

# Foundations of Bayesian NLP

MSc Artificial Intelligence

Lecturer: Wilker Aziz

Institute for Logic, Language, and Computation

2018

# The problem with MLE

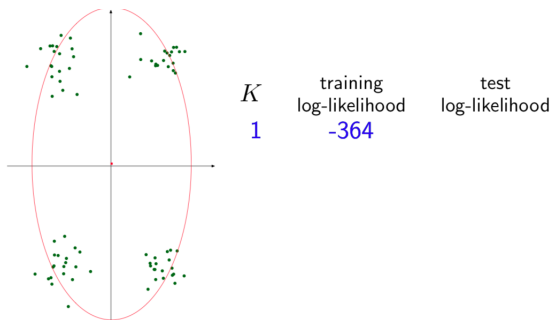
Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM

# The problem with MLE

Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM



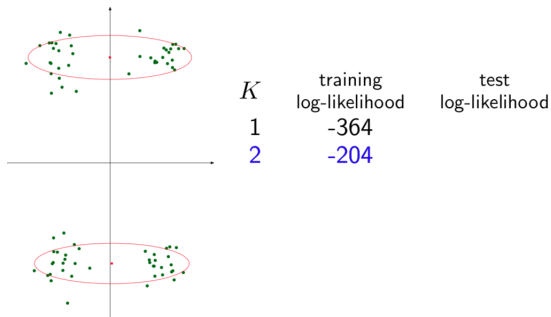
---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models

# The problem with MLE

Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM



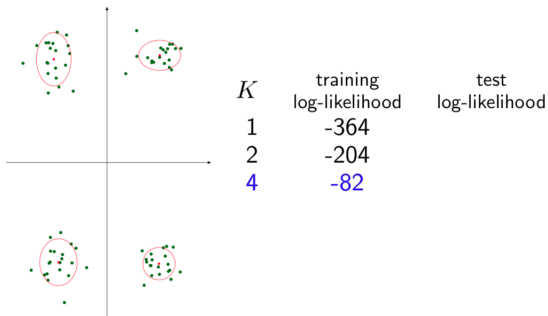
---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models

# The problem with MLE

Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM



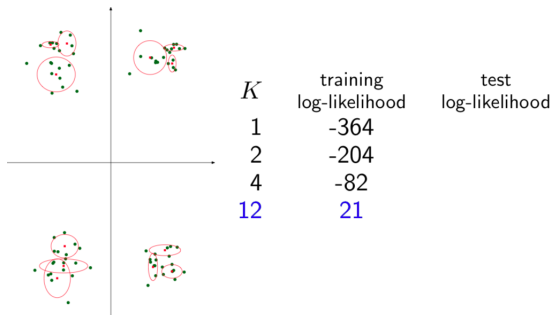
---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models

# The problem with MLE

Motivating example from Liang and Klein (2007)

- mixture of Gaussians trained via EM



- as the capacity of the model increases (more clusters), training likelihood strictly improves

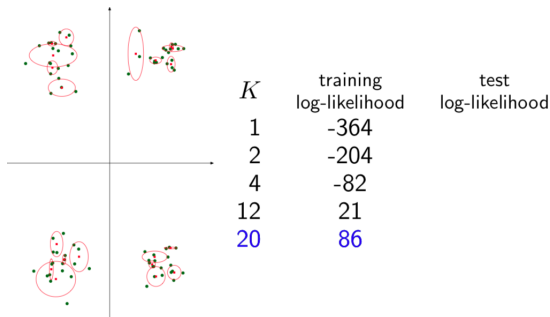
---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models

# The problem with MLE

Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM



- ▶ as the capacity of the model increases (more clusters), training likelihood strictly improves
- ▶ but what happens with test likelihood?

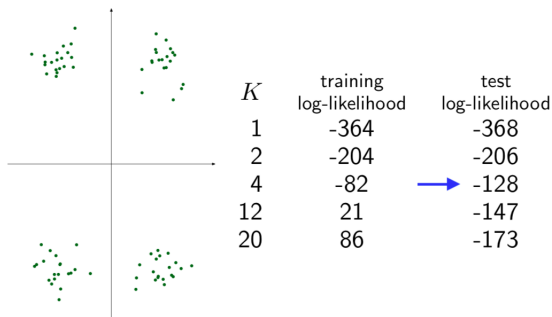
---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models

# The problem with MLE

Motivating example from Liang and Klein (2007)

- ▶ mixture of Gaussians trained via EM



- ▶ as the capacity of the model increases (more clusters), training likelihood strictly improves
- ▶ but what happens with test likelihood?

---

Example from Liang and Klein (2007): ACL tutorial on Structured Bayesian Nonparametric Models



# The problem with MLE

That's why you were told to always do model selection

- ▶ on heldout set
- ▶ preferably via cross-validation

# The problem with MLE

That's why you were told to always do model selection

- ▶ on heldout set
- ▶ preferably via cross-validation

Can you see limitations of this approach?

# The problem with MLE

That's why you were told to always do model selection

- ▶ on heldout set
- ▶ preferably via cross-validation

Can you see limitations of this approach?

- ▶ availability of data
- ▶ representativeness of heldout set
- ▶ discrete optimisation: combinatorial search over models

# NLP1

Preliminaries

Bayesian modelling

Applications

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$
- ▶  $N$  cluster indicators  
 $\mathbf{z} = \langle z_1, \dots, z_N \rangle$



# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$
- ▶  $N$  cluster indicators  
 $\mathbf{z} = \langle z_1, \dots, z_N \rangle$
- ▶  $i$ th cluster indicator  $z_i \in \{1, \dots, C\}$

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$
- ▶  $N$  cluster indicators  
 $\mathbf{z} = \langle z_1, \dots, z_N \rangle$
- ▶  $i$ th cluster indicator  $z_i \in \{1, \dots, C\}$
- ▶ all but the  $i$ th cluster assignment  $\mathbf{z}_{-i}$

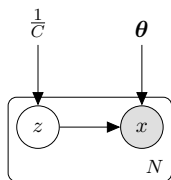
# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$
- ▶  $N$  cluster indicators  
 $\mathbf{z} = \langle z_1, \dots, z_N \rangle$
- ▶  $i$ th cluster indicator  $z_i \in \{1, \dots, C\}$
- ▶ all but the  $i$ th cluster assignment  $\mathbf{z}_{-i}$
- ▶ Parameter vector  
 $\theta = \langle \theta_1, \dots, \theta_K \rangle$

# Conventions

- ▶  $N$  observations  
 $\mathbf{x} = \langle x_1, \dots, x_N \rangle$
- ▶  $i$ th observation  $x_i \in \{1, \dots, K\}$
- ▶ all but the  $i$ th observation  $\mathbf{x}_{-i}$
- ▶  $N$  cluster indicators  
 $\mathbf{z} = \langle z_1, \dots, z_N \rangle$
- ▶  $i$ th cluster indicator  $z_i \in \{1, \dots, C\}$
- ▶ all but the  $i$ th cluster assignment  $\mathbf{z}_{-i}$
- ▶ Parameter vector  
 $\theta = \langle \theta_1, \dots, \theta_K \rangle$
- ▶ Collection of parameter vectors  
 $\boldsymbol{\theta} = \langle \theta^{(1)}, \dots, \theta^{(C)} \rangle$

