} \frac{\partial x^d}{\partial h_k^d} \frac{\partial h_k^d}{\partial \theta_k^d}$$

$$= \frac{\partial \ell(y_j, x^d)}{\partial x^d} \frac{\partial \sigma(h_k^d)}{\partial h_k^d} x^{d-1} = \delta_k^d x^{d-1}$$
- With

$$\delta_k^d = \frac{\partial \ell(y_j, x^d)}{\partial x^d} \frac{\partial \sigma(h_k^d)}{\partial h_k^d}$$



- Stochastic gradient for hidden units for instance x_j :

$$\frac{\partial \ell(y_j, x^d)}{\partial \theta_k^i} = \frac{\partial \ell(y_j, x^d)}{\partial h_k^i} \frac{\partial h_k^i}{\partial \theta_k^i} = \delta_k^i x^{i-1}$$
- With

$$\delta_k^i = \frac{\partial \ell(y_j, x^d)}{\partial h_k^i} \frac{\partial h_k^i}{\partial \theta_k^i}$$

$$= \frac{\partial \ell(y_j, x^d)}{\partial (x_1^{i+1}, \dots, x_{n_{i+1}}^{i+1})} \frac{\partial (x_1^{i+1}, \dots, x_{n_{i+1}}^{i+1})}{\partial h_k^i}$$

$$= \sum_{l=1}^{n_{i+1}} \frac{\partial \ell(y_j, x^d)}{\partial h_l^{i+1}} \frac{\partial h_l^{i+1}}{\partial x_k^i} \frac{\partial x_k^i}{\partial h_k^i}$$

$$= \sum_{l=1}^{n_{i+1}} \delta_l^{i+1} \theta_{lk}^{i+1} \frac{\partial \sigma(h_k^i)}{\partial h_k^i}$$

Back Propagation: Algorithm

- Iterate over training instances (x, y) :
 - Forward propagation: for $i=0 \dots d$:
 - For $k=1 \dots n_i$: $h_k^i = \theta_{k1}^i x^{i-1} + \theta_{k0}^i$
 - $x^i = \sigma(h^i)$
 - Back propagation:
 - For $k=1 \dots n_i$: $\delta_k^i = \frac{\partial \ell(y, x^d)}{\partial h_k^i} \frac{\partial h_k^i}{\partial x_k^i}$
 - $\theta_{k1}^{i+1} = \theta_{k1}^i - \alpha \delta_k^i x^{i-1}$
 - For $k=1 \dots n_i$: $\delta_k^i = \sigma'(h_k^i) \sum_{l=1}^{n_{i+1}} \delta_l^{i+1} \theta_{lk}^{i+1}$
 - $\theta_{k0}^{i+1} = \theta_{k0}^i - \alpha \delta_k^i x^{i-1}$
- Until convergence

Normal initialization with:

- Draw from $N\left[0, \sqrt{\frac{2}{n_{i+1} + n_i}}\right]$.

n is number of layers

Uniform initialization (Glorot initialization):

- Draw from $U\left[-\frac{6}{n_{i+1} + n_i}, \frac{6}{n_{i+1} + n_i}\right]$.

Parallel Inference - weight calculation

CNN

$$h^i = \theta^i x^{i-1}$$

$$\begin{bmatrix} h_1^i \\ \vdots \\ h_{n_i}^i \end{bmatrix} = \begin{bmatrix} \theta_{11}^i & \dots & \theta_{1n_{i-1}}^i \\ \vdots & & \vdots \\ \theta_{n_i1}^i & \dots & \theta_{n_i n_{i-1}}^i \end{bmatrix} \begin{bmatrix} x_1^{i-1} \\ \vdots \\ x_{n_{i-1}}^{i-1} \end{bmatrix}$$

Keep $[x_1^{i-1}, \dots, x_{n_{i-1}}^{i-1}]$ in cache.

