

432 Class 10 Slides

github.com/THOMASELOVE/2020-432

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Setup

```
library(magrittr); library(janitor); library(here)
```

```
library(skimr)
```

```
library(naniar)
```

```
library(simputation)
```

```
library(broom)
```

```
library(rms) # note: also loads Hmisc
```

```
library(tidyverse)
```

Today's Data

Heart and Estrogen/Progestin Study (HERS)

- Clinical trial of hormone therapy for the prevention of recurrent heart attacks and deaths among 2763 post-menopausal women with existing coronary heart disease (see Hulley et al 1998 and many subsequent references, including Vittinghoff, Chapter 4.)
- We're excluding the women in the trial with a diabetes diagnosis.

```
hers_raw <- read_csv(here("data/hersdata.csv")) %>%  
  clean_names()
```

```
hers1 <- hers_raw %>%  
  filter(diabetes == "no") %>%  
  select(subject, ldl, ht, age, smoking, drinkany, sbp,  
         physact, bmi, diabetes)
```

The Codebook (n = 2032)

Variable	Description
subject	subject code
HT	factor: hormone therapy or placebo
diabetes	yes or no (all are no in our sample)
ldl	LDL cholesterol in mg/dl
age	age in years
smoking	yes or no
drinkany	yes or no
sbp	systolic BP in mm Hg
physact	5-level factor, details next slide
bmi	body-mass index in kg/m^2

Goal Predict ldl using age, smoking, drinkany, sbp, physact and bmi, across both HT levels but restricted to women without diabetes.

The physact variable

```
hers1 %>% count(physact)
```

```
# A tibble: 5 x 2
```

	physact	n
	<chr>	<int>
1	about as active	674
2	much less active	107
3	much more active	252
4	somewhat less active	322
5	somewhat more active	677

Comparison is to activity levels for these women just before menopause.

Any missing data?

```
miss_var_summary(hers1)
```

```
# A tibble: 10 x 3
  variable n_miss pct_miss
  <chr>      <int>    <dbl>
1 ldl          7    0.344
2 drinkany     2    0.0984
3 bmi          2    0.0984
4 subject      0     0
5 ht           0     0
6 age          0     0
7 smoking      0     0
8 sbp          0     0
9 physact      0     0
10 diabetes    0     0
```

Single Imputation for drinkany, bmi and ldl

Since drinkany is a factor, we have to do some extra work to impute.

```
set.seed(432092)

hers2 <- hers1 %>%
  mutate(drinkany_n =
    ifelse(drinkany == "yes", 1, 0)) %>%
  impute_pmm(drinkany_n ~ age + smoking) %>%
  mutate(drinkany =
    ifelse(drinkany_n == 1, "yes", "no")) %>%
  impute_rlm(bmi ~ age + smoking + sbp) %>%
  impute_rlm(ldl ~ age + smoking + sbp + bmi)
```

Now, check missingness...

```
miss_var_summary(hers2)
```

```
# A tibble: 11 x 3
```

	variable <chr>	n_miss <int>	pct_miss <dbl>
1	subject	0	0
2	ldl	0	0
3	ht	0	0
4	age	0	0
5	smoking	0	0
6	drinkany	0	0
7	sbp	0	0
8	physact	0	0
9	bmi	0	0
10	diabetes	0	0
11	drinkany_n	0	0

Multiple Imputation using aregImpute from Hmisc

Model to predict all missing values of any variables, using additive regression bootstrapping and predictive mean matching.

Steps are:

- 1 aregImpute draws a sample with replacement from the observations where the target variable is observed, not missing.
- 2 It then fits a flexible additive model to predict this target variable while finding the optimum transformation of it.
- 3 It then uses this fitted flexible model to predict the target variable in all of the original observations.
- 4 Finally, it imputes each missing value of the target variable with the observed value whose predicted transformed value is closest to the predicted transformed value of the missing value.

Fitting a Multiple Imputation Model

```
set.seed(4320132)
dd <- datadist(hers1)
options(datadist = "dd")
fit3 <- aregImpute(~ ldl + age + smoking + drinkany +
                  sbp + physact + bmi,
                  nk = c(0, 3:5), tlinear = FALSE,
                  data = hers1, B = 10, n.impute = 20)
```

Iteration 1 Iteration 2 Iteration 3 Iteration 4 Iteration 5 It

Multiple Imputation using `aregImpute` from `Hmisc`

`aregImpute` requires specifications of all variables, and several other details:

- `n.impute` = number of imputations, we'll run 20
- `nk` = number of knots to describe level of complexity, with our choice `nk = c(0, 3:5)` we'll fit both linear models and models with restricted cubic splines with 3, 4, and 5 knots
- `tlinear = FALSE` allows the target variable to have a non-linear transformation when `nk` is 3 or more
- `B = 10` specifies 10 bootstrap samples will be used
- `data` specifies the source of the variables

aregImpute Imputation Results (1 of 4)

```
fit3
```

Multiple Imputation using Bootstrap and PMM

```
aregImpute(formula = ~ldl + age + smoking + drinkany + sbp +  
  physact + bmi, data = hers1, n.impute = 20, nk = c(0, 3:5),  
  tlinear = FALSE, B = 10)
```

```
n: 2032      p: 7      Imputations: 20      nk: 0
```

Number of NAs:

ldl	age	smoking	drinkany	sbp	physact	bmi
7	0	0	2	0	0	2

fit3 Imputation Results (2 of 4)

R-squares for Predicting Non-Missing Values for Each Variable Using Last Imputations of Predictors

ldl	drinkany	bmi
0.041	0.014	0.109

fit3 Imputation Results (3 of 4)

Resampling results for determining the complexity of imputation models

Variable being imputed: ldl

Bootstrap bias-corrected summaries:

Statistic	nk = 0	nk = 3	nk = 4	nk = 5
R ²	0.0139	0.0149	0.00776	0.0124
mean absolute error	28.3594	42.9139	44.09937	39.8266
median abs. error	22.8301	35.5441	38.85302	32.6386

10-fold cross-validated:

Statistic	nk = 0	nk = 3	nk = 4	nk = 5
R ²	0.0214	0.0180	0.01517	0.0191
mean absolute error	145.7176	43.5007	45.02428	44.2456
median abs. error	141.4238	36.4102	38.88053	37.3141

fit3 Imputation Results (4 of 4)

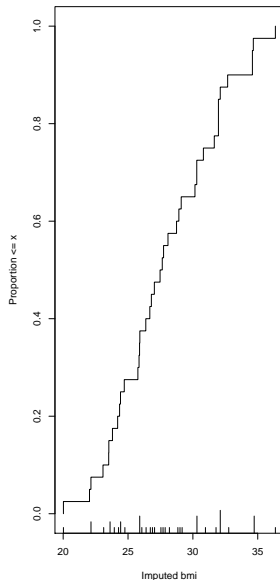
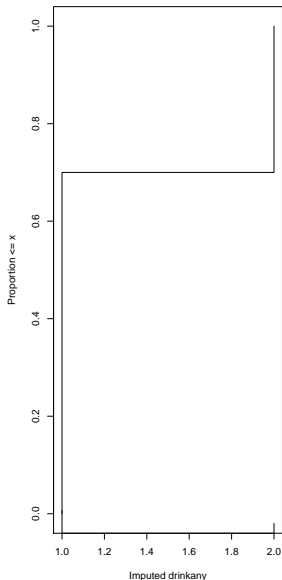
Variable being imputed: drinkany

	nk=0	nk=3	nk=4	nk=5
Bootstrap R^2	0.0163	0.0113	0.0102	0.00986
10-fold cv R^2	0.0205	0.0249	0.0163	0.01358
Bootstrap mean error	0.4470	0.4568	0.4558	0.46624
10-fold cv mean error	0.4450	0.4454	0.4476	0.44676
Bootstrap median error	0.0000	0.0000	0.0000	0.00000
10-fold cv median error	0.0000	0.0500	0.1000	0.00000

Variable being imputed: bmi

	nk=0	nk=3	nk=4	nk=5
Bootstrap R^2	0.0845	0.0932	0.0946	0.0847
10-fold cv R^2	0.0864	0.0903	0.0968	0.0899
Bootstrap mean error	3.7829	4.8119	4.9226	5.1775
10-fold cv mean error	27.6776	4.8359	4.9390	5.1136
Bootstrap median error	2.9955	3.9704	3.9371	4.2634
10-fold cv median error	27.0143	3.9894	3.9431	4.1876

A plot of the imputed values... (results)



A plot of the imputed values... (code)

```
par(mfrow = c(1,3))  
plot(fit3)  
par(mfrow = c(1,1))
```

- For `ldl`, we imputed most of the 7 missing subjects in most of the 20 imputation runs to values within a range of around 120 through 200, but occasionally, we imputed values that were substantially lower than 100.
- For `drinkany` we imputed about 70% no and 30% yes.
- For `bmi`, we imputed values ranging from about 23 to 27 in many cases, and up near 40 in other cases.
- This method never imputes a value for a variable that doesn't already exist in the data.

