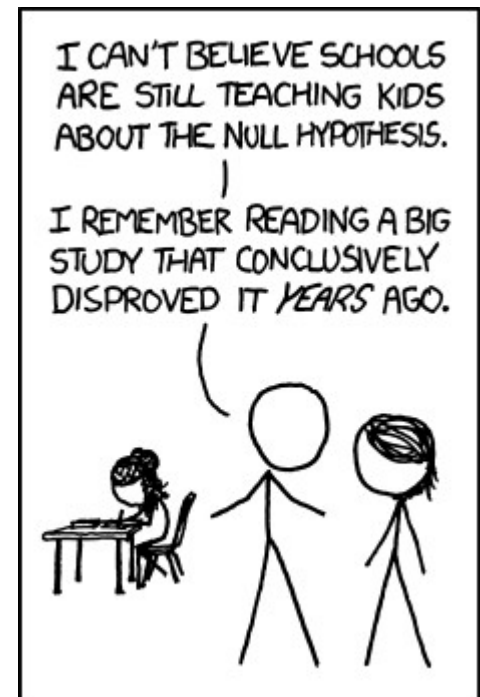


Modern Particle Physics Experiments



Many aspects of particle physics are of stochastic nature, e.g.:

- **outputs of experiments**

The experiment results are test for various theoretical explanations.

How can we decide between theories?

→ use statistical hypothesis testing paradigm

“Theory” – this can be a different explanation of some phenomena:

- “SUSY particles exist” vs “SUSY particles do not exist”

or a statement about a value of some parameter:

- “electron neutrino mass is lower than $0.9 \text{ eV}/c^2$ ” vs

“electron neutrino mass is lower than $0.1 \text{ eV}/c^2$ ”

There is an everlasting confusion about answer of question

“Probability of what?” Two points of view exist:

- frequentist (“classic”): **probability = a fraction of desired outcomes out of many (imagined) trials**
- one **can not** talk about “probability that a theory is correct”, or “mass of particle X is within limits” as many trials on those would require many Universes
- one **can talk** about “probability that given (random) range contains a true value”

Let us use standard ruler, with 0.1cm ticks, to measure length of a pencil.

Measurement tool reading: 15 cm

Measurement uncertainty: $0.1/\sqrt{12} = 0.03$ cm

What is the length of a pencil? 15 ± 0.03 cm

The full sentence: “There is gaussian 1σ probability (68.2%) that a range $[15-0.03, 15+0.03]$ cm covers the true length of a pencil”.

Never, say: “The probability that the length of a pencil ...” as there is only one pencil.

- Bayesian: **probability = a subjective degree of belief**
 - one **can talk** about probability of whatever, since this is your degree of belief that given statement is true
 - can be coherent between different people, give rigid prescription of assigning the probability

Bayesian reasoning usually uses the Bayes theorem:

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)}$$

posterior probability \nearrow $P(B|A)$ \nwarrow prior probability $P(B)$

Assume some experiment result: x_0

Assume statistical model, or

hypothesis $H(\theta)$, for probability density function (pdf) of x :

$$P(x|H) = P(x|\theta)$$

Likelihood: the value of the probability for given experiment result (this is **not a probability!**):

$$\mathcal{L}(x_0, \theta) = P(x_0|H) = P(x_0|\theta)$$

Assume some there are two models, or hypotheses:

H_0 – null hypothesis

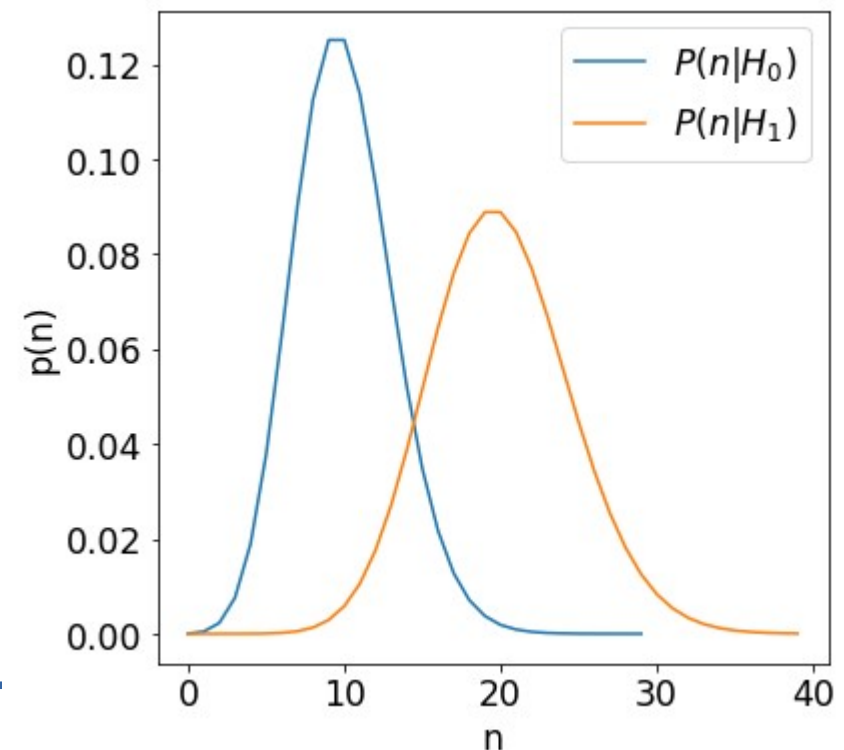
H_1 – alternative hypothesis

Hypotheses provide statistical models for experiment outcome:

$$P(x|H_0)$$

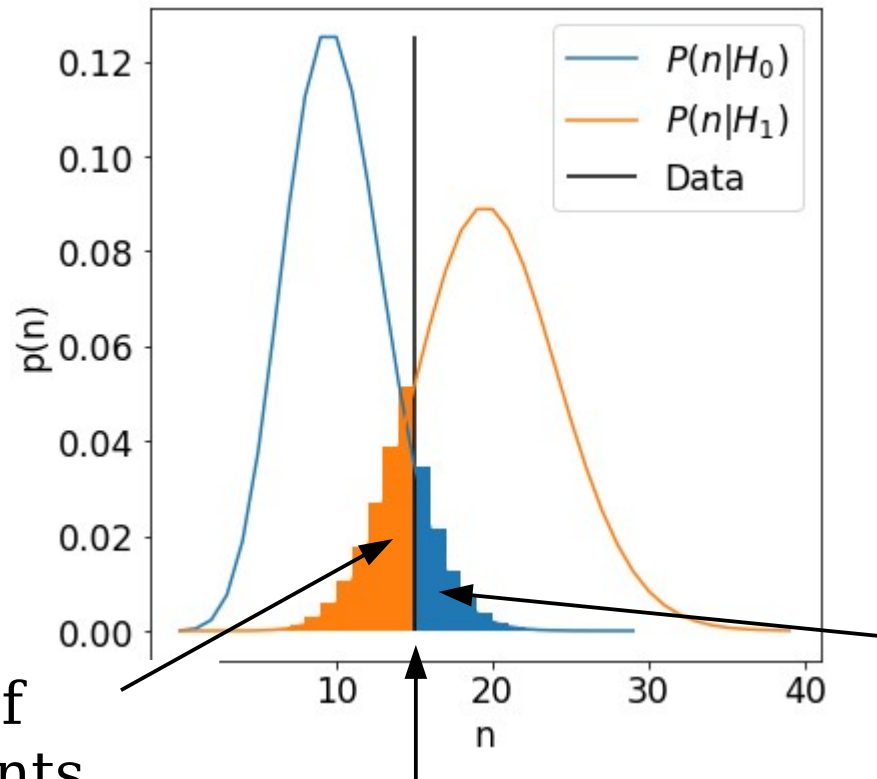
$$P(x|H_1)$$

Question: can we reject H_0 given the observed experiment outcome?



Definitions:

- α (Type I error) – probability of rejecting H_0 when it is true. α is a **significance of the test**.
- β (Type II error) – probability of accepting H_0 when it is false



α - probability of getting more events than some threshold, when H_0 when is true

β - probability of getting less events than some threshold, when H_1 when is true

decision
threshold

Optimization task:

- minimize β (Type II error) for given value of α (Type I error)

Solution:

- do not look at observable (e.g. event count) p.d.f, but at the p.d.f for the observable likelihood ratio:

$$\lambda = \frac{\mathcal{L}(x_0|H_0)}{\mathcal{L}(x_0|H_1)}$$

- or more computationally convenient:

$$t = -2 \ln \frac{\mathcal{L}(x_0|H_0)}{\mathcal{L}(x_0|H_1)}$$

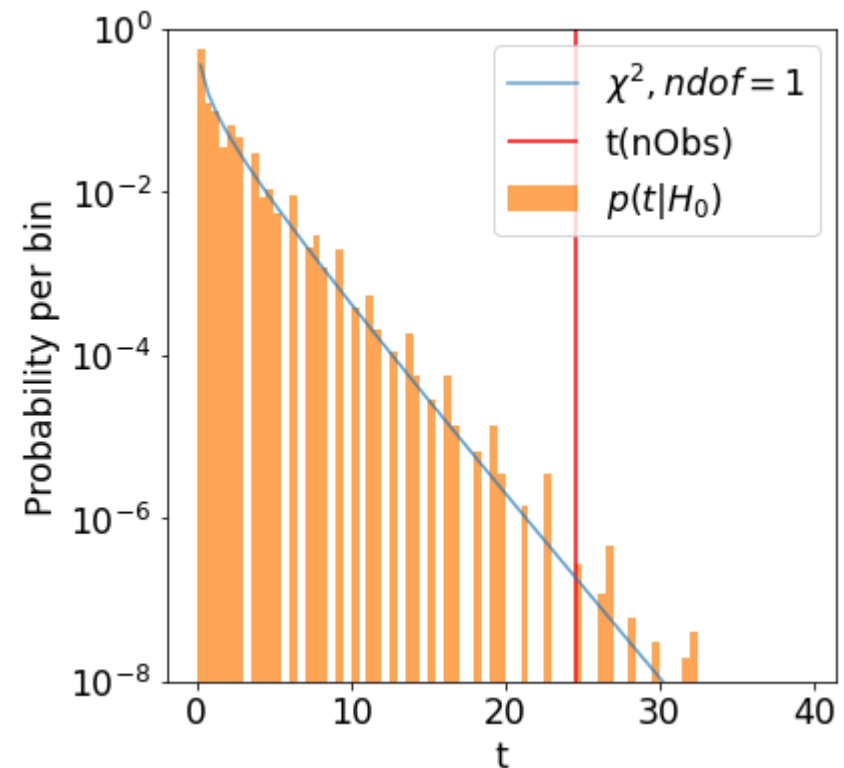
Algorithm:

- choose the α value (Type I error) **before** the experiment.
- calculate the likelihood ratio distribution, and find t_0 :

$$t_0 : \int_{t_0}^{\infty} P(t|H_0) dt = \alpha$$

(luckily, often the distribution of t can be approximated by Gaussian distribution, and one can use analytical approximations)

- look at the data, and calculate t_{DATA}
- reject H_0 if $t_{\text{DATA}} > t_0$



Wilks theorem: in the limit of large amount of data the likelihood ratio for data drawn from H_0 model approaches the χ^2 distribution with number of degrees equal to difference of number of free parameters in H_0 and H_1 :

$$P(t|H_0) \rightarrow \chi^2_{dof H_1 - dof H_0}$$

If there is only one free parameter (number of signal events, S) we have:

$$P(t|H_0) \rightarrow \chi^2_1$$

$$p_0 = \int_{t_{obs}}^{\infty} P(t|H_0) dt = 2(1 - \Phi(\sqrt{t_{obs}}))$$

↑
often called
p-value

↑
Gaussian
cumulative
probability
function (cdf)

The t variable is sensitive for both up- and -downward fluctuations. In case of looking for a new signal, one is interested only in additional events on top of known background. In that case a modified variable is better:

$$q_0 = \begin{cases} -2 \ln \frac{\mathcal{L}(x_0 | H_0 = B)}{\mathcal{L}(x_0 | H_1 = \hat{S} + B)} & \hat{S} > 0 \\ 0 & \hat{S} < 0 \end{cases}$$

$$p_0 = \int_{q_{0,obs}}^{\infty} P(q | H_0) dt = 1 - \Phi(\sqrt{q_{0,obs}})$$

$$z = \Phi^{-1}(1 - p_0) = \sqrt{q_0}$$

Number of single sided Gaussian sigma corresponding to p_0

Gaussian cumulative probability function (cdf)

Poisson case:

H_0 - expected number of events is B

H_1 - expected number of events is μ_{H_1} - the data best fit value: $\mu_{H_1} = nObs$

$$q_0(nObs) = -2 \ln \frac{\mathcal{L}(x_0|H_0)}{\mathcal{L}(x_0|H_1)} = -2 \ln \left(\frac{B^{nObs} e^{-B}}{nObs!} \middle/ \frac{(\mu_{H_1})^{nObs} e^{-\mu_{H_1}}}{nObs!} \right) =$$

$$-2 \ln \left(\frac{B^{nObs}}{(nObs)^{nObs}} e^{-B+nObs} \right) = -2 [nObs (\ln(B) - \ln(nObs)) + nObs - B] =$$

$$2 [nObs (\ln(nObs) - \ln(B)) + B - nObs]$$

- if B is known, then $S = nObs - B$
- if we plan an experiment for some known S and B , then expected $nObs = S + B$

$$\begin{aligned}
 z &\simeq \sqrt{q_0} = \sqrt{2[nObs[\ln(nObs) - \ln(B)] + B - nObs]} = |nObs = S + B| = \\
 &= \sqrt{2[(S+B)\ln(\frac{S+B}{B}) + B - B - S]} = \sqrt{2[(S+B)\ln(1+\frac{S}{B}) - S]} = \\
 &\sqrt{2[B(1+\frac{S}{B})\ln(1+\frac{S}{B}) - B\frac{S}{B}]} \simeq |x = \frac{S}{B}, S \ll B \rightarrow x \ll 1| \simeq \\
 &\sqrt{2B[(1+x)\log(1+x) - x]} \simeq \sqrt{2B[(1+x)(x - \frac{x^2}{2}) - x]} \simeq |\text{terms up to } x^2| \\
 &\sqrt{2B[x - \frac{x^2}{2} + x^2 - x]} = \sqrt{2B[\frac{x^2}{2}]} = \sqrt{\frac{S^2}{B}} = \frac{S}{\sqrt{B}}
 \end{aligned}$$

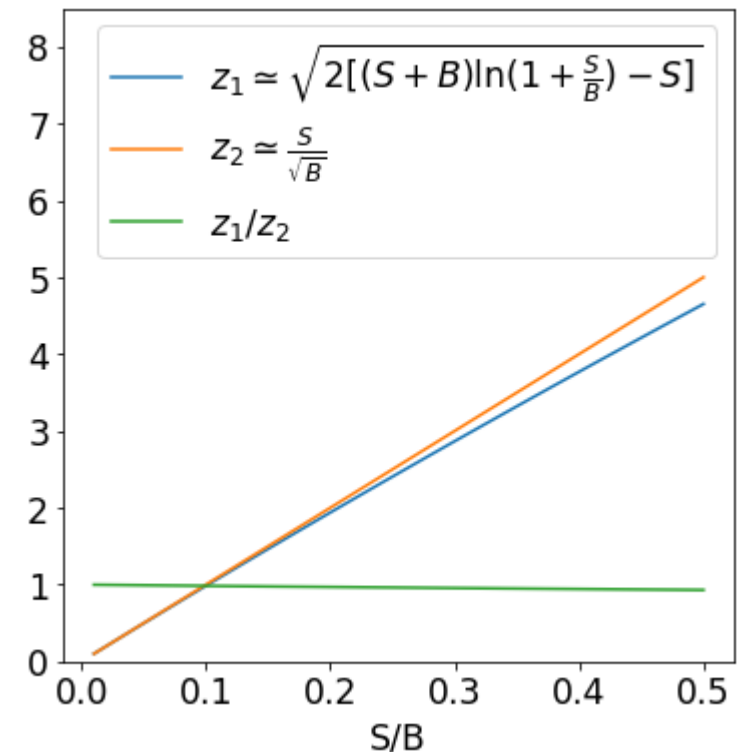
• large amount of data: $z \simeq \sqrt{q_0} = \sqrt{2[nObs[\ln(nObs) - \ln(B)] + B - nObs]}$

• in addition $S \ll B$: $z \simeq \frac{S}{\sqrt{B}}$

A dataset for which the best fit values reproduce the assumed values, e.g $S_{\text{best fit}} = S$ is called

“Asimov dataset”

Similarly putting $nObs = S + B$ is called using Asimov values for observation.