For all rows $j = 1..n_i$ (in parallel):

- Load $[\theta_{j1}^i, \dots, \theta_{jn_{i-1}}^i]$ into cache.
- For all $k = 1..n_{i-1}$ (in parallel): multiply and sum $\theta_{jk}^i x_k^{i-1}$.

Convolutional Layers

- Convolution, $k \times k \times d$, stride > 1 .
 - Input size: $x \times y \times d'$, stride s .
 - output size: $\frac{(x-k+1)}{s} \times \frac{(y-k+1)}{s} \times d$.
- Convolutional layer has
 - $k \times k \times d'$ parameters.
- Decreases the spatial resolution.

Bayesian Learning

Bayes' equation:

$$P(\theta|X, y) = \frac{P(y|X, \theta)P(\theta)}{P(y|X)}$$

Likelihood of observing $y|X$ when model parameter is θ .

A priori ("prior") probability of nature choosing θ

Probability of observing $y|X$; independent of θ .

A posteriori ("posterior") probability that θ is the correct parameter given observations $y|X$.

- Maximum-likelihood (ML) model:
 - $\theta_{ML} = \arg \max_{\theta} P(y|X, \theta)$.
- Maximum-a-posteriori (MAP) model:
 - $\theta_{MAP} = \arg \max_{\theta} P(\theta|y, X) = \arg \max_{\theta} \frac{P(y|X, \theta)P(\theta)}{P(y|X)}$

$$= \arg \max_{\theta} P(y|X, \theta)P(\theta)$$

Posterior \propto likelihood \times prior

- Most likely value y^* for new input x^* (Bayes-optimal decision):
 - $y^* = \arg \max_y P(y|x^*, y, X)$
 - $P(y^*|x^*, y, X) = \int P(y^*, \theta|x^*, y, X)d\theta$

$$= \int P(y^*|x^*, \theta)P(\theta|y, X)d\theta$$

Predictive distribution

"Bayesian model averaging". Often computationally infeasible, but has a closed-form solution in some cases.

Linear Regression:

Maximum Likelihood Model

- Assumption 1: Nature generates parameter θ^* of a linear function $f_{\theta^*}(x) = x^T \theta^*$ according to $p(\theta)$.
- Assumption 2: Given inputs X , nature generates outputs y :
 - $y_i = f_{\theta^*}(x_i) + \epsilon_i$ with $\epsilon_i \sim N(\epsilon|0, \sigma^2)$.
 - $p(y_i|x_i, \theta^*) = N(y_i|x_i^T \theta^*, \sigma^2)$

Maximum-likelihood (ML) model:

$$\theta_{ML} = \arg \min_{\theta} \sum_{i=1}^n (y_i - x_i^T \theta)^2$$

- Known as least-squares method in statistics.

- Maximum-likelihood (ML) model:
 - $\theta_{ML} = \arg \max_{\theta} P(y|X, \theta) = \arg \max_{\theta} \prod_{i=1}^n N(y_i|x_i^T \theta, \sigma^2)$

$$= \arg \max_{\theta} \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2\sigma^2}(y_i - x_i^T \theta)^2\right\}$$

$$= \arg \min_{\theta} \sum_{i=1}^n (y_i - x_i^T \theta)^2$$
- Log is a monotononic transformation:
 - $\arg \max_{\theta} P(y|X, \theta) = \arg \max_{\theta} \log P(y|X, \theta)$
- Also constant terms (constant in θ) can be dropped

Unregularized linear regression with squared loss

Maximum-a-posteriori model

$$\theta_{MAP} = \arg \max_{\theta} P(\theta|y, X) = \arg \max_{\theta} \frac{P(y|X, \theta)P(\theta)}{P(y|X)}$$

Maximum-a-posteriori (MAP) model:

$$\theta_{MAP} = \arg \min_{\theta} \sum_{i=1}^n (y_i - x_i^T \theta)^2 - \frac{\sigma^2}{\sigma_p^2} \theta^T \theta$$

Same optimization criterion as ridge regression.