Kitchen Sink Model (Main Effects only)

```
mod_ks <- ols(ldl ~ age + smoking + drinkany + sbp +  
              physact + bmi, data = hers2)  
anova(mod_ks)
```

Analysis of Variance

Response: ldl

Factor	d.f.	Partial SS	MS	F	P
age	1	9330.911	9330.911	6.93	0.0085
smoking	1	8199.755	8199.755	6.09	0.0137
drinkany	1	6444.424	6444.424	4.79	0.0288
sbp	1	9274.287	9274.287	6.89	0.0087
physact	4	10874.528	2718.632	2.02	0.0891
bmi	1	15876.957	15876.957	11.80	0.0006
REGRESSION	9	60077.708	6675.301	4.96	<.0001
ERROR	2022	2721037.890	1345.716		

Spearman ρ^2 Plot

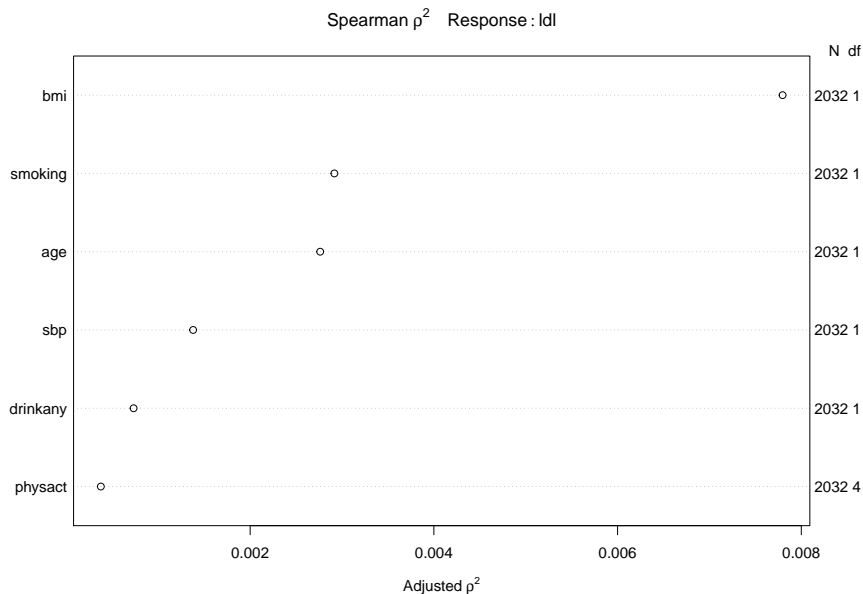
How should we prioritize the degrees of freedom we spend on non-linearity?

```
plot(spearman2(ldl ~ age + smoking + drinkany + sbp +  
              physact + bmi, data = hers2))
```

Plot's on the next page.

- Note the use of the simple imputation hers2 data here. Why?

Spearman ρ^2 Plot Result



Spending Degrees of Freedom

We're spending 9 degrees of freedom in our kitchen sink model. (We can verify this with anova or the plot.)

- Each quantitative main effect costs 1 df to estimate
- Each binary categorical variable also costs 1 df
- Multi-categorical variables with L levels cost $L-1$ df to estimate

Suppose we're willing to spend up to a total of **14** degrees of freedom (i.e. a combined 5 more on interaction terms and other ways to capture non-linearity.)

What should we choose?

What did we see in the Spearman ρ^2 Plot?

Group 1 (largest adjusted ρ^2)

- bmi, a quantitative predictor, is furthest to the right

Group 2 (next largest)

- smoking, a binary predictor, is next, followed closely by
- age, a quantitative predictor

Other predictors (rest of the group)

- sbp, quantitative
- drinkany, binary
- physact, multi-categorical (5 levels)

Impact of Adding Non-Linearity on Degrees of Freedom Spent

What happens when we add a non-linear term?

- Adding a polynomial of degree D costs D degrees of freedom.
 - So a polynomial of degree 2 (quadratic) costs 2 df, or 1 more than the main effect alone.
- Adding a restricted cubic spline with K knots costs $K-1$ df.
 - So adding a rcs with 4 knots uses 3 df, or 2 more than the main effect alone.
 - We restrict ourselves to considering splines with 3, 4, or 5 knots.
- Adding an interaction (product term) depends on the main effects of the predictors we are interacting
 - If the product term's predictors have $df1$ and $df2$ degrees of freedom, the product term adds $df1 \times df2$ degrees of freedom to the main effects model.
 - An interaction of a binary and quantitative variable adds $1 \times 1 = 1$ additional degree of freedom to the main effects model.
 - When we use a quantitative variable in a spline and interaction, we'll do the interaction on the main effect, not the spline.

