## TD1: Non parametric regression

## 1 Bias and variance

Consider the regression problem of dimension N:

$$X_k = \theta_k + \varepsilon_k, k = 1, \dots, N,$$

where  $X_1, \ldots, X_N$  are i.i.d random variables in  $\mathbb{R}$ ,

 $\theta \in \mathbb{R}^N$  the unknown parameter and  $\varepsilon_1, \dots, \varepsilon_N$  are i.i.d samples following a centered Gaussian with known variance  $\sigma^2$ .

We want to find an estimator of  $\theta$  of the form  $\widehat{\theta}(\lambda) = \lambda X$  with  $\lambda \in [0, 1]$ .

- 1. Compute its bias  $\mathbb{E}(\widehat{\theta}(\lambda) \theta)$  and the sum of the variance its coordinates:  $\sum_{k=1}^{N} \mathbb{V}(\widehat{\theta}(\lambda)_k)$ . What is the effect of  $\lambda$  on the estimator?
- 2. Compute  $\lambda^*$  the minimizer of the oracle risk of  $\widehat{\theta}(\lambda)$  given by:

$$R(\lambda) = \mathbb{E}(L(\lambda))$$

where 
$$L(\lambda) = \sum_{k=1}^{N} (\widehat{\theta}(\lambda)_k - \theta_k)^2$$
.

- 3. Propose an unbiased estimator of  $\theta_k^2$ , given the observed sample  $x_k$ .
- 4. Deduce an unbiased estimator  $C(\lambda)$  of  $R(\lambda) \sum_{k=1}^N \theta_k^2$

## 2 Regression

Let  $f: \mathbb{R} \to \mathbb{R}$  be an unknown function. Suppose we observe noisy values of f:

$$Y_i = f(X_i) + \epsilon_i,$$

where  $(X, \epsilon), (X_i, \epsilon_i)_{i=1,\dots,n} \subset \mathbb{R} \times \mathbb{R}$  is an i.i.d. collection of random variables such that  $\epsilon \perp X$ ,  $\mathbb{E}[\epsilon] = 0$ ,  $\mathbb{E}[\epsilon^2] = \sigma^2$  and  $\mathbb{E}[f(X)^2] < \infty$ . To estimate the function f, we shall use a sequence of functions  $(h_k)_{k\geq 1}$  where each  $h_k : \mathbb{R} \to \mathbb{R}$  and  $G = \mathbb{E}[h(X)h(X)^T]$  is invertible. The procedure is as follows. We choose K, we set  $h(x) = (h_1(x), \dots h_K(x))^T$  and then compute

$$\theta_{K,n} \in \operatorname*{arg\,min}_{\theta \in \mathbb{R}^K} \sum_{i=1}^n (Y_i - h(X_i)^T \theta)^2.$$

The estimate of f is given by  $f_{K,n}(x) = h(x)^T \theta_{K,n}$ . Define the risk  $R(\theta) = \mathbb{E}[(f(X) - h(X)^T \theta)^2]$ . The aim is to study the  $R(\theta_{n,K})$ .

- 1. Let  $\theta_K^* \in \arg\min_{\theta \in \mathbb{R}^K} R(\theta)$ . Give the normal equations satisfied by  $\theta_K^*$  and deduce an expression for  $\theta_K^*$ .
- 2. Give the expression for  $\theta_{K,n}$ .
- 3. Show that the estimated function  $f_{K,n}$  is invariant by any linear invertible transform on the set of functions h, i.e. h is replaced by Ah where  $A \in \mathbb{R}^{K \times K}$  is invertible. In the following we shall assume that  $G = I_K$ .

- 4. Show that for any  $\theta \in \mathbb{R}^K$ ,  $R(\theta) = R(\theta_K^*) + \|\theta \theta_K^*\|^2$  where the norm  $\|\cdot\|$  should be specified.
- 5. From now on, we suppose that the smallest eigenvalue of  $n^{-1} \sum_{i=1}^{n} h(X_i) h(X_i)^T$  is lower bounded by  $\lambda > 0$ . Show that  $\|\theta_{K,n} \theta_K^*\| \le \lambda^{-1/2} \|n^{-1} \sum_{i=1}^{n} \xi_i h(X_i)\|$  where  $\xi_i \ne \epsilon_i$  should be specified.
- 6. Show that  $n^{-1} \sum_{i=1}^{n} \xi_i h(X_i) \to 0$  almost surely (hint: use the Cauchy-Schwarz inequality).
- 7. Suppose that  $f \in \text{span}((h_k)_{k \geq 1})$ . Conclude that choosing K large enough,  $\limsup_n R(\theta_{K,n})$  can be made arbitrarily small.