Bayesian Machine Learning Course 67564

Solution To Exercise 3: Evidence and Kernels

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Contents

Th	eoretical
1.1	Input-Specific Noise
	1.1.1 Q1 $p(y \theta) \sim \mathcal{N}(H\theta, \Gamma)$
	1.1.2 Q2 $\hat{\theta}_{MLE} = (H^T \Gamma^{-1} H)^{-1} H^T \Gamma^{-1} y \dots$
	1.1.3 Q3 $\theta \sim \mathcal{N}(\mu_0, \Sigma_0) \rightarrow p(\theta y) = ?$
1.2	
	1.2.1 Q4 $k(x,y) = \sum_{i} \sum_{j} g_{i}(x) f_{j}(x) f_{j}(y) g_{i}(y) \dots$
	1.2.2 Q5 $k(x,y) = \overline{k_1(x,y)} \cdot k_2(x,y)$ Is A Valid Kernel
1.3	Kernel Functions
	Kernel Functions
	1.3.2 Q7 $k(x,y) = k_1(x,y) - k_2(x,y)$ Is Not Valid
	1.3.3 Q8 $k(x,y) = k_a(x_a, y_a) + k_b(x_b, y_b)$ Is Valid
	1.3.4 Q9 $k(x,y) = \sqrt{\ell^T(x)\ell(y)}$ Is Not Valid
1.4	Evidence in the Dual Space
	1.4.1 Q10 $p(y k(\cdot,\cdot),\sigma_2) = \mathcal{N}(y 0,K+I\sigma^2)$
Pr	actical
2.1	Evidence for Artificial Functions
	2.1.1 $f_1(x) = x^2 - 1$
	$2.1.2 f_2(x) = -x^4 + 3x^2 + 50\sin\left(\frac{x}{6}\right) \dots \dots \dots \dots \dots \dots \dots \dots \dots $
	$2.1.3 f_3(x) = \frac{1}{2}x^6 - 0.75x^4 + 2.75x^2 \dots \dots \dots \dots \dots \dots \dots \dots \dots $
	$2.1.4 f_4(x) = \frac{5}{1+e^{-4x}} - \begin{cases} x & x-2>0\\ 0 & o.w. \end{cases}$
	2.1.5 $f_5(x) = \cos 4x + 4 x-2 $
2.2	Estimating the Sample Noise for Temperature Prediction
	2.2.1 Q7 No

1 Theoretical

1.1 Input-Specific Noise

$$y(x_i) = \theta h(x_i) + \eta_i \qquad \eta_i \sim \mathcal{N}\left(0, \sigma_i^2\right)$$

1.1.1 Q1 $p(y|\theta) \sim \mathcal{N}(H\theta, \Gamma)$

For each i:

$$p(y_i|\theta) \sim \mathcal{N}\left(y_i \mid \theta^T h(x_i), I\sigma_i^2\right)$$

Explicitly:

$$p(y_i|\theta) = \frac{1}{\sqrt{2\pi\sigma_i^2}} \exp\left[-\frac{1}{2\sigma_i^2} \left(h(x_i)^T \theta - y_i\right)^2\right]$$

Denote $\eta = \mathcal{N}(0, \Gamma)$ where:

$$\Gamma = daig(\left\{\sigma_i^2\right\}_{i=1}^n) = \begin{bmatrix} \sigma_1^2 & 0 \\ & \ddots & \\ 0 & \sigma_n^2 \end{bmatrix}$$

As y is an affine transformation of θ and η we get:

$$p(y|\theta) \sim \mathcal{N}(H\theta, \Gamma)$$

Explicitly:

$$p(y|\theta) = \frac{1}{\sqrt{(2\pi)^d |\Gamma|}} \exp\left[-\frac{1}{2} (H\theta - y)^T \Gamma^{-1} (H\theta - y)\right]$$

1.1.2 Q2
$$\hat{\theta}_{MLE} = (H^T \Gamma^{-1} H)^{-1} H^T \Gamma^{-1} y$$

We write the log-likelihood:

$$\ell(y|\theta) = \ln \mathcal{N} (H\theta, \Gamma) = -\frac{1}{2} (H\theta - y)^T \Gamma^{-1} (H\theta - y) + const$$

So in fact we try to minimize:

$$L = \frac{1}{2} (H\theta - y)^T \Gamma^{-1} (H\theta - y)$$

And so:

$$\begin{split} \frac{\partial L}{\partial \theta} &= H^T \Gamma^{-1} (H\theta - y) \stackrel{!}{=} 0 \\ iff H^T \Gamma^{-1} y &= H^T \Gamma^{-1} H\theta \\ iff \hat{\theta}_{MLE} &= \left(H^T \Gamma^{-1} H\right)^{-1} H^T \Gamma^{-1} y \end{split}$$

