# Interconnectedness of Sovereign Yield curves

Badics Milan Csaba, Balazs Kotro
Draft version

#### Abstract

**Abstract:** This paper is examining the linkages between the whole tenor structure of the yield curves of 12 sovereigns from all over the globe. The curves got decomposed to level, slope and curvature factors by the Nelson Siegel model. TBC...

### 1 Methodology

### 1.1 The Nelson and Siegel framework

Among the statistical models for interest rate, the influential model designed by Diebold-Li [Diebold & Li, 2006] is widely used in market applications. This model is a dynamic extension of the Nelson-Siegel model ([Nelson & Siegel, 1987]) for the cross-section fit for the yield curve. The Nelson-Siegel model corresponds to fitting the following equation for the yield curve observed in the market on a specific date:

$$y_{it}(m_{it}) = \beta_{1t} + \beta_{2t} \left( \frac{1 - e^{-\lambda \tau_t}}{\lambda \tau_t} \right) + \beta_{3t} \left( \frac{1 - e^{-\lambda \tau_t}}{\lambda \tau_t} - e^{-\lambda \tau_t} \right) + \epsilon_{it}$$
 (1)

where  $y_{it}(m_{it})$  are the observed rates on a given date i and maturity t, and  $\beta_{1t}$ ,  $\beta_{2t}$ ,  $\beta_{3t}$  and  $\tau_t$  are parameters. The Nelson-Siegel model is a parsimonious way of fitting the yield curve while managing to capture a part of the stylized facts in interest rate process, such as the exponential formats present in the yield curves. The parameters  $\beta_{it}$  have economic interpretations, where textbeta<sub>1t</sub> presents a long-term level interpretation,  $\beta_{2t}$  short-term components, and  $\beta_{3t}$  medium-term components. It may also be interpreted as decompositions of Level, Slope and Curvature of the yield curve, according to the terminology developed by [Litterman & Scheinkman, 1991]. These components may be used directly in the immunization process of interest rate portfolios.

The purpose of these models is to allow fitting, and subsequent interpolations and extrapolations of the yield curve based on a parametric structure, which concurs with othernon-parametric fitting models such as smoothing-splines. Besides the parsimonious estimation, the [Nelson & Siegel, 1987] model has two additional advantages over non-parametric models. The first advantage is that the extrapolation of the curve has a better performance due to the exponential nature of this model. The second advantage is that the parameters  $\beta_{1t}$ ,  $\beta_{2t}$  and  $\beta_{3t}$  have interpretation of level, slope and curvature compatible with the interpretation of three factors proposed by [Litterman & Scheinkman, 1991], a benchmark in literature. This makes theinterpretation and comparison of the results obtained in the curve fitting easier. The extension formulated by [Diebold & Li, 2006] renders the [Nelson & Siegel, 1987] model dynamic (adjusting the several days observed for the yield curve) by means of a procedure in 3 stages:

- The Nelson-Siegel model (with  $\tau$ fixed, thus making the model linear in the parameters) is fitted by Ordinary Least Squares for each date, estimating the parameters  $\beta_{1t}$ ,  $\beta_{2t}$ ,  $\beta_{3t}$ .
- The dynamics of the system is modelled by a vector autoregressive (VAR) model for the parameters  $\beta_{1t}$ ,  $\beta_{2t}$  and  $\beta_{3t}$ , estimated in the first stage.
- Forecasts for these parameters are made through the VAR model estimated for vectors  $\beta_{1t}$ ,  $\beta_{2t}$  and  $\beta_{3t}$ . By substituting the forecasted parameters in Nelson-Siegel model given by equation '1' it is possible to forecast future interest rate curves.

According to [Diebold & Li, 2006], this dynamic formulation has the purpose of capturing the set of existing stylized facts in the term structure of interest rates, such as the fact that while the yield curve is crescent and concave, it may also assume inverted shapes like decreasing curves and slope changes. Other stylized facts captured by [Diebold & Li, 2006] models are the high persistence in the temporal dynamics (rates with same maturity are highly dependent on the past), and the fact that persistence in the long-term rates is higher than in the short-term rates.

