Generalized linear models

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```
## it's nice to include packages at the top
## (and NOT automatically install them)
## try not to carry over packages you don't use
library(ggplot2); theme_set(theme_bw())
## diagnostics
library(performance)
library(DHARMa)
## downstream model evaluation
library(broom)
library(dotwhisker)
library(emmeans)
library(effects)
library(marginaleffects)
library(parameters)
## library(ggeffects)
```

References

Faraway (2016), McCullagh and Nelder (1989) (classic), Wood (2017, very rapid review)

Basics

- assume $\mathbf{y}_i \sim \mathrm{Dist}(g^{-1}((\mathbf{X}\beta)_i))$
- g = link function
- $\eta = X\beta = linear predictor$
- link scale or linear predictor scale vs. data or response scale
- GLMs inverse-transform η , they don't transform y
- allows:

- separate control of heteroscedasticity and nonlinearity
- almost as convenient/efficient as LMs
- equivalent to MLE in many cases
- the vast majority of GLMs are logistic (binary data) or Poisson
- lots of inference, diagnostics, etc. inherited from LM framework

Exponential family

- $f(x|\theta) = h(x)g(\theta) \exp(\eta(\theta)T(x))$
- e.g. Poisson: $f(x|\theta) = \theta^x \exp(-\theta)/x! = (1/x!) \exp(-\theta) \exp(x \log(\theta))$
- h(x) = 1/x!; $g(\theta) = \exp(-\theta)$; $\eta(\theta) = \log(\theta)$; T(x) = x
- models with T(x) = x are in the **exponential dispersion** (sub) family (Jørgensen 1987),
- $\eta(\theta)$ is the **canonical link** function for the family (nice mathematical properties, e.g. observed information = expected information [*Fisher*(*scoring*])
- binomial, Poisson, Gamma (inverse Gaussian, von Mises distribution ...)

Computation

- iteratively reweighted least squares
 - if we know μ_i , we know V_i (up to a proportion)
 - we can do weighted least squares on the link scale
 - use new β estimates to recalculate μ_i and V
- needs starting values, but almost always robust to them
- alternatives/more complex versions: glmmTMB, VGAM packages
- lots more detail on the computational issues (McCullagh and Nelder 1989; Marschner 2011; Myers et al. 2010; Dobson and Barnett 2008; Mount 2012; Gelman 2011; Robinson 2010)
- basics (from Faraway (2016) p. 155)

IRLS

Can show (Faraway p. 154) from definition of exponential (dispersion) family that the score equations are equivalent to

$$\sum_i \frac{(y_i - \mu_i)}{V(\mu_i)} \frac{\partial \mu_i}{\partial \beta_j} = 0$$

which is equivalent to minimizing

$$\sum_{i} \frac{(y_i - \mu_i)^2}{V(\mu_i)}$$

• compute adjusted dependent variate:

$$Z_{(i)} = \hat{\eta}_{(i)} + (Y - \hat{\mu}_{(i)}) \left(\frac{d\eta}{d\mu}\right)_{(i)}$$

(note: $\frac{d\eta}{d\mu} = \frac{d\eta}{dg(\eta)} = 1/g'(\eta)$: translate from raw to linear predictor scale)

• compute weights

$$W_{(i)}^{-1} = \left(\frac{d\eta}{d\mu}\right)_{(i)}^{2} V(\hat{\mu}_{(i)})$$

(translate variance from raw to linear predictor scale). This is the inverse variance of $Z_{(i)}$.

• regress $z_{(i)}$ on the covariates with weights $W_{(i)}$ to get new β estimates (\rightarrow new η , μ , $V(\mu)$...)

Tricky bits: starting values, non-convergence, etc.. (We will worry about these later!)

```
myglmfit <- function(y, X, family, tol=1e-8, maxit=50) {</pre>
    mu <- y ## set initial values
    ## set up 'oldbeta' and 'beta' so they're not identical
    oldbeta <- rep(0, ncol(X))</pre>
    beta <- rep(1, ncol(X))
    it <- 1 ## number of iterations
    while (it < maxit && max(abs((1-beta/oldbeta)))>tol) {
        oldbeta <- beta
        eta <- family$linkfun(mu) ## calc. linear predictor
        mm <- familymu.eta(eta) ## calc. d(mu)/d(eta)
        adjdev <- eta + (y-mu)/mm ## adjusted response
        W \leftarrow c(1/(mm^2*family*variance(mu))) ## weights
        beta <- lm.wfit(X, adjdev, W)$coefficients ## weighted least-squares
        mu <- family$linkinv(X %*% beta)</pre>
                                                   ## compute new mu
                                                   ## update
        it <- it+1
    }
    beta
X <- model.matrix(~wool*tension, data=warpbreaks)</pre>
y <- warpbreaks$breaks
myglmfit(y,X,poisson())
```

