

Integrated Nested Laplace Approximations

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Introduction

- In almost all uses of statistics,
 - 1 major interest centres on inference about characteristics of a population through observations made in a representative sample from that population.
 - Estimation
 - Hypothesis testing
 - Prediction
 - 2 Two approaches: Frequentist and Bayesian.
 - Frequentists: An orthodox view that sampling is infinite and decision rules can be sharp.
 - Bayesians: Unknown quantities are treated probabilistically and the state of the world can always be updated

Review Linear Models I

① Linear regression:

$$y \sim N(\mu, \sigma^2)$$

$$\mu = \mathbf{X}_i^T \beta + \epsilon$$

② Generalized Linear model: $y \sim \text{Exponentialfamily}$

- Generalized linear model (a unified framework): Random effects, Hierarchical models, Missing variables, Nested and Non-nested models; all handled in the same framework.
- The three components of generalized linear model:

i Distributions of response variables:

$$f(y_i; \theta_i) = \exp\{y_i b(\theta_i) + c(\theta_i) + d(y_i)\} \text{ and denoting}$$

$$\mu_i = E(\mathbf{Y}_i)$$

Review Linear Models II

- ii A linear predictor:

$$g(\mu) = \eta_i = \mathbf{X}_i^T \beta, \text{ where } \mathbf{X} = \begin{pmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_n^T \end{pmatrix} \text{ and } \beta = (\beta_1, \dots, \beta_p)$$

- iii A monotonic link function: $g(\mu_i) = \eta_i = \mathbf{X}_i^T \beta$

- ③ Generalized Additive Model: $y \sim \text{exponential family}$

$$\mu = E(y)$$

$$g(\mu) = \alpha_0 + f(x_1) + f(x_2) + \dots + f(x_p)$$

- i incorporate non-linear forms of the predictors
- ii Now the linear predictor incorporates smooth functions $f(x)$ of at least some (possibly all) covariates

Review Linear Models III

- iii maximize the quality of prediction of a dependent variable Y from various distributions, by estimating unspecific (non-parametric) functions of the predictor variables which are "connected" to the dependent variable via a link function.

4 Generalized Mixed Models:

- i Mixed models include both the usual fixed population effects and subject- or cluster-specific random effects in the linear predictor
- ii the likelihood function under a GLMM typically involves integrals with no analytic expressions, and therefore is difficult to evaluate

Latent Gaussian Models

- Core idea: Unobserved multivariate Gaussian random variable \mathbf{x} , with density $\pi(\mathbf{x}|\theta)$.
- Some of the elements in the the random vector \mathbf{x} are indirectly observed through the data \mathbf{y}
- The observed data are assumed to be conditionally independent given the latent field \mathbf{x} , ie.
$$\pi(\mathbf{x}, \theta|\mathbf{y}) \propto \pi(\theta)\pi(\mathbf{x}|\theta) \prod \pi(y_i|x_i).$$
- Main interest for inference: posterior marginals for x_i and possibly, posterior marginals of θ or some θ_i .
- A wide range of models well known from the literature can be formulated as special cases of latent Gaussian models, for example: generalised additive models, generalised additive mixed models, geoadditive models, univariate and multivariate stochastic volatility models.

Latent Gaussian models

- Hierarchical models are used when the data are structured in groups. e.g. demographically, temporally, spatially
- Latent Gaussian can be represented by a hierarchical structure containing three stages.
 - 1 The first stage is formed by conditionally independent likelihood function. That is, $\pi(\mathbf{y}|\mathbf{x}, \theta) = \prod_{i=1}^n \pi(y_i|\eta_i, \theta)$ (\mathbf{y} is vector of response variable, \mathbf{x} is latent field, θ is hyper-parameter vector, and η is linear predictor).
 - 2 Second stage is formed by the latent Gaussian distribution with mean $\mu(\theta)$ and precision matrix $\mathbf{Q}(\theta)$ to the latent field conditional on the hyper-parameter. That is, $\mathbf{x}|\theta \sim N(\mu(\theta), \mathbf{Q}(\theta)^{-1})$.
 - 3 The third stage is formed by the posterior distribution assigned to the hyper-parameters. That is $\theta \sim \pi(\theta)$.

