COMP 540 Final Note

1 Linear models for regression

1.1 Least squares regression

- Loss function: $J(\theta) = \frac{1}{2m} \sum_{i=1}^{m} (y^{(i)} h_{\theta}(x^{(i)}))^2$ Vectorized form: $J(\theta) = \frac{1}{2m} (X\theta - y)^T (X\theta - y)$
- Gradient: $\nabla_{\theta} J(\theta) = \frac{1}{m} X^T (X\theta y)$
- Closed form solution: $\hat{\theta} = (X^T X)^{-1} X^T y$
- Assume $y^{(i)} = (\theta^*)^T x^{(i)} + \epsilon$, where $E[\epsilon] = 0, Var[\epsilon] = \sigma^2$: $E[\hat{\theta}] = \theta^*, Var[\hat{\theta}] = (X^T X)^{-1} \sigma^2$.
- \equiv MLE on θ with normal distributed error ϵ .

1.2 L2-Regularization: ridge regression

• Loss function:

$$J(\theta) = \frac{1}{2m} \sum_{i=1}^{m} (y^{(i)} - h_{\theta}(x^{(i)}))^2 + \frac{\lambda}{2m} \sum_{j=1}^{d} \theta_j^2$$

- Closed form solutionn: $\hat{\theta} = (X^T X + \lambda I)^{-1} X^T y$
- \equiv MAP on θ with prior: $\theta \sim \mathcal{N}(0, \alpha^2 I)$, with $\frac{\lambda}{m} = \frac{\sigma^2}{\alpha^2}$.

1.3 L1-Regularization: lasso regression

• Loss function:

$$J(\theta) = \frac{1}{2m} \sum_{i=1}^{m} (y^{(i)} - h_{\theta}(x^{(i)}))^2 + \frac{\lambda}{2m} \sum_{j=1}^{d} |\theta_j|$$

• \equiv MAP on θ with prior: $\theta_i \sim Laplace(0, \alpha), \forall j$

1.4 Locally weighted linear regression

- Loss function: $J(\theta) = \frac{1}{2m} \sum_{i=0}^{m} w^{(i)} (y^{(i)} \theta^T x^{(i)})^2$ where $w^{(i)} = \exp\left(-\frac{(x-x^{(i)})^T (x-x^{(i)})}{2\sigma^2}\right)$ Vectorized form: $J(\theta) = \frac{1}{2m} (X\theta y)^T W(X\theta y)$
- Gradient: $\nabla_{\theta} J(\theta) = \frac{1}{m} X^T W(X\theta y)$
- Closed form solution: $\hat{\theta} = (X^T W X)^{-1} X^T W y$
- Non-parametric method.

2 Linear models for classification

2.1 Discriminative models for classification

- $P(y = 1|x) = h_{\theta}(x) = \frac{1}{1 + \exp(-\theta^T x)}$ $\Rightarrow \log \frac{P(y=1|x)}{P(y=0|x)} = \theta^T x$
- Loss (cross-entropy) function: (convex & has a global minimum) $J(\theta) = -\frac{1}{m} \sum_{i=0}^{m} y^{(i)} \log h_{\theta}(x^{(i)}) + (1 y^{(i)}) \log(1 h_{\theta}(x^{(i)}))$
- L2-regularizion: $J_{reg}(\theta) = J(\theta) + \frac{\lambda}{m} \sum_{j=1}^{d} \theta_j^2$ L1-regularizion: $J_{reg}(\theta) = J(\theta) + \frac{\lambda}{m} \sum_{j=1}^{d} |\theta_j|$

2.2 Generative models for classification

2.2.1 Gaussian discriminant analysis (GDA)

- Assumptions: $y \sim Bernoulli(\phi)$ $x|y = 0 \sim \mathcal{N}(\mu_0, \Sigma) \& x|y = 1 \sim \mathcal{N}(\mu_1, \Sigma)$
- Likelihood: $\mathcal{L}(D) = \prod_{i=1}^{m} P(x^{(i)}, y^{(i)}; \phi, \mu_0, \mu_1, \Sigma)$ = $\prod_{i=1}^{m} \phi^{y^{(i)}} (1 - \phi)^{(1-y^{(i)})} \mathcal{N}(x^{(i)}|\mu_1, \Sigma)^{y^{(i)}} \mathcal{N}(x^{(i)}|\mu_0, \Sigma)^{(1-y^{(i)})}$
- Estimation for parameters:

$$\begin{split} & \phi = \frac{1}{m} \sum_{i=1}^{m} y(i) \\ & \mu_1 = \frac{\sum_{i=1}^{m} y^{(i)} x^{(i)}}{\sum_{i=1}^{m} y^{(i)}}, \mu_0 = \frac{\sum_{i=1}^{m} (1 - y^{(i)}) x^{(i)}}{\sum_{i=1}^{m} (1 - y^{(i)})} \\ & \Sigma = \frac{1}{m} \sum_{i=1}^{m} (x - \mu_{y(i)}) (x - \mu_{y(i)})^T \end{split}$$

 Linear decision boundaries when same Σ; quadratic boundaries when each class has its own Σ.