Analytic solution (see lecture on ridge regression):

$$\theta_{MAP} = \left(X^T X + \frac{\sigma^2}{\sigma_p^2} I \right)^{-1} X^T y$$

Bayes Optimal

Bayes-optimal decision is made by a weighted sum over all model parameters:

$$P(y|x^*, y, X) = \int P(y|x^*, \theta)P(\theta|y, X)d\theta$$

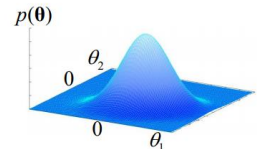
Nature generates parameter θ^* of linear model $f_{\theta^*}(x) = x^T \theta^*$ according to $p(\theta)$.

For convenience, assume $p(\theta) = N(\theta|0, \sigma_p^2 I)$.

$$p(\theta) = N(\theta|0, \sigma_p^2 I)$$

$$= \frac{1}{2\pi^{m/2} \sigma_p^m} \exp\left(-\frac{1}{2\sigma_p^2} |\theta|^2\right)$$

$\sigma_p^2 \in \square$ controls strength of prior



Sequential Learning

- Training examples arrive sequentially.
- Each training example (x_i, y_i) changes prior $p_{i-1}(\theta)$ into posterior $p_i(\theta)$ which becomes the new prior $p_i(\theta)$
 - $P(\theta|y, X) = \frac{1}{Z} P_0(\theta) P(y|X, \theta)$

$$= \frac{1}{Z} P_0(\theta) \prod_{i=1}^n P(y_i|x_i, \theta)$$

$$= \frac{1}{Z} \underbrace{P_0(\theta) P(y_1|x_1, \theta)}_{P_1(\theta)} \underbrace{P(y_2|x_2, \theta) P(y_3|x_3, \theta) \dots P(y_n|x_n, \theta)}_{P_n(\theta)}$$

Prediction

- Predictive distribution for linear regression
 - $P(y|x^*, y, X) = \int P(y|x^*, \theta)P(\theta|y, X)d\theta$

$$= \int N(y|x^*, \theta)N(\theta|\bar{\theta}, A^{-1})d\theta$$

$$= N(y|\bar{\theta}^T x^*, \sigma^2 + x^{*T} A^{-1} x^*)$$
 - With $\bar{\theta} = \left(X^T X + \frac{\sigma^2}{\sigma_p^2} I \right)^{-1} X^T y$
 - And $A^{-1} = \sigma^{-2} X^T X + \sigma_p^{-2} I$.
- Bayes-optimal prediction:
 - $y^* = \arg \max_y P(y|x^*, y, X) = \bar{\theta}^T x^*$

Linear Classification

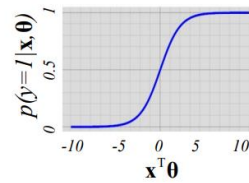
Written jointly for both classes:

$$\diamond P(y|\mathbf{x}, \boldsymbol{\theta}) = \sigma(\mathbf{y}\mathbf{x}^T\boldsymbol{\theta}) = \frac{1}{1+e^{-\mathbf{y}\mathbf{x}^T\boldsymbol{\theta}}}$$

$$\diamond P(y = +1|\mathbf{x}, \boldsymbol{\theta}) = \sigma(\mathbf{x}^T\boldsymbol{\theta}) = \frac{1}{1+e^{-\mathbf{x}^T\boldsymbol{\theta}}}$$

$$\diamond P(y = -1|\mathbf{x}, \boldsymbol{\theta}) = 1 - P(y = +1|\mathbf{x}, \boldsymbol{\theta})$$

Sigmoid function maps $[-\infty, +\infty] \rightarrow [0, 1]$.



Maximum-likelihood model

- Maximum-likelihood model:

$$\begin{aligned} \diamond \boldsymbol{\theta}_{ML} &= \arg \max_{\boldsymbol{\theta}} P(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta}) \\ &= \arg \max_{\boldsymbol{\theta}} \prod_{i=1}^n \frac{1}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} \\ &= \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n -\log \frac{1}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} \\ &= \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) \end{aligned}$$

- No analytic solution; numeric optimization, for instance, using (stochastic) gradient descent.