#### 1.2 The Toda-Yamamoto model

The Toda–Yamamoto procedure begins from the following premise: The implementation of the classic Granger Causality test from a VAR (Vector AutoRegressive) model can lead to non-stationarity problems in the series, as it is necessary to confirm the type of existing cointegration. The authors of ... point out that the "conventional" Wald test produces integrated or cointegrated causal VAR models, which would inevitably lead to obtaining spurious Granger causality relationships. However, the Toda–Yamamoto procedure drastically avoids this handicap by developing a Modified Wald test (MWALD) for restrictions on the parameters of a VAR (p) model. This test is generated on a  $\chi_p$  distribution, with  $p = p + d_{max}$  (or number of time lags). In Wolfe-Rufael's words, the fundamental idea underlying this procedure is to "artificially augment the correct VAR order, p, by the maximal order of integration, say  $d_{max}$ . Once this is done, a  $(p + d_{max})$ -th order of VAR is estimated and

the coefficients of the last lagged dmax vector are ignored". The resulting VAR  $(p + d_{max})$  model is formulated in Equations (3) and (4):

$$Y_{t} = \alpha_{0} + \sum_{i=1}^{k} \delta_{1i} Y_{t-i} + \sum_{j=k+1}^{d_{max}} \alpha_{1j} Y_{t-j} + \sum_{j=1}^{k} \theta_{1j} X_{t-j} + \sum_{j=k+1}^{d_{max}} \beta_{1j} X_{t-j} + \omega_{1t}$$

$$(2)$$

$$X_{t} = \alpha_{1} + \sum_{i=1}^{k} \delta_{2i} Y_{t-i} + \sum_{j=k+1}^{d_{max}} \alpha_{2j} Y_{t-j} + \sum_{j=1}^{k} \theta_{2j} X_{t-j} + \sum_{j=k+1}^{d_{max}} \beta_{2j} X_{t-j} + \omega_{2t}$$

$$(3)$$

where  $\omega_{1t}$  and  $\omega_{2t}$  are the VAR error terms and  $d_{max}$  is the maximum order of integration, according to the original specification of the Toda–Yamamoto procedure. Therefore, in Equation (3), causality in the sense of Granger between X and Y will be detected, provided that  $\delta_{1i} \neq 0$  for every i, and, on an identical basis, Equation (4) will imply causality in the sense of Granger between X and Y, if  $\delta_{2i} \neq 0$  for every i.

Once the VAR  $(p + d_{max})$  model is obtained, the implementation of the Toda–Yamamoto procedure in practice requires the realization of three differentiated steps:

- Testing each time-series to conclude the maximum order of integration d<sub>max</sub> of the variables by using, individually or jointly, the following tests: ADF (Augmented Dickey–Fuller), KPSS (Kwiatkowski–Phillips–Schmidt–Shin), and/or PPE (Phillips-Perron).
- Next, the optimal lag length (p) should be obtained based on the criteria: AIC (Akaike Information Criterion), FPE (Akaike's Final Prediction Error), SC (Schwartz), HQ (Hannan and Quinn), and LR (Likelihood-Ratio), seeking, as much as possible, an optimal length supported by the maximum degree of unanimity between criteria.
- Finally, the Granger causality test between the variables X and Y (in both directions) is properly performed by considering that the rejection of the null hypothesis implies the existence of causality in the sense of Granger according to the Toda–Yamamoto procedure and that a reciprocal rejection would indicate a bilateral causal relationship between the analyzed variables.

