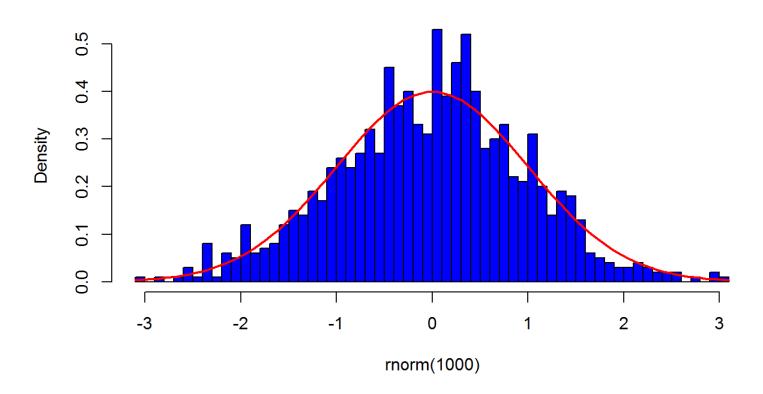
Introduction to Statistical Tests in R

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Outline

- Normal Distribution & Normality Test
- Welch's Two-sample T-test
- · Wilcoxon Rank Sum Test
- · Pearson Correlation
- Multivariate Regression
- Negative Binomial & DESeq
- Principal Component Analysis

Histogram of N(0,1), n=1000



```
set.seed(234)
hist(rnorm(1000), breaks=80, prob=TRUE, col="blue", main="Histogram of N(0,1), n=1000")
curve(dnorm(x, mean=0, sd=1), col="red", lwd=2, add=TRUE, yaxt="n")
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```

Normal distribution: a continuous distribution

$$X \sim N(\mu, \sigma^2)$$

Density function:

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

- Fully specified by mean μ and variance σ^2
- $e \approx 2.718$; $\pi \approx 3.142$.
- Bell shaped
- Symmetric around μ (mean/median/mode)
- Histogram proportion limit, $n \to +\infty$

Empirical Rule:

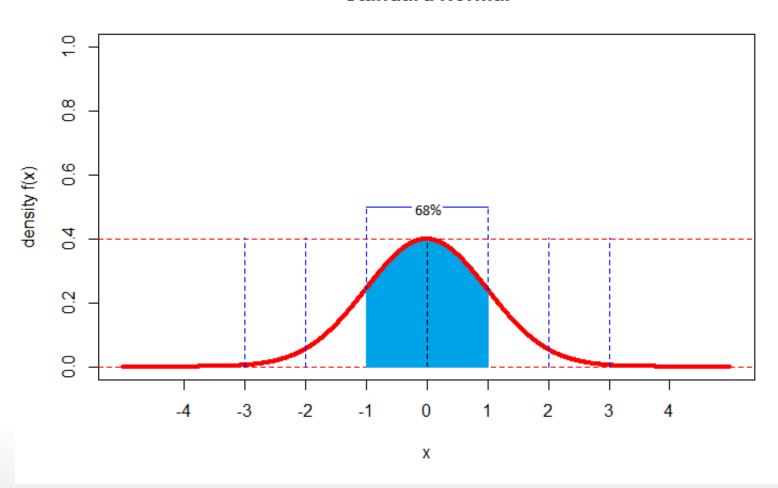
Given *any* $N(\mu, \sigma^2)$:

- 68% data $\in [\mu \sigma, \mu + \sigma]$
- 95% data $\in [\mu 2\sigma, \mu + 2\sigma]$
- 99.7% data $\in [\mu 3\sigma, \mu + 3\sigma]$

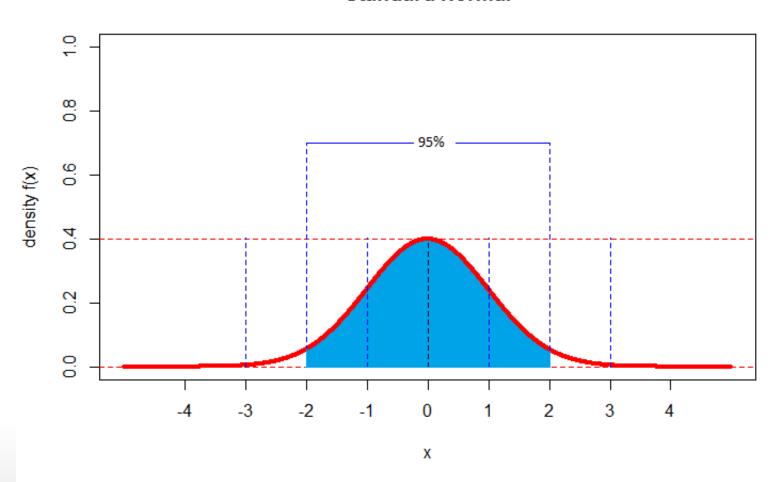
Standard Normal N(0,1): $\mu = 0$, $\sigma^2 = 1$

- 68% data $\in [-1, 1]$
- 95% data \in [-2, 2]
- 99.7% data $\in [-3, 3]$

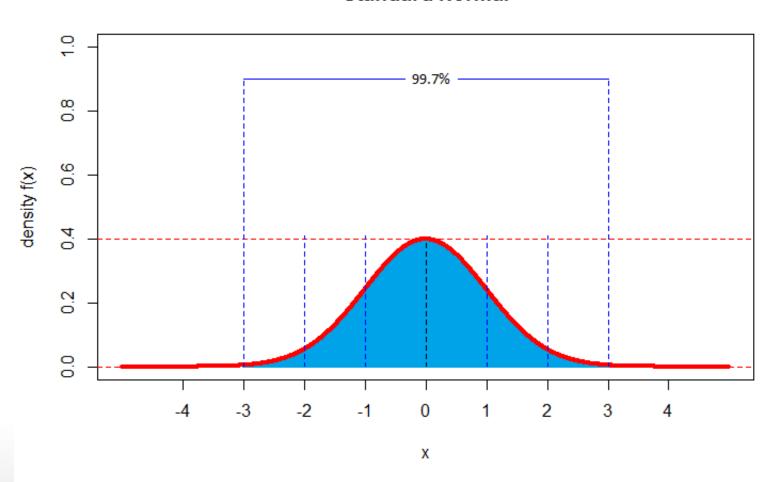
Standard Normal



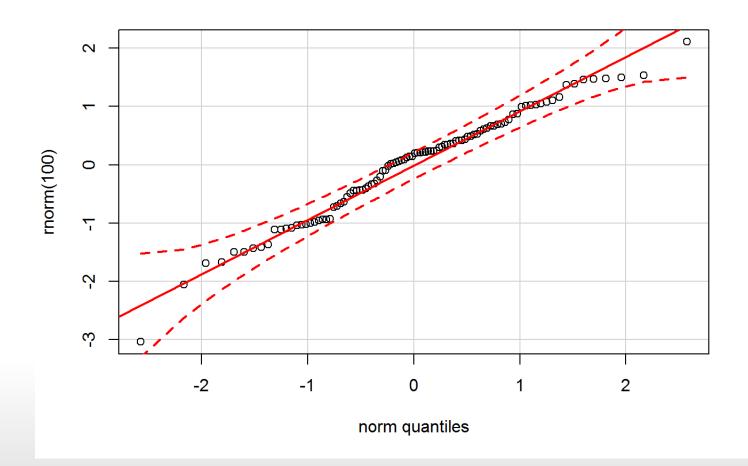
Standard Normal



Standard Normal



set.seed(234); car::qqPlot(rnorm(100)) # QQ-plot with 95% CI