2.2.2 Naive Bayes models

- Assumptions: $P(x|y) = \prod_{i=1}^{d} P(x_i|y)$
- Bernoulli Naive Bayes models are estimated using counts, regularize using Beta prior (

 pre-count).

2.3 Model criteria

- False negative: positive predicted to be negative.
- False positive: negative predicted to be positive.
- specificity = $P(y_{pred} = 0|y = 0) = \frac{TN}{FP + TN}$
- sensitivity = $P(y_{pred} = 1|y = 1) = \frac{TP}{FN+TP}$
- true positive rate (TPR) = sensitivity
- false positive rate (FPR) = 1 specificity
- ROC curve represents FPR and TPR as a function of classification threshold. 0.5 ≤ (area under curve) ≤ 1.0.

2.4 Multiclass classification

2.4.1 One vs. All (OVA) & One vs. One (OVO)

- OVA: not theoretically justified; simple and widely used.
- OVO: needs $O(K^2)$ classifiers for K classes; overfitting!

2.4.2 Softmax

Log-likelyhood:

$$\ell(\mathcal{D}) = \frac{1}{m} \sum_{i=1}^{m} \sum_{c=1}^{K} I(y^{(i)} = c) \log \frac{\exp(\theta^{(c)} T_x(i))}{\sum_{c'} \exp(\theta^{(c')} T_x(i))}$$

- Regularized loss: $J(\theta) = -\ell(\mathcal{D}) + \frac{\lambda}{2m} \sum_{j=1}^d \sum_{c=1}^K \theta_j^{(c)}^2$
- Gradient: $\nabla_{\theta} J(\theta) = -\frac{1}{m} \sum_{i=1}^{m} x^{(i)} (I\{y^{(i)} = c\} P(y^{(i)} = c|x^{(i)};\theta)) + \frac{\lambda}{m} \sum_{j=1}^{d} \theta_{j}^{(c)}$

3 Kernel methods

3.1 Kernel functions

- Gaussian kernel: $\kappa(x, x') = \exp\left(-\frac{\|x x'\|^2}{2\sigma^2}\right)$
- Polynomial kernel: $\kappa(x, x') = (1 + x^T x')^p$
- Mercer's theorem (\Leftrightarrow valid kernel): Gram matrix K whose elements are $\kappa(x^{(i)}, x^{(j)}), 1 \leq i, j \leq m$, should be positive definite for all possible $\{x^{(i)}|1 \leq i \leq m\}$ $\Leftrightarrow \exists \phi \text{ s.t. } \kappa(x, x') = \phi(x)^T \phi(x')$

3.2 Perceptron

- Prediction: $h_{\theta}(x) = \operatorname{sign}(\theta^T x)$
- Update rule: $\theta \leftarrow \theta + \eta x^{(i)} y^{(i)}$, when $h_{\theta}(x^{(i)}) y^{(i)} = -1$
- Convergence bounds: Let $\|x^{(i)}\| \le R, \forall 1 \le i \le m$, the perceptron converges in at most $\frac{R^2 \|\theta^*\|^2}{\gamma^2}$ updates, where $\gamma > 0, \ y^{(i)}(\theta^T x^{(i)}) \ge \ \gamma, \forall 1 \le i \le m. \ (Appendix)$
- Kernalized version: (if $\eta = 1$, $\theta = \sum_{(x,y) \in \mathcal{D}_{mistake}} xy$) $\hat{y} = \text{sign}(\sum_{(x^{(i)},y^{(i)}) \in \mathcal{D}} \alpha^{(i)} \langle x^{(i)}, x \rangle$), in training: update $\alpha^{(i)} \leftarrow \alpha^{(i)} + y^{(i)}$, when $y\hat{y} = -1$

3.3 Support vector machine (SVM)

3.3.1 Maximize margin

- Optimization problem: $\min_{\theta,\theta_0} \frac{1}{2} \|\theta\|^2$, subject to $y^{(i)}(\theta^T x^{(i)} + \theta_0) \geq 1, \forall 1 \leq i \leq m$
- Dual problem: $\max_{\alpha} \sum_{i=1}^{m} \alpha_{i} \frac{1}{2} \sum_{i=1}^{m} \sum_{j=1}^{m} \alpha_{i} \alpha_{j} y^{(i)} y^{(j)} \langle x^{(i)}, x^{(j)} \rangle$ subject to: $\alpha_{i} \geq 0, \forall 1 \leq i \leq m; \sum_{i=1}^{m} \alpha_{i} y^{(i)} = 0$
- Results of solving Lagrange: $\theta = \sum_{i=1}^m \alpha^{(i)} x^{(i)} y^{(i)} \ \& \ \sum_{i=1}^m \alpha^{(i)} y^{(i)} = 0$
- KTT condition yields: $\alpha^{(i)}[y^{(i)}(\theta^Tx^{(i)} + \theta_0) 1] = 0, \forall 1 \le i \le m$
- Prediction: $h_{\theta}(x) = \operatorname{sign}(\sum_{i=1}^{m} \alpha^{(i)} y^{(i)} \langle x^{(i)}, x \rangle + \theta_0)$

