

# Econometrics II: Assignment 2

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January 12, 2021

## Question 1

First use pooled OLS to check the impact of including and excluding ASVABC on the estimate of  $\alpha_1$ . Present and explain the result.

	<i>Dependent variable:</i>	
	(1)	(2)
ASVABC	0.123*** (0.005)	
AGE	0.436*** (0.061)	0.389*** (0.061)
AGESQ	-0.001 (0.001)	-0.000 (0.001)
S	0.789*** (0.020)	1.042*** (0.017)
ETHBLACK	-1.218*** (0.124)	-2.299*** (0.115)
URBAN	1.301*** (0.086)	1.354*** (0.087)
REGNE	-2.301*** (0.189)	-1.499*** (0.187)
REGNC	-3.810*** (0.184)	-3.062*** (0.182)
REGW	-2.956*** (0.191)	-2.223*** (0.189)
REGS	-3.889*** (0.182)	-3.308*** (0.181)
Observations	40,043	40,043
$R^2$	0.225	0.215
Adjusted $R^2$	0.225	0.215
Residual Std. Error	7.127(df = 40033)	7.173(df = 40034)
F Statistic	1293.484*** (df = 9.0; 40033.0)	1372.448*** (df = 8.0; 40034.0)
<i>Note:</i>		*p<0.1; **p<0.05; ***p<0.01

The inclusion of the proxy ability decreases the estimate for the coefficient of schooling. Hence, given all other standard assumptions, ability and schooling are positively correlated, and the omission of a proxy for ability overestimates the impact of schooling.

## Question 2

Perform a pooled OLS analysis to obtain insight in the heterogeneity of returns to schooling by ethnicity. Present the results and comment on the outcomes: what are the conclusions based on this?

	<i>Dependent variable:</i>		
	Interaction	Not Black	Black
	(1)	(2)	(3)
BLACKxS	-0.076 (0.053)		
ASVABC	0.123*** (0.005)	0.117*** (0.006)	0.165*** (0.011)
AGE	0.434*** (0.061)	0.457*** (0.067)	0.140 (0.117)
AGESQ	-0.001 (0.001)	-0.001 (0.001)	0.001 (0.002)
S	0.796*** (0.021)	0.791*** (0.022)	0.732*** (0.041)
ETHBLACK	-0.242 (0.688)	0.000*** (0.000)	-8.262*** (1.358)
URBAN	1.298*** (0.086)	1.431*** (0.093)	-0.079 (0.210)
REGNE	-2.318*** (0.190)	-2.402*** (0.208)	-1.232*** (0.372)
REGNC	-3.825*** (0.184)	-3.921*** (0.202)	-2.391*** (0.372)
REGW	-2.970*** (0.191)	-3.085*** (0.209)	-1.494*** (0.408)
REGS	-3.904*** (0.182)	-3.942*** (0.201)	-3.144*** (0.352)
Observations	40,043	35,223	4,820
$R^2$	0.225	0.213	0.304
Adjusted $R^2$	0.225	0.213	0.303
<i>Note:</i> *p<0.1; **p<0.05; ***p<0.01			

We can see that the interaction effect is insignificant: that is to say, there is no significant difference between blacks and non-black in the influence of schooling on earnings. When we split up the sample into blacks and non-black, we get a slightly different view: the point estimate for the effect of schooling seems to be slightly lower for black people than for non-black people. However, as seen in the pooled regression with interaction effect, the differential impact is not statistically significant.

### Question 3

Perform the analysis for heterogenous schooling effects using the random effects model. Present the results and compare the outcomes with the pooled OLS results obtained before. Interpret the outcomes.

Table 1:

	<i>Dependent variable:</i>
	EARNINGS
ASVABC	0.124*** (0.011)
AGE	0.433*** (0.051)
AGESQ	-0.001 (0.001)
S	0.826*** (0.036)
ETHBLACK	1.908 (1.236)
URBAN	0.559*** (0.103)
REGNE	-15.959*** (0.865)
REGNC	-17.416*** (0.855)
REGW	-16.295*** (0.862)
REGS	-17.305*** (0.851)
BLACKxS	-0.231** (0.095)
Observations	40,043
R <sup>2</sup>	0.163
Adjusted R <sup>2</sup>	0.163
F Statistic	35,260.370***

*Note:* \*p<0.1; \*\*p<0.05; \*\*\*p<0.01

The point estimate for schooling is now close to the point estimate for schooling in the pooled OLS regression including the proxy for ability. Hence, the random effects estimator looks a lot like the pooled estimator, indicating that the contribution from the within group estimator is marginal. This can also be observed when looking at the decomposition of the explained variance: the between R-squared is larger than the

within R-squared, indicating the model does a better job explaining the changes between individuals rather than individuals over time.

(Still need to explain: black ethnicity · schooling is significant)

#### Question 4

**A priori, would you plead for using fixed effects estimation or random effects estimation? Explain your answer.**

A priori, it would make more sense to use fixed-effects rather than random effects, because it is very likely that the unobservable individual components  $\eta_i$  are correlated to the predictor variables  $X$  rather than being random. For example,  $\eta_i$  can be interpreted as being some measure of ability or innate willingness to exert effort, and that is likely related to age, schooling and test score. A possible correlation would violate the randomness of  $\eta_i$  required by random effects, and hence, fixed effects would be preferred.

## Question 5

Apply the fixed effects estimator to analyze the heterogenous schooling effect. Interpret the outcomes.

Table 2:	
	<i>Dependent variable:</i>
	EARNINGS
AGE	0.428*** (0.052)
AGESQ	−0.0004 (0.001)
S	0.852*** (0.071)
URBAN	0.204* (0.118)
REGNE	0.672** (0.300)
REGNC	−0.447* (0.262)
REGW	1.060*** (0.300)
BLACKxS	−1.054*** (0.231)
Observations	40,043
R <sup>2</sup>	0.134
Adjusted R <sup>2</sup>	0.017
F Statistic	682.917*** (df = 8; 35270)
<i>Note:</i>	*p<0.1; **p<0.05; ***p<0.01

## Question 6

Fixed effects estimation may not be as efficient as random effects estimation, but is robust to correlation between regressors and the random effect. Can we perform a Hausman test in this context? Perform the test you propose.

The test tests the null hypothesis that the unique errors are not correlated with the regressors.

```
phtest(fixed_effects, random_effects)
```

```
##  
## Hausman Test  
##  
## data: formula  
## chisq = 54.565, df = 8, p-value = 5.364e-09  
## alternative hypothesis: one model is inconsistent
```

The null hypothesis is rejected, implying that the unique parts are correlated with the regressors, and hence, random effects is an inconsistent estimator.

## Question 7