Erasmus School of Economics

Econometrics I

Lecture 5: Endogeneity

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Today

1. Endogenous regressors

2. Instrumental variables – IV/2SLS

3. Testing for exogeneity and validity of instruments

Endogenous regressors

Endogenous regressors and instrumental variables

We start with the linear regression model

$$y_i = x_i'\beta + \varepsilon_i, \qquad i = 1, \ldots, n,$$

where assumptions 2, 3, 4, 5, and 6 hold.

In matrix notation:

$$y = X\beta + \varepsilon$$
.

The regressors x_i are now **stochastic**, assumption 1 (fixed X) is replaced with Assumption 1*: $plim\left(\frac{1}{n}X'X\right) = Q$ of rank k.

Endogenity

When the orthogonality condition

$$p\lim\left(\frac{1}{n}X'\varepsilon\right)=0$$

holds, the regressors X are called **exogenous**.

There is endogeneity or endogenous regressors when this condition does not hold, that is,

$$\operatorname{plim}\left(\frac{1}{n}X'\varepsilon\right)\neq 0,$$

or in other words $cov(x_{ji}, \varepsilon_i) \neq 0$ for some j.

Examples and intuition

Possible reasons for correlation between X and ε are:

1. Omitted variables that influence X and y (through ε):

Example

Suppose we model the grade for Econometrics I using the number of followed lectures as explanatory factor.

An omitted variable is prior knowledge of the student. This is possibly correlated with the lectures taken and the final grade.

 $\rightarrow \, \mathsf{Endogeneity!}$

Examples and intuition

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- \rightarrow Endogeneity!
 - 2. Strategic behavior based on unobserved factors:

Example

Model sales vs. prices. If the salesperson determines the price and has extra knowledge (on ε), the price will be correlated with the error.

 \rightarrow Endogeneity!

Example 5.1: Endogeneity

Consequences of endogeneity

The OLS estimator equals $b = \beta + (X'X)^{-1}X'\varepsilon = \beta + (\frac{1}{n}X'X)^{-1}(\frac{1}{n}X'\varepsilon)$, therefore

$$\mathsf{plim}(b) = \beta + \mathsf{plim}\left(\frac{1}{n}X'X\right)^{-1}\mathsf{plim}\left(\frac{1}{n}X'\varepsilon\right) = \beta + Q^{-1}\mathsf{plim}\left(\frac{1}{n}X'\varepsilon\right).$$

If $p\lim_{n} \frac{1}{n} X' \varepsilon \neq 0$ then we have endogeneity

- In general we then have correlation between X and ε (Note: no correlation does not necessarily mean independent)
- When there is endogeneity $plim(b) \neq \beta$, in other words the OLS estimator is **inconsistent**
- Testing/estimation is no longer useful!

There are three important causes of endogeneity:

- Omitted variables
- Measurement errors
- Simultaneity

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1. Omitted variables

Suppose that in reality y_i depends on two variables, x_{2i} and x_{3i} , but we only use x_{2i} in the model for y_i , that is,

DGP :
$$y_i = \beta_1 + \beta_2 x_{2i} + \beta_3 x_{3i} + u_i$$
,
Model : $y_i = \beta_1 + \beta_2 x_{2i} + \varepsilon_i$,

such that $\varepsilon_i = \beta_3 x_{3i} + u_i$. When $\beta_3 \neq 0$ and $\text{cov}(x_{2i}, x_{3i}) \neq 0$, it follows that $\text{cov}(x_{2i}, \varepsilon_i) = \beta_3 \text{cov}(x_{2i}, x_{3i}) \neq 0$, or x_{2i} is endogenous.

2. Measurement error

• Suppose that in reality y_i depends on x_{2i}^*

DGP:
$$y_i = \beta_1 + \beta_2 x_{2i}^* + u_i$$
.