3.3.2 Non-separable case

- Optimization problem: $\min_{\theta,\theta_0} \frac{1}{2} \|\theta\|^2 + C \sum_{i=1}^m \xi^{(i)},$ subject to $y^{(i)}(\theta^T x^{(i)} + \theta_0) \ge 1 \xi^{(i)}, \forall 1 \le i \le m;$ and $\xi^{(i)} \ge 0, \forall 1 \le i \le m$
- Dual problem: $\max_{\alpha} \sum_{i=1}^{m} \alpha_i \frac{1}{2} \sum_{i=1}^{m} \sum_{j=1}^{m} \alpha_i \alpha_j y^{(i)} y^{(j)} \langle x^{(i)}, x^{(j)} \rangle$ subject to: $0 \le \alpha_i \le C, \forall 1 \le i \le m; \sum_{i=1}^{m} \alpha_i y^{(i)} = 0$

3.3.3 Hinge loss - another view of SVM

- Define $h_{\theta}(x) = \theta^T x + \theta_0$, rewrite constraints into $\xi^{(i)} = \max(0, 1 y^{(i)} h_{\theta}(x))$.
- Loss function: $J(\theta) = C \sum_{i=1}^{m} \max(0, 1 y^{(i)} h_{\theta}(x)) + \frac{1}{2} ||\theta||^2$

3.3.4 Multiclass SVM

• Loss function:

$$J(\theta) = \sum_{i=1}^{m} \sum_{i \neq y(i)} \max(0, \theta^{(j)^{T}} x^{(i)} - \theta^{y^{(j)^{T}}} x^{(i)} + \Delta)$$

Δ is some fixed margin.

3.3.5 SVM for regression

- Loss function: $L_{\epsilon}(y, \hat{y}) = \max(0, |y \hat{y}| \epsilon)$ in $J(\theta) = C \sum_{i=1}^{m} L_{\epsilon}(y^{(i)}, \theta^T x^{(i)} + \theta_0) + \frac{1}{2} \theta^T \theta$
- $J(\theta)$ is convex but not differentiable, the optimization problem is unconstrained.
- Quadratic programming (QP) problem: $\min_{\theta} C \sum_{i=1}^{m} (\xi_{+}^{(i)} + \xi_{-}^{(i)}) + \frac{1}{2} \theta^{T} \theta, \text{ subject to}$ $\xi_{+}^{(i)} \geq 0, \xi_{-}^{(i)} \geq 0, \forall 1 \leq i \leq m;$ $\theta^{T} x^{(i)} + \theta_{0} \xi_{-}^{(i)} \epsilon \leq y^{(i)} \leq \theta^{T} x^{(i)} + \theta_{0} + \xi_{+}^{(i)} + \epsilon$
- Lagrange yields: $\theta = \sum_{i=1}^{m} (\alpha_{+}^{(i)} \alpha_{-}^{(i)}) x^{(i)}$, where $\alpha_{+}^{(i)}$ and $\alpha_{-}^{(i)}$ are Lagrange multipliers of two constraints.
- Prediction: $h_{\theta}(x) = \sum_{i=1}^{m} (\alpha_{+}^{(i)} \alpha_{-}^{(i)}) \langle x^{(i)}, x \rangle + \theta_{0}$

4 Neural Networks

Note: A feed forward network with a linear output layer and at least one hidden layer with any squashing activation function (e.g. sigmoid, ReLU), can approaximate any function from $\mathbb{R}^d \to \mathbb{R}$ to arbitrary percision with enough hidden units.

4.1 Activation functions

- $sigmoid(x) = \sigma(x) = \frac{1}{1 + \exp(-x)}$ $\frac{d}{dx}\sigma(x) = \sigma(x)(1 - \sigma(x))$
- $\tanh(x) = \frac{\exp(x) \exp(-x)}{\exp(x) + \exp(-x)}$ (unbias, better than sigmoid)
- ReLU(x) = max(0, x)
- Softmax function: $g(z_i) = \frac{\exp(z_i)}{\sum_j \exp(z_j)}$ (output layer)

4.2 Backpropagation

- Using chain rule, propagate derivatives in inverse order.
- Gradients need to be added up at forks (accumulative)!

4.3 Convolutional neural networks

- Let K be filter number, F be filter size, S be stride, P be padding.
- A conv layer takes as input of a volume $W_1 \times H_1 \times D_1$, produces an output volume $W_2 \times H_2 \times D_2$:

$$W_2 = \frac{W_1 - F + 2P}{S} + 1$$

$$H_2 = \frac{H_1 - F + 2P}{S} + 1$$

$$D_2 = K$$

- Total number of parameters: $F \times F \times D_1$ weights per filter; $F \times F \times D_1 \times K$ weights, K biases.
- Pooling layer (subsampling): input $W_1 \times H_1 \times D$, output $W_2 \times H_2 \times D$:

$$W_2 \times H_2 \times D$$
.
 $W_2 = \frac{W_1 - F}{S} + 1$
 $H_2 = \frac{H_1 - F}{S} + 1$

