We develop **two new inference strategies** for **performing statistically correct a**nd **computational efficient inference** in complicated real-world simulators.

Efficient Bayesian Inference for Nested Simulators

Bradley J. Gram-Hansen 1,2,* Christian Schroeder de Witt 1,* Phil H.S. Torr 1 Yee Whye Teh 2 Atılım Güneş Baydin 1 Tom Rainforth 2

Department of Engineering Science, Department of Statistics, University of Oxford

Equal Contribution

Background

- Simulators arise in a number of industrial and scientific domains, encoding sophisticated generative models^{[3][4]}.
- Probabilistic programming provides a way to perform statistical inference over simulations of events in a programmatic way.
- Simulators contain many nested sub-procedures that generate *priors* that the outer simulator program is dependent upon.

Problem

 In order leverage scalable variational inference (VI) and Markov Chain Monte Carlo (MCMC) inference methods, we need to construct the program density for the entire simulator, nested sub-procedures inhibit this.

Solution

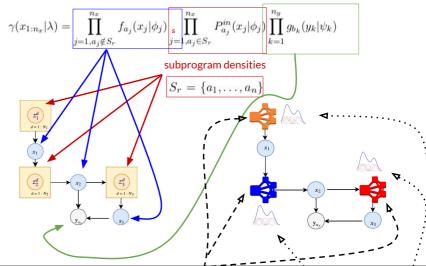
- \circ Introduce a way of replacing sub-procedures by a variational surrogate which turns an intractable program density, into a tractable density (Method 1)^[2]
- Our second method learns an unbiased regressor to predict
 the marliginal of the sub-procedure. Once learnt, we
 combine that with the output of evaluating the
 sub-procedure to learn a approximation that can be used to
 construct the full program density (Method 2)^[2]

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How does it work?

Can write the unnormalized simulator program density as:



Method 1: Learn the Surrogate via a Variational Objective

Replace a sampling sub-procedures by using input-output pairs in order to learn an approximate surrogate:

$$\begin{split} P_{pr}(x_{1:n_x}|\lambda) &:= \prod_{j=1}^{n_x} f_{a_j}(x_j|\phi_j) \simeq q(x_{1:n_x}|\lambda;\kappa) \\ &= \prod_{\substack{j=1\\a_j \notin S_r}}^{n_x} f_{a_j}(x_j|\phi_j) \prod_{\substack{j=1\\a_j \in S_r}}^{n_x} q_{a_j}^{in}(x_j|\phi_j;\eta_{a_j}) \end{split}$$

Here, the variational parameters can be learned by optimising a variational objective:

$$\kappa^* = \underset{\sim}{\operatorname{argmin}} \ J(\eta) = \underset{\sim}{\operatorname{argmin}} \ \mathbb{E}_{X \sim P_{pr}} \left[-\log q(X = x_{1:n_x} | \phi; \kappa) \right]$$

where the variational parameters for each sub-procedure are in the set:

$$\kappa = \{\eta_{a_j}; a_j \in S_r\}$$

with corresponding gradient update:

$$\nabla_{\eta_r} J(\kappa) \approx \frac{1}{N} \sum_{n=1}^N \sum_{j=1}^{n_x} \mathbb{I}(r=a_j) \nabla_{\eta_r} \log(q_r^{in}(x_j^n | \phi_j; \eta_r))$$

where: $x_{1:n_x}^n \stackrel{i.i.d.}{\sim} P_{pr}(x_{1:n_x|\lambda})$

Method 2: Learn the normalisation constant of the nested sub-procedure.

Alternatively, we can rewrite the unnormalized simulator program density as:

$$\begin{split} \gamma(x_{1:n_x}|\lambda) &= \prod_{i=1,a_i \notin S_R}^{n_x} f_{a_i}(x_i;\phi_i) \overline{\prod_{j=1,a_j \in S_R}^{n_x} \frac{\gamma_{a_j}^{in}(x_j|\phi_j)}{I_{a_j}^{in}(\phi_j)}} \prod_{k=1}^{n_y} g_{b_k}(y_k;\phi_k) \\ \text{with the normalisation constant (marginal)} \\ I_{a_j}^{in}(\phi_j) &= \int \gamma_{a_j}^{in}(x_j = z_{1:n_x}^{a_j}|\phi_j) dz_{1:n_x}^{a_j} \end{split}$$

The marginal is not a normalized density, hence approximate it using a regressor:

$$P_{a_j}^{in}(x_j|\phi_j) \simeq \frac{\gamma_{a_j}^{in}(x_j|\phi_j)}{R_{a_j}(\phi_j;\tau)}$$

We learn an unbiased mean using the L2-Norm^[1]:

$$\begin{split} \mathcal{L} &= \mathbb{E}_{\phi_{j}} \left\{ \mathbb{E}_{\hat{I}_{a_{i}}} \left[\left\| R_{a_{j}}(\phi_{j}; \tau) - \hat{I}_{a_{i}}(\phi_{j}) \right\|_{2}^{2} \ \middle| \ \phi_{j} \right] \right\} \\ \nabla_{\kappa} \mathcal{L} &= \mathbb{E}_{\phi_{j}} \left\{ \mathbb{E}_{\hat{I}_{a_{i}}} \left[\nabla_{\kappa} \left\| R_{a_{j}}(\phi_{j}; \tau) - \hat{I}_{a_{i}}(\phi_{j}) \right\|_{2}^{2} \ \middle| \ \phi_{j} \right] \right\} \end{split}$$

Once trained, we can run inference on the approximate, unnormalised, target program density:

$$\gamma(1:x_{n_x}|\lambda) = \prod_{i=1,a_i \notin S_R}^{n_x} f_{a_i}(x_i;\phi_i) \prod_{k=1,b_k}^{n_y} g_{b_k}(y_k;\phi_k) \prod_{j=1,a_i \in S_R}^{n_x} \frac{\gamma_{a_j}^{in}(x_j|\phi_j)}{R_{a_j}(\phi_j;\tau)}$$

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