- x_{2i}^* is observed with an error e_i (with mean 0 and variance σ_e^2). We observe $x_{2i} = x_{2i}^* + e_i$.
- x_{2i} is endogenous in the regression of x_{2i} on y_i

Simultaneity

DGP:

$$y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$

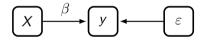
$$x_i = \gamma_1 + \gamma_2 \varepsilon_i + \eta_i$$

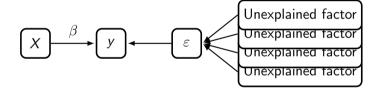
Model: $y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$

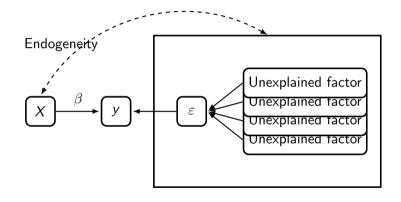
Now it holds "by definition" that x_i is correlated with ε_i .

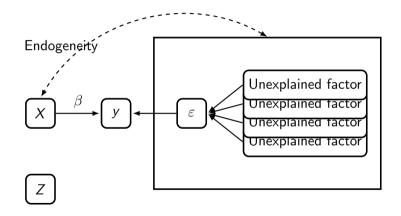
 \rightarrow Both variables y_i and x_i are determined at the same time (simultaneously)

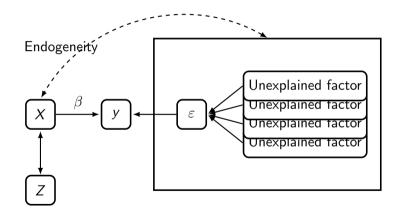
Instrumental variables

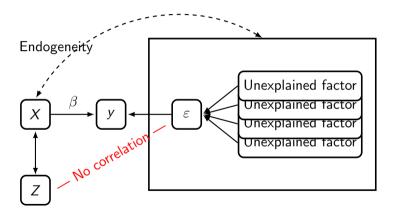












Important: $corr(Z, \varepsilon) = 0$ does not imply corr(y, Z) = 0

"Solving endogeneity": Instrumental variables

Given the model

$$y_i = x_i'\beta + \varepsilon_i,$$
 with $cov(x_i, \varepsilon_i) \neq 0$

and with k explanatory variables

Idea:

- Find m variables (z_i) that are correlated with x_i , but not with ε_i
- Part of x_i that is explained by z_i is by definition not correlated with ε_i

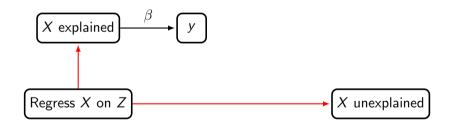
Procedure:

- Explain X with $Z \to \text{explained part uncorrelated with } \varepsilon$.
- Use explained part to explain y.
- \rightarrow 2-stage least squares (2SLS)

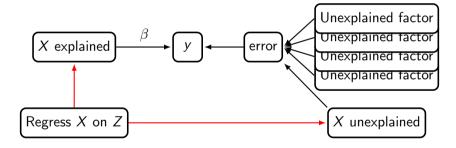
1. Use Z to decompose X in explained and unexplained part



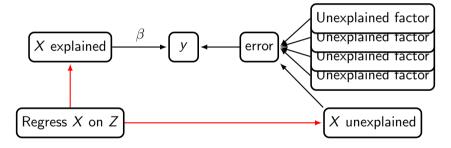
- 1. Use Z to decompose X in explained and unexplained part
- 2. Effect size of explained part on y equals β



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Endogeneity is solved as

- X unexplained not correlated with X explained
- X explained is exogenous

Two conditions for the instruments

(1) Z and ε uncorrelated

$$p\lim \frac{1}{n}Z'\varepsilon=0$$

validity of the instrument

(2) Z (sufficiently) correlated with X

$$p\lim \frac{1}{n}Z'X = Q_{ZX}$$

with $rank(Q_{ZX}) = k$ (rank condition) relevance of the instrument

 \rightarrow If the instrument fails condition 1 (Z and ε uncorrelated) it is invalid, if it fails condition 2 (Z correlated with X) it is irrelevant.

Finding instruments

What are good instruments?