Maximum-likelihood model:

$$\diamond \boldsymbol{\theta}_{ML} = \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}})$$

Gradient:

$$\begin{aligned} \diamond \frac{\partial}{\partial \boldsymbol{\theta}} \sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) \\ &= \sum_{i=1}^n \frac{\partial}{\partial (1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}})} \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) \frac{\partial}{\partial (-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta})} (1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) \frac{\partial}{\partial \boldsymbol{\theta}} (-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}) \\ &= \sum_{i=1}^n \frac{1}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}} (-\mathbf{y}_i \mathbf{x}_i^T) = \sum_{i=1}^n -\mathbf{y}_i \mathbf{x}_i^T \frac{e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} \\ &= \sum_{i=1}^n -\mathbf{y}_i \mathbf{x}_i^T \frac{1}{1 + e^{\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} \\ &= \sum_{i=1}^n \mathbf{y}_i \mathbf{x}_i (1 - \sigma(\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta})) \end{aligned}$$

Maximum-a-posteriori model

- Maximum-a-posteriori model with prior $P(\boldsymbol{\theta}) = N(\boldsymbol{\theta}|\mathbf{0}, \sigma^2 \mathbf{I})$:

$$\begin{aligned} \diamond \boldsymbol{\theta}_{MAP} &= \arg \max_{\boldsymbol{\theta}} P(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta}) P(\boldsymbol{\theta}) \\ &= \arg \max_{\boldsymbol{\theta}} \prod_{i=1}^n \frac{1}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} N(\boldsymbol{\theta}|\mathbf{0}, \sigma^2 \mathbf{I}) \\ &= \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n -\log \frac{1}{1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}} - \log N(\boldsymbol{\theta}|\mathbf{0}, \sigma^2 \mathbf{I}) \\ &= \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) + \frac{1}{2\sigma^2} \boldsymbol{\theta}^T \boldsymbol{\theta} \end{aligned}$$

- No analytic solution; numeric optimization, for instance, using (stochastic) gradient descent.

- Maximum-a-posteriori model with prior $P(\boldsymbol{\theta}) = N(\boldsymbol{\theta}|\mathbf{0}, \sigma^2 \mathbf{I})$:

$$\diamond \boldsymbol{\theta}_{MAP} = \arg \min_{\boldsymbol{\theta}} \sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) + \frac{1}{2\sigma_p^2} \boldsymbol{\theta}^T \boldsymbol{\theta}$$

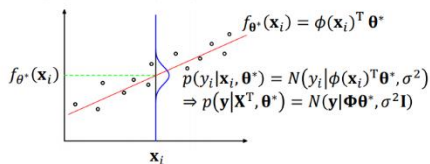
- Gradient:

$$\begin{aligned} \diamond \frac{\partial}{\partial \boldsymbol{\theta}} \left(\sum_{i=1}^n \log(1 + e^{-\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta}}) + \frac{1}{2\sigma_p^2} \boldsymbol{\theta}^T \boldsymbol{\theta} \right) \\ = \mathbf{y}_i \mathbf{x}_i (1 - \sigma(\mathbf{y}_i \mathbf{x}_i^T \boldsymbol{\theta})) + \frac{1}{2\sigma_p^2} \boldsymbol{\theta} \end{aligned}$$

Gaussian processes

Generalized Linear Regression (Finite-Dimensional Case)