#### 1.3 Multilayer causality based network

We denote the proposed mulitilayer causality based networks as  $\overline{\Omega} = \{G^{[1]}, G^{[2]}, \dots G^{[L]}\}$  with L layers and N nodes, where  $\{G^{[a]}\} = G(V, A^{[a]})$  is layer a of multilayer causality based networks,  $V = \{1, 2, \dots, N\}$  is the set of nodes, and  $A^{[a]}$  is the set of edges of layer a. On each layer, nodes represent a yield curve factor, and a directed edge indicates that there is a corresponding causality effect from the staring node to the terminal one. In our case L=3, and we assume that the first layer, the second layer and the third layer corresponds to level, slope and curvature layers, respectively. For any two factors  $i,j \in V$ , we draw a direct edge from i to j on the first (second, third) layer, if node i has a level (slope, curvature) causing effect on node j.  $A^{[a]}=\{a_{ij}^{[a]}\}_{N\times N}$  is a directed binary connection matrix for all pairs of nodes i and j olayer a, where the element  $a_{ij}^{[a]}$  in the matrix  $A^{[a]}$  is defined as

$$\alpha_{ij} = \begin{cases} 1 & \text{if } i \neq j \text{ and } i \text{ has a corresponding causality effect on } j \text{ layer } \alpha \\ 0 & \text{otherwise} \end{cases}$$
 (4)

Thus, multilayer information spillover networks are simplified to a 3 dimensional  $N \times N$  adjacency matrix by mathematical notation. Considering the unpredictability of the financial system and the dynamic changes of interconnectedness amon the yield curve nodes, we build time-varying multilayer causality based networks using rolling window analysis. TBC...