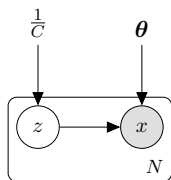
# Mixture model



Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

► categorical likelihood

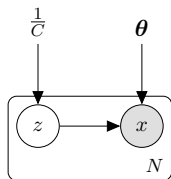
# Mixture model



Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

- ▶ categorical likelihood
- ▶ uniform prior over mixture components, i.e. mixing weights are fixed and uniform

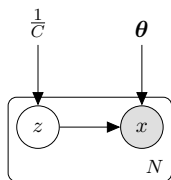
# Mixture model



Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

- ▶ categorical likelihood
- ▶ uniform prior over mixture components, i.e. mixing weights are fixed and uniform
- ▶  $\theta^{(c)} \in \Delta_{K-1}$

# Mixture model



For  $i = 1, \dots, N$

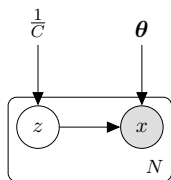
Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

- ▶ categorical likelihood
- ▶ uniform prior over mixture components, i.e. mixing weights are fixed and uniform
- ▶  $\theta^{(c)} \in \Delta_{K-1}$

$$Z_i \sim \mathcal{U}(C)$$



# Mixture model



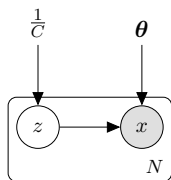
For  $i = 1, \dots, N$

Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

- ▶ categorical likelihood
- ▶ uniform prior over mixture components, i.e. mixing weights are fixed and uniform
- ▶  $\theta^{(c)} \in \Delta_{K-1}$

$$\begin{aligned} Z_i &\sim \mathcal{U}(C) \\ X_i | \theta, \mathbf{z}_{-i}, z_i = c &\sim \text{Cat}(\theta^{(c)}) \end{aligned} \tag{1}$$

# Mixture model



For  $i = 1, \dots, N$

Let's assume  $x$  to be 1 of  $K$ , and  $z$  to be 1 of  $C$

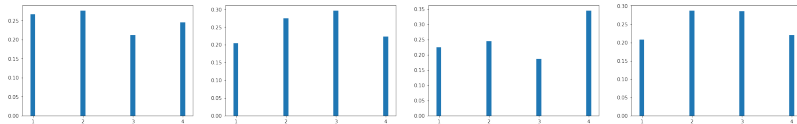
- ▶ categorical likelihood
- ▶ uniform prior over mixture components, i.e. mixing weights are fixed and uniform
- ▶  $\theta^{(c)} \in \Delta_{K-1}$

$$\begin{aligned} Z_i &\sim \mathcal{U}(C) \\ X_i | \theta, \mathbf{z}_{-i}, z_i = c &\sim \text{Cat}(\theta^{(c)}) \end{aligned} \tag{1}$$

What is a sensible conditional distribution  $X | \theta^{(c)} \sim \text{Cat}(\theta^{(c)})$ ?

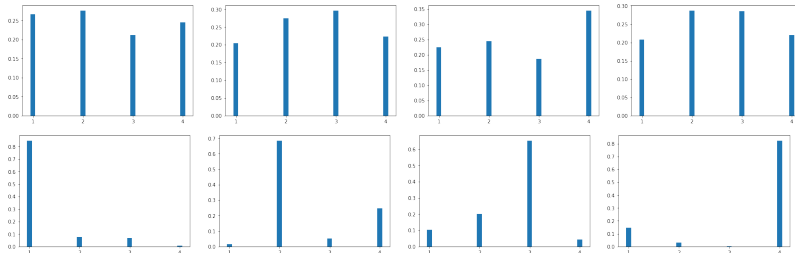
# What makes a good conditional?

$c = 1$  (the blue cluster),  $K = 4$



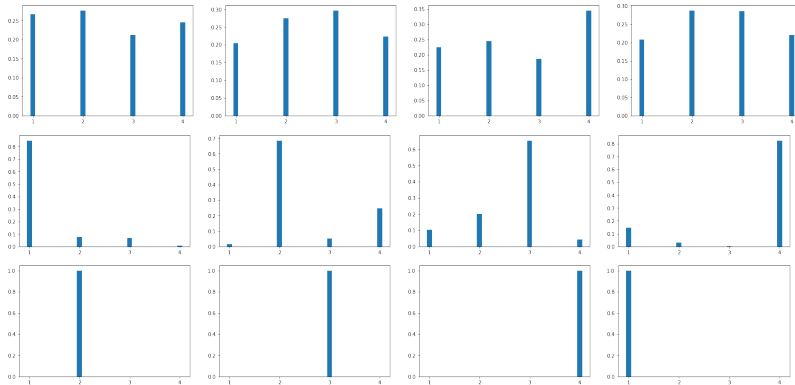
# What makes a good conditional?

$c = 1$  (the blue cluster),  $K = 4$



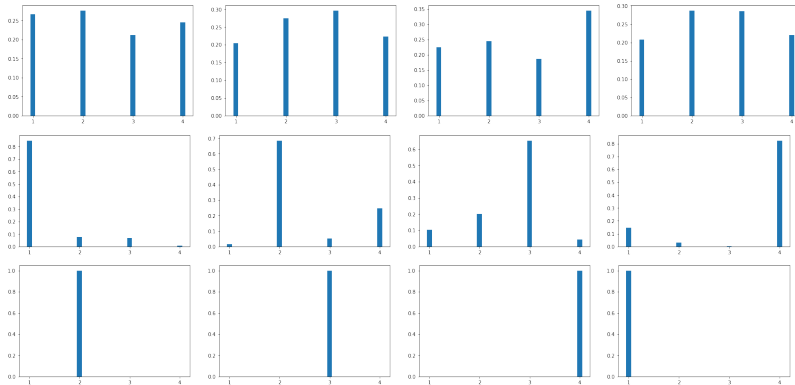
# What makes a good conditional?

$c = 1$  (the blue cluster),  $K = 4$



# What makes a good conditional?

$c = 1$  (the blue cluster),  $K = 4$



Can you make any assumptions before observing data?

# Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} =$$

# Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}}$$



# Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}} \propto P(d|h)P(h) \quad (2)$$

# Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}} \propto P(d|h)P(h) \quad (2)$$

- ▶ the likelihood tells you how well a hypothesis  $h$  explains the observed data  $d$ ;

# Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}} \propto P(d|h)P(h) \quad (2)$$

- ▶ the likelihood tells you how well a hypothesis  $h$  explains the observed data  $d$ ;
- ▶ the prior tells you how much  $h$  conforms to expectations about what a good hypothesis looks like **regardless** of observed data;

## Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}} \propto P(d|h)P(h) \quad (2)$$

- ▶ the likelihood tells you how well a hypothesis  $h$  explains the observed data  $d$ ;
- ▶ the prior tells you how much  $h$  conforms to expectations about what a good hypothesis looks like **regardless** of observed data;
- ▶ the evidence tells you how well your model  $\mathcal{M}$  explains the data, i.e.  $P(d)$  is actually  $P(d|\mathcal{M})$

## Bayes rule

What does Bayes rule tell you?

$$\underbrace{P(h|d)}_{\text{posterior}} = \frac{\overbrace{P(d|h)}^{\text{likelihood}} \overbrace{P(h)}^{\text{prior}}}{\underbrace{P(d)}_{\text{evidence}}} \propto P(d|h)P(h) \quad (2)$$

- ▶ the likelihood tells you how well a hypothesis  $h$  explains the observed data  $d$ ;
- ▶ the prior tells you how much  $h$  conforms to expectations about what a good hypothesis looks like **regardless** of observed data;
- ▶ the evidence tells you how well your model  $\mathcal{M}$  explains the data, i.e.  $P(d)$  is actually  $P(d|\mathcal{M})$
- ▶ the posterior updates our beliefs about hypotheses **in light of** observed data.

# Maximum likelihood estimation

An optimisation problem based on the *(log-)likelihood function*

$$h^{\star} = \arg \max_h P(d|h)$$

# Maximum likelihood estimation

An optimisation problem based on the *(log-)likelihood function*

$$h^{\star} = \arg \max_h P(d|h) = \arg \max_h \underbrace{\log P(d|h)}_{\mathcal{L}(h)} \quad (3)$$

# Maximum likelihood estimation

An optimisation problem based on the *(log-)likelihood function*

$$h^* = \arg \max_h P(d|h) = \arg \max_h \underbrace{\log P(d|h)}_{\mathcal{L}(h)} \quad (3)$$

- ▶ all hypotheses are **equally likely a priori**;



# Maximum likelihood estimation

An optimisation problem based on the *(log-)likelihood function*

$$h^* = \arg \max_h P(d|h) = \arg \max_h \underbrace{\log P(d|h)}_{\mathcal{L}(h)} \quad (3)$$

- ▶ all hypotheses are **equally likely a priori**;
- ▶ can be approached by coordinate ascent methods;

# Maximum likelihood estimation

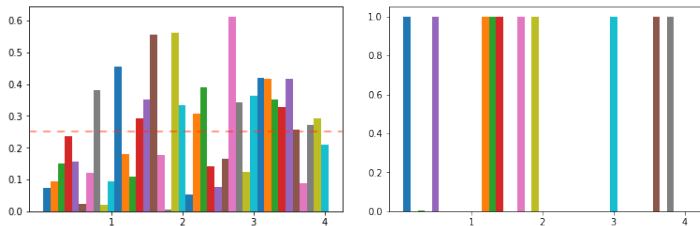
An optimisation problem based on the *(log-)likelihood function*

$$h^* = \arg \max_h P(d|h) = \arg \max_h \underbrace{\log P(d|h)}_{\mathcal{L}(h)} \quad (3)$$

- ▶ all hypotheses are **equally likely a priori**;
- ▶ can be approached by coordinate ascent methods;
- ▶ local optimality guarantees;

# All the same a priori

Before data, MLE is equally happy with the hypotheses on the left



# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak

# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder

# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder
- ▶ still a point estimate, teaches us very little about the overall model (set of assumptions)

# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder
- ▶ still a point estimate, teaches us very little about the overall model (set of assumptions)

“I read before that Bayesian priors are just like regularisers,

# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder
- ▶ still a point estimate, teaches us very little about the overall model (set of assumptions)

“I read before that Bayesian priors are just like regularisers, I even know that a Gaussian prior is just  $L_2$  regularisation”



# Constraining MLE

Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder
- ▶ still a point estimate, teaches us very little about the overall model (set of assumptions)

“I read before that Bayesian priors are just like regularisers, I even know that a Gaussian prior is just  $L_2$  regularisation”

- ▶ that only covers the specification of a prior

# Constraining MLE

## Maximum a posteriori

$$\begin{aligned} h^* &= \arg \max_h P(d|h)P(h) \\ &= \arg \max_h \log P(d|h) + \log P(h) \end{aligned} \tag{4}$$

- ▶ perhaps fine if  $P(h)$  has a single narrow peak
- ▶ priors often indicate preference for a subset of hypotheses over another, multiple peaks make optimisation considerably harder
- ▶ still a point estimate, teaches us very little about the overall model (set of assumptions)

“I read before that Bayesian priors are just like regularisers, I even know that a Gaussian prior is just  $L_2$  regularisation”

- ▶ that only covers the specification of a prior
- ▶ Bayesian modelling does not end at prior specification  
you need the crucial part: posterior inference

# NLP1

Preliminaries

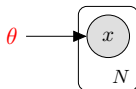
Bayesian modelling

Dirichlet-Multinomial model

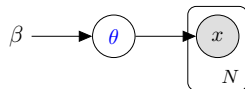
Applications

# A Bayesian model

Frequentist



Bayesian

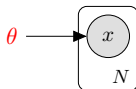


In a Bayesian model, parameters are no different from data

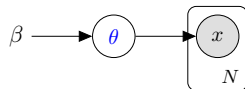
- ▶ they are random variables much like data

# A Bayesian model

Frequentist



Bayesian

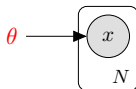


In a Bayesian model, parameters are no different from data

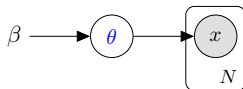
- ▶ they are random variables much like data
- ▶ only they are not observed

# A Bayesian model

Frequentist



Bayesian



In a Bayesian model, parameters are no different from data

- ▶ they are random variables much like data
- ▶ only they are not observed

Bayesians do condition on deterministic quantities

- ▶  $\beta$  here are called *hyperparameters*

# A Bayesian model



In a Bayesian model, parameters are no different from data

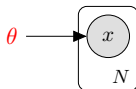
- ▶ they are random variables much like data
- ▶ only they are not observed

Bayesians do condition on deterministic quantities

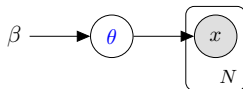
- ▶  $\beta$  here are called *hyperparameters*
- ▶ but most Bayesians leave those fixed (no search!)

# A Bayesian model

Frequentist



Bayesian



In a Bayesian model, parameters are no different from data

- ▶ they are random variables much like data
- ▶ only they are not observed

Bayesians do condition on deterministic quantities

- ▶  $\beta$  here are called *hyperparameters*
- ▶ but most Bayesians leave those fixed (no search!)