Review Bayesian

- Bayesian inference is based on computing the posterior distribution of a vector of the model parameters \mathbf{x} conditioned on the vector of the observed data \mathbf{y}
- Posterior distribution can be written as:

$$\pi(\mathbf{x}/\mathbf{y}) = \frac{\pi(\mathbf{y}|\mathbf{x})p(\mathbf{x})}{\int \pi(\mathbf{y}|\mathbf{x})p(\mathbf{x})d\mathbf{x}} \quad (1)$$

$$\pi(\mathbf{x}/\mathbf{y}) \propto \pi(\mathbf{y}|\mathbf{x})p(\mathbf{x}) \quad (2)$$

$\pi(\mathbf{y}|\mathbf{x})$: likelihood of the model

$p(\mathbf{x})$: prior distribution of the model parameters

- Usually $\pi(\mathbf{x}/\mathbf{y})$ is highly multivariate \Rightarrow difficult to obtain. Rarely computed in closed form
- Here is where computational approaches are needed.
- Bayesian methods are becoming increasingly popular as techniques for modelling “systems” since the advent of

Review ...

Broadly speaking, there are three general steps to Bayesian data analysis:

- 1 Setting up of a full joint probability distribution for both observable, \mathbf{y} and parameters, \mathbf{x} ;

$$\pi(\mathbf{y}|\mathbf{x}) = \pi(\mathbf{y}|\mathbf{x})p(\mathbf{x})$$

- 2 Update knowledge about the unknown parameters by conditioning this probability model on observed data, $\pi(\mathbf{x}|\mathbf{y})$
- 3 Evaluate the fit of the model to the data and the sensitivity of the conclusions to the assumptions.

Review...

- Three key quantities of interest are:
 - i Prior predictive, $p(y)$: The normalizing constant $p(y)$ in Bayes Theorem is a very important quantity defined by:

$$p(y) = \int \pi(\mathbf{y}, \mathbf{x}) d\mathbf{x} = \int \pi(\mathbf{y}|\mathbf{x}) p(\mathbf{x}) d\mathbf{x}$$

It represents the probability of observing the data that was observed before it was observed.

- ii Marginal effects of a subset of parameters in a multivariate model.

Let $\mathbf{x} = (x_1, \dots, x_p)$ denote a p dimensional model. Suppose we are interested in $\pi(x_i|\mathbf{y})$, for some subset $x_i \in \mathbf{x}$, then

$$\pi(x_i|\mathbf{y}) = \int \pi(x_i, \mathbf{x}_{-i}|\mathbf{y}) d\mathbf{x}_{-i} = \int \pi(x_i|\mathbf{x}_{-i}, \mathbf{y}) \pi(\mathbf{x}_{-i}|\mathbf{y}) d\mathbf{x}_{-i}$$

where $\mathbf{x}_{-i} = \mathbf{x} \setminus x_i$ denotes the vector \mathbf{x} with x_i removed.

- iii Posterior predictions: Let $\tilde{\mathbf{y}}$ denote some future unobserved response of the system, then the posterior predictive $\pi(\tilde{\mathbf{y}}, \mathbf{y})$ is:



$(\tilde{\mathbf{y}}, \mathbf{y})$ are conditionally independent given \mathbf{x} ; but clearly, $\pi(\tilde{\mathbf{y}}, \mathbf{y})$ are dependent.

Computational methods... I

- ① Markov Chain Monte Carlo Simulation (MCMC): the most widely used method to estimate posterior distribution
 - Can effectively be applied to any model
 - MCMC can provide (nearly) exact inference, given perfect convergence and MCMC error goes to 0.
 - Their efficiency could be limited by complexity of the model,(eg hierarchical models).
 - Possible solutions:
 - More complex model specification (eg Blocking)
 - More complex sampling schemes (eg Hamiltonian Monte Carlo,No U-turn sampling
 - Alternative methods of inference (eg Approximate Bayesian Computation (ABC), INLA)
- ② Approximate Bayesian Computation (ABC)

Computational methods... II

- ③ Integrated Nested Laplace Approximation (INLA): computer performance of Bayesian inference for latent Gaussian models (LGM).