```
shapiro.test(rnorm(100)) # Shapiro-Wilk test of normality

## The calculation of the p value is exact for n = 3, otherwise
## approximations are used, separately for 3 < n <12 and n >11

## Missing values are allowed, but the number of non-missing values
## must be between 3 and 5000.
```

```
##
## Shapiro-Wilk normality test
##
## data: rnorm(100)
## W = 0.98084, p-value = 0.1544
```

Welch's Two-sample T-test

$$H_{0}: \mu_{x} = \mu_{y} \quad vs. \quad H_{1}: \mu_{x} \neq \mu_{y}$$

$$T = \frac{\bar{x} - \bar{y}}{\sqrt{\frac{s_{x}^{2}}{n_{x}} + \frac{s_{y}^{2}}{n_{y}}}} \sim t_{d.f.}$$

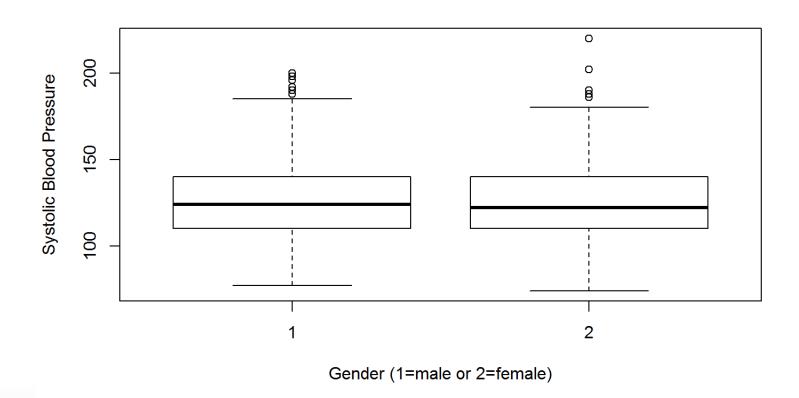
$$d.f. = \left(\frac{s_{x}^{2}}{n_{x}} + \frac{s_{y}^{2}}{n_{y}}\right)^{2} / \left[\frac{\left(\frac{s_{x}^{2}}{n_{x}}\right)^{2} + \left(\frac{s_{y}^{2}}{n_{y}}\right)^{2}}{n_{x} - 1} + \frac{\left(\frac{s_{y}^{2}}{n_{y}}\right)^{2}}{n_{y} - 1}}\right]$$

Assumption:

- Sample x and y are independent.
- For small samples, both x and y population need to be normal.

Welch's Two-sample T-test

· Example Data:



Welch's Two-sample T-test

· Two-sided test in R:

```
t.test(SYSBP~SEX)
```

```
##
## Welch Two Sample t-test
##
## data: SYSBP by SEX
## t = 1.6755, df = 2506.5, p-value = 0.09395
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.1634916 2.0833293
## sample estimates:
## mean in group 1 mean in group 2
## 126.0106 125.0507
```

One-sided Two-sample T-test

t.test(SYSBP[SEX==1], SYSBP[SEX==2], "greater")

```
H_0: \mu_M = \mu_F \quad vs. \quad H_1: \mu_M > \mu_F
```

 H_0 : two populations follow the same distribution

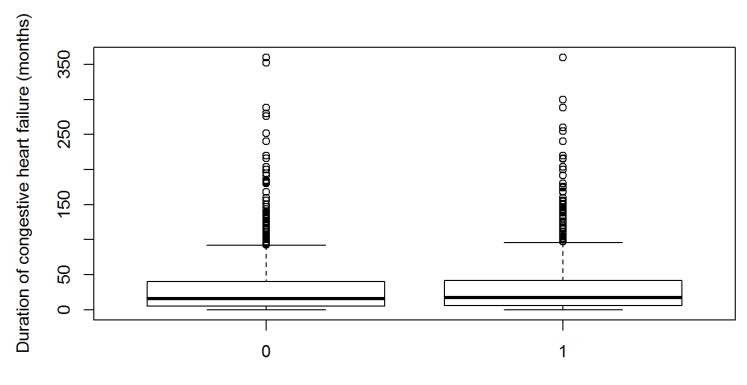
 H_1 : two populations follow different distributions

T =Sum of ranks of sample x in the combined data

Assumption:

- Sample x and y are independent.
- The population follows a continuous distribution.

· Example Data:



Treatment (0=Placebo, 1=Treatment)

· Two-sided test in R:

```
wilcox.test(CHFDUR~TRTMT)
```

```
##
## Wilcoxon rank sum test with continuity correction
##
## data: CHFDUR by TRTMT
## W = 5665900, p-value = 0.2627
## alternative hypothesis: true location shift is not equal to 0
```

· One-sided test in R:

```
wilcox.test(CHFDUR[TRTMT==0], CHFDUR[TRTMT==1], "less")