Model we'll fit with `ols`

Fitting a model to predict `ldl` using

- `bmi` with a restricted cubic spline, 5 knots
- `age` with a quadratic polynomial
- `sbp` as a linear term
- `drinkany` indicator
- `physact` factor
- `smoking` indicator and its interaction with the main effect of `bmi`

We can fit this to the data

- restricted to complete cases (`hers1`, effectively)
- after simple imputation (`hers2`)
- after our multiple imputation (`fit3`)

Fitting the model to the complete cases

```
d <- datadist(hers1)
options(datadist = "d")

m1 <- ols(ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +
          drinkany + physact + smoking +
          smoking %ia% bmi, data = hers1,
          x = TRUE, y = TRUE)
```

where %ia% identifies the linear interaction alone.

m1 results (slide 1 of 3)

```
> m1
```

Frequencies of Missing Values Due to Each Variable

ldl	bmi	age	sbp	drinkany	physact	smoking
7	2	0	0	2	0	0

Linear Regression Model

```
ols(formula = ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +  
      drinkany + physact + smoking + smoking %ia% bmi,  
      data = hers1, x = TRUE, y = TRUE)
```

		Model Likelihood Ratio Test		Discrimination Indexes	
Obs	2021	LR chi2	52.61	R2	0.026
sigma	36.7430	d.f.	14	R2 adj	0.019
d.f.	2006	Pr(> chi2)	0.0000	g	6.629

m1 results (slide 2 of 3)

	Coef	S.E.	t	Pr(> t)
Intercept	121.6057	68.2000	1.78	0.0747
bmi	1.5687	1.0107	1.55	0.1208
bmi'	-8.6685	9.1577	-0.95	0.3440
bmi''	40.5712	37.4468	1.08	0.2787
bmi'''	-55.8872	44.5946	-1.25	0.2103
age	-0.5791	1.9657	-0.29	0.7683
age^2	0.0018	0.0149	0.12	0.9024
sbp	0.1221	0.0453	2.69	0.0072
drinkany=yes	-3.7427	1.6629	-2.25	0.0245
physact=much less active	-4.5660	3.8904	-1.17	0.2407
physact=much more active	-0.3291	2.7521	-0.12	0.9048
physact=somewhat less active	-0.0160	2.5270	-0.01	0.9950
physact=somewhat more active	3.7731	2.0293	1.86	0.0631
smoking=yes	-7.0832	12.0586	-0.59	0.5570
smoking=yes * bmi	0.4961	0.4391	1.13	0.2587

m1 results (slide 3 of 3)

Residuals

Min	1Q	Median	3Q	Max
-113.440	-24.519	-3.778	20.940	197.087

Fitting the model after simple imputation

```
dd <- datadist(hers2)
options(datadist = "dd")

m2 <- ols(ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +
          drinkany + physact + smoking +
          smoking %ia% bmi, data = hers2,
          x = TRUE, y = TRUE)
```

where, again, %ia% identifies the linear interaction alone.

m2 results (slide 1 of 2)

```
> m2
Linear Regression Model

ols(formula = ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp + drinkany +
      physact + smoking + smoking %ia% bmi, data = hers2, x = TRUE,
      y = TRUE)
```

		Model Likelihood		Discrimination	
		Ratio Test		Indexes	
Obs	2032	LR chi2	53.14	R2	0.026
sigma	36.6503	d.f.	14	R2 adj	0.019
d.f.	2017	Pr(> chi2)	0.0000	g	6.631

Residuals

	Min	1Q	Median	3Q	Max
	-113.379	-24.326	-3.835	20.832	197.097

m2 results (slide 2 of 2)

	Coef	S.E.	t	Pr(> t)
Intercept	120.2662	67.6113	1.78	0.0754
bmi	1.5508	1.0071	1.54	0.1237
bmi'	-8.4486	9.0978	-0.93	0.3532
bmi''	39.6413	37.1378	1.07	0.2859
bmi'''	-54.8924	44.2677	-1.24	0.2151
age	-0.5249	1.9490	-0.27	0.7877
age^2	0.0014	0.0148	0.10	0.9233
sbp	0.1209	0.0451	2.68	0.0074
drinkany=yes	-3.7023	1.6544	-2.24	0.0253
physact=much less active	-4.7408	3.8621	-1.23	0.2198
physact=much more active	-0.2635	2.7391	-0.10	0.9234
physact=somewhat less active	0.0130	2.5101	0.01	0.9959
physact=somewhat more active	3.8031	2.0193	1.88	0.0598
smoking=yes	-6.8961	12.0196	-0.57	0.5662
smoking=yes * bmi	0.4892	0.4375	1.12	0.2636

