Review of linear models

4 Sep 2023

```
## it's nice to include packages at the top
## (and NOT automatically install them)
library(faux)
library(grid)
library(ggplot2); theme_set(theme_bw())
```

Basics

- assume $\mathbf{y} \sim \text{Normal}(\mathbf{X}\beta, \sigma)^1$
- **X** is the *model matrix*, can be anything we want it to be
- the *Gauss-Markov theorem* (Wikipedia) makes weaker assumptions: $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \epsilon$; as long as ϵ is mean-zero, homoscedastic with finite variance, and uncorrelated ... then the OLS solution

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^{\top} \mathbf{X})^{-1} \mathbf{X}^{\top} \mathbf{y}$$

is the BLUE (or MVUE).

• we'll embrace the assumptions (which are needed for inference!)

Computation

- matrix decompositions (QR with pivoting)
- big problems: biglm, speedglm, RcppEigen::fastLm
 - optimized BLAS, kernel trick, etc.
 - memory vs speed vs robustness ...
 - -p vs. n vs. many-small-regressions vs. ...

¹Notation-abuse warning ...

Inference

- σ^2 (residual variance) is RSS/(n-p)
- The covariance matrix is $\Sigma = \sigma^2(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}$.
- Individual coefficients are t-distributed
- Linear combinations of coefficients (contrasts or predictions) are t-distributed with covariance matrix $\mathbf{C}^{\top}\Sigma^{-1}\mathbf{C}$
- Joint hypotheses on coefficients are *F*-distributed
- Wald and likelihood ratio test comparisons are equivalent (but need to be careful about marginality)

Model matrices

- model definition converted to **X** before we start
- input variables vs predictor variables (Schielzeth (2010), Gelman and Hill (2006), CV)
 - transformations
 - encoding of categorical variables: contrasts
 - interactions
 - basis expansions (e.g. polynomials)

Wilkinson-Rogers formulas

- Wilkinson and Rogers (1973), updated by Chambers and Hastie (1991, ch. 2)
- operators: +, *, :, /, -, ^
- I()

Contrasts

treatment contrasts

- intercept = baseline, subsequent values are differences

- contrast matrix:

$$\mathbf{C}\beta = \begin{pmatrix} 1 & 0 & 0 & \dots \\ 1 & 1 & 0 & \dots \\ 1 & 0 & 1 & \dots \\ \vdots & \vdots & \vdots & \ddots \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \dots \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \mu_3 \\ \dots \end{pmatrix} \quad .$$

Maybe easier to start from the **inverse** contrast matrix: $\beta = \mathbf{C}^{-1}\mu$.

```
C <- cbind(1, contr.treatment(3)) ## R omits the intercept by default
solve(C)

1 2 3
1 0 0
2 -1 1 0
3 -1 0 1</pre>
```

We have to specify the baseline level (contr.treatment uses first level of a factor; contr.SAS() uses the last level).

It's nice when contrasts are *orthogonal*, i.e. all rows are independent $\to \mathbf{C}^\top \mathbf{C}$ is diagonal.

Sum-to-zero contrasts

- intercept is the (unweighted!) average rather than baseline value $(\sum \mu_i/n)$
- \bullet other parameters are differences between mean of level i and intercept $(\mu_i \sum \mu_i/n)$
- last level is dropped

```
mfun <- function(C) MASS::fractions(solve(C))</pre>
  (C <- cbind(1,contr.sum(3)))
  [,1] [,2] [,3]
1
    1
      1
2
    1
         0
    1
       -1 -1
  mfun(C)
         2
[1,] 1/3 1/3 1/3
[2,] 2/3 -1/3 -1/3
[3,] -1/3 2/3 -1/3
```

Helmert contrasts

- Weird but orthogonal
- intercept, diff of first two levels, diff of level 3 from 1 & 2, ...

```
(C <- cbind(1,contr.helmert(3)))</pre>
  [,1] [,2] [,3]
1
    1
        -1
             -1
             -1
    1
         1
3
    1
            2
         0
  mfun(C)
         2
[1,] 1/3 1/3 1/3
[2,] -1/2 1/2
[3,] -1/6 -1/6 1/3
```

others

- MASS::contr.sdif() (successive-differences)
- contr.poly() (orthogonal polynomial contrasts)

practical issues

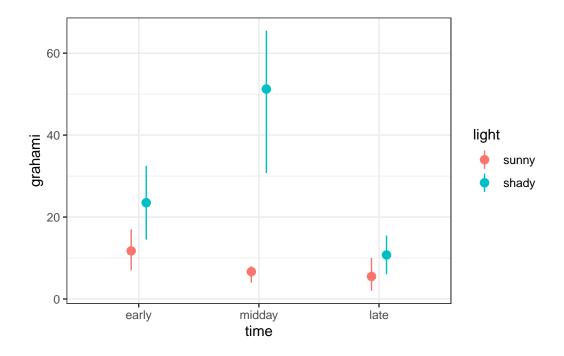
- too many ways to set contrasts (options(), contrasts(f) <-, lm(..., contrasts = list(...))
- terrible naming conventions: you can get used to it or use the faux package
- OK to fit models and later use emmeans to recover desired contrasts (switching linear bases)

```
mtcars$fcyl <- factor(mtcars$cyl)
lm0 <- lm(mpg ~ fcyl, mtcars)
cn <- function(x) names(coef(x))
cn(lm0)</pre>
```