##

## Wilcoxon rank sum test with continuity correction

##

## data: CHFDUR[TRTMT == 0] and CHFDUR[TRTMT == 1]

## W = 5665900, p-value = 0.1314

## alternative hypothesis: true location shift is less than 0
```

· Pearson Correlation Coefficient:

$$r = \hat{\rho} = \frac{\sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^{n} (x_i - \bar{x})^2} \sqrt{\sum_{i=1}^{n} (y_i - \bar{y})^2}}$$

- $r \in [-1, 1]$
- r = 0: no linear association; does not imply independence
- r = 1: fit perfectly by an increasing line
- r = -1: fit perfectly by a decreasing line

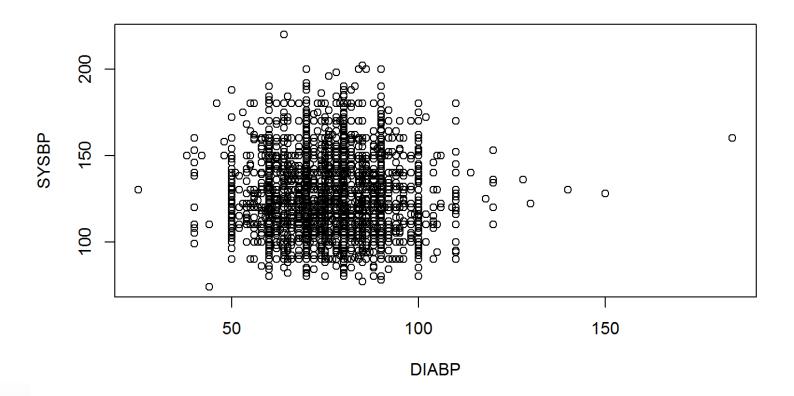
• Hypothesis Test of $\rho = 0$:

Assumption: both sample drawn from normal distributions

$$T = r \sqrt{\frac{n-2}{1-r^2}} \sim t_{n-2}$$

- Holds approximately for non-normal distributions if sample sizes are large.
- Asymptotic confidence interval by Fisher's Ztransformation.

• Example Data:



 $H_0: \rho = 0 \quad vs. \quad H_1: \rho \neq 0$

```
cor.test(DIABP, SYSBP)
```

```
##
## Pearson's product-moment correlation
##
## data: DIABP and SYSBP
## t = 0.83224, df = 6790, p-value = 0.4053
## alternative hypothesis: true correlation is not equal to 0
## 95 percent confidence interval:
## -0.01368677 0.03387401
## sample estimates:
## cor
## 0.01009933
```

 $H_0: \rho = 0$ vs. $H_1: \rho > 0$

cor.test(DIABP, SYSBP, "greater")

```
##
## Pearson's product-moment correlation
##
## data: DIABP and SYSBP
## t = 0.83224, df = 6790, p-value = 0.2027
## alternative hypothesis: true correlation is greater than 0
## 95 percent confidence interval:
## -0.00986294 1.00000000
## sample estimates:
## cor
## 0.01009933
```

Model Setup

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \dots + \beta_p x_p + \epsilon$$
$$\epsilon \sim N(0, \sigma^2), \quad i.i.d.$$

Equivalently

$$y = X\beta + \epsilon$$

$$\epsilon \sim MVN(0, \sigma^2 I)$$

or

$$y|X \sim MVN(X\beta, \sigma^2 I)$$

Coefficient Test:

$$T = \frac{\hat{\beta}}{\sqrt{Var(\hat{\beta})}} \sim t_{n-p} \quad or \quad N(0,1)$$

$$\widehat{\beta} = (X'X)^{-1}X'y$$

$$Var(\widehat{\boldsymbol{\beta}}) = (X'X)^{-1}X' \ Var(\boldsymbol{y}|\boldsymbol{X}) \ X(X'X)^{-1}$$
$$= (X'X)^{-1}X' \ (\sigma^2\boldsymbol{I}) \ X(X'X)^{-1}$$
$$= \sigma^2(X'X)^{-1}$$

Overall Significance Test

 H_0 : The fit of the full model is the same as the intercept-only model

 H_1 : The full model fits the data better than the intercept-only model

Equivalently,

$$H_0$$
: $\beta_1 = \beta_2 = \dots = \beta_p = 0$ vs. H_1 : $\beta_j \neq 0$ for some j

$$F = \frac{\sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2 / (p)}{\sum_{i=1}^{n} (y_i - \hat{y}_i)^2 / (n - p - 1)} \sim F_{(p, n - p - 1)}$$

Goodness of fit:

Coefficient of determination

$$R^{2} = 1 - \frac{\sum_{i=1}^{n} (y_{i} - \hat{y}_{i})^{2}}{\sum_{i}^{n} (y_{i} - \bar{y})^{2}}$$

Proportion of the response variance that is explained by the regression, although not bounded between 0 and 1.

Adjusted R squared

$$R_{adj}^2 = 1 - \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2 / (n - p - 1)}{\sum_{i=1}^n (y_i - \bar{y})^2 / (n - 1)} < R^2$$

 R_{adj}^2 adjusts for number of X's relative to sample size.

· Example Data: Response

```
## Min. 1st Qu. Median Mean 3rd Qu. Max. NA's ## 0.000 4.100 4.300 4.397 4.600 434.000 801

summary(KLEVEL[- which(KLEVEL==434)])

## Min. 1st Qu. Median Mean 3rd Qu. Max. NA's ## 0.000 4.100 4.300 4.326 4.600 6.300 801

length(KLEVEL)

## [1] 6800
```

· Example Data: Response

[1] 6799

```
## The following objects are masked from DIGdata (pos = 6):

##
## ACEINHIB, ACTLIMIT, AGE, ANGINA, BMI, CHESTX, CHFDUR,
CHFETIOL, CREAT, CREV, CREVDAYS, CVD, CVDDAYS, DEATH,
DEATHDAY, DIABETES, DIABP, DIG, DIGDAYS, DIGDOSE, DIGDOSER,
DIGUSE, DIURET, DIURETK, DWHF, DWHFDAYS, EJFMETH, EJFPER,
## ELEVJVP, EXERTDYS, FUNCTCLS, HEARTRIE, HOSP, HOSPDAYS, HYDRAL,
HYPERTEN, ID, KLEVEL, KSUPP, MI, MIDAYS, NHOSP, NITRATES,
## NSYM, OCVD, OCVDDAYS, OTH, OTHDAYS, PEDEMA, PREVMI, PULCONG,
RACE, RALES, REASON, RESTDYS, RINF, RINFDAYS, S3, SEX, STRK,
## STRKDAYS, SVA, SVADAYS, SYSBP, TRTMT, UANG, UANGDAYS, VASOD,
VENA, VENADAYS, WHF, WHFDAYS

length(KLEVEL)
```

· Example Data: Regressors

```
table(DIG) # Digoxin Toxicity (0=No, 1=Yes)
## DIG
## 6701 98
table(DIABETES) # History of Diabetes (0=No or Unkonw, 1=Yes)
## DIABETES
## 0 1
## 4866 1933
table(SEX) # 1=Male, 2=Female
## SEX
## 1 2
## 5280 1519
```

Example Data: fit regression with only main effects