Examples of latent Gaussian models:

- Most of (generalised) linear models
- Smoothing spline models
- state space models,
- Semi-parametric regression,
- spatial and spatio-temporal models

....etc

Introduction to INLA

- Integrated nested Laplace approximation (INLA) is a computational approach to statistical inference for Latent Gaussian Markov Random field (GMRF), introduced by Rue and Martino (2007).
- It was proposed as an alternative to the usually time consuming MCMC methods.
- Issues of convergence and mixing that are inherent to MCMC are no more problems with INLA.
- Perform fast Bayesian inference in the broad class of latent Gaussian models.
- The concept of LGM is intended for the modelling stage in a unified way using algorithm and software tool.

latent Gaussian Markov random Field models

- A latent GMRF model is a hierarchical model
 - First: assume a probability model for the observations \mathbf{y} given some latent parameters \mathbf{x} and some additional parameters $\boldsymbol{\theta}$

$$\mathbf{y}|\mathbf{x}, \boldsymbol{\theta} \sim \pi(\mathbf{y}|\mathbf{x}, \boldsymbol{\theta}) = \prod_j \pi(y_j|x_j, \boldsymbol{\theta}) \quad (3)$$

- Second: assume the hyper-parameters are described by a Gaussian Markov random Field

$$\mathbf{x}|\boldsymbol{\theta} \sim \text{Normal}(\mathbf{0}, Q(\boldsymbol{\theta})) \quad (4)$$

$$\mathbf{x}_l \perp\!\!\!\perp \mathbf{x}_m | \mathbf{x}_{-lm} \quad (5)$$

- In general, we can partition $\boldsymbol{\theta} = (\boldsymbol{\theta}_1, \boldsymbol{\theta}_2)$, where $\boldsymbol{\theta}_1$ are the hyper parameter and $\boldsymbol{\theta}_2$ are the nuisance parameter and re-express the model as follows:

| | |
|---|-------------|
| $\boldsymbol{\theta} \sim p(\boldsymbol{\theta})$ | hyper prior |
| $\mathbf{x} \boldsymbol{\theta} \sim p(\mathbf{x} \boldsymbol{\theta}) = \text{Normal}(\mathbf{0}, \boldsymbol{\Sigma}(\boldsymbol{\theta}_1))$ | GMRF prior |
| $\mathbf{y} \mathbf{x}, \boldsymbol{\theta} \sim \prod_j \pi(y_j x_j, \boldsymbol{\theta}_2)$ | data model |

LGMRF model as a general framework

- $\mathbf{y} = \{y_i\}_{i=1}^n$ is assumed to belong to an exponential family
- A very general way of formulation this problem is by modelling the mean μ_i for the i-th unit by means of a structured additive linear predictor η_i through a link function $g(\cdot)$, $g(\mu_i) = \eta_i$

$$\eta_i = \alpha + \sum_{j=1}^{n_f} f^{(j)}(u_{ji}) + \sum_{k=1}^{n_\beta} \beta_k z_{ki} + \epsilon_i$$

where,

- α is the intercept
- $f^{(j)}$ is a set of unknown functions of the covariates \mathbf{u}
- β_k 's represent the linear effect of variates \mathbf{z}
- ϵ_i 's are unstructured terms
- Denote the vector of all **latent** Gaussian variables to be $\mathbf{x} = (\{\eta_i\}, \alpha, \{f^{(j)}\}, \{\beta_k\}) \sim \text{Gaussian}(\mathbf{0}, Q(\theta_1))$
 - where θ_1 is vector of hyper-parameters.

Example: Dynamic models/ State space models

Dynamic models/ State space models

$$y_t = F_t' x_t + v_t \quad (6)$$

$$x_t = G_t' x_{t-1} + w_t \quad (7)$$

$$(8)$$

where

$$v_t \sim N(0, V_t) \quad (9)$$

$$w_t \sim N(0, W_t) \quad (10)$$

$$(11)$$

y_t is a time sequence of scalar observations and x_t is a sequence of state(latent) parameters describing locally the system. F_t is a vector of explanatory variables, while G_t represents a matrix describing the states evolution.

Aim

- The INLA approach provides a fast way to do Bayesian inference using accurate approximations to $\pi(x_i|\mathbf{y})$ and $\pi(\theta_j|\mathbf{y})$ for $\forall i$.

$$\pi(\theta_j|\mathbf{y}) = \int \pi(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta}_{-j} \quad (12)$$

$$\pi(x_i|\mathbf{y}) = \int \pi(x_i|\boldsymbol{\theta}, \mathbf{y}) \pi(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \quad (13)$$