anova(m2) results

```
> anova(m2)
```

Analysis of Variance

Response: ldl

Factor	d.f.	Partial SS	MS	F	P
bmi (Factor+Higher Order Factors)	5	2.758824e+04	5517.64861	4.11	0.0010
All Interactions	1	1.679813e+03	1679.81344	1.25	0.2636
Nonlinear	3	9.735452e+03	3245.15068	2.42	0.0647
age	2	9.175762e+03	4587.88077	3.42	0.0330
Nonlinear	1	1.244351e+01	12.44351	0.01	0.9233
sbp	1	9.657476e+03	9657.47569	7.19	0.0074
drinkany	1	6.726918e+03	6726.91809	5.01	0.0253
physact	4	9.709992e+03	2427.49791	1.81	0.1247
smoking (Factor+Higher Order Factors)	2	1.085405e+04	5427.02463	4.04	0.0177
All Interactions	1	1.679813e+03	1679.81344	1.25	0.2636
smoking * bmi (Factor+Higher Order Factors)	1	1.679813e+03	1679.81344	1.25	0.2636
TOTAL NONLINEAR	4	9.738807e+03	2434.70175	1.81	0.1237
TOTAL NONLINEAR + INTERACTION	5	1.171134e+04	2342.26845	1.74	0.1214
REGRESSION	14	7.178905e+04	5127.78931	3.82	<.0001
ERROR	2017	2.709327e+06	1343.24569		

Validation of summary statistics

```
validate(m2)
```

	index.orig	training	test	optimism
R-square	0.0258	0.0322	0.0193	0.0129
MSE	1333.3300	1331.2658	1342.2160	-10.9502
g	6.6306	7.3074	5.9494	1.3581
Intercept	0.0000	0.0000	25.8305	-25.8305
Slope	1.0000	1.0000	0.8225	0.1775

	index.corrected	n
R-square	0.0129	40
MSE	1344.2801	40
g	5.2726	40
Intercept	25.8305	40
Slope	0.8225	40

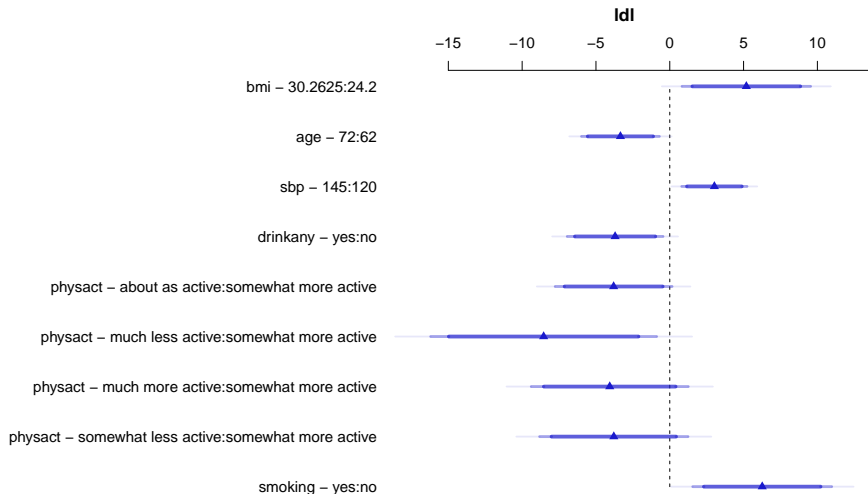
summary(m2) results

```
> summary(m2)
```

Effects	Response : ldl						
Factor	Low	High	Diff.	Effect	S.E.	Lower 0.95	Upper 0.95
bmi	24.2	30.263	6.0625	5.1862	2.2217	0.82921	9.54330
age	62.0	72.000	10.0000	-3.3412	1.3450	-5.97890	-0.70357
sbp	120.0	145.000	25.0000	3.0218	1.1270	0.81165	5.23190
drinkany - yes:no	1.0	2.000	NA	-3.7023	1.6544	-6.94690	-0.45779
physact - about as active:somewhat more active	5.0	1.000	NA	-3.8031	2.0193	-7.76310	0.15695
physact - much less active:somewhat more active	5.0	2.000	NA	-8.5439	3.9035	-16.19900	-0.88862
physact - much more active:somewhat more active	5.0	3.000	NA	-4.0666	2.7125	-9.38630	1.25310
physact - somewhat less active:somewhat more active	5.0	4.000	NA	-3.7901	2.5633	-8.81720	1.23690
smoking - yes:no	1.0	2.000	NA	6.2635	2.4009	1.55500	10.97200