• Classical architecture: [(CONV - RELU)*N - POOL]*M - (FC - RELU)*K - SOFTMAX

4.4 Optimization for training deep models

4.4.1 Momentum

- Require: learning rate α, initial parameter θ, batch size m, momentum parameter μ, initial velocity v.
- Sample a mini batch of m examples $(x^{(i)}, y^{(i)})$
- Compute gradient estimate $\hat{g} \leftarrow \frac{1}{m} \nabla_{\theta} \sum_{i} L(h_{\theta}(x^{(i)}), y^{(i)})$
- Compute velocity update $v \leftarrow \mu v \alpha \hat{q}$
- Apply update $\theta \leftarrow \theta + v$

4.4.2 Nesterov momentum

- Require: learning rate α , initial parameter θ , batch size m, momentum parameter μ , initial velocity v.
- Sample a mini batch of m examples $(x^{(i)}, y^{(i)})$
- Apply interim update $\tilde{\theta} \leftarrow \tilde{\theta} + \mu v$
- Compute gradient at interim point $\hat{g} = \frac{1}{m} \nabla_{\tilde{\theta}} \sum_{i} L(h_{\tilde{\theta}(x^{(i)}),y^{(i)}})$
- Compute velocity update $v \leftarrow \mu v \alpha \hat{g}$
- Apply update $\theta \leftarrow \theta + v$

4.4.3 AdaGrad

- Require: step size α , initial parameter θ , batch size m, $\delta = 10^{-7}$ (constant for numerical stability)
- Compute gradient on minibatch $\hat{g}_{ij}^{(l)} \leftarrow \frac{1}{m} \nabla_{\theta_{ij}}^{(l)} \sum_{i} L(h_{\theta}(x^{(i)}), y^{(i)})$
- Accumulate squared gradient $r_{ij}^{(l)} \leftarrow r_{ij}^{(l)} + \hat{g}_{ij}^{(l)} * \hat{g}_{ij}^{(l)}$
- Compute gradient $\Delta \theta_{ij}^{(l)} \leftarrow -\frac{\alpha}{\delta + \sqrt{r_{ij}^{(l)}}} \hat{g}_{ij}^{(l)}$
- Apply update $\theta_{ij}^{(l)} \leftarrow \theta_{ij}^{(l)} + \Delta \theta_{ij}^{(l)}$

4.4.4 RMSprop

- Require: step size α , initial parameter θ , batch size m, $\delta = 10^{-7}$ (constant for numerical stability), exponential decay rate ρ .
- Compute gradient on minibatch $\hat{g}_{ij}^{(l)} \leftarrow \frac{1}{m} \nabla_{\theta_{ij}}^{(l)} \sum_{i} L(h_{\theta}(x^{(i)}), y^{(i)})$
- Accumulate squared gradient $r_{ij}^{(l)} \leftarrow \rho r_{ij}^{(l)} + (1-\rho)\hat{g}_{ij}^{(l)} * \hat{g}_{ij}^{(l)}$
- Compute gradient $\Delta \theta_{ij}^{(l)} \leftarrow -\frac{\alpha}{\delta + \sqrt{r_{ij}^{(l)}}} \hat{g}_{ij}^{(l)}$
- Apply update $\theta_{ij}^{(l)} \leftarrow \theta_{ij}^{(l)} + \Delta \theta_{ij}^{(l)}$

4.4.5 Adam (adaptive moments)

- Require: step size α (10⁻³ default), initial parameter θ , batch size m, exponential decay rates for moment estimates ρ_1 (0.99 default) & ρ_2 (0.999 default), $\delta = 10^{-7}$ (constant for numerical stability)
- Initialize time step t = 0, first and second moment s = 0, r = 0
- Compute gradient on minibatch $\hat{g}_{ij}^{(l)} \leftarrow \frac{1}{m} \nabla_{\theta_{ij}}^{(l)} \sum_{i} L(h_{\theta}(x^{(i)}), y^{(i)})$
- $t \leftarrow t + 1$
- Update biased first moment $s_{ij}^{(l)} \leftarrow \rho_1 s_{ij}^{(l)} + (1 \rho_1) \hat{g}_{ij}^{(l)}$
- Update biased second moment $r_{ij}^{(l)} \leftarrow \rho_2 r_{ij}^{(l)} + (1 \rho_2) \hat{g}_{ij}^{(l)} * \hat{g}_{ij}^{(l)}$
- Correct bias in first moment $\hat{s}_{ij}^{(l)} = \frac{s_{ij}^{(l)}}{1 - \rho_1^t}$
- Correct bias in second moment $\hat{r}_{ij}^{(l)} = \frac{r_{ij}^{(l)}}{1-\rho_{ij}^{l}}$
- Compute gradient $\Delta \theta_{ij}^{(l)} \leftarrow -\frac{\alpha \hat{s}_{ij}^{(l)}}{\delta + \sqrt{\hat{r}_{ij}^{(l)}}}$
- Apply update $\theta_{ij}^{(l)} \leftarrow \theta_{ij}^{(l)} + \Delta \theta_{ij}^{(l)}$

5 Decision trees

5.1 Cost functions

- Misclassification rate: $(\hat{y} = \text{the majority label in } \mathcal{D})$ $cost(\mathcal{D}) = \frac{1}{|\mathcal{D}|} \sum_{(x,y) \in \mathcal{D}} I(y \neq \hat{y})$
- Entropy: $(p = \text{fraction of positive examples in } \mathcal{D})$ $cost(\mathcal{D}) = -p \log_2 p - (1-p) \log_2 (1-p)$
- Gini index: (same as entropy) $cost(\mathcal{D}) = 2p(1-p)$