- The key feature is to use this form to construct **nested** approximations

$$\tilde{\pi}(\theta_j|\mathbf{y}) = \int \tilde{\pi}(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta}_{-j} \quad (14)$$

$$\tilde{\pi}(x_i|\mathbf{y}) = \int \tilde{\pi}(x_i|\boldsymbol{\theta}, \mathbf{y}) \tilde{\pi}(\boldsymbol{\theta}|\mathbf{y}) d\boldsymbol{\theta} \quad (15)$$

where $\tilde{\pi}(.|.)$ is an approximated conditional density of it

Steps in the INLA project

- 1 Find a Laplace approximation to $\pi(\boldsymbol{\theta}|\mathbf{y})$. [detail](#) [▶ LA](#)
- 2 Find an approximation to $\pi(x_i|\boldsymbol{\theta}, \mathbf{y})$ [detail](#)
 - Gaussian approximation
 - fast but inaccurate
 - Laplace approximation
 - fast but computationally demanding
 - Simplified Laplace approximation
 - default in R-INLA
 - trade-off between speed and accuracy
- 3 Numerical integration [detail](#)

$$\tilde{\pi}_i(x_i|\mathbf{y}) = \sum_k \tilde{\pi}_i(x_i|\boldsymbol{\theta}_k, \mathbf{y}) \tilde{\pi}_i(\boldsymbol{\theta}_k|\mathbf{y}) \Delta_k \quad (16)$$

where the sum is over values of $\boldsymbol{\theta}$ with area-weights Δ_k .

- Grid strategy
- Central composite design (CCD)

Step 1: Find a Laplace approximation to $\pi(\boldsymbol{\theta}|\mathbf{y})$

- $\pi(\boldsymbol{\theta}|\mathbf{y})$ can be easily obtained by

$$\pi(\boldsymbol{\theta}|\mathbf{y}) = \frac{\pi(\mathbf{x}, \boldsymbol{\theta}|\mathbf{y})}{\pi(\mathbf{x}|\boldsymbol{\theta}, \mathbf{y})} \quad (17)$$

$$= \frac{\pi(\mathbf{x}, \boldsymbol{\theta}, \mathbf{y})}{\pi(\mathbf{y})} \frac{1}{\pi(\mathbf{x}|\boldsymbol{\theta}, \mathbf{y})} \quad (18)$$

$$= \frac{\pi(\boldsymbol{\theta})\pi(\mathbf{x}|\boldsymbol{\theta})\pi(\mathbf{y}|\mathbf{x})}{\pi(\mathbf{y})} \frac{1}{\pi(\mathbf{x}|\boldsymbol{\theta}, \mathbf{y})} \quad (19)$$

$$\propto \frac{\pi(\boldsymbol{\theta})\pi(\mathbf{x}|\boldsymbol{\theta})\pi(\mathbf{y}|\mathbf{x})}{\pi(\mathbf{x}|\boldsymbol{\theta}, \mathbf{y})} \quad (20)$$

$$\approx \frac{\pi(\mathbf{x}, \boldsymbol{\theta}, \mathbf{y})}{\tilde{\pi}_G(\mathbf{x}|\boldsymbol{\theta}, \mathbf{y})} \Big|_{\mathbf{x}=\mathbf{x}^*(\boldsymbol{\theta})} \doteq \tilde{\pi}(\boldsymbol{\theta}|\mathbf{y}) \quad (21)$$

where

- $\mathbf{x} = \mathbf{x}^*(\boldsymbol{\theta})$ is the mode of the full conditional for \mathbf{x} , for a given $\boldsymbol{\theta}$

Step 1: Find a Laplace approximation to $\pi(\boldsymbol{\theta}|\mathbf{y})$

- ① find the mode $\tilde{\boldsymbol{\theta}}$ by optimising $\log\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ with respect to $\boldsymbol{\theta}$
 - Newton-like algorithm
- ② Compute the Hessian at the modal configuration $\boldsymbol{\theta}^*$
- ③ Explore $\log\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ with respect to $\boldsymbol{\theta}$ using \mathbf{z} -parametrisation
 - Define $\boldsymbol{\theta}$ through \mathbf{z}

$$\boldsymbol{\theta}(\mathbf{z}) = \boldsymbol{\theta}^* + \mathbf{V}\boldsymbol{\Lambda}^{1/2}\mathbf{z} \quad (22)$$

If $\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ is a Gaussian density, then \mathbf{z} is $N(\mathbf{0}, \mathbf{I})$. And $\boldsymbol{\Sigma} = \mathbf{V}\boldsymbol{\Lambda}\mathbf{V}'$ be the eigen-decomposition of $\boldsymbol{\Sigma}$.