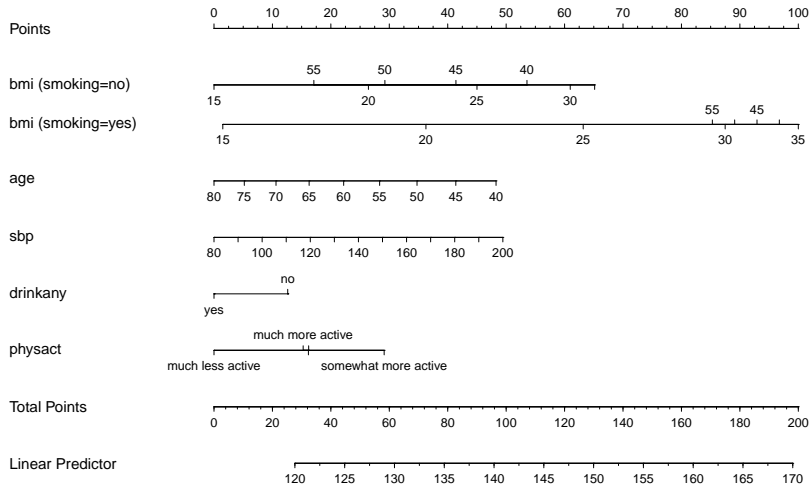
Adjusted to: bmi=26.9 smoking=no

plot(summary(m2)) results



Adjusted to: bmi=26.9 smoking=no

plot(nomogram(m2))



Making Predictions for an Individual

Suppose now that we want to use R to get a prediction for a new individual subject with $\text{bmi} = 30$, $\text{age} = 50$, $\text{smoking} = \text{yes}$ and $\text{physact} = \text{about as active}$, $\text{drinkany} = \text{yes}$ and sbp of 150.

```
predict(m2, expand.grid(bmi = 30, age = 50, smoking = "yes",  
                        physact = "about as active",  
                        drinkany = "yes", sbp = 150),  
        conf.int = 0.95, conf.type = "individual")
```

\$linear.predictors	\$lower	\$upper
160.9399	88.48615	233.3936

Making Predictions for a Long-Run Mean

The other kind of prediction we might wish to make is for the mean of a series of subjects whose `bmi = 30`, `age = 50`, `smoking = yes` and `physact = about as active`, `drinkany = yes` and `sbp` of 150.

```
predict(m2, expand.grid(bmi = 30, age = 50, smoking = "yes",  
                        physact = "about as active",  
                        drinkany = "yes", sbp = 150),  
        conf.int = 0.95, conf.type = "mean")
```

\$linear.predictors	\$lower	\$upper
160.9399	151.8119	170.0679

Of course, the confidence interval will always be narrower than the prediction interval given the same predictor values.

Influential Points?

```
which.influence(m2, cutoff = 0.4)
```

```
$Intercept
```

```
[1] 1135
```

```
$age
```

```
[1] 1135
```

```
$smoking
```

```
[1] 132
```

```
$`smoking * bmi`
```

```
[1] 132
```

Fitting the Model using Multiple Imputation

What do we have now?

- An imputation model `fit3`

```
fit3 <- aregImpute(~ ldl + age + smoking + drinkany + sbp +  
  physact + bmi, nk = c(0, 3:5), tlinear = FALSE,  
  data = hers1, B = 10, n.impute = 20, x = TRUE)
```

- A prediction model

```
m1 <- ols(ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +  
  drinkany + physact + smoking + smoking %ia% bmi,  
  x = TRUE, y = TRUE)
```

Now we put them together with the `fit.mult.impute` function...

Linear Regression & Imputation Model

```
m3imp <-  
  fit.mult.impute(ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +  
    drinkany + physact + smoking +  
    smoking %ia% bmi,  
    fitter = ols, xtrans = fit3,  
    data = hers1)
```

Variance Inflation Factors Due to Imputation:

Intercept	bmi
1.00	1.00
bmi'	bmi''
1.00	1.01
bmi'''	age
1.01	1.00
age^2	sbp
1.00	1.01

m3imp results (1 of 2)

```
> m3imp
Linear Regression Model

fit.mult.impute(formula = ldl ~ rcs(bmi, 5) + pol(age, 2) + sbp +
  drinkany + physact + smoking + smoking %ia% bmi, fitter = ols,
  xtrans = fit3, data = hers1)
```

		Model Likelihood		Discrimination	
		Ratio Test		Indexes	
Obs	2032	LR chi2	53.30	R2	0.026
sigma	36.7128	d.f.	14	R2 adj	0.019
d.f.	2017	Pr(> chi2)	0.0000	g	6.652

Residuals

Min	1Q	Median	3Q	Max
-113.10	-24.46	-3.81	20.92	197.42

m3imp results (2 of 2)

	Coef	S.E.	t	Pr(> t)
Intercept	121.1499	67.7998	1.79	0.0741
bmi	1.5445	1.0097	1.53	0.1263
bmi'	-8.2945	9.1027	-0.91	0.3623
bmi''	39.0890	37.3055	1.05	0.2949
bmi'''	-54.2119	44.4779	-1.22	0.2230
age	-0.5521	1.9547	-0.28	0.7776
age^2	0.0016	0.0148	0.11	0.9119
sbp	0.1216	0.0453	2.69	0.0073
drinkany=yes	-3.7404	1.6625	-2.25	0.0246
physact=much less active	-4.7426	3.8692	-1.23	0.2204
physact=much more active	-0.2665	2.7455	-0.10	0.9227
physact=somewhat less active	0.0313	2.5214	0.01	0.9901
physact=somewhat more active	3.8060	2.0257	1.88	0.0604
smoking=yes	-6.9198	12.0472	-0.57	0.5658
smoking=yes * bmi	0.4917	0.4388	1.12	0.2626

anova(m3imp)