5.2 Decision tree for regression

• Define $cost(\mathcal{D}) = \sum_{i=1}^{m} (y^{(i)} - \overline{y})^2$ where $\overline{y} = \frac{1}{|\mathcal{D}|} \sum_{i=1}^{m} y^{(i)}$

5.3 Node is not worth splitting when

- Node is pure
- Depth exceeds max depth
- $|\mathcal{D}_{left}|$ or $|\mathcal{D}_{right}|$ is too small
- Reduction in cost is too small

5.4 Avoid overfitting

- Early stopping stop growing the tree when the decrease in error is not sufficient to justify the complexity of an additional level.
- Post pruning grow the full tree and then prune using a validation set to guide subtree removal. Evaluate CV error on each subtree and pick tree whose error is within 1 standard deviation of minimum.

6 Ensembles

6.1 Bagging

- Assume errors of individual members are uncorrelated.
- For regression: when $h_l(x) = f(x) + \epsilon_l(x)$ for $1 \le l \le L$, and $\epsilon_l \sim \mathcal{N}(0, \sigma_l^2)$ $\Rightarrow E_{bag} = \frac{1}{L} E_{avg}$ (expected squared error)
- For classification: $\epsilon = \text{error of each classifier, and } \epsilon < \frac{1}{2}$ $\Rightarrow E_{bag} = \sum_{i=\frac{L}{n}+1}^{L} {i \choose L} \epsilon^{i} (1-\epsilon)^{L-i}$

6.2 Boosting

- Loss function: $J_l = \sum_{i=1}^m w_l^{(i)} I(h_l(x^{(i)}) \neq y^{(i)})$
- Prediction: $h(x) = \operatorname{sign}(\sum_{l=1}^{L} \alpha_l h_l(x))$
- Adaboost algorithm: initialize $w_1^{(i)} = \frac{1}{m}$, for $1 \le i \le m$
- 1. fit h_l to minimize $J_l = \frac{1}{m} \sum_{i=1}^m w_l^{(i)} L(y^{(i)}, h_l(x^{(i)}))$
- 2. calculate error rate $\epsilon_l = \frac{\sum_{i=1}^m w_l^{(i)} \mathrm{I}(h_l(x^{(i)} \neq y^{(i)}))}{\sum_{i=1}^m w_l^{(i)}}$
- 3. calculate $\alpha_l = \frac{1}{2} \log \left(\frac{1 \epsilon_l}{\epsilon_l} \right)$, stop if $\epsilon_l \ge \frac{1}{2}$
- 4. update $w_{l+1}^{(i)} = \begin{cases} w_l^{(i)} \exp{(\alpha_l)}, \text{ incorrect on } x^{(i)} \\ w_l^{(i)} \exp{(-\alpha_l)}, \text{ correct on } x^{(i)} \end{cases}$

6.3 Gradient boosting

6.3.1 Gradient boosting for regression

• Residuals are negative gradients (squared error loss): $J = \sum_{i=1}^{m} \frac{1}{2} (y - h(x))^2 \Rightarrow \frac{\partial J}{\partial h(x^{(i)})} = h(x^{(i)}) - y^{(i)}$

6.3.1 Gradient boosting for classification

- Loss function: $J = \frac{1}{m} \sum_{i=1}^{m} D_{KL}(y^{(i)}, h(x^{(i)}))$
- Gradient boosting algorithm (k classes):
- Start with an initial $h^0...h^k$ for $x^{(1)}...x^{(m)}$
- Repeat until convergence: Calculate matrix of gradients, it each h_{add} to the negative gradient, $h \leftarrow h + h_{add}$

7 Probabilistic graphical models

7.1 Directed models – Bayesian network

- $P(X) = \prod_{i} P(x_i | Parents(x_i))$
- Reduce number of parameters O(kⁿ) → O(nk^m) if each variable in graph has no more than m parents.

7.2 Undirected models - Markov network

- $\tilde{P}(X) = \prod_{C \in G} \phi(C)$ (C is a clique in graph)
- Partition function: $Z = \int_X \tilde{P}(X) dX$ or $Z = \sum_X \tilde{P}(X)$
- $P(X) = \frac{1}{7}\tilde{P}(X)$
- Energy function $E: \tilde{P}(X) = \exp(-E(X))$ high (low) energy \Leftrightarrow low (high) $\tilde{P}(X)$

7.3 Sampling

7.3.1 Ancestral sampling

- For directed graphical models, polynomial time.
- Algorithm:
 - Sort variables in topological order
 - Sample x_i from distribution $P(x_i|Parents(x_i))$

7.3.2 Gibbs sampling

- · For undirected graphical models.
- Algorithm:
 - Start with randomly generated values $x_1, ..., x_n$
- Iteratively visit each x_i and sample a value for it based on $P(x_i|Neighbors(x_i))$
- Repeat previous step, generate stream of samples