- produce a grid of H points $\{\boldsymbol{\theta}_k^*\}$ associate with mass and area weights $\{\triangle_k\}$

Step 1: Find a Laplace approximation to $\pi(\boldsymbol{\theta}|\mathbf{y})$

The procedure when $\log\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ is unimodal

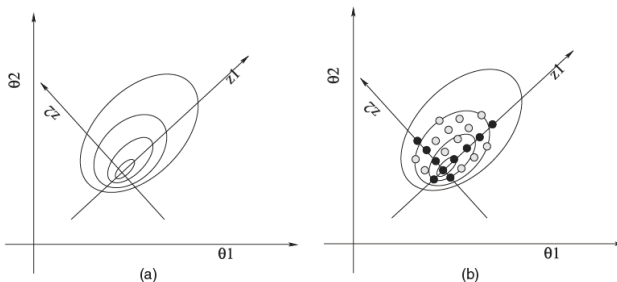


Fig. 1. Illustration of the exploration of the posterior marginal for θ : in (a) the mode is located and the Hessian and the co-ordinate system for \mathbf{z} are computed; in (b) each co-ordinate direction is explored (●) until the log-density drops below a certain limit; finally the new points (●) are explored

- This re-parametrisation corrects for scale and rotation.
- This re-parametrisation simplifies numerical integration.

Step 2: Find an approximation to $\pi(x_i|\theta, \mathbf{y})$

The marginals for components x_i of the latent field

$$\pi(x_i|\mathbf{y}) = \int \pi(x_i|\theta, \mathbf{y})\pi(\theta|\mathbf{y})d\theta \quad (23)$$

- this step is a bit complex than step 1 because in general \mathbf{x} contain more elements than θ .

Gaussian Approximation

- directly using normal distribution

$$\pi(\mathbf{x}) = (2\pi)^{-\frac{n}{2}} |Q|^{-\frac{1}{2}} e^{-\frac{1}{2}(\mathbf{x}-\mu)'Q(\mathbf{x}-\mu)} \quad (24)$$

- based on the Cholesky decomposition of the precision matrix $Q(.) = LL'$
- fast but not very good due to the lack of skewness (Rue and Martino, 2007)

Step 2: Find an approximation to $\pi(x_i|\theta, \mathbf{y})$

Laplace Approximation

$$\pi_{LA}(x_i|\theta, \mathbf{y}) = \frac{\pi(\{x_j, \mathbf{x}_{-j}\}|\theta, \mathbf{y})}{\pi(\mathbf{x}_{-j}|x_j, \theta, \mathbf{y})} \quad (25)$$

$$= \frac{\pi(\{x_j, \mathbf{x}_{-j}\}, \theta|\mathbf{y})}{\pi(\theta|\mathbf{y})} \frac{1}{\pi(\mathbf{x}_{-j}|x_j, \theta, \mathbf{y})} \quad (26)$$

$$\propto \frac{\pi(\mathbf{x}, \theta|\mathbf{y})}{\pi(\mathbf{x}_{-j}|x_j, \theta, \mathbf{y})} \propto \frac{\pi(\theta)\pi(\mathbf{x}|\theta)\pi(\mathbf{y}|\mathbf{x})}{\pi(\mathbf{x}_{-j}, x_j, \theta, \mathbf{y})} \quad (27)$$

$$\propto \frac{\pi(\mathbf{x}, \theta, \mathbf{y})}{\tilde{\pi}_{GG}(\mathbf{x}_{-j}|x_i, \theta, \mathbf{y})} \Big|_{\mathbf{x}_{-j}=\mathbf{x}_{-j}^*(x_i, \theta)} \quad (28)$$

$$\doteq \tilde{\pi}_{LA}(x_j|\theta, \mathbf{y}) \quad (29)$$

where $\tilde{\pi}_{GG}$ is the Gaussian approximation to $\mathbf{x}_{-j}|x_i, \theta, \mathbf{y}$; and $(\mathbf{x}_{-j}^*(x_j, \theta))$ is the modal configuration. Note that

- very accurate as $\mathbf{x}_{-j}|x_i, \theta, \mathbf{y}$ is reasonably normal
- computational expensive

Step 2: Find an approximation to $\pi(x_i|\theta, \mathbf{y})$

Simplified Laplace Approximation.