```
summary(lm(KLEVEL ~ DIG + DIABETES + SEX))
## Call:
## lm(formula = KLEVEL ~ DIG + DIABETES + SEX)
## Residuals:
                10 Median
      Min
                                       Max
  -4.3445 - 0.2700 - 0.0113 0.2887 1.9887
## Coefficients:
## Estimate Sto
## (Intercept) 4.27755
              Estimate Std. Error t value Pr(>|t|
                          0.02066 207.029 < 2e-16
## DIG
               -0.18386
                           0.05463 -3.365 0.000769 ***
## DIABETES
                           0.01446 2.292 0.021948 *
              0.03315
## SEX
                           0.01566
               0.03380
                                     2.158 0.030978 *
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.5057 on 5994 degrees of freedom
  (801 observations deleted due to missingness)
## Multiple R-squared: 0.003505, Adjusted R-squared: 0.003007
## F-statistic: 7.029 on 3 and 5994 DF, p-value: 0.0001028
```

Example Data: fit regression with interaction effects

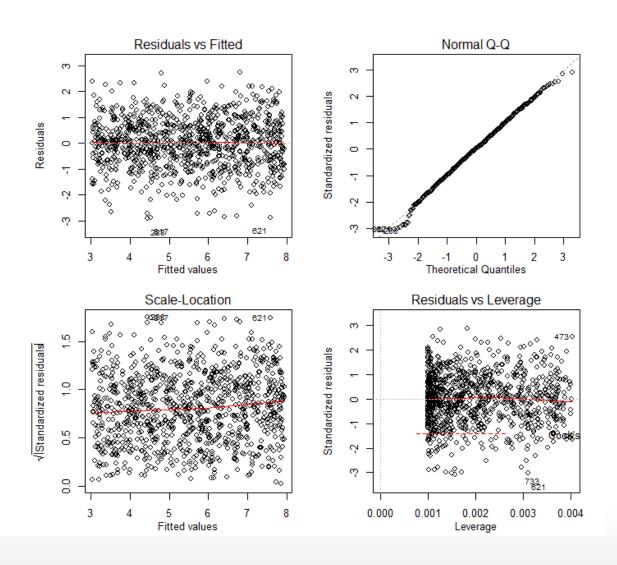
```
summary(lm(KLEVEL ~.^3, data=data.frame(DIG, DIABETES, SEX)))
## Call:
## lm(formula = KLEVEL ~ .^3, data = data.frame(DIG, DIABETES, SEX))
## Residuals:
               10 Median
      Min
## -4.3502 -0.2595 -0.0100 0.2900 1.9900
## Coefficients:
##
## (Intercept)
                   Estimate Std. Error t value Pr(>|t|
                               0.02417 176.685
                   4.27056
                                                 <2e-16
## DIG
                   -0.39556
                               0.19004 - 2.082
                                                 0.0374 *
## DIABETES
                               0.04516 1.559
                   0.07042
                                                 0.1190
## SEX
                               0.01872
                   0.03945
                                       2.107
                                                 0.0352 *
## DIG:DIABETES
                 -0.04542
                               0.37968 - 0.120
                                                 0.9048
## DIG:SEX
                   0.16555
                               0.13721 1.207
                                                 0.2277
               -0.03021
## DIABETES:SEX
                               0.03492 - 0.865
                                                 0.3870
## DIG:DIABETES:SEX 0.02521
                               0.27985
                                       0.090
                                                 0.9282
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.5058 on 5990 degrees of freedom
  (801 observations deleted due to missingness)
## Multiple R-squared: 0.003983, Adjusted R-squared: 0.002819
## F-statistic: 3.422 on 7 and 5990 DF, p-value: 0.00118
```

Diagnostic Plots:

- 1. Check linearity and homoscedasticity:
 - * Residuals vs. Fitted Plot
 - * Scale-Location Plot
- 2. Check residual normality:
 - * Q-Q Plot
- 3. Identify influential data points:
 - * Residuals vs. Leverage Plot

· Simulation with uniform regressor in linear form:

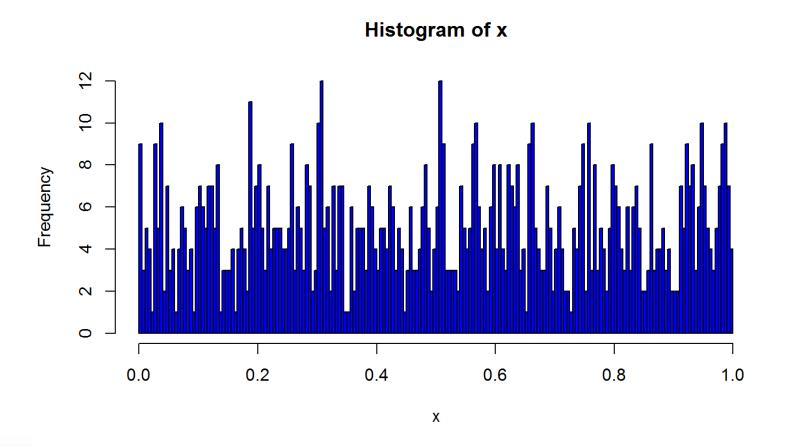
```
set.seed(1234)
x <- runif(1000)
y <- 3+5*x+rnorm(1000)
par(mfrow=c(2,2)); plot(lm(y~x));</pre>
```



· Fitting summary:

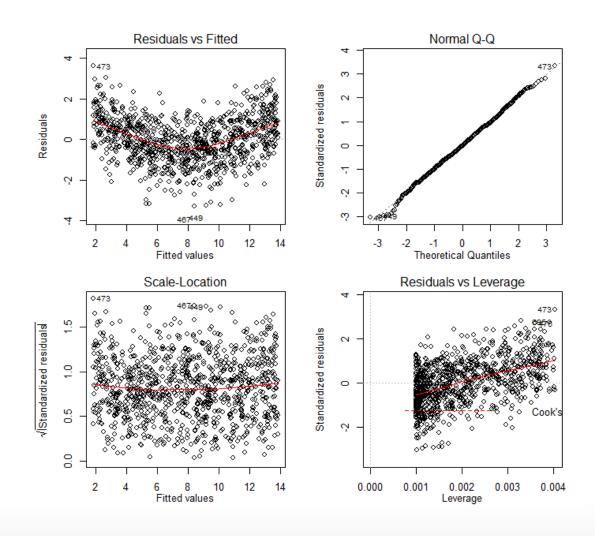
```
##
## Call:
## lm(formula = y \sim x)
##
## Residuals:
               10 Median
       Min
                                         Max
## -2.91933 -0.62956 0.01084 0.63819 2.73178
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 3.04449 0.06028 50.51 <2e-16 ***
          4.88928 0.10306 47.44 <2e-16 ***
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9486 on 998 degrees of freedom
## Multiple R-squared: 0.6928, Adjusted R-squared: 0.6925
## F-statistic: 2251 on 1 and 998 DF, p-value: < 2.2e-16
```

• Uniform regressor:



· Simulation with uniform regressor in quadratic form: case 1