- All exogenous variables in x_i (incl. constant) (These have to be included in Z)
- Other instruments are always needed. Use the properties of the problem at hand.
 - \rightarrow These additional instruments should not directly influence y.
- The stronger the correlation between Z and X the better (as long as there is no correlation between Z and ε)
- \rightarrow For every endogenous variable we need to find at least one instrument

Examples of instruments

- In example attendance/grade: policy change to obligatory attendance
- In example price/sales: prices of raw materials

Requirements IV/2SLS

To perform 2SLS we need:

1. Z and ε uncorrelated

$$p\lim_{n}\frac{1}{n}Z'\varepsilon=0$$

2. Z (sufficiently) correlated with X

$$plim \frac{1}{n}Z'X = Q_{ZX}$$

with $rank(Q_{ZX}) = k$ (rank condition)

3. Z "stable" and not "multi collinear"

$$p\lim \frac{1}{n}Z'Z = Q_{ZZ}$$

with rank(Q_{ZZ}) = m

4. Enough instruments (Z) (order condition)

$$m \ge k$$

IV estimator

case: m = k

Given model

$$y = X\beta + \varepsilon$$

with all assumptions for IV (see earlier slide).

IV estimator:

$$b_{IV} = (Z'X)^{-1}Z'y$$

Intuitively for single regression (one variable and one instrument):

$$b_{IV} = \frac{\delta y/\delta Z}{\delta X/\delta Z}$$
$$= \frac{(Z'Z)^{-1}Z'y}{(Z'Z)^{-1}Z'X}$$
$$= (Z'X)^{-1}Z'y$$

2-stage least squares

case: $m \ge k$

Given model

$$y = X\beta + \varepsilon$$

with all assumptions for IV (see earlier slide).

Steps:

- 1. regress X on Z $\rightarrow \hat{X} = Z(Z'Z)^{-1}Z'X = P_ZX$
- 2. regress y on \hat{X}

2SLS estimator:

$$b_{2SLS} = (\hat{X}'\hat{X})^{-1}\hat{X}'y$$

= $(X'P'_{Z}P_{Z}X)^{-1}X'P'_{Z}y$
= $(X'P_{Z}X)^{-1}X'P_{Z}y$

If m = k: $b_{IV} = (Z'X)^{-1}Z'y$ the 2SLS estimator simplifies to the IV estimator

Instrumental variables as GMM estimator

The assumptions on ε_i , x_i and z_i can be summarized in m moment conditions

$$E[z_{i}\varepsilon_{i}] = E\begin{pmatrix} z_{1i}\varepsilon_{i} \\ z_{2i}\varepsilon_{i} \\ \vdots \\ z_{mi}\varepsilon_{i} \end{pmatrix} = 0$$

As $\varepsilon_i = y_i - x_i'\beta$ we have $E[z_i\varepsilon_i] = E[z_i(y_i - x_i'\beta)]$ If $m \ge k$ the m conditions provide enough information to estimate the k parameters (if m = k exactly enough info).

For a specific dataset we need to solve

$$G_n(\beta) = \sum_{i=1}^n g_i(\beta) = \sum_{i=1}^n z_i(y_i - x_i'\beta) = Z'(y - X\beta) = 0$$

If m > k this cannot be done exactly

IV as GMM estimator

"Solving": $G_n(\beta) = Z'(y - X\beta) = 0$ \rightarrow Try to come as close as possible to 0

Solve: $\min_{\beta} G'_n W G_n$

The optimal weight matrix is

$$W = J_0^{-1} = (\mathsf{E}[g_i(\beta)g_i(\beta)'])^{-1} = (\mathsf{E}[(z_i\varepsilon_i)(z_i\varepsilon_i)'])^{-1} = (\mathsf{E}[\varepsilon_i^2z_iz_i'])^{-1}$$

For a specific dataset we estimate W with $(\hat{\sigma}^2 \frac{1}{n} Z'Z)^{-1}$

So

$$\min(y - X\beta)'Z(Z'Z)^{-1}Z'(y - X\beta) = \min(y - X\beta)'P_Z(y - X\beta)$$

$$\rightarrow b_{IV} = (X'P_ZX)^{-1}X'P_Zy$$

Properties of two stage least squares

The IV/2SLS estimator b_{IV}/b_{2SLS} has the following properties:

- 1. consistency
- 2. asymptotic normality:

$$b_{2SLS} pprox N\left(eta, \sigma^2(\hat{X}'\hat{X})^{-1}
ight).$$

Important remarks

- 1. In the "first-stage" regression you have to include all exogenous variables
- 2. The variance of the residuals ε_i , σ^2 (necessary for the standard errors of b_{2SLS}), is to be estimated using the "2SLS"-residuals: $e_{2SLS} = y Xb_{2SLS}$.