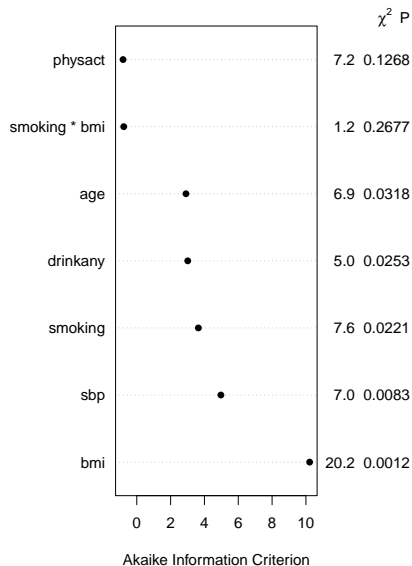
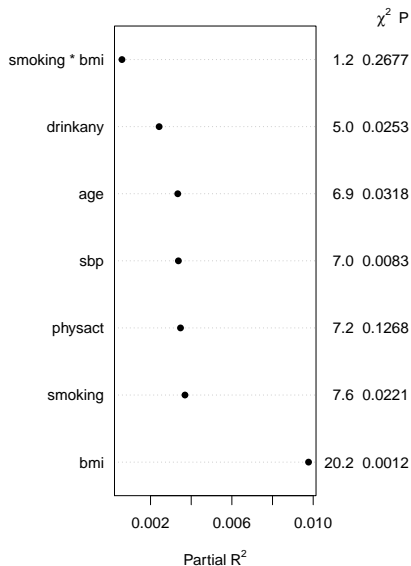
```
> anova(m3imp)
```

Analysis of Variance

Response: ldl

Factor	d.f.	Partial SS	MS	F	P
bmi (Factor+Higher Order Factors)	5	27514.6406	5502.9281	4.08	0.0011
All Interactions	1	1692.6044	1692.6044	1.26	0.2626
Nonlinear	3	9741.6194	3247.2065	2.41	0.0653
age	2	9078.9851	4539.4926	3.37	0.0347
Nonlinear	1	16.5032	16.5032	0.01	0.9119
sbp	1	9721.1667	9721.1667	7.21	0.0073
drinkany	1	6822.3861	6822.3861	5.06	0.0246
physact	4	9690.3632	2422.5908	1.80	0.1267
smoking (Factor+Higher Order Factors)	2	10845.6127	5422.8063	4.02	0.0180
All Interactions	1	1692.6044	1692.6044	1.26	0.2626
smoking * bmi (Factor+Higher Order Factors)	1	1692.6044	1692.6044	1.26	0.2626
TOTAL NONLINEAR	4	9747.0966	2436.7741	1.81	0.1246
TOTAL NONLINEAR + INTERACTION	5	11717.3715	2343.4743	1.74	0.1225
REGRESSION	14	71571.1297	5112.2236	3.79	<.0001
ERROR	2017	2718570.0412	1347.8285		

Evaluation via Partial R² and AIC (result)



Evaluation via Partial R^2 and AIC (code)

```
par(mfrow = c(1,2))  
plot(anova(m3imp), what="partial R2")  
plot(anova(m3imp), what="aic")  
par(mfrow = c(1,1))
```

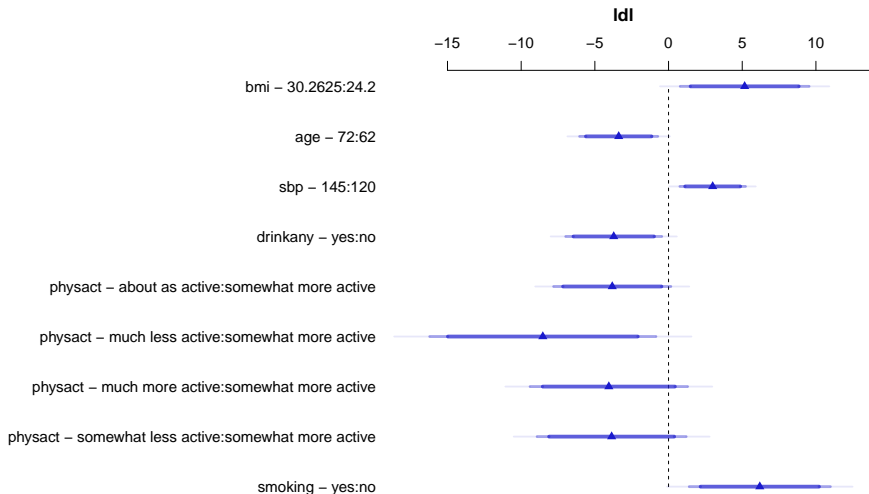
summary(m3imp)

```
> summary(m3imp)
```

