# Assignment 3

### Walter Verwer & Bas Machielsen

## 1/17/2021

## Question 1

### 1.1

```
data <- tribble(
    ~ color, ~ no_treated, ~ no_contr, ~ avg_treated, ~avg_control,
    "purple", 100, 100, 9, 7,
    "blue", 75, 25, 13, 8,
    "green", 25, 75, 10, 9
)

data %>%
    group_by(color) %>%
    summarize(treatment_effect = avg_treated - avg_control) %>%
    kable(caption = "Treatment effect per color")
```

Table 1: Treatment effect per color

color	treatment_effect
blue	5
green	1
purple	2

### 1.2

The ATE is defined as  $\mathbb{E}[\delta] = \mathbb{E}[Y_1^*] - \mathbb{E}[Y_0^*]$  which are the expectations of the potential outcomes. In general, these two variables are not observed. Under the *random assignment* assumption, we assume that  $\mathbb{E}[Y_1^*] = \mathbb{E}[Y_1^*|D=1]$  and  $\mathbb{E}[Y_0^*] = \mathbb{E}[Y_0^*|D=0]$ , which can be estimated by their sample by their sample equivalents:

```
data %>%
    summarize(
        e_y1_d_is_1 = (9*100 + 13 * 75 + 10 * 25) / sum(no_treated),
        e_y0_d_is_0 = (7 * 100 + 8 * 25 + 9 * 75) / sum(no_contr)) %>%
    summarize(ate = e_y1_d_is_1 - e_y0_d_is_0) %>%
    kable(caption = "ATE")
```

```
Table 2: ATE
```

 $\frac{\text{ate}}{2.75}$ 

1.3

The ATT is defined as  $\mathbb{E}[\delta|D=1] = \mathbb{E}[Y^1|D=1] - \mathbb{E}[Y^0|D=1]$ . The first term is readily observable. The second term is estimated by us as  $\hat{\mathbb{E}}[Y^0|D=1] = \mathbb{E}[Y^0|D=0]$ . Hence:

```
data_ate <- data %>%
   mutate(n = no_treated + no_contr) %>%
    summarize(e_y1_d_is_1 = (9 * 100 + 13 * 75 + 10 * 25) / sum(no_treated),
              e_y0_dis_0 = (7 * 100 + 8 * 25 + 9 * 75) / sum(no_contr))
data_ate
## # A tibble: 1 x 2
    e_y1_d_is_1 e_y0_d_is_0
##
           <dbl>
                       <dbl>
                       7.88
            10.6
## 1
data_ate %>%
    summarize(att = e_y1_d_is_1 - e_y0_d_is_0) %>%
   kable(caption = "ATT")
```

Table 3: ATT

 $\frac{\overline{\text{att}}}{2.75}$ 

So the ATE = ATT (because of randomization).

## **Question 2**

2.1

Compute the fraction of students in all three groups (control, low-reward and high-reward) that complete all first-year courses before the start of the second academic year. Show within a table that background characteristics are balanced over the treatment groups.

```
bonus_clean %>%
   group_by(kind_treatment) %>%
   summarize(fraction_pass = sum(pass)/n()) %>%
   kable(caption = "Fraction passed per treatment", digits = 3)
```

Table 4: Fraction passed per treatment

kind_treatment	fraction_pass
bonus0	0.195
bonus1500	0.241
bonus500	0.202

Table 5: Means and SDs according to treatment

var	bonus0_mean	bonus0_sd	bonus1500_mean	bonus1500_sd	bonus500_mean	bonus500_sd
p0	0.553	0.265	0.573	0.248	0.530	0.251
job	0.760	0.430	0.805	0.399	0.829	0.379
myeduc	12.293	3.041	12.590	2.992	12.119	3.316
fyeduc	13.378	3.416	13.422	3.596	13.524	3.273
effort	19.549	9.460	18.303	10.592	18.477	10.475
math	5.476	1.468	5.388	1.258	5.386	1.360

### 2.2

Use the linear probability model to regress the dummy variable for completing all courses on the assignment of the three treatment groups. Interpret the treatment effects. Next include as additional regressors father's education, high-school math score and the subjective assessment about the pass probability.

```
# We need to regress pass on the treatment group assignment. This is
# a simple linear probability model. Model is denoted by prob_1.
# Note, we need to omit bonus0, cause of a dummy variable trap.
prob_1 <- lm(pass ~ bonus500 + bonus1500, data=bonus)
# Same as before, but now with some extra variables
prob_2 <- lm(pass ~ bonus500 + bonus1500 + fyeduc + math + p0, data=bonus)
# for q2.3:
prob_3 <- lm(pass ~ bonus500 + bonus1500 + fyeduc + math + p0 + job + effort, data=bonus)</pre>
```

```
# Implement HC-robust standard errors (standard in linear prob model)
#library(sandwich)
#cov <- vcovHC(reg.model, type = "HC")
#robust.se <- sqrt(diag(cov))

# Create table:
stargazer(prob_1, prob_2, prob_3, header=FALSE, style='aer')</pre>
```