We will study an example that illustrates important concepts

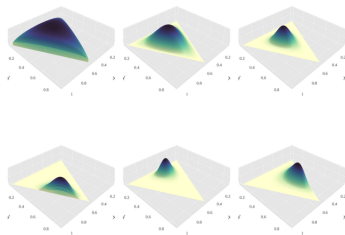
Dirichlet-Multinomial model



# Dirichlet distribution

A distribution over the open simplex of  $K$ -dimensional vectors we denote the simplex by

$$\Delta_{K-1} = \left\{ \theta \in \mathbb{R}_{>0}^K : \sum_{k=1}^K \theta_k = 1 \right\} \subseteq \mathbb{R}_{>0}^K \quad (5)$$



Use this [notebook](#) and this [wikipedia](#) to learn more

## Count vector

For observations  $\mathbf{x}$ , where  $x_i$  is 1 of  $K$   
define  $n^{(\mathbf{x})}$  as the  $K$ -dimensional vector such that

$$n_k = \sum_{i=1}^N [x_i = k] \quad (6)$$

# Count vector

For observations  $\mathbf{x}$ , where  $x_i$  is 1 of  $K$   
define  $n^{(\mathbf{x})}$  as the  $K$ -dimensional vector such that

$$n_k = \sum_{i=1}^N [x_i = k] \quad (6)$$

Example: for  $K = 3$  and  $N = 6$

$$\mathbf{x} = \langle x_1 = \textcolor{red}{2}, x_2 = \textcolor{teal}{3}, x_3 = \textcolor{blue}{1}, x_4 = \textcolor{red}{2}, x_5 = \textcolor{red}{2}, x_6 = \textcolor{teal}{3} \rangle$$
$$n^{(\mathbf{x})} =$$

# Count vector

For observations  $\mathbf{x}$ , where  $x_i$  is 1 of  $K$   
define  $n^{(\mathbf{x})}$  as the  $K$ -dimensional vector such that

$$n_k = \sum_{i=1}^N [x_i = k] \quad (6)$$

Example: for  $K = 3$  and  $N = 6$

$$\begin{aligned} \mathbf{x} &= \langle x_1 = 2, x_2 = 3, x_3 = 1, x_4 = 2, x_5 = 2, x_6 = 3 \rangle \\ n^{(\mathbf{x})} &= \langle n_1 = 1, n_2 = 3, n_3 = 2 \rangle \end{aligned}$$

# Gamma function

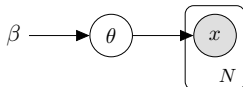
A generalisation of the factorial function to  $\mathbb{R}$

$$\Gamma(z) = \int_0^{\infty} \epsilon^{z-1} \exp(-\epsilon) d\epsilon \quad (7)$$

Properties

- ▶  $\Gamma(n) = (n-1)!$  for positive integer  $n$
- ▶  $\Gamma(z) = (z-1)\Gamma(z-1)$

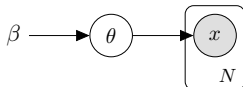
# Dirchlet-Multinomial



Model

$$\begin{aligned}\theta|\beta &\sim \text{Dir}(\beta) \\ X_i|\theta &\sim \text{Cat}(\theta) \quad \text{for } i = 1, \dots, N\end{aligned}\tag{8}$$

# Dirchlet-Multinomial



Model

$$\begin{aligned}\theta|\beta &\sim \text{Dir}(\beta) \\ X_i|\theta &\sim \text{Cat}(\theta) \quad \text{for } i = 1, \dots, N\end{aligned}\tag{8}$$

Joint distribution

$$\begin{aligned}P(\mathbf{x}, \theta|\beta) &= P(\theta)P(\mathbf{x}|\theta) \\ &= \text{Dir}(\theta|\beta) \text{Mult}(n^{(\mathbf{x})}|\theta, N)\end{aligned}\tag{9}$$

# Multinomial likelihood

For  $\theta \in \Delta_{K-1}$

$$P(\mathbf{x}|\theta) = \text{Mult}(n^{(\mathbf{x})}|\theta, N)$$



# Multinomial likelihood

For  $\theta \in \Delta_{K-1}$

$$\begin{aligned} P(\mathbf{x}|\theta) &= \text{Mult}(n^{(\mathbf{x})}|\theta, N) \\ &= \frac{N!}{\prod_{k=1}^K n_k!} \prod_{k=1}^K \theta_k^{n_k} \end{aligned}$$

# Multinomial likelihood

For  $\theta \in \Delta_{K-1}$

$$\begin{aligned} P(\mathbf{x}|\theta) &= \text{Mult}(n^{(\mathbf{x})}|\theta, N) \\ &= \frac{N!}{\prod_{k=1}^K n_k!} \prod_{k=1}^K \theta_k^{n_k} \\ &= \frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k} \end{aligned}$$

# Multinomial likelihood

For  $\theta \in \Delta_{K-1}$

$$\begin{aligned} P(\mathbf{x}|\theta) &= \text{Mult}(n^{(\mathbf{x})}|\theta, N) \\ &= \frac{N!}{\prod_{k=1}^K n_k!} \prod_{k=1}^K \theta_k^{n_k} \\ &= \frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k} \end{aligned} \tag{10}$$

Example: for  $K = 3$  and  $N = 6$

$$\theta = \langle \theta_1 = 0.2, \theta_2 = 0.3, \theta_3 = 0.5 \rangle$$

$$\mathbf{x} = \langle x_1 = 2, x_2 = 3, x_3 = 1, x_4 = 2, x_5 = 2, x_6 = 3 \rangle$$

$$n^{(\mathbf{x})} = \langle n_1 = 1, n_2 = 3, n_3 = 2 \rangle$$

$$P(\mathbf{x}|\theta) = \frac{\Gamma(\dots)}{\prod \dots} \theta_1^1 \times \theta_2^3 \times \theta_3^2$$

# Dirichlet prior

For  $\beta \in \mathbb{R}_{>0}^K$

$$\text{Dir}(\theta|\beta) = \frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1}$$

# Dirichlet prior

For  $\beta \in \mathbb{R}_{>0}^K$

$$\begin{aligned}\text{Dir}(\theta|\beta) &= \frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1} \\ &\propto \prod_{k=1}^K \theta_k^{\beta_k-1}\end{aligned}\tag{11}$$

# Dirichlet prior

For  $\beta \in \mathbb{R}_{>0}^K$

$$\begin{aligned}\text{Dir}(\theta|\beta) &= \frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1} \\ &\propto \prod_{k=1}^K \theta_k^{\beta_k-1}\end{aligned}\tag{11}$$

We call

$$\int_{\Delta_{K-1}} \prod_{k=1}^K \theta_k^{\beta_k-1} = \frac{\prod_{k=1}^K \Gamma(\beta_k)}{\Gamma(\sum_{k=1}^K \beta_k)}$$

the *Dirichlet normaliser*

# Posterior

$$P(\theta|\mathbf{x},\beta) \propto$$

# Posterior

$$P(\theta|\mathbf{x},\beta) \propto P(\mathbf{x}|\theta)P(\theta|\beta)$$



# Posterior

$$P(\theta|\mathbf{x}, \beta) \propto P(\mathbf{x}|\theta)P(\theta|\beta)$$
$$\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times$$