Table 1: This is the eaxemple table

Maturity	Mean	Std. Dev	Minimum	Maximum	ρ(1)	ρ(10)
Germany					,	,
1 year	0.0256	1.6275	-0.9690	4.6900	1.0000	0.9960
5 years	0.0257	1.6324	-0.9410	4.7670	0.9990	0.9930
10 years	0.0243	1.5451	-0.7220	4.6860	0.9930	0.9640
30 years	0.0228	1.4482	-0.2440	5.1950	0.9990	0.9880
Italy						
1 year	0.0241	1.5327	-0.4840	8.3940	0.9980	0.9860
5 years	0.0239	1.5199	0.2370	7.8950	0.9980	0.9850
10 years	0.0218	1.3876	0.8750	7.4920	0.9980	0.9860
30 years	0.0188	1.1971	2.0430	7.5840	0.9980	0.9850
France						
1 year	0.0250	1.5891	-0.8010	4.6570	1.0000	0.9960
5 years	0.0247	1.5730	-0.7730	4.9100	0.9990	0.9930
10 years	0.0231	1.4670	-0.4150	4.8510	0.9990	0.9910
30 years	0.0194	1.2351	0.4190	5.1160	0.9990	0.9860
USA						
1 year	0.0254	1.6128	0.0540	5.3230	1.0000	0.9980
5 years	0.0194	1.2315	0.5590	5.3010	0.9990	0.9890
10 years	0.0163	1.0394	1.3890	5.3880	0.9980	0.9830
30 years	0.0151	0.9618	1.9920	5.8390	0.9970	0.9760
Canada	0.0100	1.0001	0.2000	4.0000	1 0000	0.0050
1 year	0.0193	1.2301	0.3000	4.8090	1.0000	0.9950
5 years	0.0180	1.1463	0.4840	4.8010	0.9990	0.9890
10 years	0.0173	1.0976	0.9830	5.0760	0.9990	0.9870
30 years	0.0159	1.0118	1.3060	5.6120	0.9980	0.9850
Mexico						
1 year	0.0318	2.0250	1.5120	10.5700	0.9790	0.9880
5 years	0.0238	1.5108	3.7860	10.8970	0.9860	0.9810
10 years	0.0213	1.3524	4.6190	12.4130	0.9920	0.9700
30 years	0.0195	1.2384	5.8730	12.7260	0.9930	0.9470
Japan 1 year	0.0043	0.2715	-0.3710	0.8500	0.9990	0.9920
		$0.2713 \\ 0.4852$			0.9990	0.9920 $0.9900$
5 years	0.0076		-0.3960 -0.2850	1.6310		
10 years 30 years	$0.0103 \\ 0.0120$	$0.6563 \\ 0.7606$	0.2830 $0.0530$	$\frac{2.0500}{3.2950}$	$0.9990 \\ 0.9980$	$0.9900 \\ 0.9850$
•	0.0120	0.7000	0.0550	3.2930	0.9960	0.9650
China	0.0115	0.7296	0.9570	4.3820	0.9920	0.9660
1 year	$0.0115 \\ 0.0093$	$0.7290 \\ 0.5939$	1.7820	4.3620 $4.8740$	0.9920 $0.9970$	0.9000 $0.9730$
5 years	0.0095 $0.0090$	0.5959 $0.5702$	2.4810			0.9730 $0.9640$
10 years 30 years	0.0090 $0.0097$	$0.5702 \\ 0.6149$	2.4810 $2.4700$	$5.5030 \\ 6.0090$	0.9930	0.9040
	0.0097	0.0149	2.4700	0.0090		
Australia	0.0270	1 7740	0.6750	7 2760	0.0000	0.0020
1 year	0.0279	1.7749	0.6750	7.3760	0.9990	0.9920
5 years	0.0264	1.6769	0.6390	6.9600	0.9990	0.9900
10 years 30 years	$0.0235 \\ 0.0190$	$1.4971 \\ 1.2069$	$0.8850 \\ 1.5580$	$6.8730 \\ 6.8880$	$0.9990 \\ 0.9980$	$0.9880 \\ 0.9830$
· ·	0.0190	1.2009	1.5560	0.8880	0.9960	0.9650
Norway	0.0000	1 4100	0.1000	6 9490	0.0000	0.0040
1 year	0.0222	1.4100	0.1990	6.2430	0.9990	0.9940
5 years	0.0200	1.2734	0.5450	5.3350	0.9990	0.9910
10 years	0.0188	1.1935	0.8880	5.2760	0.9990	0.9890
30 years	0.0175	1.1127	0.8820	5.2730	0.9990	0.9870
United Kingdom	U USU6	1.0559	0.0940	5 0090	0.0000	0.0040
1 year	0.0308	1.9558	0.0240	5.8830	0.9990	0.9950
5 years	0.0259	1.6444	0.1610	5.8210	0.9990	0.9910
10 years 30 years	$0.0223 \\ 0.0175$	$1.4181 \\ 1.1099$	$0.4000 \\ 0.9390$	$5.5430 \\ 5.0700$	$0.9990 \\ 0.9990$	$0.9890 \\ 0.9870$
Switzerland	0.02.0	000	2.2000		2.2000	0.00.0
1 year	0.0195	1.2405	-1.1650	3.3750	1.0000	0.9960
	0.0190					
	0.0187	1 1921	-1 1960	3 2000	(),9990	(),9930
5 years 10 years	$0.0187 \\ 0.0188$	1.1921 $1.1976$	-1.1960 -1.1380	$3.2000 \\ 3.4550$	$0.9990 \\ 0.9990$	$0.9930 \\ 0.9910$