```
set.seed(1234)
x <- runif(1000)
y <- 3+5*x+7*x^2+rnorm(1000)
par(mfrow=c(2,2)); plot(lm(y~x));</pre>
```

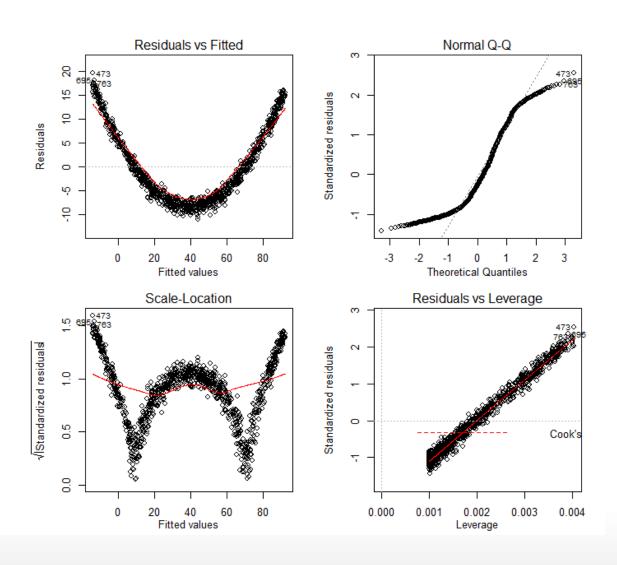


Fitting summary:

```
##
## Call:
## lm(formula = y \sim x)
##
## Residuals:
              10 Median
      Min
                              30
                                     Max
## -3.3057 -0.6886 -0.0230 0.7566 3.6079
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.83989 0.06924 26.57 <2e-16 ***
             11.98389 0.11839 101.22 <2e-16 ***
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1.09 on 998 degrees of freedom
## Multiple R-squared: 0.9112, Adjusted R-squared: 0.9112
## F-statistic: 1.025e+04 on 1 and 998 DF, p-value: < 2.2e-16
```

· Simulation with uniform regressor in quadratic form: case 2

```
set.seed(1234)
x <- runif(1000)
y <- 3+5*x+100*x^2+rnorm(1000)
par(mfrow=c(2,2)); plot(lm(y~x));</pre>
```

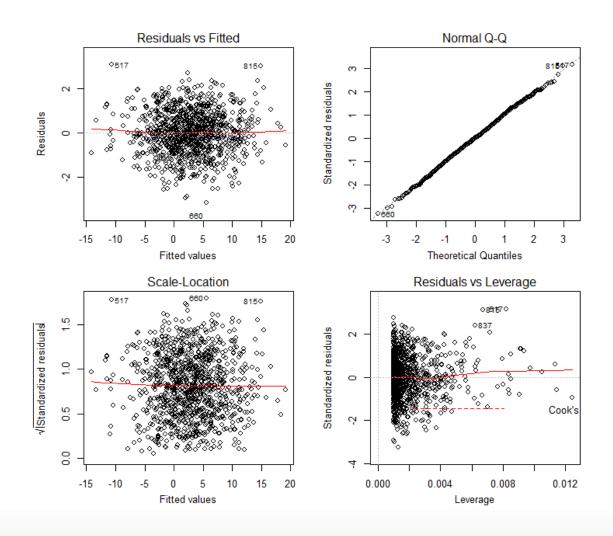


Fitting summary:

```
##
## Call:
## lm(formula = y \sim x)
##
## Residuals:
            10 Median 30
     Min
                              Max
## -11.037 -6.767 -2.016 6.182 19.554
##
## Coefficients:
##
           Estimate Std. Error t value Pr(>|t|)
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7.707 on 998 degrees of freedom
## Multiple R-squared: 0.9416, Adjusted R-squared: 0.9416
## F-statistic: 1.61e+04 on 1 and 998 DF, p-value: < 2.2e-16
```

· Simulation with normal regressor in linear form:

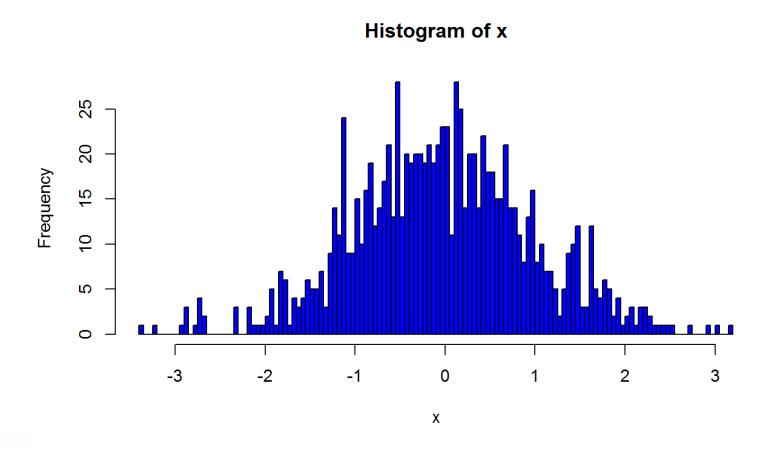
```
set.seed(1234)
x <- rnorm(1000)
y <- 3+5*x+rnorm(1000)
par(mfrow=c(2,2)); plot(lm(y~x));</pre>
```



Fitting summary:

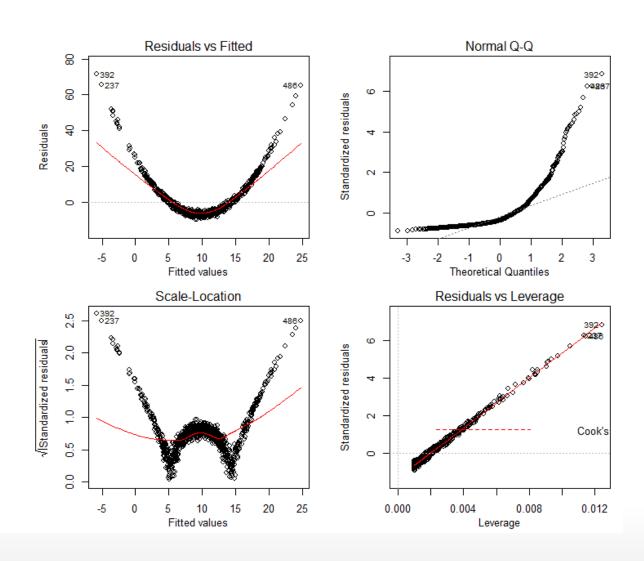
```
##
## Call:
## lm(formula = y \sim x)
##
## Residuals:
               10 Median
      Min
                              30
                                     Max
## -3.1661 -0.6439 0.0145 0.6537 3.0684
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 3.01599 0.03100 97.28 <2e-16 ***
              5.05571 0.03109 162.60 <2e-16 ***
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9801 on 998 degrees of freedom
## Multiple R-squared: 0.9636, Adjusted R-squared: 0.9636
## F-statistic: 2.644e+04 on 1 and 998 DF, p-value: < 2.2e-16
```

· Normal regressor:



· Simulation with normal regressor in quadratic form:

```
set.seed(1234)
x <- rnorm(1000)
y <- 3+5*x+7*x^2+rnorm(1000)
par(mfrow=c(2,2)); plot(lm(y~x));</pre>
```

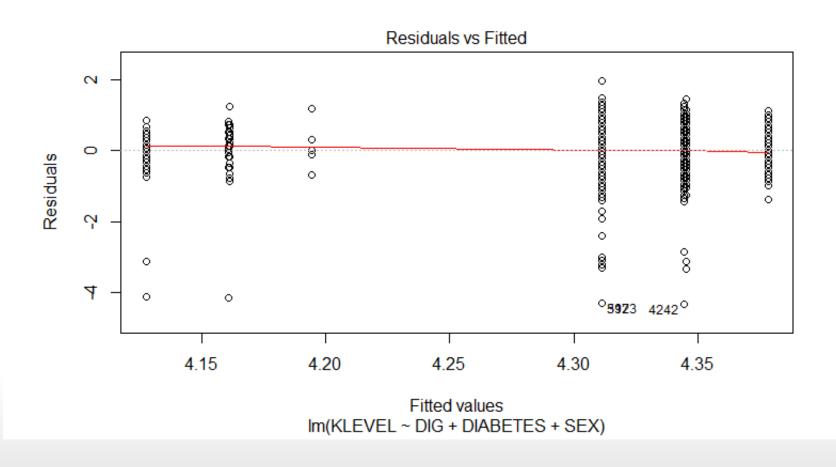


Fitting summary:

```
##
## Call:
## lm(formula = y \sim x)
##
## Residuals:
     Min 10 Median
                          30
                                Max
## -9.727 -6.123 -3.983 1.715 71.486
##
## Coefficients:
##
             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 9.9659 0.3328 29.94 <2e-16 ***
                4.6471 0.3338 13.92 <2e-16 ***
## x
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 10.52 on 998 degrees of freedom
## Multiple R-squared: 0.1627, Adjusted R-squared: 0.1618
## F-statistic: 193.9 on 1 and 998 DF, p-value: < 2.2e-16
```

· Example Data: diagnostic plots of regression with only main effects

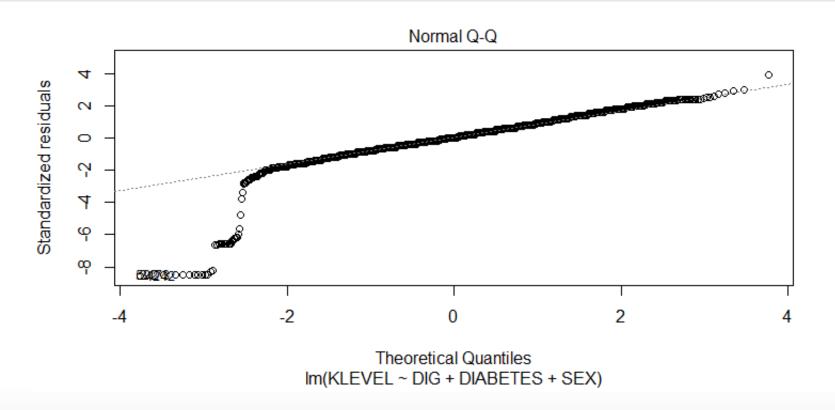
```
plot(lm(KLEVEL ~ DIG + DIABETES + SEX))
```



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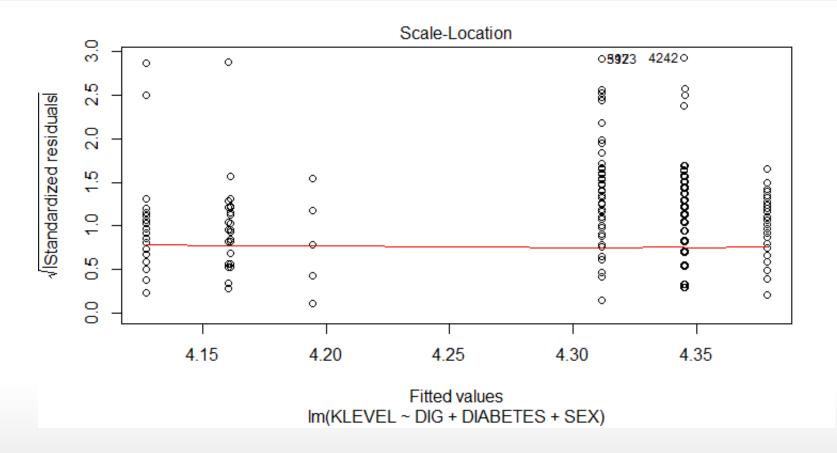
· Example Data: diagnostic plots of regression with only main effects

Hit <Return> to see next plot:



· Example Data: diagnostic plots of regression with only main effects

Hit <Return> to see next plot:



• Example Data: diagnostic plots of regression with only main effects

Hit <Return> to see next plot:

• Tests for checking constant variance assumption:

Multicollinearity

- Linear relationship exists among predictors, without the response variable.
- Evaluation of $(X'X)^{-1}$ is unstable.
- Variance inflation factor (VIF) > 10.

$$VIF(\hat{\beta}_i) = \frac{1}{1 - R_i^2}$$

• Condition number $\kappa > 15$.

$$k = \sqrt{\frac{|\lambda_{max}|}{|\lambda_{min}|}}, \quad \lambda_{min} \neq 0$$

Check multicollinearity:

```
car::vif(lm(KLEVEL ~., data=data.frame(DIG, DIABETES, SEX)))
        DIG DIABETES
## 1.000550 1.000082 1.000481
car::vif(lm(KLEVEL ~.^3, data=data.frame(DIG, DIABETES, SEX)))
                DIG
                            DIABETES
                                                  SEX
                                                           DIG:DIABETES
                                                              12.352165
          12.103871
                            9.747427
                                             1.429230
                        DIABETES:SEX DIG:DIABETES:SEX
            DIG:SEX
          11.986131
                           10.179364
                                            12.206203
kappa(lm(KLEVEL ~., data=data.frame(DIG, DIABETES, SEX)))^0.5
## [1] 3.174667
kappa(lm(KLEVEL ~.^3, data=data.frame(DIG, DIABETES, SEX)))^0.5
```

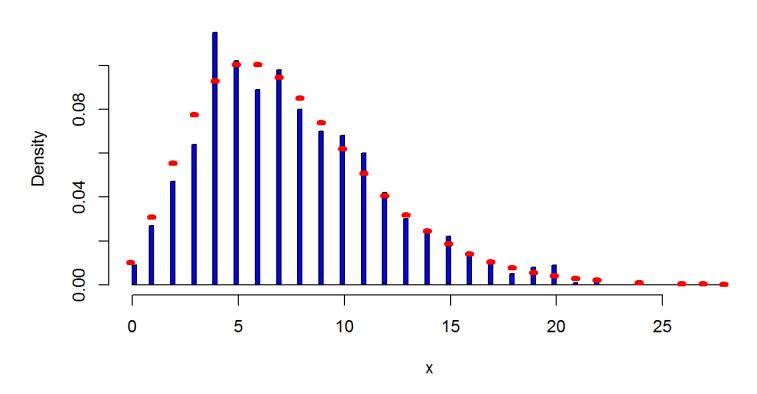
[1] 11**.**33542 57/82

Simulation with uniform regressor in quadratic form:

· Fitting summary:

```
##
## Call:
## lm(formula = y \sim x + I(x^2))
##
## Residuals:
               10 Median
       Min
                                        Max
## -2.91890 -0.62897 0.01039 0.63878 2.73195
##
## Coefficients:
##
        Estimate Std. Error t value Pr(>|t|)
## (Intercept) 3.04651 0.09059 33.63 <2e-16 ***
             4.87741 0.41127 11.86 <2e-16 ***
## x
## I(x^2) 100.01171 0.39283 254.59 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9491 on 997 degrees of freedom
## Multiple R-squared: 0.9991, Adjusted R-squared: 0.9991
## F-statistic: 5.632e+05 on 2 and 997 DF, p-value: < 2.2e-16
```

Histogram of NB(5,0.6), n=1000



Negative Binomial: a <u>discrete</u> distribution

$$X \sim NB(r, p)$$

$$\mu = E[X] = \frac{rp}{1-p}, \qquad \sigma^2 = Var[X] = \frac{rp}{(1-p)^2}$$

- X: number of successes until r of failures have occurred
- p: probability of success in each trial

Density function:

$$f(x) = \Pr(X = x | r, p) = {x + r - 1 \choose x} p^x (1 - p)^r$$

- Fully specified by r, p
- Asymmetric, mean≠median≠mode, in general
- Histogram proportion limit, $n \to +\infty$

DESeq

<u>Function:</u> typically used to decide whether, for a given gene/class, under multiple conditions, an observed difference in read counts is significant

- RNA-Seq: class → target transcript
- ChIP-Seq: class → binding region

Assumption: read counts follow negative binomial distributions

- Discrete, positive, skewed → (log-)normal is unsuitable
- Small number of replicates → insufficient for rank based or permutation methods

DESeq Read Counts Model

$$K_{ij} \sim NB(\mu_{ij}, \sigma_{ij}^{2})$$

$$\mu_{ij} = q_{i\rho} \times s_{j}$$

$$\sigma_{ij}^{2} = \mu_{ij} + s_{j}^{2} \times v_{i\rho} = \mu_{ij} + s_{j}^{2} \times f(q_{i\rho})$$

$$i = 1, ..., n \qquad j = 1, ..., m$$

 K_{ij} : number of reads in sample j assigned to gene i $q_{i\rho}$: averaged read counts of gene i in condition ρ s_j : size factor, adjusting for the coverage or sampling depth of library j

Goal of PCA:

- Summarize a n × p data matrix containing p variables by uncorrelated linear combinations of the original p variables.
- These uncorrelated linear combinations are called principal components or principal axes.
- The principal components are ordered; the 1st component has the largest variance, corresponding to the maximum possible portion of the variation in the original data matrix that can be explained by a linear combination of the original p variables.

Assumption of PCA:

 Relationships among the variables are approximately linear.

Data Preparation:

 Normalization (centering and rescaling) the variable measurements are often needed prior to application of PCA, if the goal of the analysis is to decompose variation in data matrix induced by relationships between variables.

Simulation 1:

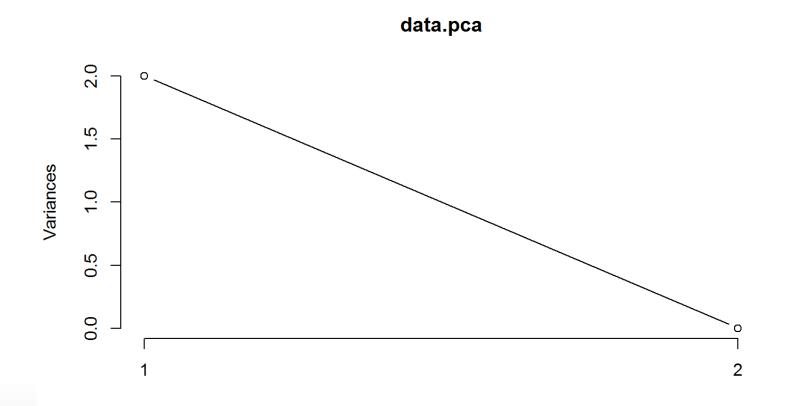
```
set.seed(1234)
x \leftarrow runif(1000)
y < -3+5*x
data \leftarrow cbind(x,y)
data.pca <- prcomp(data, center=T, scale=T)</pre>
print(data.pca)
## Standard deviations:
## [1] 1.414214e+00 8.918918e-16
##
## Rotation:
##
            PC1
                        PC2
## \times 0.7071068 -0.7071068
## y 0.7071068 0.7071068
```

Simulation 1:

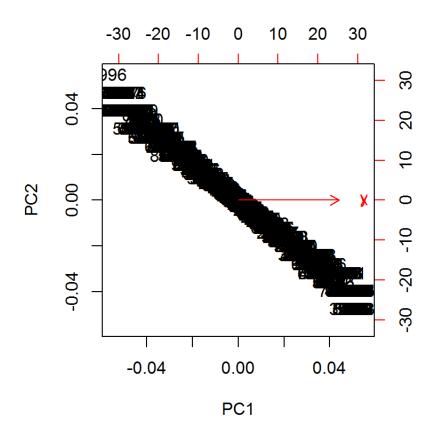
```
summary(data.pca)
```