# Posterior

$$P(\theta|\mathbf{x}, \beta) \propto P(\mathbf{x}|\theta)P(\theta|\beta)$$
$$\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k - 1}}_{\text{Dir}(\theta|\beta)}$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k - 1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \end{aligned}$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k - 1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \prod_{k=1}^K \theta_k^{\beta_k - 1} \end{aligned}$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k - 1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \prod_{k=1}^K \theta_k^{\beta_k - 1} \\ &= \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} \end{aligned}$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}^{(\mathbf{x})}|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \prod_{k=1}^K \theta_k^{\beta_k-1} \\ &= \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} \propto \text{Dir}(\theta|\mathbf{n}^{(\mathbf{x})} + \beta) \end{aligned}$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(n^{(\mathbf{x})}|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \prod_{k=1}^K \theta_k^{\beta_k-1} \\ &= \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} \propto \text{Dir}(\theta|n^{(\mathbf{x})} + \beta) \end{aligned}$$

Thus

$$P(\theta|\mathbf{x}, \beta) = \underbrace{\prod_{k=1}^K \theta_k^{n_k + \beta_k - 1}}_{\frac{1}{\text{normaliser}} \text{ of } \text{Dir}(n^{(\mathbf{x})} + \beta)} \quad (12)$$

# Posterior

$$\begin{aligned} P(\theta|\mathbf{x}, \beta) &\propto P(\mathbf{x}|\theta)P(\theta|\beta) \\ &\propto \underbrace{\frac{\Gamma(\sum_{k=1}^K n_k + 1)}{\prod_{k=1}^K \Gamma(n_k + 1)} \prod_{k=1}^K \theta_k^{n_k}}_{\text{Mult}(\mathbf{n}(\mathbf{x})|\theta)} \times \underbrace{\frac{\Gamma(\sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(\beta_k)} \prod_{k=1}^K \theta_k^{\beta_k-1}}_{\text{Dir}(\theta|\beta)} \\ &\propto \prod_{k=1}^K \theta_k^{n_k} \times \prod_{k=1}^K \theta_k^{\beta_k-1} \\ &= \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} \propto \text{Dir}(\theta|\mathbf{n}(\mathbf{x}) + \beta) \end{aligned}$$

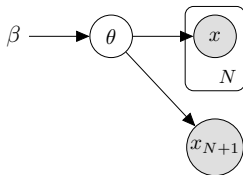
Thus

$$P(\theta|\mathbf{x}, \beta) = \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\substack{1 \\ \text{normaliser}} \text{ of } \text{Dir}(\mathbf{n}(\mathbf{x}) + \beta)} \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} \quad (12)$$



# Posterior predictive distribution

Suppose a new data point  $x_{N+1} = j$  is available



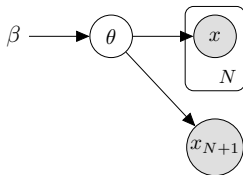
$$P(x_{N+1} = j | \mathbf{x}, \beta) = \int_{\Delta_{K-1}} P(\theta, x_{N+1} | \mathbf{x}, \beta) d\theta$$

---

$x_{N+1}$  is independent of  $\mathbf{x}$  given  $\theta$

# Posterior predictive distribution

Suppose a new data point  $x_{N+1} = j$  is available



$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} P(\theta, x_{N+1} | \mathbf{x}, \beta) d\theta \\ &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \end{aligned}$$

---

$x_{N+1}$  is independent of  $\mathbf{x}$  given  $\theta$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta$$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \int_{\Delta_{K-1}} \theta_j \times d\theta \end{aligned}$$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \int_{\Delta_{K-1}} \theta_j \times \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} d\theta \end{aligned}$$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \int_{\Delta_{K-1}} \theta_j \times \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} d\theta \\ &= \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \int_{\Delta_{K-1}} \theta_j \times d\theta \end{aligned}$$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \int_{\Delta_{K-1}} \theta_j \times \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} d\theta \\ &= \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \int_{\Delta_{K-1}} \theta_j \times \underbrace{\theta_j^{n_j + \beta_j - 1} \prod_{\substack{k=1 \\ k \neq j}}^K \theta_k^{n_k + \beta_k - 1}}_{\prod_{k=1}^K \theta_k^{n_k + \beta_k - 1}} d\theta \end{aligned}$$

## Posterior predictive distribution (cont.)

Suppose a new data point  $x_{N+1} = j$  is available

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \int_{\Delta_{K-1}} \theta_j \times \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \prod_{k=1}^K \theta_k^{n_k + \beta_k - 1} d\theta \\ &= \underbrace{\frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)}}_{\text{constant wrt } \theta} \int_{\Delta_{K-1}} \theta_j \times \underbrace{\theta_j^{n_j + \beta_j - 1} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1}}_{\prod_{k=1}^K \theta_k^{n_k + \beta_k - 1}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta \end{aligned}$$



## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{integral}} \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{1}{\Gamma(N + \sum_{k=1}^K \beta_k)} \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{\Gamma(N + \sum_{k=1}^K \beta_k + 1)}{\Gamma(N + \sum_{k=1}^K \beta_k + 1)} \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{\Gamma(n_j + \beta_j + 1) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{\Gamma(N + \sum_{k=1}^K \beta_k + 1)} \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned} P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{\Gamma(n_j + \beta_j + 1) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{\Gamma(N + \sum_{k=1}^K \beta_k + 1)} \\ &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{(n_j + \beta_j) \Gamma(n_j + \beta_j) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{(N + \sum_{k=1}^K \beta_k) \Gamma(N + \sum_{k=1}^K \beta_k)} \end{aligned}$$

## Posterior predictive distribution (cont.)

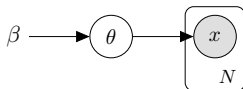
$$\begin{aligned}
 P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{\Gamma(n_j + \beta_j + 1) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{\Gamma(N + \sum_{k=1}^K \beta_k + 1)} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{(n_j + \beta_j) \Gamma(n_j + \beta_j) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{(N + \sum_{k=1}^K \beta_k) \Gamma(N + \sum_{k=1}^K \beta_k)} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{(n_j + \beta_j) \Gamma(n_j + \beta_j) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{(N + \sum_{k=1}^K \beta_k) \Gamma(N + \sum_{k=1}^K \beta_k)}
 \end{aligned}$$