Table 6:

		pass	
	(1)	(2)	(3)
bonus500	0.007	0.015	0.023
	(0.064)	(0.057)	(0.058)
bonus1500	0.046	0.048	0.056
	(0.064)	(0.058)	(0.059)
fyeduc		-0.001	0.000
•		(0.007)	(0.007)
math		0.119***	0.124***
		(0.018)	(0.018)
p0		0.249***	$0.170^{*}$
•		(0.095)	(0.098)
job			-0.062
•			(0.060)
effort			0.008***
			(0.002)
Constant	0.195***	-0.576***	-0.678***
	(0.045)	(0.128)	(0.144)
Observations	249	245	230
$\mathbb{R}^2$	0.002	0.209	0.264
Adjusted R <sup>2</sup>	-0.006	0.192	0.240
Residual Std. Error	0.411 (df = 246)	0.366 (df = 239)	0.358 (df = 222)
F Statistic	0.297 (df = 2; 246)	12.629*** (df = 5; 239)	11.359*** (df = 7; 222)
Notes:	***Significant at the	e 1 nercent level	

Notes:

In the first model, we find that students who receive the 500 bonus are 0.007 percentage points more likely to pass the first year, and students who receive the f 1,500 bonus are 0.046 percentage points more likely to pass the first year, both relative to the group that receives no bonus. This effect is, however, not statistically significant, indicating that there is a high variance in passing within treatment groups, or, alternatively, that the sample size is too sample to statistically detect a relatively small effect size.

<sup>\*\*\*</sup>Significant at the 1 percent level.

<sup>\*\*</sup>Significant at the 5 percent level.

<sup>\*</sup>Significant at the 10 percent level.

The point estimates change only slightly when including a vector of control variables, indicating that the treatment conditional on these (potential) confounders does not significantly increase the probability of passing.

### 2.3

Next also include as regressors in your model whether a student has a job and the amount of study effort. Comment on this approach. Do you consider this an improvement over (ii)?

In the next model, we observe that the point estimates for the treatment effects increase a bit, indicating that effort was a confounder. Hence, we do consider this to be an improvement, although the treatment effects are still not significantly different from zero.

### 2.4

Use your preferred model specification to estimate the effects of the financial incentives on some other outcomes: dropping out and credit points collected (in the first year and after three years).

% Table created by stargazer v.5.2.2 by Marek Hlavac, Harvard University. E-mail: hlavac at fas.harvard.edu % Date and time: Wed, Jan 20, 2021 - 10:58:00

### 2.5

Given the sample size and the estimates you have obtained above, what would be the minimum detectable effect size of this experiment?

The proportion of treated subjects  $p = \frac{83+84}{82+83+84} = 0.67$ .

```
bonus_clean %>%
   group_by(kind_treatment) %>%
   summarize(test = n()) %>%
   kable()
```

## 'summarise()' ungrouping output (override with '.groups' argument)

test
82
83
84

The estimated variance of the error term in the full model is

We know that the minimum effect size as a function of p, power q, level  $\alpha$ ,

## 2.6

Initially, the researchers were aiming at an increases in the pass rate of 10%- points. How large should the sample size of the experiment have been in that case?

Table 7:

	dropout	stp2001	stp2004
	(1)	(2)	(3)
bonus500	-0.048	-0.144	-2.388
	(0.068)	(2.690)	(7.999)
bonus1500	-0.057	0.227	0.294
	(0.070)	(2.744)	(8.160)
fyeduc	0.009	-0.309	-1.231
	(0.009)	(0.337)	(1.002)
math	-0.077***	6.565***	15.561***
	(0.022)	(0.856)	(2.545)
p0	-0.150	13.582***	18.312
•	(0.116)	(4.567)	(13.579)
job	0.034	-0.904	5.628
	(0.071)	(2.791)	(8.298)
effort	-0.018***	0.928***	2.920***
	(0.003)	(0.111)	(0.331)
Constant	1.071***	-22.455***	-50.458**
	(0.170)	(6.713)	(19.962)
Observations	230	230	230
$\mathbb{R}^2$	0.232	0.445	0.388
Adjusted R <sup>2</sup>	0.207	0.427	0.369
Residual Std. Error (df = 222)	0.424	16.695	49.642
F Statistic (df = 7; 222)	9.563***	25.388***	20.112***

Notes:

<sup>\*\*\*</sup>Significant at the 1 percent level.
\*\*Significant at the 5 percent level.
\*Significant at the 10 percent level.