- based on a Taylor's series expansion up to the third order of both numerator and denominator for $\tilde{\pi}(x_i|\theta, \mathbf{y})$. i.e. calculate

$$\log \tilde{\pi}_{SLA}(x_i^s|\theta, \mathbf{y}) = \text{const.} - \frac{1}{2}(x_i^s)^2 + \gamma_i^1(\theta)x_i^s + \frac{1}{6}(x_i^s)^3\gamma_i^3(\theta) + \dots \quad (30)$$

And fit a skew normal density

- this effectively corrects the Gaussian approximation for location and skewness to increase the fit to the required distributions
- implemented by default by R-INLA

back

Step 3: Numerical integration

$$\tilde{\pi}_i(x_i|\mathbf{y}) = \sum_k \tilde{\pi}_i(x_i|\boldsymbol{\theta}_k, \mathbf{y}) \tilde{\pi}_i(\boldsymbol{\theta}_k|\mathbf{y}) \Delta_k \quad (31)$$

where the sum is over values of $\boldsymbol{\theta}$ with area-weights Δ_k .

Integration with respect to $\boldsymbol{\theta}$

1 Grid strategy

- Find mode / Compute Hessian / Grid search
- Note: accurate but may be time consuming if $m=\dim(\boldsymbol{\theta})$ is large

2 Central composite design

- "Use small amount of support points in the m-dimensional space of $\boldsymbol{\theta}$ "
- "Augment each center point with a group of points used to estimate the curvature of $\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ "
- implemented by default by R-INLA

INLA algorithm (operationally)

- ① marginal joint posterior for the hyper-parameters $\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$
 - find the mode $\tilde{\boldsymbol{\theta}}$ by optimising $\log\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ with respect to $\boldsymbol{\theta}$
 - Newton-like algorithm
 - Compute the Hessian at $\boldsymbol{\theta}^*$
 - Explore $\log\tilde{\pi}(\boldsymbol{\theta}|\mathbf{y})$ with respect to $\boldsymbol{\theta}$ using **z**-parametrisation
 - produce a grid of H points $\{\theta_k^*\}$ associate with mass and area weights $\{\Delta_k\}$
- ② For each element $\{\theta_k^*\}$ in the grid
 - find the marginal posterior $\tilde{\pi}(\theta_k^*|\mathbf{y})$
 - evaluate the conditional posterior $\tilde{\pi}(x_j|\theta_k^*, \mathbf{y})$
- ③ obtain the marginal posterior $\tilde{\pi}_i(x_i|\mathbf{y})$ using numerical integration

$$\tilde{\pi}_i(x_i|\mathbf{y}) = \sum_k \tilde{\pi}_i(x_i|\boldsymbol{\theta}_k, \mathbf{y}) \tilde{\pi}_i(\boldsymbol{\theta}_k|\mathbf{y}) \Delta_k \quad (32)$$

where the sum is over values of $\boldsymbol{\theta}$ with area-weights Δ_k .

Integrated Nested Laplace Approximation (INLA)

Integrated Nested Laplace Approximation

- Laplace is the fundamental tool to estimate the unknown distributions
- Laplace approximation are nested within one another
- Numerical integration is used to obtain marginal posterior distributions

- Proposed formulated framework to fit common types of DM using INLA through simulated data sets
- Firstly, consider simple univariate dynamic linear model
- Next, state-space models that can be directly fitted using model options in the INLA library
- Models that could not be fitted using INLA's standard tools

A Toy Example (Ruiz et al (2010))

- Simple Simulated example of a first order univariate dynamic linear model
- The observational and the system equations respectively of the model is given as:

$$y_t = x_t + v_t, \quad v_t \sim N(0, V), \quad t = 1, \dots, n \quad (33)$$

$$x_t = x_{t-1} + w_t, \quad w_t \sim N(0, W), \quad t = 2, \dots, n \quad (34)$$

assuming $F_t = G_t = 1$, $V_t = V$ and $W_t = W$, for all t
vector of hyperparameters is given by $\theta = \{V, W\}$
latent field corresponds to $\xi = \{x_1, \dots, x_n\}$