## Posterior predictive distribution (cont.)

$$\begin{aligned}
 P(x_{N+1} = j | \mathbf{x}, \beta) &= \int_{\Delta_{K-1}} \underbrace{P(x_{N+1} = j | \theta)}_{\text{likelihood}} \underbrace{P(\theta | \mathbf{x}, \beta)}_{\text{posterior}} d\theta \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \underbrace{\int_{\Delta_{K-1}} \theta_j^{n_j + \beta_j} \prod_{k \neq j} \theta_k^{n_k + \beta_k - 1} d\theta}_{\text{Dir normaliser}} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{\Gamma(n_j + \beta_j + 1) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{\Gamma(N + \sum_{k=1}^K \beta_k + 1)} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{(n_j + \beta_j) \Gamma(n_j + \beta_j) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{(N + \sum_{k=1}^K \beta_k) \Gamma(N + \sum_{k=1}^K \beta_k)} \\
 &= \frac{\Gamma(N + \sum_{k=1}^K \beta_k)}{\prod_{k=1}^K \Gamma(n_k + \beta_k)} \frac{(n_j + \beta_j) \Gamma(n_j + \beta_j) \prod_{k \neq j} \Gamma(n_k + \beta_k)}{(N + \sum_{k=1}^K \beta_k) \Gamma(N + \sum_{k=1}^K \beta_k)} \\
 &= \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k}
 \end{aligned}$$



# Dirichlet-Multinomial (overview)



Joint distribution

$$\begin{aligned} P(\mathbf{x}, \theta | \beta) &= P(\theta) P(\mathbf{x} | \theta) \\ &= \text{Dir}(\theta | \beta) \text{Mult}(n^{(\mathbf{x})} | \theta, N) \end{aligned} \tag{13}$$

Posterior

$$P(\theta | \mathbf{x}, \beta) = \text{Dir}(\theta | n^{(\mathbf{x})} + \beta) \tag{14}$$

Predictive posterior

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k} \tag{15}$$

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

- ▶ in our Dirichlet-Multinomial model any re-ordering of the observations is equally likely to occur

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

- ▶ in our Dirichlet-Multinomial model any re-ordering of the observations is equally likely to occur

Combine that fact with the predictive posterior result

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k} \quad (16)$$

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

- ▶ in our Dirichlet-Multinomial model any re-ordering of the observations is equally likely to occur

Combine that fact with the predictive posterior result

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k} \quad (16)$$

and we can single out any observation, e.g.  $\mathbf{x}_i$

$$P(\mathbf{x}_i = j | \mathbf{x}_{-i}, \beta) = \text{—————} \quad (17)$$

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

- ▶ in our Dirichlet-Multinomial model any re-ordering of the observations is equally likely to occur

Combine that fact with the predictive posterior result

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k} \quad (16)$$

and we can single out any observation, e.g.  $\mathbf{x}_i$

$$P(\mathbf{x}_i = j | \mathbf{x}_{-i}, \beta) = \frac{n_j + \beta_j}{N - 1 + \sum_{k=1}^K \beta_k} \quad (17)$$

# Exchangeability

Random variables are called **exchangeable** under a model when all permutations of the set of outcomes have the same probability

- ▶ in our Dirichlet-Multinomial model any re-ordering of the observations is equally likely to occur

Combine that fact with the predictive posterior result

$$P(x_{N+1} = j | \mathbf{x}, \beta) = \frac{n_j + \beta_j}{N + \sum_{k=1}^K \beta_k} \quad (16)$$

and we can single out any observation, e.g.  $\mathbf{x}_i$

$$P(\mathbf{x}_i = j | \mathbf{x}_{-i}, \beta) = \frac{n_j^{(\mathbf{x}_{-i})} + \beta_j}{N - 1 + \sum_{k=1}^K \beta_k} \quad (17)$$

# Summary

Friends do not let friends optimise



# Summary

Friends do not let friends optimise

- ▶ no point estimates, we use all possible model parameters

# Summary

Friends do not let friends optimise

- ▶ no point estimates, we use all possible model parameters
- ▶ this is called *Bayesian inference*, or simply, inference

# Summary

Friends do not let friends optimise

- ▶ no point estimates, we use all possible model parameters
- ▶ this is called *Bayesian inference*, or simply, inference
- ▶ Bayesian models have memory: the posterior summarises what we learnt from data

# Summary

Friends do not let friends optimise

- ▶ no point estimates, we use all possible model parameters
- ▶ this is called *Bayesian inference*, or simply, inference
- ▶ Bayesian models have memory: the posterior summarises what we learnt from data
- ▶ If we collect more data  $\mathbf{x}'$ , we can update the posterior,  
$$P(\theta|\mathbf{x}, \mathbf{x}', \beta) = \text{Dir}(\theta|n^{(\mathbf{x})} + n^{(\mathbf{x}')} + \beta)$$

# Summary

Friends do not let friends optimise

- ▶ no point estimates, we use all possible model parameters
- ▶ this is called *Bayesian inference*, or simply, inference
- ▶ Bayesian models have memory: the posterior summarises what we learnt from data
- ▶ If we collect more data  $\mathbf{x}'$ , we can update the posterior,  
$$P(\theta|\mathbf{x}, \mathbf{x}', \beta) = \text{Dir}(\theta|n^{(\mathbf{x})} + n^{(\mathbf{x}')} + \beta)$$
- ▶ MLE is memoryless: there is one fixed  $\theta$ , no matter how much more data you see,  $\theta$  will never change

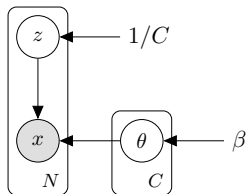
# NLP1

Preliminaries

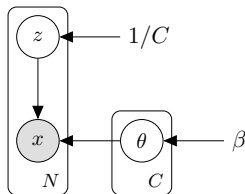
Bayesian modelling

Applications

# Bayesian mixture model with categorical observations



# Bayesian mixture model with categorical observations

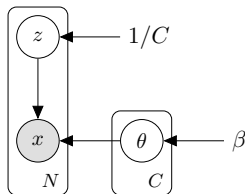


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$



# Bayesian mixture model with categorical observations



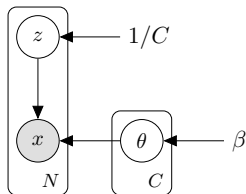
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta)$$

# Bayesian mixture model with categorical observations



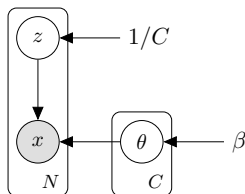
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ = P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times \end{aligned}$$

# Bayesian mixture model with categorical observations



Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

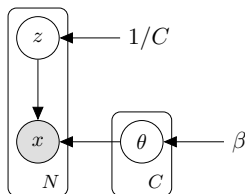
$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta)$$

$$= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c)$$

# Bayesian mixture model with categorical observations

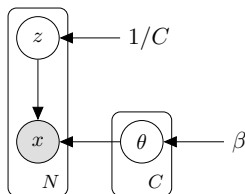


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{1}{C} \times \end{aligned}$$

# Bayesian mixture model with categorical observations

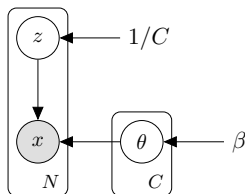


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{1}{C} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \end{aligned}$$

# Bayesian mixture model with categorical observations

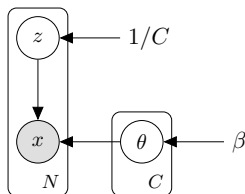


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{1}{C} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \\ P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) &\propto P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \end{aligned}$$

# Bayesian mixture model with categorical observations



Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{1}{C} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \\ P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) &\propto P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \\ &\propto \frac{n_{c,z} + \beta}{n_c + K\beta} \end{aligned}$$

## Mixing weights

What does it mean to have uniform prior over components?



## Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

## Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

- ▶ uniform prior leaves it up to the likelihood to control sparsity

## Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

- ▶ uniform prior leaves it up to the likelihood to control sparsity
- ▶ luckily we are promoting sparse likelihoods  $X|z$   
because  $\theta^{(z)} \sim \text{Dir}(\beta)$

# Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

- ▶ uniform prior leaves it up to the likelihood to control sparsity
- ▶ luckily we are promoting sparse likelihoods  $X|z$   
because  $\theta^{(z)} \sim \text{Dir}(\beta)$
- ▶ but  $P(z)$  has nothing to do with it!

## Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

- ▶ uniform prior leaves it up to the likelihood to control sparsity
- ▶ luckily we are promoting sparse likelihoods  $X|z$   
because  $\theta^{(z)} \sim \text{Dir}(\beta)$
- ▶ but  $P(z)$  has nothing to do with it!

Is there really no preference we can express about  $P(z)$ ?

## Mixing weights

What does it mean to have uniform prior over components?

- ▶ unlike it may seem, it does not mean to promote diversity!

Let's see whether the posterior is *peaked*

$$P(z|x) = \frac{\frac{1}{C} \times P(x|z)}{P(x)} \propto P(x|z)$$

- ▶ uniform prior leaves it up to the likelihood to control sparsity
- ▶ luckily we are promoting sparse likelihoods  $X|z$   
because  $\theta^{(z)} \sim \text{Dir}(\beta)$
- ▶ but  $P(z)$  has nothing to do with it!

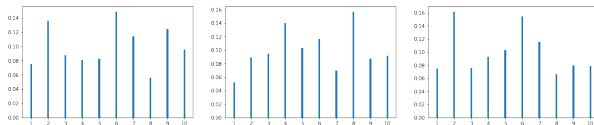
Is there really no preference we can express about  $P(z)$ ?

- ▶ what about preferring to use fewer components?

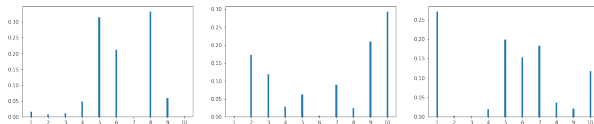
# Sparse prior over mixing weights

Say we have 10 components, how do you want to use them?

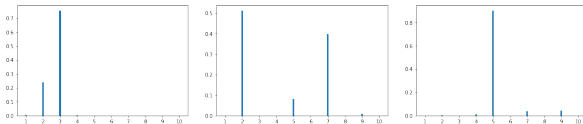
I couldn't care less



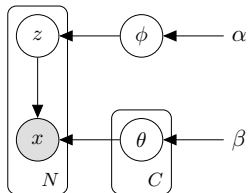
Sparingly



Like I pass students

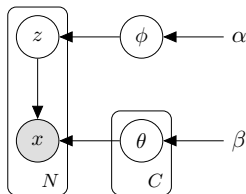


# Bayesian mixture model - Sparse prior over mixing weights





# Bayesian mixture model - Sparse prior over mixing weights

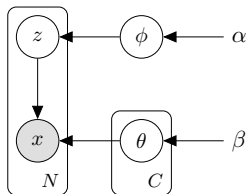


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

# Bayesian mixture model - Sparse prior over mixing weights



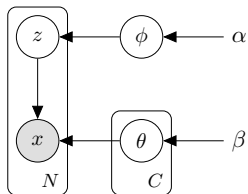
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta)$$

# Bayesian mixture model - Sparse prior over mixing weights



Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

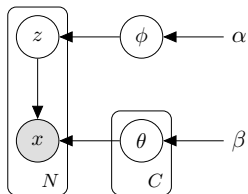
$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta)$$

$$= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times$$

# Bayesian mixture model - Sparse prior over mixing weights

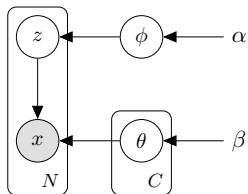


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta)$$
$$= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c)$$

# Bayesian mixture model - Sparse prior over mixing weights

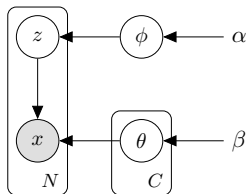


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{n_c + \alpha}{N - 1 + C\alpha} \times \end{aligned}$$

# Bayesian mixture model - Sparse prior over mixing weights



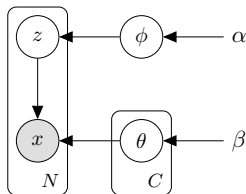
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$

$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{n_c + \alpha}{N - 1 + C\alpha} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \end{aligned}$$

# Bayesian mixture model - Sparse prior over mixing weights

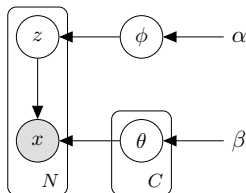


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{n_c + \alpha}{N - 1 + C\alpha} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \\ P(z_i = c | \mathbf{x}, \mathbf{z}_{-i}, \alpha, \beta) &\propto P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \end{aligned}$$

# Bayesian mixture model - Sparse prior over mixing weights



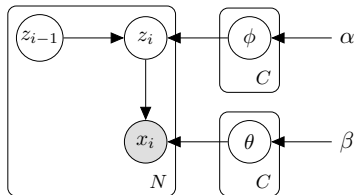
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$n_{c,k} = \sum_{j \neq i} [z_j = c][x_j = k]$$
$$n_c = \sum_{k=1}^K n_{c,k}$$

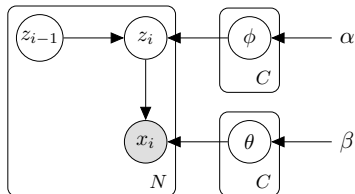
$$\begin{aligned} P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \\ &= P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta) \times P(x_i = k | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \beta, z_i = c) \\ &= \frac{n_c + \alpha}{N - 1 + C\alpha} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \\ P(z_i = c | \mathbf{x}, \mathbf{z}_{-i}, \alpha, \beta) &\propto P(x_i = k, z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \\ &\propto \frac{n_c + \alpha}{N - 1 + C\alpha} \times \frac{n_{c,z} + \beta}{n_c + K\beta} \end{aligned}$$