7.4 Hidden Markov models

- Specified by sets S (hidden states), O (observations) and probability parameters $\lambda = [\pi, a, b]$
- $-\pi$ is initial state probability
- a is hidden state transition probability
- b is emission probability
- Inference problems:
 - Filtering: $P(X_t|e_1,...,e_t)$
 - Smoothing: $P(X_k | e_1, ..., e_t), k < t$
 - Most likely state sequence:
 - $\arg \max_{X_1,...,X_t} P(X_1,...,X_t|e_1,...,e_t)$

7.4.1 Forward computation - filtering

- Define: $\alpha_t(i) = P(e_1, ..., e_t, X_t = s_i)$
- Algorithm:
- $\begin{array}{l} 3^{-1} \\ -\alpha_0(i) = \pi_i, \ 1 \leq i \leq n \ \text{where} \ |S| = n \\ -\alpha_{t+1}(i) = b_j(e_{t+1}) \sum_{i=1}^n \alpha_t(i) a_{ij}, \ 1 \leq j \leq n, \\ 0 \leq t \leq T-1 \end{array}$
- Time complexity: $O(n^2T)$

7.4.2 Backward computation - smoothing

- $P(X_k|e_1,...,e_t) \propto P(e_{k+1},...,e_t|X_k)P(X_k|e_1,...,e_t)$ where
- Define: $\beta_k(i) = P(e_{k+1}, ..., e_t | X_k = s_i)$
- Algorithm:
 - $\begin{array}{l} -\beta_T(i) = 1, \ 1 \leq i \leq n \\ -\beta_k(i) = \sum_{j=1}^n a_{ij} b_j(e_{k+1}) \beta_{k+1}(j), \ 1 \leq j \leq n, \\ 0 \leq k \leq T-1 \end{array}$
- Time complexity: $O(n^2T)$

7.4.3 Viterbi algorithm – most likely sequence

- $\delta_t(i) = \max_{X_1,...,X_{t-1}} P(X_1,...,X_{t-1},X_t = s_i,e_1,...,e_t)$
- Algorithm:
- $-\delta_0(i) = \pi(i), 1 < i < n$ $-\delta_{t+1}(j) = \max_{i} \delta_{t}(i) a_{ij} b_{i}(e_{t+1}), 1 < j < n, 0 < t < T-1$
- Time complexity: $O(n^2T)$

7.4.4 Parameters estimation

- Paired sequences: $\hat{a}_{ij} = \frac{\#s_i \to s_j}{\#s_i} \& \hat{b}_j(e_k) = \frac{\#s_j \to e_k}{\#s_i}$
- Observation sequences only Baum-Welch EM:
- Define: $\xi_t(i,j) = P(X_t = s_i, X_{t+1} = s_i | e_1, ..., e_T, \lambda)$
- $\begin{array}{l} \xi_t(i,j) = \frac{\alpha_t(i)a_{ij}b_j(e_{t+1})\beta_{t+1}(j)}{\sum_{i=1}^n \sum_{j=1}^n \alpha_t(i)a_{ij}b_j(e_{t+1})\beta_{t+1}(j)} \\ \text{Let } \gamma_t(i) = P(X_t = s_i|e_1,...,e_T,\lambda) = \sum_{j=1}^n \xi_t(i,j) \end{array}$
- Estimate $\hat{\pi}_i = \gamma_1(i)$, $\hat{a}_{ij} = \frac{\sum_{t=1}^{T-1} \xi_t(i,j)}{\sum_{t=1}^{T-1} \gamma_t(i)}$
- $\hat{b}_{j}(e_{k}) = \frac{\sum_{t=1}^{T} \gamma_{t}(j) * I(E_{t} = e_{k})}{\sum_{t=1}^{T} \gamma_{t}(j)}$
- Algorithm:
 - Guess $\lambda_0 = [\pi_0, a_0, b_0]$
 - Repeat until convergence:
 - Calculate α , β from λ
 - Re-estimate λ from α, β

8 Unsupervised learning

8.1 Principal components analysis (PCA)

- Assume: data distrubution is unimodal Gaussian (fully explained by mean & variance)
- Assume: information to be preserved is in the variance
- Project data $\mathbb{R}^d \to \mathbb{R}^k$ (k < d), maximizing variance.

8.1.1 PCA method

- Zero-mean, unit variance transform on D
- Find S = covariance matrix of transformed D
- Find $\lambda_1, ..., \lambda_k$ (the k largest eigenvalues of S) and associated eigenvectors u_1, \dots, u_k
- Project $x^{(i)} \mapsto [u_1^T x^{(i)}, \dots, u_t^T x^{(i)}]^T$, where $x^{(i)} \in \mathbb{R}^d$

8.1.2 Kernel PCA

- Idea: map $x^{(i)} \mapsto \phi(x^{(i)})$ where $\phi: \mathbb{R}^d \to \mathbb{R}^D$ (D >> d)
- The result when project back to \mathbb{R}^d will be nonlinear.
- Algorithm:
- Pick a kernel
- Construct kernel matrix K over data $x^{(1)}, ..., x^{(m)}$
- Zero mean the kernel matrix K to get \tilde{K}
- Solve the eigenvalue problem $K\alpha = \lambda \alpha$, $\alpha \in \mathbb{R}^m$
- For a new x, we project it as: $y_j = \sum_{i=1}^m \alpha_j^{(i)} \kappa(x, x^{(i)}),$ for j = 1, ..., L (# of components) where eigenvectors are ordered by value

8.2 Expectation maximization algorithm (EM)

8.2.1 K-means

- Cost function: $J = \sum_{i=1}^{m} \sum_{k=1}^{K} z_k^{(i)} ||x^{(i)} \mu_k||^2$
- E-step: cluster assignment, minimize J wrt. z, fix μ
- M-step: relocate means, minimize J wrt. μ , fix z
- Time complexity: O(mK) per iteration
- Converges to local minimum & vulnerable to outliers.