Interactions

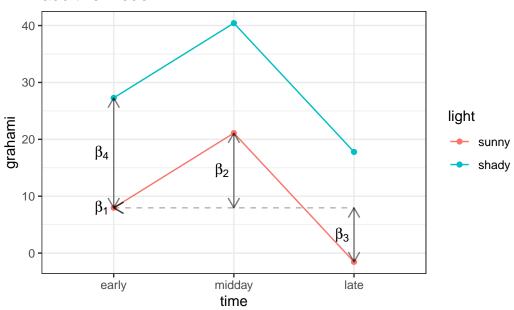
- differences in differences
- parameter values of main effects (and p values etc.) depend on contrasts/centering!
- overall model fit (R^2 , predictions, etc.) is invariant

Lizard data (Schoener 1970):



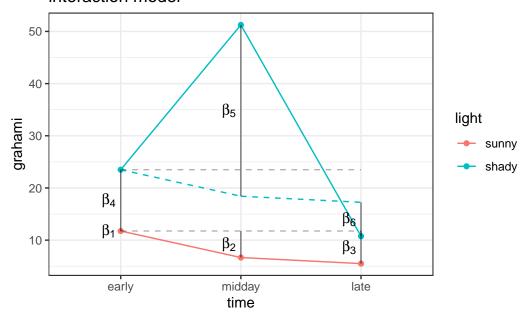
```
data("lizards", package = "brglm2")
lmTL1 <- lm(grahami~time+light,data=lizards)</pre>
pp <- with(lizards,expand.grid(time=levels(time),light=levels(light)))</pre>
pp$grahami <- predict(lmTL1,newdata=pp)</pre>
cc <- as.list(plyr::rename(coef(lmTL1),c(`(Intercept)`="int")))</pre>
labelpos <- with(cc,</pre>
  list(x=c(1,2,3,1), xend=c(1,2,3,1),
      y=c(int,int,int,int),
      yend=c(int,int+timemidday,int+timelate,int+lightshady)))
xpos <- -0.1
ggplot(pp,aes(x=time,y=grahami,colour=light))+geom_point()+
  geom_line(aes(group=light))+
  annotate("segment", x=labelpos$x, xend=labelpos$xend, y=labelpos$y,
           yend=labelpos$yend,alpha=0.5,
           arrow=arrow(length = unit(0.3, "cm"), ends="both"))+
  annotate("text", x=with(labelpos, (x+xend)/2)+xpos, y=with(labelpos, (y+yend)/2),
label=paste0("beta[",1:4,"]"),parse=TRUE)+
  annotate("segment", x=labelpos$x[1], xend=labelpos$x[3], y=labelpos$y[1],
           yend=labelpos$y[1],alpha=0.3,lty=2) +
    labs(title = "additive model")
```

additive model



```
lmTL2 <- lm(grahami~time*light,data=lizards)</pre>
gg_color_hue <- function(n) {</pre>
  hues = seq(15, 375, length=n+1)
  hcl(h=hues, l=65, c=100)[1:n]
}
pp2 <- pp
pp2$grahami <- predict(lmTL2,newdata=pp)</pre>
cc <- as.list(plyr::rename(coef(lmTL2),c(`(Intercept)`="int",</pre>
        `timemidday:lightshady`="midshady",`timelate:lightshady`="lateshady")))
labelpos <- with(cc,</pre>
  list(x=c(1,2,3,1,2,3), xend=c(1,2,3,1,2,3),
      y=c(int,int,int,int,int+lightshady+timemidday,int+lightshady+timelate),
      yend=c(int,int+timemidday,int+timelate,int+lightshady,
             int+timemidday+lightshady+midshady,int+timelate+lightshady+lateshady)))
xpos < -0.1
ggplot(pp2,aes(x=time,y=grahami,colour=light))+geom_point()+
  geom_line(aes(group=light))+
  annotate("segment", x=1:2, xend=2:3,
           y=with(cc,c(int+lightshady,int+timemidday+lightshady)),
           yend=with(cc,c(int+timemidday+lightshady,int+timelate+lightshady)),
           colour=gg_color_hue(2)[2],lty=2)+
  annotate("segment", x=labelpos$x, xend=labelpos$xend, y=labelpos$y,
```

interaction model



Marginality

- Venables (1998)
- 'type (X) sums of squares'
- scaling and centering (Schielzeth 2010) alleviates many problems; sum-to-zero contrasts (weighted or unweighted?)

Downstream methods

- prediction, effects plots
- uncertainty of predictions

- emmeans, marginaleffects, effects, sjPlot ...
- tidy(), performance, insight, etc. ...

Diagnostics

- linearity,
- base R: stats::plot.lm()
- performance::check_model()
- DHARMa (simulateResiduals(., plot = TRUE))

References

- Chambers, J. M., and T. J. Hastie, eds. 1991. *Statistical Models in S*. 1st ed. Chapman & Hall/CRC.
- Gelman, Andrew, and Jennifer Hill. 2006. *Data Analysis Using Regression and Multi-level/Hierarchical Models*. Cambridge, England: Cambridge University Press.
- Schielzeth, Holger. 2010. "Simple Means to Improve the Interpretability of Regression Coefficients: Interpretation of Regression Coefficients." *Methods in Ecology and Evolution* 1 (2): 103–13. https://doi.org/10.1111/j.2041-210X.2010.00012.x.
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- Wilkinson, G. N., and C. E. Rogers. 1973. "Symbolic Description of Factorial Models for Analysis of Variance." *Journal of the Royal Statistical Society. Series C (Applied Statistics)* 22 (3): 392–99. https://doi.org/10.2307/2346786.