# Bayesian HMM



# Bayesian HMM

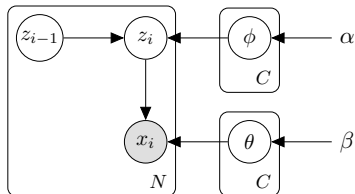


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$m_{b,c} = \sum_{j \neq i} [z_{j-1} = b][z_j = c]$$

$$m_b = \sum_{c=1}^C m_{b,c}$$

# Bayesian HMM

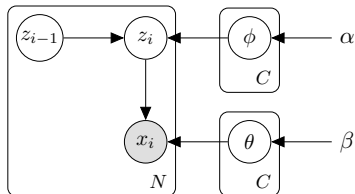


Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$m_{b,c} = \sum_{j \neq i} [z_{j-1} = b][z_j = c]$$
$$m_b = \sum_{c=1}^C m_{b,c}$$

$$P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta)$$

# Bayesian HMM



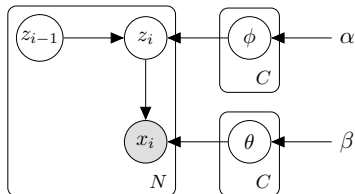
Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$m_{b,c} = \sum_{j \neq i} [z_{j-1} = b][z_j = c]$$

$$m_b = \sum_{c=1}^C m_{b,c}$$

$$P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \quad \text{note that } \begin{cases} z_{i-1} = b \\ z_{i+1} = d \end{cases} \text{ is in } \mathbf{z}_{-i}$$

# Bayesian HMM



Define counts based on joint assignments to  $\mathbf{x}_{-i}, \mathbf{z}_{-i}$

$$m_{b,c} = \sum_{j \neq i} [z_{j-1} = b][z_j = c]$$
$$m_b = \sum_{c=1}^C m_{b,c}$$

$$P(z_i = c | \mathbf{x}_{-i}, \mathbf{z}_{-i}, \alpha, \beta) \quad \text{note that} \quad \begin{cases} z_{i-1} = b \\ z_{i+1} = d \end{cases} \quad \text{is in } \mathbf{z}_{-i}$$

$$\propto \frac{m_{b,c} + \alpha}{m_b + C\alpha} \times \frac{n_{c,k} + \beta}{n_c + K\beta} \times \frac{m_{c,d} + \alpha}{m_c + C\alpha}$$

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$



# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$
- ▶ possible states are assignments to the variables in the model

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$
- ▶ possible states are assignments to the variables in the model
- ▶ by defining  $\mathbf{P}$  properly we guarantee that  $\pi$  is the true posterior

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$
- ▶ possible states are assignments to the variables in the model
- ▶ by defining  $\mathbf{P}$  properly we guarantee that  $\pi$  is the true posterior
- ▶ once the chain has converged each  $y_t$  will be a sample from the posterior

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$
- ▶ possible states are assignments to the variables in the model
- ▶ by defining  $\mathbf{P}$  properly we guarantee that  $\pi$  is the true posterior
- ▶ once the chain has converged each  $y_t$  will be a sample from the posterior
- ▶ we can design  $\mathbf{P}$  by decomposing it  $P_1 \cdots P_M$   
where each component satisfies  $P_k(y, y')\pi(y) = P_k(y', y)\pi(y')$

# Markov Chain Monte Carlo

We draw from the posterior  $P(\mathbf{z}|\mathbf{x})$  via a Markov chain of random states  $Y_1, \dots, Y_T$  where  $P(y_t|y_{<t}) = P(y_t|y_{t-1})$

- ▶ the transition probability from  $y$  to  $y'$  is coded in a matrix  $\mathbf{P}$   
 $P_{ij}$  corresponds to  $P(Y = i, Y = j)$
- ▶ under certain conditions the chain converges to a stationary distribution  $\pi$  such that  $\mathbf{P}\pi = \pi$
- ▶ possible states are assignments to the variables in the model
- ▶ by defining  $\mathbf{P}$  properly we guarantee that  $\pi$  is the true posterior
- ▶ once the chain has converged each  $y_t$  will be a sample from the posterior
- ▶ we can design  $\mathbf{P}$  by decomposing it  $P_1 \cdots P_M$   
where each component satisfies  $P_k(y, y')\pi(y) = P_k(y', y)\pi(y')$
- ▶ applying each of  $P_k$  in turn or choosing  $P_k$  at random produces a  $\mathbf{P}$  that satisfies the necessary conditions

# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

To obtain a new state we

1. start a draft state  $\mathbf{z} = \mathbf{z}^{(t-1)}$

# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

To obtain a new state we

1. start a draft state  $\mathbf{z} = \mathbf{z}^{(t-1)}$
2. repeat for  $i = 1, \dots, N$



# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

To obtain a new state we

1. start a draft state  $\mathbf{z} = \mathbf{z}^{(t-1)}$
2. repeat for  $i = 1, \dots, N$ 
  - ▶ resample  $Z_i \sim P(z_i | \mathbf{x}_{-i}, \mathbf{z}_{-i})$   
only variables in the Markov blanket of  $z_i$  play a role  
that's why this is feasible

# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

To obtain a new state we

1. start a draft state  $\mathbf{z} = \mathbf{z}^{(t-1)}$
2. repeat for  $i = 1, \dots, N$ 
  - ▶ resample  $Z_i \sim P(z_i | \mathbf{x}_{-i}, \mathbf{z}_{-i})$   
only variables in the Markov blanket of  $z_i$  play a role  
that's why this is feasible
3. after complete pass over the data we have a new state  $\mathbf{z}^{(t)}$

# Gibbs sampler

We want to sample from  $P(\mathbf{z}|\mathbf{x})$  with a Markov chain  
a state  $y_t = \mathbf{z}^{(t)}$  is the  $t$ -th assignment to  $\mathbf{z}$

To obtain a new state we

1. start a draft state  $\mathbf{z} = \mathbf{z}^{(t-1)}$
2. repeat for  $i = 1, \dots, N$ 
  - ▶ resample  $Z_i \sim P(z_i|\mathbf{x}_{-i}, \mathbf{z}_{-i})$   
only variables in the Markov blanket of  $z_i$  play a role  
that's why this is feasible
3. after complete pass over the data we have a new state  $\mathbf{z}^{(t)}$

When we have collected a large number  $T$  of samples

- ▶ we can summarise the distribution and/or make decisions

# Summary

- ▶ Friends don't let friends optimise

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates
- ▶ Comparing Bayesian models is easier



# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates
- ▶ Comparing Bayesian models is easier
- ▶ Bayesian modelling requires some maths ;)

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates
- ▶ Comparing Bayesian models is easier
- ▶ Bayesian modelling requires some maths ;)
- ▶ Some families enjoy analytically available posteriors

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates
- ▶ Comparing Bayesian models is easier
- ▶ Bayesian modelling requires some maths ;)
- ▶ Some families enjoy analytically available posteriors
- ▶ Inference can be done by simulation (MCMC)

# Summary

- ▶ Friends don't let friends optimise
- ▶ Bayesian modelling is not only about prior specification
- ▶ Bayesian modelling is about uncertainty quantification
- ▶ Bayesians compare models (a set of assumptions)  
not point estimates
- ▶ Comparing Bayesian models is easier
- ▶ Bayesian modelling requires some maths ;)
- ▶ Some families enjoy analytically available posteriors
- ▶ Inference can be done by simulation (MCMC)

For more, take ML4NLP ;D

# References I