8.2.2 Gaussian mixture model (GMM)

- Generative mode: $P(x^{(i)}) = \sum_{k=1}^{K} P(z^{(i)} = k) P(x^{(i)} | z^{(i)} = k)$ $z^{(i)} \sim \text{Multinomial}(\pi); \pi_k > 0, \sum_k \pi_k = 1$ $x^{(i)}|_{z(i)=k} \sim \mathcal{N}(\mu_k, \Sigma_k)$
- Infer $z^{(i)}$ for each $x^{(i)}$, where $\theta = \{\pi, \mu, \Sigma\}$: $P(z^{(i)} = k|x^{(i)}; \theta) = \frac{P(z^{(i)} = k)P(x^{(i)}|z^{(i)} = k; \theta)}{\sum_{k'} P(z^{(i)} = k')P(x^{(i)}|z^{(i)} = k'; \theta)}$
- Soft EM algorithm:
- Guess values of θ = {π, μ, Σ}
- E-step: calculate the responsibility of each component toward generating $x^{(i)}$: $r_k^{(i)} = P(z^{(i)} = k|x^{(i)}; \theta)$
- M-step: given $r_k^{(i)}$ and $x^{(i)}, \ 1 \leq i \leq m, \ 1 \leq k \leq K,$ re-estimate π, μ, Σ :

$$\pi_k = \frac{1}{m} \sum_{i=1}^m r_k^{(i)}, \qquad \mu_k = \frac{\sum_{i=1}^m r_k^{(i)} x^{(i)}}{\sum_{i=1}^m r_k^{(i)}},$$

$$\Sigma_k = \frac{\sum_{i=1}^m r_k^{(i)} (x^{(i)} - \mu_k) (x^{(i)} - \mu_k)^T}{\sum_{i=1}^m r_k^{(i)}}$$

9 Reinforcement learning

- 9.1 Markov decision process (MDP)
- 9.2 Model-based RL
- 9.3 Model-free RL
- 9.4 TD learning
- 9.5 Q learning

Appendix

Distributions

- Poisson distribution PMF: $P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$
- Normal distribution PDF:

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{(x-\mu)^2}{2\sigma^2}\right\}$$

• Multivariate normal distribution PDF: $f(X; \mu, \Sigma) = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} \exp\left(-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)\right)$

$$f(X; \mu, \Sigma) = \frac{1}{\sqrt{(2\pi)^k |\Sigma|}} \exp\left(-\frac{1}{2}(X - \mu)^T \Sigma^{-1}(X - \mu)\right)$$

- Laplace distribution PDF: $f(x; \mu, b) = \frac{1}{2b} \exp\left\{-\frac{|x-\mu|}{b}\right\}$
- Beta distribution PDF: $f(x; \alpha, \beta) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1}$
- Given that $X_1 \sim Pois(\lambda_1), X_2 \sim Pois(\lambda_2), X = X_1 + X_2,$ X_1 and X_2 are independent: $X \sim Pois(\lambda_1 + \lambda_2)$.
- Given that $P(X_0 = x_0) = \alpha_0 \exp\left\{-\frac{(x_0 \mu_0)^2}{2\sigma_0^2}\right\}$, and that $P(X_1 = x_1 | X_0 = x_0) = \alpha_1 \exp\left\{-\frac{(x_1 - x_0)^2}{2\sigma^2}\right\}$: $P(X_1 = x_1) = \alpha_0 \alpha_1 \sqrt{\frac{2\pi\sigma_0^2 \sigma_1^2}{\sigma_0^2 + \sigma_0^2}} \exp\left\{-\frac{1}{2} \frac{(x_1 - \mu_0)^2}{\sigma_0^2 + \sigma_0^2}\right\}$

Information theory

- Conditional information:
- $H(Y|X) = \sum_{y \in Y} \sum_{x \in X} p(x,y) \log \frac{p(x)}{p(x,y)}$
- Mutual information: $I(X;Y) = \sum_{y \in Y} \sum_{x \in X} p(x,y) \log \frac{p(x,y)}{p(x)P(y)}$
- Kullback-Leibler (KL) divergence: $D_{KL}(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)}$
- Cross entropy: $H(p,q) = E_p[-\log q] = H(p) + D_{KL}(p||q)$

Convex function

- A function f(x) is convex on a set S iff for $\lambda \in [0,1]$, and $\forall x, y \in S$: $f(\lambda x + (1 \lambda)y) \leq \lambda f(x) + (1 \lambda)f(y)$.
- A function f(x) is convex on a set S iff $\frac{d^2}{dx^2}f(x)$ is positive semidefinite everywhere in the set.