Discovery:

H_0 – background only hypothesis (B)

H_1 – signal + background hypothesis (S+B)

test statistics - q_0

Usual significance of the test: **5σ** - **$\alpha = 2.87 \cdot 10^{-7}$**

Exclusion:

H_0 – signal+background hypothesis (S+B)

H_1 – background only hypothesis (B)

test statistics - q_μ

Usual significance of the test: **95 CL** - **$\alpha = 0.05$**

Discovery:

$$q_0 = \begin{cases} -2 \ln \frac{\mathcal{L}(x_0 | H_0 = B)}{\mathcal{L}(x_0 | H_1 = \hat{S} + B)} & \hat{S} > 0 \\ 0 & \hat{S} < 0 \end{cases}$$

Exclusion:

$$q_\mu = \begin{cases} -2 \ln \frac{\mathcal{L}(x_0 | H_0 = S + B)}{\mathcal{L}(x_0 | H_1 = \hat{S} + B)} & \hat{S} < S \\ 0 & \hat{S} > S \end{cases}$$

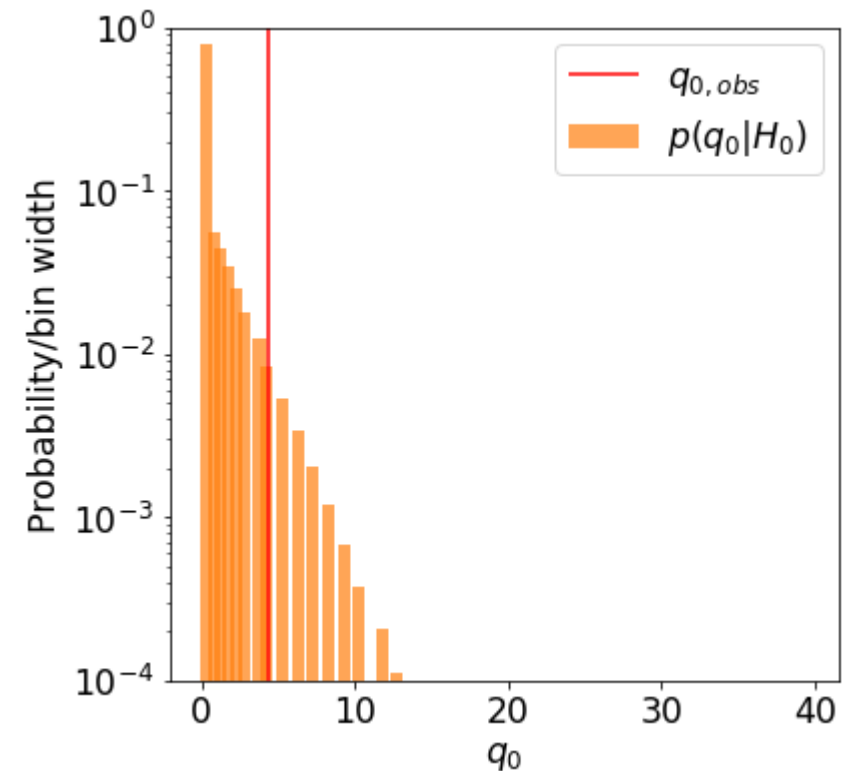
Important:

\hat{S} is the best fit S value to the observed data!

Assumptions and observations:

$$\alpha = 2.87 \cdot 10^{-7}, B = 20, N_{\text{data}} = 30$$

$$S_{\text{best fit}} = 10$$



Can we reject the null hypothesis?

`q0_Obs: 4.33`

`MC estimate for $p(q_0 > q_{0_Obs})$: 1.35E-02, single sided Gaussian sigma: 2.2`

`Gauss estimate for $p(q_0 > q_{0_Obs})$: 1.87E-02, single sided Gaussian sigma: 2.1`

`alpha: 2.87E-07`

`Null hypothesis (B only) is NOT REJECTED`

Assumptions:

$$\alpha = 2.87 \cdot 10^{-7}, B = 20, S = 10$$

What would be significance for the nominal S?