Table 2: This is the eaxemple table

Factor	Mean	Std. Dev	Minimum	Maximum	Jarque-Bera t-statistics	P values
Germany Level Slope Curvature	2.915 -1.856 -3.723	1.535 1.062 1.720	-0.343 -4.544 -7.147	5.413 0.140 0.732	$347.041 \\ 210.098 \\ 234.599$	0 0 0
Italy Level Slope Curvature	4.784 -3.427 -4.251	1.292 1.570 2.236	1.980 -7.009 -8.604	7.998 -0.435 4.752	106.101 183.385 194.464	0 0 0
France Level Slope Curvature	3.391 -2.237 -4.294	1.367 1.191 1.956	0.263 -4.731 -7.820	5.484 0.016 1.073	417.881 162.987 210.076	0 0 0
USA Level Slope Curvature	3.957 -2.408 -3.633	0.988 1.550 2.498	1.881 -5.519 -9.577	5.867 0.710 0.723	323.945 139.147 228.960	0 0 0
Canada Level Slope Curvature	3.431 -1.726 -2.493	1.102 1.228 1.627	1.232 -4.839 -6.257	5.897 0.583 1.307	274.286 251.083 206.333	0 0 0
Mexico Level Slope Curvature	8.579 -2.356 -4.157	1.282 1.825 2.852	5.550 -6.135 -14.835	13.413 0.674 0.489	834.740 340.695 321.622	0 0 0
Japan Level Slope Curvature	1.703 -1.280 -3.694	0.828 $0.625$ $1.283$	-0.018 -2.827 -6.033	3.256 -0.019 -0.874	406.313 155.924 278.292	0 0 0
China Level Slope Curvature	4.024 -1.531 -1.243	0.616 0.809 0.924	2.704 -3.869 -5.198	6.523 1.648 1.259	2047.454 275.315 1128.360	0 0 0
Australia Level Slope Curvature	4.629 -0.898 -2.081	1.230 0.980 1.839	1.398 -3.868 -6.590	6.773 1.003 2.247	262.680 105.808 296.577	0 0 0
Norway Level Slope Curvature	3.215 -1.223 -1.586	1.124 1.072 1.223	1.024 -4.041 -4.681	5.228 2.262 1.726	333.917 167.102 337.556	0 0 0
United Kingdom Level Slope Curvature	3.663 -1.759 -3.329	1.232 1.746 3.021	0.867 -5.418 -8.767	5.798 1.346 3.647	329.213 211.589 159.960	0 0 0
Switzerland Level Slope Curvature	1.773 -1.181 -2.940	1.176 0.699 1.211	-0.756 -3.323 -7.766	3.856 0.909 0.616	289.824 219.716 543.782	0 0 0

Tal	1	3:	Α.	DF
Tak	വമ	٦٠.	Δ	I ) H

Country	GE	ER	IT	Ά	FR	tΑ	U	S	CA	N	MΣ	KN
	value	Р	value	Р	value	Ρ	value	Р	value	Р	value	P
Level Slope Curvature	-2.3040 -2.0878 -2.2176	$0.4496 \\ 0.5412 \\ 0.4862$	-1.4617 -2.2616 -3.5924	$0.8063 \\ 0.4676 \\ 0.0334$	-2.1607 -1.6409 -2.3566	0.5103 $0.7304$ $0.4273$	-3.6572 -1.5000 -1.6211	0.0271 $0.7900$ $0.7388$	-3.0599 -1.5719 -2.3508	$\begin{array}{c} 0.1295 \\ 0.7596 \\ 0.4298 \end{array}$	-3.9732 -1.8809 -2.7054	0.0103 $0.6288$ $0.2796$
Country	JP	Υ	CE	IN	AU	JS	NE	EK	U	K	SV	VI
Country	JP value	PY P	CH	IN P	AU value	JS P	NE value	EK P	value	K P	SV value	VI P

#### Table 4: KPSS

Country	GE	ER	IT	Α	FR	A	U	S	CA	N	MX	KN
	value	P	value	P	value	Ρ	value	P	value	P	value	P
Level Slope Curvature	32.897 $4.255$ $8.958$	$0.010 \\ 0.010 \\ 0.010$	$14.577 \\ 7.608 \\ 10.208$	$0.010 \\ 0.010 \\ 0.010$	$29.872 \\ 4.501 \\ 15.213$	$0.010 \\ 0.010 \\ 0.010$	26.849 5.301 7.339	$0.010 \\ 0.010 \\ 0.010$	$32.182 \\ 5.403 \\ 4.590$	$0.010 \\ 0.010 \\ 0.010$	$15.235 \\ 4.780 \\ 4.926$	$0.010 \\ 0.010 \\ 0.010$
Country	JP value	Y P	CH value	IN P	AU value	JS P	NE value	EK P	Ul value	K P	SV value	VI P