Convergence of perceptron

Proof. Let $\theta^{(k-1)}$ be the parameter vector when the algorithm makes a mistake on (x,y).

$$\theta^{(k)} = \theta^{(k-1)} + \eta xy$$

Take dot product on both sides with θ^* (some separating hyperplane)

$$\begin{aligned} \boldsymbol{\theta}^{*T} \boldsymbol{\theta}^{(k)} &= \boldsymbol{\theta}^{*T} (\boldsymbol{\theta}^{(k-1)} + \eta x y) \\ &= \boldsymbol{\theta}^{*T} \boldsymbol{\theta}^{(k-1)} + \eta y (\boldsymbol{\theta}^{*T} x) \\ &\geq \boldsymbol{\theta}^{*T} \boldsymbol{\theta}^{(k-1)} + \eta \gamma \end{aligned}$$

If $\theta^{(0)}$ is all zeros vector, then

$$\theta^{*T}\theta^{(k)} > \eta k \gamma$$

From the update rule,

$$\begin{aligned} \boldsymbol{\theta}^{(k)} &= \boldsymbol{\theta}^{(k-1)} + \eta x y \\ \|\boldsymbol{\theta}^{(k)}\|^2 &= \|\boldsymbol{\theta}^{(k-1)} + \eta x y\|^2 \\ &= \|\boldsymbol{\theta}^{(k-1)}\|^2 + \eta^2 y^2 \|x\|^2 + 2\eta y (\boldsymbol{\theta}^{(k)})^T x) \\ &\leq \|\boldsymbol{\theta}^{(k-1)}\|^2 + \eta^2 \|x\|^2 \\ &\leq \|\boldsymbol{\theta}^{(k-1)}\|^2 + \eta^2 R^2 \end{aligned}$$

Starting with $\theta^{(0)}$ of all zeros,

$$\|\theta^{(k-1)}\|^2 \le k\eta^2 R^2 \tag{2}$$

Putting (1) and (2) together, $k \leq \frac{R^2 \|\theta^*\|^2}{\gamma^2}$.

Mercer's theorem

Proof. Since K is positive definite, $K = u^T \Lambda u$, where Λ is a diagonal matrix with entries $\lambda^{(i)} > 0$. Consider an element $\kappa(x^{(i)}, x^{(j)})$ of K. We can construct this element as follows

$$\kappa(x^{(i)}, x^{(j)}) = (\Lambda^{\frac{1}{2}} u_{:,i})^T (\Lambda^{\frac{1}{2}} u_{:,j})$$

Now define $\phi(x^{(i)}) = \Lambda^{\frac{1}{2}} u_{:,i}$, then we have $\kappa(x^{(i)}, x^{(j)}) = \phi(x^{(i)})^T \phi(x^{(j)})$.

1) Adaboost principle

Kernel PCA derivation

Correctness of EM

Proof. For any **Z** with non-zero probability $p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})$, we can write:

$$\log p(\mathbf{X}|\boldsymbol{\theta}) = \log p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta}) - \log p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})$$

We take the expectation over possible values of the unknown data **Z** under the current parameter estimate $\theta^{(t)}$ by multiplying both

sides by $p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(t)})$ and summing over \mathbf{Z} . The left-hand side is the expectation of a constant, so we get:

$$\log p(\mathbf{X}|\boldsymbol{\theta}) = \sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(t)}) \log p(\mathbf{X}, \mathbf{Z}|\boldsymbol{\theta})$$
$$- \sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta}^{(t)}) \log p(\mathbf{Z}|\mathbf{X}, \boldsymbol{\theta})$$
$$= Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) + H(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$$

where $H(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$ is defined by the negated sum it is replacing. This last equation holds for any value of $\boldsymbol{\theta}$ including $\boldsymbol{\theta} = \boldsymbol{\theta}^{(t)}$,

$$\log p(\mathbf{X}|\boldsymbol{\theta}^{(t)}) = Q(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)}) + H(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)})$$

and subtracting this last equation from the previous equation gives

$$\begin{split} \log p(\mathbf{X}|\boldsymbol{\theta}) - \log p(\mathbf{X}|\boldsymbol{\theta}^{(t)}) = & Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) - Q(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)}) \\ & + H(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) - H(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)}) \end{split}$$

However, Gibbs' inequality tells us that $H(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) \geq H(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)})$, so we can conclude that

$$\log p(\mathbf{X}|\boldsymbol{\theta}) - \log p(\mathbf{X}|\boldsymbol{\theta}^{(t)}) \ge Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) - Q(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)})$$

In words, choosing $\boldsymbol{\theta}$ to improve $Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})$ beyond $Q(\boldsymbol{\theta}^{(t)}|\boldsymbol{\theta}^{(t)})$ cannot cause $\log p(\mathbf{X}|\boldsymbol{\theta})$ to decrease below $\log p(\mathbf{X}|\boldsymbol{\theta}^{(t)})$, and so the marginal likelihood of the data is non-decreasing.