- Assumption 1: Nature generates parameter $\boldsymbol{\theta}^*$ of a linear function $f_{\boldsymbol{\theta}^*}(\mathbf{x}) = \phi(\mathbf{x})^T \boldsymbol{\theta}^*$ according to $p(\boldsymbol{\theta}) = N(\boldsymbol{\theta}|\mathbf{0}, \sigma_p^2 \mathbf{I})$.
- Assumption 2: Inputs are \mathbf{X} with feature representation Φ ; line i of Φ contains row vector $\phi(\mathbf{x}_i)^T$. Nature generates outputs \mathbf{y} :
 - $y_i = f_{\boldsymbol{\theta}^*}(\mathbf{x}_i) + \epsilon_i$ with $\epsilon_i \sim N(\epsilon|0, \sigma^2)$.
 - $p(y_i|\mathbf{x}_i, \boldsymbol{\theta}^*) = N(y_i|\phi(\mathbf{x}_i)^T \boldsymbol{\theta}^*, \sigma^2)$



Data generation assumptions:

- Given inputs \mathbf{X} , nature generates target values $\bar{\mathbf{y}} \sim N(\bar{\mathbf{y}}|0, \sigma_p^2 \mathbf{K})$.
- Then, nature generates observations $y_i = \bar{y}_i + \epsilon_i$ with noise $\epsilon_i \sim N(\epsilon|0, \sigma^2)$.

Bayes Optimal

Bayes-Optimal Prediction for Classification

- Predictive distribution given the data

$$\begin{aligned} \diamond P(y|\mathbf{x}^*, \mathbf{y}, \mathbf{X}) &= \int P(y|\mathbf{x}^*, \boldsymbol{\theta}) P(\boldsymbol{\theta}|\mathbf{y}, \mathbf{X}) d\boldsymbol{\theta} \\ &= \int \frac{1}{1 + e^{-\mathbf{y}^* \mathbf{x}^{*T} \boldsymbol{\theta}}} N(\boldsymbol{\theta}|\mathbf{0}, \sigma^2 \mathbf{I}) d\boldsymbol{\theta} \end{aligned}$$

- No closed-form solution for logistic regression.
- Possible to approximate by sampling from the posterior.
- Standard approximation: **use only MAP model** instead of integrating over model space.

Bayes-optimal prediction:

$$\diamond \mathbf{y}^* = \arg \max_{\mathbf{y}} P(\mathbf{y}|\mathbf{x}^*, \mathbf{y}, \mathbf{X}) = \bar{\boldsymbol{\theta}}^T \mathbf{x}^*$$

$$\diamond \text{With } \bar{\boldsymbol{\theta}} = (\mathbf{X}^T \mathbf{X} + \frac{\sigma^2}{\sigma_p^2} \mathbf{I})^{-1} \mathbf{X}^T \mathbf{y}.$$

Number of parameters θ_j = number of attributes in \mathbf{x} .

Summary

- Linear regression:
 - Maximum-likelihood model $\arg \max_{\boldsymbol{\theta}} P(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta})$.
 - Maximum-a-posteriori model $\arg \max_{\boldsymbol{\theta}} P(\boldsymbol{\theta}|\mathbf{y}, \mathbf{X})$.
 - Posterior distribution over models $P(\boldsymbol{\theta}|\mathbf{y}, \mathbf{X})$.
 - Bayesian prediction, predictive distribution $\arg \max_{\mathbf{y}} P(\mathbf{y}^*|\mathbf{x}^*, \mathbf{y}, \mathbf{X})$.
- Linear classification (logistic regression):
 - Predictive distribution $P(\mathbf{y}^*|\mathbf{x}^*, \boldsymbol{\theta})$.
 - Maximum-likelihood model $\arg \max_{\boldsymbol{\theta}} P(\mathbf{y}|\mathbf{X}, \boldsymbol{\theta})$.
 - Maximum-a-posteriori model $\arg \max_{\boldsymbol{\theta}} P(\boldsymbol{\theta}|\mathbf{y}, \mathbf{X})$.
 - Bayesian Prediction $\arg \max_{\mathbf{y}} P(\mathbf{y}|\mathbf{x}^*, \mathbf{y}, \mathbf{X})$.
- Nonlinear models: Gaussian processes.