1.1.3 Q3 $\theta \sim \mathcal{N}(\mu_0, \Sigma_0) \rightarrow p(\theta|y) = ?$

With Bayes' law we know: $p(\theta|y) \propto p(\theta)p(y|\theta)$ and so $p(\theta|y) \sim \mathcal{N}\left(\mu_{\theta|y}, C_{\theta|y}\right)$ as the product of 2 Gaussian. Also

$$\mu_{\theta|y} = \underset{\theta}{\operatorname{argmax}} p(\theta) p(y|\theta) = \underset{\theta}{\operatorname{argmax}} \ln p(\theta) p(y|\theta)$$

$$C_{\theta|y} = -\left(\frac{\partial^2}{\partial \theta^2} p(\theta) p(y|\theta)\right)^{-1}$$

So we can directly compute the values:

$$\begin{split} \frac{\partial}{\partial \theta} \ln p(\theta) p(y|\theta) + const &= \frac{\partial}{\partial \theta} \left[-\frac{1}{2} \left(\theta - \mu_0 \right)^T \Sigma_0^{-1} \left(\theta - \mu_0 \right) - \frac{1}{2} \left(H \theta - y \right)^T \Gamma^{-1} \left(H \theta - y \right) \right] \\ &= -\Sigma_0^{-1} (\theta - \mu_0) - H^T \Gamma^{-1} (H \theta - y) \\ &= -\Sigma_0^{-1} \theta + \Sigma_0^{-1} \mu_0 - H^T \Gamma^{-1} H \theta + H^T \Gamma^{-1} y \\ &= \Sigma_0^{-1} \mu_0 + H^T \Gamma^{-1} y - \left(\Sigma_0^{-1} + H^T \Gamma^{-1} H \right) \theta \stackrel{!}{=} 0 \\ ⇔ \ \Sigma_0^{-1} \mu_0 + H^T \Gamma^{-1} y = \left(\Sigma_0^{-1} + H^T \Gamma^{-1} H \right) \theta \\ ⇔ \ \mu_{\theta|y} = \left(\Sigma_0^{-1} + H^T \Gamma^{-1} H \right)^{-1} \left(\Sigma_0^{-1} \mu_0 + H^T \Gamma^{-1} y \right) \end{split}$$

And:

$$\frac{\partial^2}{\partial \theta^2} p(\theta) p(y|\theta) = \frac{\partial}{\partial \theta} \left[-\Sigma_0^{-1} (\theta - \mu_0) - H^T \Gamma^{-1} H \theta + H^T \Gamma^{-1} y \right]$$
$$= -\Sigma_0^{-1} - H^T \Gamma^{-1} H$$

So:

$$C_{\theta|y} = \left(\Sigma_0^{-1} + H^T \Gamma^{-1} H\right)^{-1}$$

Finally:

$$p(\theta|y) \sim \mathcal{N}\left(\left(\Sigma_0^{-1} + H^T \Gamma^{-1} H\right)^{-1} \left(\Sigma_0^{-1} \mu_0 + H^T \Gamma^{-1} y\right), \left(\Sigma_0^{-1} + H^T \Gamma^{-1} H\right)^{-1}\right)$$

1.2 Product of Kernels

1.2.1 Q4
$$k(x,y) = \sum_{i} \sum_{j} g_i(x) f_j(x) f_j(y) g_i(y)$$

Let $\{x_i\}_{i=1}^N\subseteq\mathbb{R}^N$. Denote $X=span\left(\{x_i\}_{i=1}^N\right)$ and $K=[k(x_i,x_j)]_{i,j=1}^N$ Gram's matrix for $k(\cdot,\cdot)$ and $\{x_i\}_{i=1}^N$. Since $k(\cdot,\cdot)$ is a valid kernel function, we can write $K=RR^T$ for some R and U is a basis transformation matrix such that $Ux_i=e_i$. Now:

$$f(x) \stackrel{\triangle}{=} R^T U x$$
$$f^T(x_i) f(x_j) = \left(R^T U x_i\right)^T R^T U x_j = x^T U^T R R^T U x = e_i K e_j = [K]_{i,j} = k(x_i, x_j)$$

Now let f,g be functions which hold the above equality for k_1,k_2 respectively. So:

$$k(x,y) = k_1(x,y) \cdot k_2(x,y) = f^T(x)f(y) \cdot g^T(x)g(y)$$

Denote $f^T(x) = [\cdots, f_i(x), \cdots]$ and the same for g. Now:

$$k(x,y) = \left(\sum_{i} f_i(x)f_i(y)\right) \cdot \left(\sum_{j} g_j(x)g_j(y)\right) = \sum_{j} \left(\sum_{i} f_i(x)f_i(y)g_j(x)g_j(y)\right)$$

Q.E.D.