### Table 5: ADF(1)

Country	GE	R	IT	A	FR	A	US	S	CA	N	MX	N
	value	P	value	P								
Level Slope Curvature	-17.051 -15.899 -18.514	$0.010 \\ 0.010 \\ 0.010$	-16.015 -14.628 -18.671	$0.010 \\ 0.010 \\ 0.010$	-15.986 -14.593 -17.646	$0.010 \\ 0.010 \\ 0.010$	-15.788 -16.547 -17.454	$0.010 \\ 0.010 \\ 0.010$	-16.353 -16.323 -16.016	$0.010 \\ 0.010 \\ 0.010$	-16.557 -15.194 -18.880	0.010 0.010 0.010
Country	JP	_	СН		AU		NE		UI	_	SW	_
	value	Р	value	P								
Level Slope	-16.555	0.010	-16.990 -18.902	$0.010 \\ 0.010$	-15.609 -17.872	$0.010 \\ 0.010$	-16.432 -17.605	$0.010 \\ 0.010$	-16.484 -16.116	$0.010 \\ 0.010$	-15.573 -14.598	$0.010 \\ 0.010$
Curvature	-17.512	0.010			-18.061	0.010	-17.041	0.010	-16.451	0.010	-14.266	0.010

## Table 6: KPSS(1)

Country	Gl	ER	II	îA.	FF	RA	U	JS	CA	AN	M	XN
	value	Р	value	Ρ	value	Р	value	Р	value	Ρ	value	Ρ
Level	0.039	0.100	0.127	0.100	0.058	0.100	0.026	0.100	0.060	0.100	0.083	0.100
Slope	0.083	0.100	0.073	0.100	0.118	0.100	0.158	0.100	0.171	0.100	0.091	0.100
Curvature	0.050	0.100	0.015	0.100	0.048	0.100	0.084	0.100	0.076	0.100	0.011	0.100
Country	JF	PΥ	CI	ΗN	Al	US	NI	EK	U	K	SV	VI
	value	Р										
Level	0.026	0.100	0.156	0.100	0.049	0.100	0.036	0.100	0.073	0.100	0.042	0.100
Slope	0.027	0.100	0.029	0.100	0.042	0.100	0.084	0.100	0.242	0.100	0.137	0.100
Curvature	0.059	0.100	0.045	0.100	0.037	0.100	0.030	0.100	0.160	0.100	0.045	0.100

Table 7: Engle-Granger test

	Level	Slope	Curvature
Level	75.694%	44.444%	64.583%
Slope	29.167%	74.306%	71.528%
Curvature	29.167%	78.472%	75.000%

Table 8: Edge counts

	Level	Slope	Curvature	All
Level	29	38	52	119
Slope	30	41	46	117
Curvature	25	34	23	82
All	84	113	121	318

Table 9: Edge ratio

	Table 9: Edge ratio										
	Level	Slope	Curvature	All							
Level	21.970%	26.389%	36.111%	84.470%							
Slope	20.833%	31.061%	31.944%	83.838%							
Curvature	17.361%	23.611%	17.424%	58.396%							
All	60.164%	81.061%	85.480%	25.238%							

Table 10: top nodes

	Top 5	all		Top 5	in	Top 5	out	Top 5	net
Node	All	In	Out	Node	In	Node	Out	Node	Net
$USA_L$	30	7	23	$FRA_C$	19	$USA_L$	23	$USA_L$	16
$\mathrm{AUS}\_\mathrm{S}$	28	11	17	$NOR\_S$	16	$USA\_S$	18	$USA\_S$	12
$FRA\_S$	28	14	14	$NOR_{-}C$	16	$USA\_C$	17	$USA_{-}C$	10
$NOR_{-}S$	27	16	11	$\mathrm{MEX}_{-}\mathrm{L}$	14	$\mathrm{AUS}_{ ext{-}}\!\mathrm{S}$	17	$\mathrm{DEU}_{ extsf{-}\!L}$	9
$FRA_C$	27	19	8	$FRA_S$	14	$\mathrm{DEU}_{-}\!\mathrm{L}$	16	$CAN_L$	7