Example 5.1: IV estimation

Testing for exogeneity and validity of instruments

Testing for exogeneity and validity of instruments

Testing for exogeneity

1. Hausman test (Durbin, Wu, & Hausman): comparison of covariance matrices of OLS and IV estimator under H_0 of exogeneity

Testing the validity of instruments

- 1. Sargan test (look at correlation of residuals with Z)
- 2. Look at correlation of endogenous variables and Z

Hausman test

$$H_0: plim \frac{1}{n} X' \varepsilon = 0$$

- under H_0 : OLS is consistent + efficient and IV is consistent
- under H_0 (exogeneity) OLS and IV give comparable results (both consistent), under H_1 we expect a difference
- ullet Let $d=b_{IV}-b$. Under $H_0: \mathsf{E}(d)pprox 0$ and, if errors are iid , $\mathsf{Var}(d)pprox \mathsf{Var}(b_{IV})-\mathsf{Var}(b)$
- Test statistic: $(b_{IV} b)'(\text{Var}(b_{IV}) \text{Var}(b))^{-1}(b_{IV} b) \approx \chi^2(k)$ under H_0
- Problem: *iid* is a strong assumption and in small samples one can have problems as $Var(b_{IV}) Var(b)$ is not always positive definite.

Alternative testing procedure

Use $\frac{1}{n}\sum_{i=1}^{n}x_{ji}\varepsilon_{i}$ with an estimate of ε_{i} , and next test whether this significantly differs from 0.

• Note that we cannot do this using OLS residuals $e_{OLS,i} = y_i - x_i' b_{OLS}$, as it holds that

$$\frac{1}{n}\sum_{i=1}^{n}x_{ji}e_{OLS,i}=0.$$

• Split x_{jj} in an exogenous and an endogenous part, by performing the "first-stage" regression:

$$x_{ji}=z_i'\gamma_j+v_{ji}.$$

 \Rightarrow It holds $E[x_{ji}\varepsilon_i] = E[v_{ji}\varepsilon_i]$ as we assume that $E[z_i\varepsilon_i] = 0$.

Testing exogeneity

• We can now test if $E[v_{ji}\varepsilon_i]=0$ by testing $H_0:\alpha=0$ in

$$\varepsilon_i = \alpha v_{ji} + w_i,$$

or

$$y_i = x_i'\beta + \alpha(x_{ji} - z_i'\gamma) + w_i,$$

Procedure (for k_0 endogenous variables)

- 1. Regress y on $X \rightarrow e = y Xb$
- 2. Regress all endogenous variables x_j on $Z o v_j = x_j Z \hat{\gamma}_j$
- 3. Regress e on X and v_j , $j = 1, \ldots, k_0$
- 4. $nR^2 \approx \chi^2(k_0)$ under H_0 (exogeneity)

Testing the validity of instruments

A crucial assumption for IV/2SLS is exogeneity of the instruments z_{1i}, \ldots, z_{mi} , that is, $E[z_{li}\varepsilon_i] = 0, \quad l = 1, \ldots, m$.

If this does not hold, then IV is not consistent and the Hausman test is not valid.

We can test this assumption by testing $H_0: \gamma = 0$ in

$$\varepsilon_i = z_i' \gamma + \eta_i,$$

where for implementation it is important that we use a consistent estimate of ε_i . This is given by the "IV" - or "2SLS" -residuals

$$e_{2SLS,i}=y_i-x_i'b_{2SLS}.$$

This leads to the so-called Sargan-test.

Testing the validity of instruments

Procedure Sargan test

- 1. Perform IV for $y = X\beta + \varepsilon$ with instruments $Z \rightarrow e_{IV} = y Xb_{IV}$
- 2. Regress e_{IV} on Z
- 3. $nR^2 \approx \chi^2(m-k)$ under H_0 (valid instruments)

Notes:

- this test only works if the model is overidentified (m > k)
- "Identifying assumptions" must still hold!

Example 5.1: Testing for exogeneity and validity

Let's do a small quiz