```
## Importance of components:
## PC1 PC2
## Standard deviation 1.414 8.919e-16
## Proportion of Variance 1.000 0.000e+00
## Cumulative Proportion 1.000 1.000e+00
```

plot(data.pca, type = "1")



biplot(data.pca)



Simulation 2:

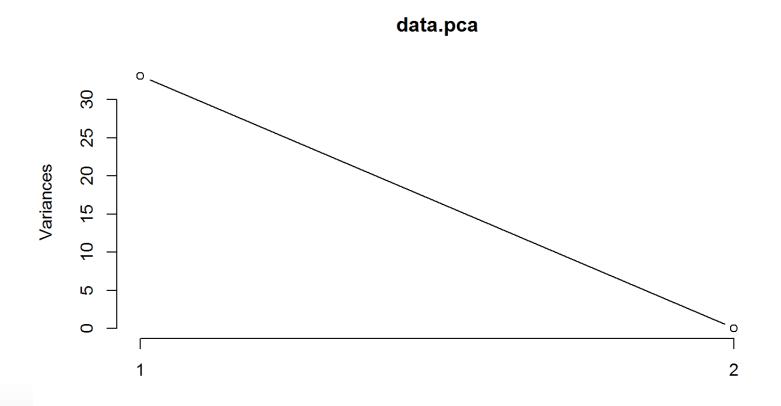
```
set.seed(1234)
x \leftarrow runif(1000)
y < -3+5*x
data \leftarrow cbind(x,y)
data.pca <- prcomp(data, center=F, scale=F)</pre>
print(data.pca)
## Standard deviations:
## [1] 5.7551213 0.1518755
##
## Rotation:
##
             PC1
                          PC2
## x -0.09822225 0.99516450
## y -0.99516450 -0.09822225
```

Simulation 2:

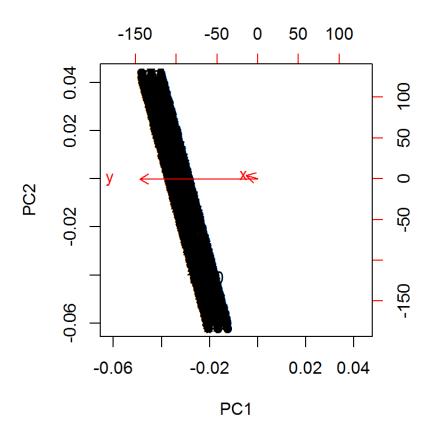
```
summary(data.pca)
```

```
## Importance of components:
## PC1 PC2
## Standard deviation 5.7551 0.1519
## Proportion of Variance 0.9993 0.0007
## Cumulative Proportion 0.9993 1.0000
```

plot(data.pca, type = "1")



biplot(data.pca)



Simulation 3:

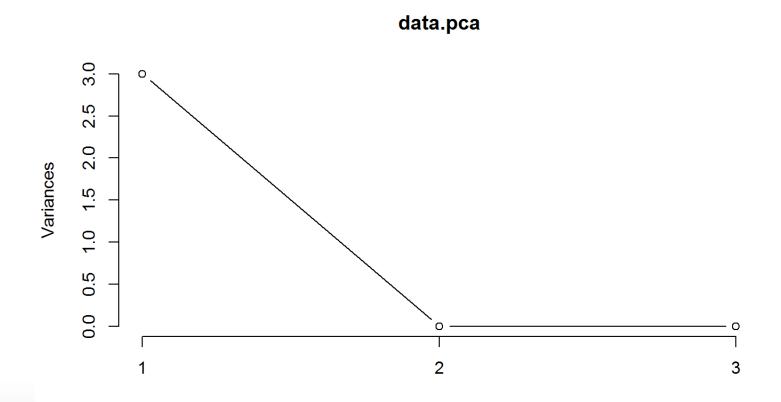
```
set.seed(1234)
x \leftarrow runif(1000)
y < -3+5*x
z < 1+4*x
data \leftarrow cbind(x,y,z)
data.pca <- prcomp(data, center=T, scale=T)</pre>
print(data.pca)
## Standard deviations:
## [1] 1.732051e+00 1.059140e-15 5.623615e-17
##
## Rotation:
## PC1 PC2
                                  PC3
## x 0.5773503 0.8155967 -0.03832384
## y 0.5773503 -0.3746089 0.72548937
                                                                        75/82
## z 0.5773503 -0.4409878 -0.68716553
```

Simulation 3:

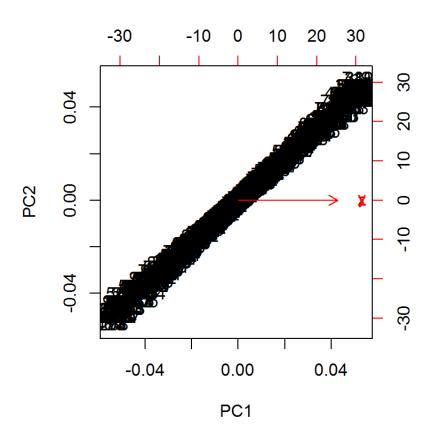
```
summary(data.pca)
```

```
## Importance of components:
## PC1 PC2 PC3
## Standard deviation 1.732 1.059e-15 5.624e-17
## Proportion of Variance 1.000 0.000e+00 0.000e+00
## Cumulative Proportion 1.000 1.000e+00 1.000e+00
```

plot(data.pca, type = "1")



biplot(data.pca)



Simulation 4:

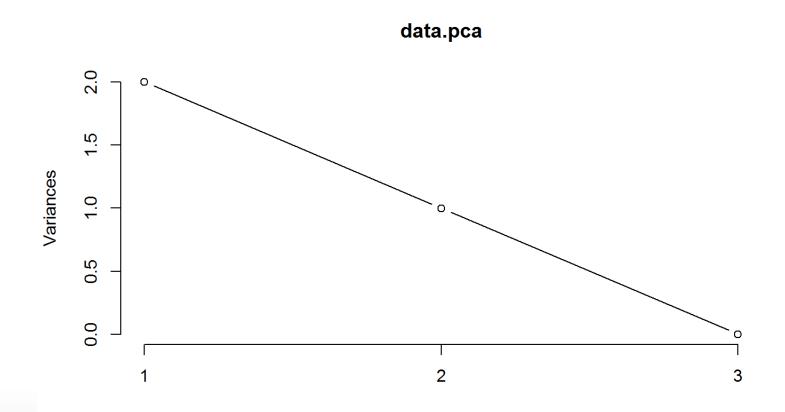
```
set.seed(1234)
x \leftarrow runif(1000)
y < -3+5*x
z \leftarrow runif(1000)
data \leftarrow cbind(x,y,z)
data.pca <- prcomp(data, center=T, scale=T)</pre>
print(data.pca)
## Standard deviations:
## [1] 1.414745e+00 9.992479e-01 8.938641e-16
##
## Rotation:
## PC1 PC2
                                       PC3
## x 0.7065765 -0.02737854 -7.071068e-01
## y 0.7065765 -0.02737854 7.071068e-01
                                                                           79/82
## z 0.0387191 0.99925013 3.816392e-17
```

Simulation 4:

```
summary(data.pca)
```

```
## Importance of components:
## PC1 PC2 PC3
## Standard deviation 1.4147 0.9992 8.939e-16
## Proportion of Variance 0.6672 0.3328 0.000e+00
## Cumulative Proportion 0.6672 1.0000 1.000e+00
```

plot(data.pca, type = "1")



biplot(data.pca)