$$z \simeq \sqrt{q_0}, \quad q_{0,Asimov} = q_0(nObs = S + B)$$

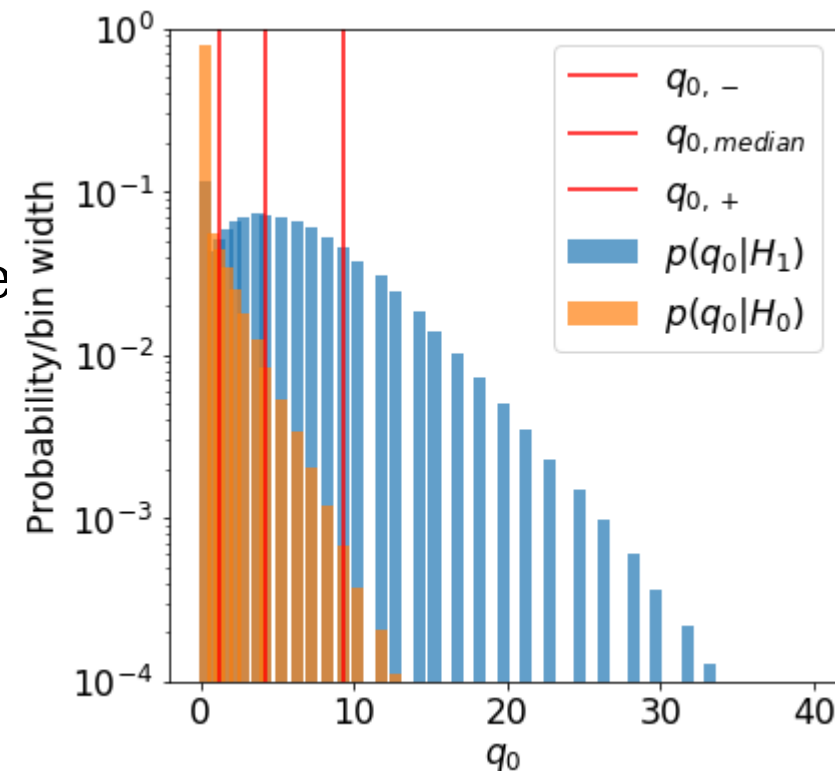
What would we observe in experiment?

q_0 values fluctuate according to number of observed events (assuming S+B), so we define:

$q_{0,\text{median}}$ – 50% of experiments will observe this or greater value

$q_{0,-}$ – 84% of experiments will observe this or greater value

$q_{0,+}$ – 16% of experiments will observe this or greater value



What significance of planned experiment we can expect?

Take the $q_{0,-/median/+}$ values calculated for S+B data, and check $p(q_{0,-/median/+} > t)$ assuming H_0 (B only)

`q0 median from Asimov dataset = 4.33`

`MC estimate for p(q0>q0_Asimov) under H1: 4.52E-01 (expected: 0.50)`

`q0 minus = 1.25, q0 median = 4.25, q0 plus = 9.25`

`MC estimate for p(q0>q0_minus) under H1: 7.93E-01 (expected: 0.84)`

`MC estimate for p(q0>q0_median) under H1: 4.52E-01 (expected: 0.50)`

`MC estimate for p(q0>q0_plus) under H1: 1.57E-01 (expected: 0.16)`

<code>Gauss estimate for p(q0>q0_minus):</code>	<code>1.32E-01,</code>	<code>single sided Gaussian sigma:</code>	<code>1.1</code>
<code>Gauss estimate for p(q0>q0_median):</code>	<code>1.96E-02,</code>	<code>single sided Gaussian sigma:</code>	<code>2.1</code>
<code>Gauss estimate for p(q0>q0_Asimov):</code>	<code>1.87E-02,</code>	<code>single sided Gaussian sigma:</code>	<code>2.1</code>
<code>Gauss estimate for p(q0>q0_plus):</code>	<code>1.18E-03,</code>	<code>single sided Gaussian sigma:</code>	<code>3.0</code>

Assumptions:

$$\alpha = 0.05, B = 200$$

What minimal number of signal events can be rejected for 50% of hypothetical experiments?

Expected answer:

Expected median exclusion on $S = 25$

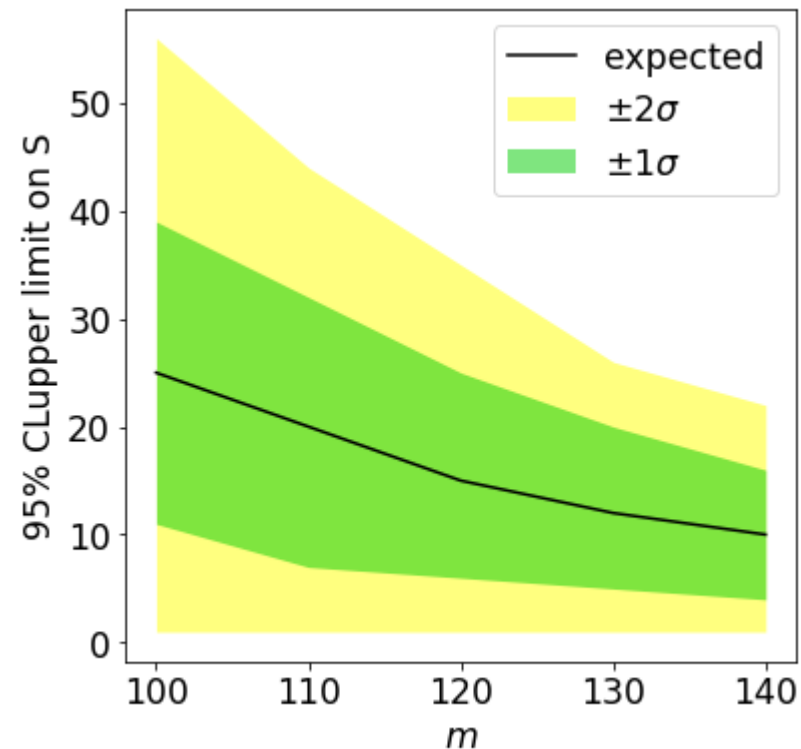
Assumptions:

- experiment is looking for a particle of unknown mass: $100 < m_x < 150$
- expected background is: $B(m_x) = 200 \cdot e^{-(X-100)/20}$

Task:

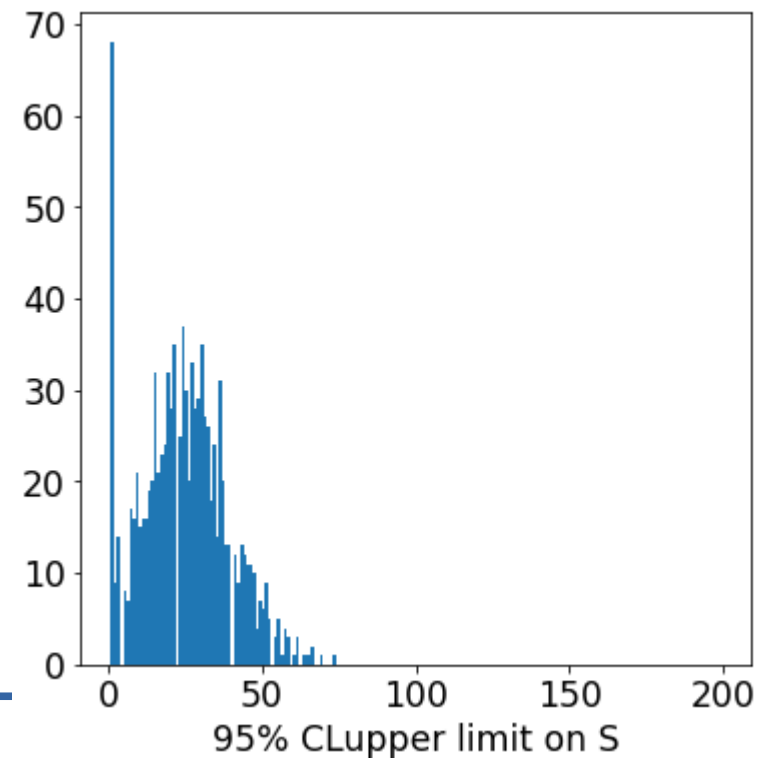
Plot excluded number of signal events for 97.6%, 84%, 50%, 16% 2.4% of experiments, as a function of m_x

Expected result:



Hints:

- generate 1000 pseudo experiments according to H_1 The number of observed events in each pseudo experiment follows $\text{Poisson}(B)$
- for each pseudo experiment test a number of S hypotheses, until S is excluded. The upper limit should be calculated using $H_1 = S + B$, where **B is the nominal B**
- create a histogram of excluded S values
- from histogram find the $\pm 2\sigma$, $\pm 1\sigma$ and median values of S



Review articles of probability and statistics on PDG portal:

- *The Review of Particle Physics*

Lectures on Statistics:

- B. Cousins, *Lectures on Statistics: in Theory: Prelude to Statistics in Practice*

More details on asymptotic formulas:

- G. Cowan, et al, *Asymptotic formulae for likelihood-based tests of new physics*