```
(Intercept) woolB tensionM tensionH woolB:tensionM 3.7967368 -0.4566272 -0.6186830 -0.5957987 0.6381768 woolB:tensionH 0.1883632
```

coef(glm(breaks~wool*tension, data=warpbreaks, family=poisson))

```
(Intercept) woolB tensionM tensionH woolB:tensionM 3.7967368 -0.4566272 -0.6186830 -0.5957987 0.6381768 woolB:tensionH 0.1883632
```

Mean-variance relations

• can show that we need only the link function and the **variance function** $V=f(\mu)$ for computation (may also depend multiplicatively on a **scale** or **dispersion parameter**, e.g. $V=\mu$ for Poisson, $V=\sigma^2$

Link functions

- canonical doesn't always work best (e.g. Gamma/inverse link)
- probit vs logit; nearly the same shape, interpretational difference
- cloglog; *log-hazard* scale
- inverse link: linear changes in the *rate* of events

Log-hazards and log-hazard offsets

- if hazard is h, probability is $1 \exp(-h)$
- $\bullet \ C(\mu) = \log(-\log(1-\mu))$
- $C^{-1}(\eta) = 1 \exp(-\exp(\eta))$
- $C^{-1}(\eta + \log(\Delta t)) = 1 \exp(-\exp(\eta) \cdot \Delta t)$
- $\bullet \rightarrow 1 (1 \mu_0)^{\Delta t}$

in R

- "family" functions contain all of the components needed for GLM fitting, prediction, etc.
- some of the components are weird (e.g. \$aic)
- canonical link is used by default

names(binomial())

```
[1] "family" "link" "linkfun" "linkinv" "variance"
[6] "dev.resids" "aic" "mu.eta" "initialize" "validmu"
[11] "valideta" "simulate" "dispersion"
```

Offsets

- allow for differential search effort, ratios, etc.
- typically add log(e)
- e.g. $\mathbf{y} \sim \operatorname{Poisson}(\mathbf{X}\beta + \log(A))$ is equivalent to modeling the response \mathbf{y}/A , but without messing up the mean-variance relationship

Offset/link tricks

- fit an exponential curve with constant variance: family = gaussian(link = "log")
- Ricker function $y = ax \exp(-bx)$: log-link, y ~ x + offset(log(x)
- Michaelis-Menten $y = ax/(b+x) \rightarrow 1/y = (b/a) \cdot (1/x) + 1/a$: inverse-link, y ~ I(1/x)

Model interpretation, visualization, testing

Parameter interpretation

- log scale: easy
- logit scale: \approx log for low baseline, \approx log(1-x) for high baseline, slope $\beta/4$ for intermediate values
- cloglog: log-hazard scale

Inference

- Wald tests (no finite-size corrections!)
- approximate Wald CIs (compute then back-transform)
- profile CIs

Overdispersion (diagnosis)

- too much variance
- SSQ of Pearson residuals $\sim \chi^2(n-p)$
 - but note Venables and Ripley (2002) p. 209 caution that this is approximate
 - simulation-based tests (e.g. DHARMa)

Overdispersion (solutions)

- quasi-likelihood (also handles underdispersion)
- compounded models (negative binomial, beta-binomial)
- observation-level random effects (== lognormal-Poisson)

Extended distributions

• VGAM, glmmTMB packages

Complete separation

- there is some linear combination of predictors that separates 0 from 1 responses (or 0 from non-zero responses in the case of count models)
- infinite MLE
- Hauck-Donner effect screws up Wald tests
- likelihood ratio tests still OK (sort of)
- Firth logistic regression (brglm2 package), Bayesian priors (arm::bayesglm)
- refs: Firth (1992), Heinze and Schemper (2002), Greenland and Mansournia (2015), Kosmidis and Firth (2021)

Zero-inflation/hurdle models

• finite mixture models

Most common GLM problems

- binomial/Poisson models with non-integer data
- failing to specify family (default Gaussian: → linear model); using glm() for linear models (unnecessary)
- predictions on effect scale
- using (k, N) rather than (k, N k) with family=binomial
- back-transforming SEs rather than CIs
- neglecting overdispersion
- Poisson for underdispersed responses
- equating negative binomial with binomial rather than Poisson
- worrying about overdispersion unnecessarily (binary/Gamma)
- ignoring random effects

Overdispersion in Bernoulli models?

I think many analysts read that binary models cannot be overdispersed and just do not question it. This happened with the deviance dispersion being the appropriate statistic to measure count model extra-dispersion. Some analysts simply took this on faith, so to speak. But they were mistaken.

Joseph Hilbe (2013)

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