1.2.2 Q5 $k(x,y) = k_1(x,y) \cdot k_2(x,y)$ **Is A Valid Kernel**

Finally, given such kernels and their respective functions f and g (as in Q4) we define:

$$h(x) = reshape \left(f(x)g^{T}(x), (, N \times N) \right) = reshape \left(\begin{bmatrix} f_{1}(x)g_{1}(x) & \cdots & f_{1}(x)g_{n}(x) \\ \vdots & \ddots & \vdots \\ f_{n}(x)g_{1}(x) & \cdots & f_{n}(x)g_{n}(x) \end{bmatrix}, (, N \times N) \right)$$

$$= \left[f_{1}(x)g_{1}(x), \dots, f_{1}(x)g_{n}(x), f_{2}(x)g_{1}(x), \dots, f_{2}(x)g_{n}(x), f_{n}(x)g_{1}(x), \dots, f_{n}(x)g_{n}(x) \right]^{T}$$

So we get:

$$h^{T}(x)h(y) = \sum_{i} \sum_{i} f_{i}(x)g_{i}(x)f_{j}(y)g_{j}(y)$$

and $k(\cdot,\cdot)$ is a valid kernel as a dot product of to vector functions. Q.E.D

1.3 Kernel Functions

1.3.1 Q6
$$k(x,y) = \exp\left[\beta \left\|g(x) - g(y)\right\|^2\right]$$
 Is Not Valid

Let $\beta = 1$ and $g(x) \stackrel{\Delta}{=} x$ so, for $D = \{x, y\}$ we get:

$$K = \left[\begin{array}{cc} 1 & \exp\left[\|x - y\|\right] \\ \exp\left[\|x - y\|\right] & 1 \end{array} \right]$$

and

$$det(K) = 1 - \exp[2||x - y||]$$

So, det(K) < 0 iff $1 < \exp[2\|x - y\|]$ iff $0 < 2\|x - y\|$ iff $0 < \|x - y\|$. I.e. for every strictly different x, y K is not PSD. If it where, we'd get $0 \le det(K)$ which isn't the case here. Q.E.D.

Note $p(x,y) \stackrel{\Delta}{=} \|g(x) - g(y)\|^2 = (g(x) - g(y))^T I(g(x) - g(y))$ is a valid kernel as I is PD. So, for any constant $\beta > 0$ it holds that $q \stackrel{\Delta}{=} \beta \cdot p$ is a valid kernel. Lastly, $k = \exp(q(x,y))$ is a valid kernel .Q.E.D.

1.3.2 Q7
$$k(x,y) = k_1(x,y) - k_2(x,y)$$
 Is Not Valid

Let $k_1(x,y) \stackrel{\triangle}{=} xy$. K is a valid kernel as a product of the identity function of real numbers. Also $k_2(x,y) = 2 \cdot k_1(x,y)$ is a valid kernel as the multiplication of a valid kernel by a positive constant. Now $k(x,y) = k_1(x,y) - k_2(x,y) = -xy$ is not a valid kernel because $\forall x \to k(x,x) \le 0$, specifically for the unit vector e_1 and k's Gram matrix K we get:

$$e_1^T K e_1 = -1$$

So K is not a PSD.

1.3.3 Q8 $k(x,y) = k_a(x_a, y_a) + k_b(x_b, y_b)$ Is Valid

Note that in this case k's Gram matrix $K = K_a + K_b$. For each vector v it holds that $0 \le v^T K_x v$ (x=a,b) as k_x is a valid kernel and so:

$$0 < v^T K_a v + v^T K_b v = v^T K v$$

Q.E.D.

1.3.4 Q9
$$k(x,y) = \sqrt{\ell^T(x)\ell(y)}$$
 Is Not Valid

Proof online...

1.4 Evidence in the Dual Space

1.4.1 Q10
$$p(y|k(\cdot,\cdot),\sigma_2) = \mathcal{N}(y|0,K+I\sigma^2)$$

Note $y = K\alpha + \eta$ and as such is an affine transformation of a Gaussian and so - a Gaussian. We just need to find its mean and cov:

$$\mathbb{E}\left[y|k(\cdot,\cdot),\sigma_2\right] = \mathbb{E}\left[K\alpha + \eta\right] = K\mathbb{E}\left[\alpha\right] + \mathbb{E}\left[\eta\right] = K0 + 0 = 0$$

$$\operatorname{cov}\left[y|k(\cdot,\cdot),\sigma_{2}\right]=\operatorname{cov}\left[K\alpha+\eta\right]=\operatorname{cov}\left[K\alpha\right]+\operatorname{cov}\left[\eta\right]=K^{T}\operatorname{cov}\left[\alpha\right]K+I\sigma^{2}=K^{T}K^{-1}K+I\sigma^{2}=K+I\sigma^{2}$$