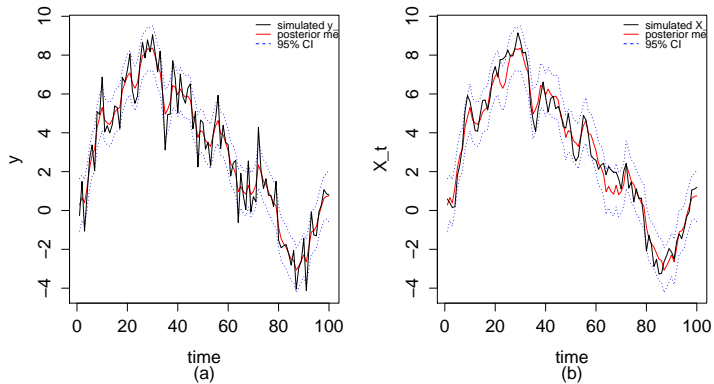


Figure : Simulated and predicted values (posterior mean and 95% credibility interval) for the observations (a) and states (b) in the toy example

Example (Ruiz et al (2010))

- Simulated data from a multiple Poisson model with two regressors, Z_{1t} and Z_{2t}
- The model has the following observational and System equations

$$(y_t | \mu_t) \sim \text{Poisson}(\mu_t) \quad (35)$$

$$\log(\mu_t) = \lambda_t = \beta_{0t} + \beta_{1t}Z_1 + \beta_{2t}Z_2, \quad t = 1, \dots, n$$

$$\beta_{0t} = \beta_{0,t-1} + \omega_{0t}, \quad \omega_{0t} \sim N(0, W_0), \quad t = 2, \dots, n$$

$$\beta_{1t} = \beta_{1,t-1} + \omega_{1t}, \quad \omega_{1t} \sim N(0, W_1), \quad t = 2, \dots, n$$

$$\beta_{2t} = \beta_{2,t-1} + \omega_{2t}, \quad \omega_{2t} \sim N(0, W_2), \quad t = 2, \dots, n$$

The linear predictor is given by $\lambda_t = F_t x_t$, where $F_t = (1, Z_{1t}, Z_{2t})$ and the regression coefficients $x_t = (\beta_{0t}, \beta_{1t}, \beta_{2t})$

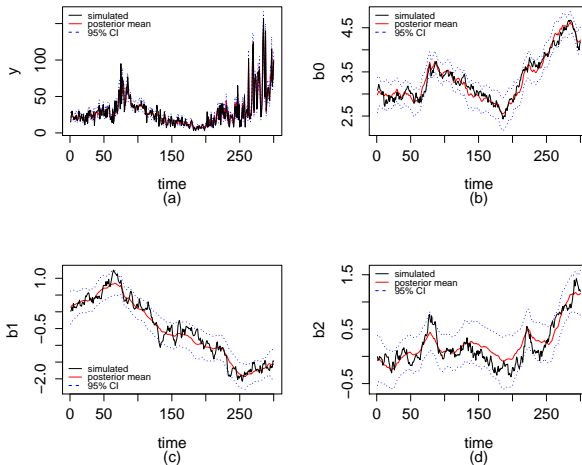


Figure : Simulated and predicted values (posterior mean and 95% credibility interval) for the observations (a) and regression coefficients, β_0 , β_1 and β_2 (b-d) in the generalized dynamic regression example

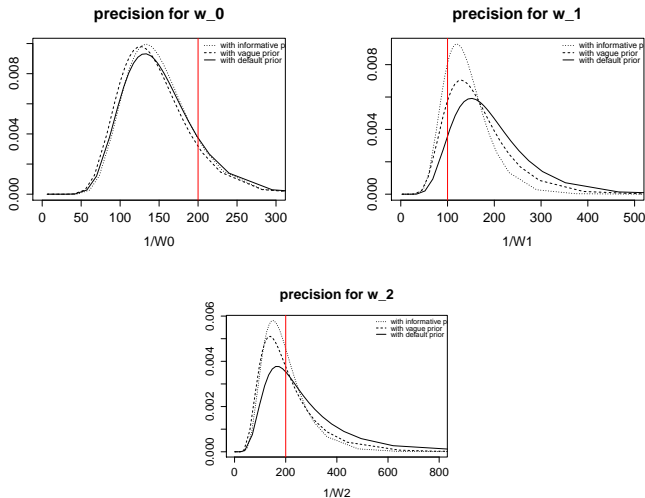


Figure : Posterior densities for the hyperparameters in the generalized dynamic regression example. Red lines indicate true simulated values

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Thank you!