Effects	Response : ldl							
Factor	Low	High	Diff.	Effect	S.E.	Lower 0.95	Upper 0.95	
bmi	24.2	30.263	6.0625	5.2108	2.2283	0.84072	9.58080	
age	62.0	72.000	10.0000	-3.3219	1.3498	-5.96910	-0.67463	
sbp	120.0	145.000	25.0000	3.0394	1.1317	0.81989	5.25880	
drinkany - yes:no	1.0	2.000	NA	-3.7404	1.6625	-7.00080	-0.47996	
physact - about as active:somewhat more active	5.0	1.000	NA	-3.8060	2.0257	-7.77860	0.16663	
physact - much less active:somewhat more active	5.0	2.000	NA	-8.5486	3.9114	-16.21900	-0.87779	
physact - much more active:somewhat more active	5.0	3.000	NA	-4.0724	2.7198	-9.40640	1.26160	
physact - somewhat less active:somewhat more active	5.0	4.000	NA	-3.7746	2.5773	-8.82900	1.27980	
smoking - yes:no	1.0	2.000	NA	6.3067	2.4196	1.56150	11.05200	

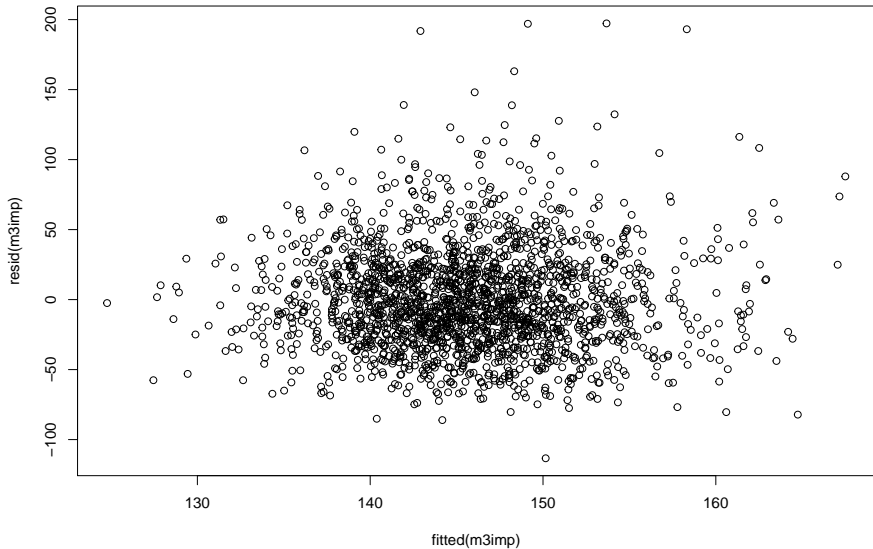
Adjusted to: bmi=26.9 smoking=no

```
plot(summary(m3imp))
```

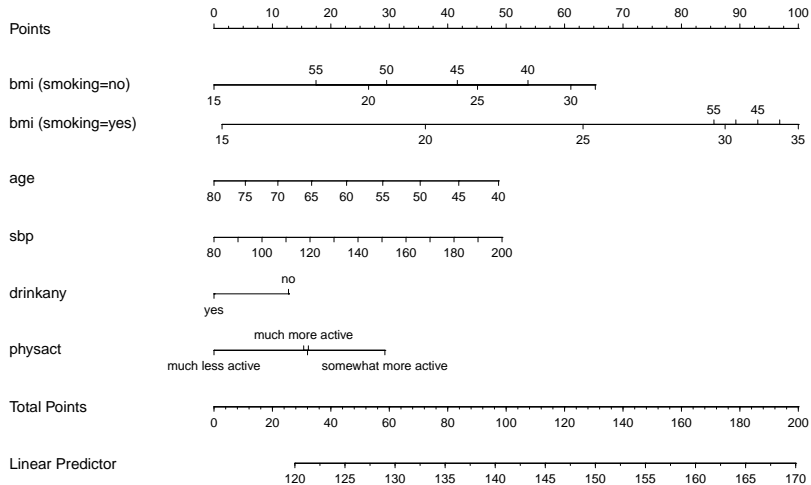


Adjusted to:bmi=26.9 smoking=no


```
plot(resid(m3imp) ~ fitted(m3imp))
```



plot(nomogram(m3imp))



Next Step

Can we do all of this for a logistic regression model?