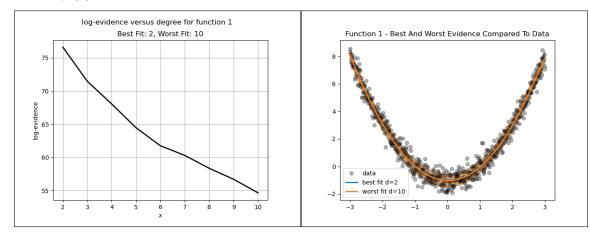
Where the last transition is due to K's symmetry. Hence:

$$p(y|k(\cdot,\cdot),\sigma_2) = \mathcal{N}(y|0,K+I\sigma^2)$$

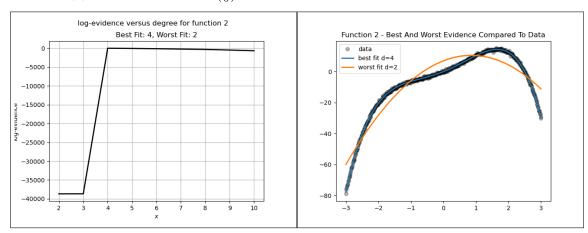
2 Practical

2.1 Evidence for Artificial Functions

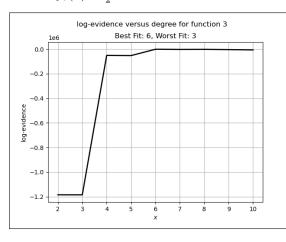
2.1.1 $f_1(x) = x^2 - 1$

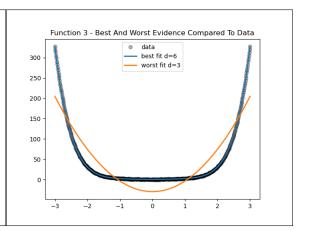


2.1.2 $f_2(x) = -x^4 + 3x^2 + 50\sin\left(\frac{x}{6}\right)$

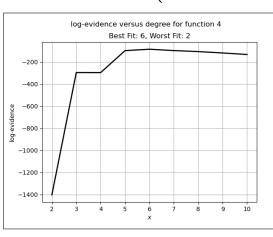


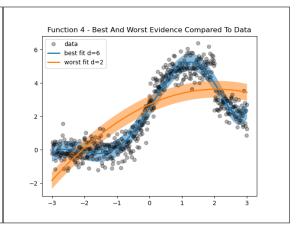
2.1.3 $f_3(x) = \frac{1}{2}x^6 - 0.75x^4 + 2.75x^2$



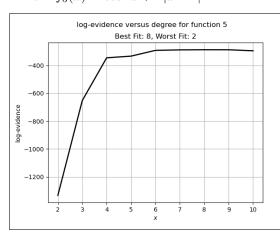


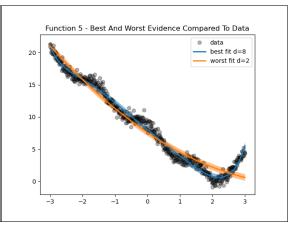
2.1.4
$$f_4(x) = \frac{5}{1+e^{-4x}} - \begin{cases} x & x-2 > 0 \\ 0 & o.w. \end{cases}$$





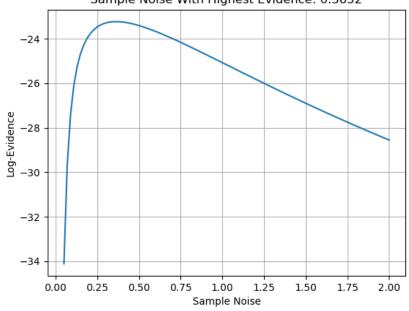
2.1.5 $f_5(x) = \cos 4x + 4|x-2|$





2.2 Estimating the Sample Noise for Temperature Prediction

Log-Evidence Score As A Function Of Sample Noise Sample Noise With Highest Evidence: 0.3652



2.2.1 Q7 No

The sample noise with the highest evidence is not necessarily the sample noise of the original measurements. When creating our model we assume a Gaussian prior on θ , we assume the noise is also a Gaussian and we assume the temperatures were decided upon using a polynomial with degree ≤ 7 . In case those assumptions hold we cold say we've found the noise. As it's extremely unlikely though, the noise with the high-test evidence is not necessarily the sample noise of the original measurements.