Homework 2

Oregon State University

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Problem 1

Part A

Given that $y_1, \dots, y_n | \theta \sim \text{iid Poisson}(\theta)$ and $\theta \sim \text{Gamma}(a, b)$, we can express the likelihood and prior as follows,

The likelihood of observing y_1, \dots, y_n given θ for a Poisson distribution is given as,

$$L(\theta|y_1,\ldots,y_n) = \prod_{i=1}^n \frac{e^{-\theta}\theta^{y_i}}{y_i!} = \frac{e^{-n\theta}\theta^{\sum_{i=1}^n y_i}}{\prod_{i=1}^n y_i!}$$

The prior distribution of θ is,

$$p(\theta) = \frac{b^a}{\Gamma(a)} \theta^{a-1} e^{-b\theta}$$

Applying Bayes' theorem, the posterior distribution is,

$$\begin{split} p(\theta|y_1,\dots,y_n) &\propto L(\theta|y_1,\dots,y_n) \times p(\theta) = \frac{e^{-n\theta}\theta^{\sum_{i=1}^n y_i}}{\prod_{i=1}^n y_i!} \times \frac{b^a}{\Gamma(a)}\theta^{a-1}e^{-b\theta} \\ &= \frac{b^a}{\Gamma(a)\prod_{i=1}^n y_i!}\theta^{(\sum_{i=1}^n y_i+a-1)}e^{-(b+n)\theta} \end{split}$$

This simplifies to a Gamma distribution,

$$\theta|y_1,\dots,y_n \sim \operatorname{Gamma}\left(\sum_{i=1}^n y_i + a, b + n\right)$$

Hence, our posterior distribution of θ given y_1, \dots, y_n is a Gamma with parameters $\sum_{i=1}^n y_i + a$ and b+n.





The posterior predictive distribution for a new observation \tilde{y} , given y_1, \dots, y_n , is,

$$p(\tilde{y}|y_1,\dots,y_n) = \int_0^\infty p(\tilde{y}|\theta) p(\theta|y_1,\dots,y_n) d\theta$$

where $p(\tilde{y}|\theta)=e^{-\theta}\theta^{\tilde{y}}/\tilde{y}!,$ and $p(\theta|y_1,\ldots,y_n)$ is $\mathrm{Gamma}(\sum_{i=1}^n y_i+a,b+n).$

The predictive distribution simplifies to a negative binomial distribution,

$$p(\tilde{y}|y_1,\dots,y_n) = \frac{\Gamma(\tilde{y} + \sum_{i=1}^n y_i + a)}{\tilde{y}!\Gamma(\sum_{i=1}^n y_i + a)} \left(\frac{b}{b+n}\right)^{a + \sum_{i=1}^n y_i} \left(\frac{n}{b+n}\right)^{\tilde{y}}$$

Therefore, the posterior predictive distribution of a new observation \tilde{y} given y_1, \dots, y_n is a negative binomial distribution with parameters $\sum_{i=1}^n y_i + a$ and $\frac{n}{b+n}$.

Problem 2

Part A

For θ_A and θ_B , using their respective priors and the sums of tumor counts,

Sum of Tumor Counts:

$$\sum y_A = 117, \quad \sum y_B = 104$$

Posterior for θ_A :

$$\theta_A|y_A \sim \text{Gamma}(237, 20)$$

• Mean: $\mu_A = 11.85$

• Variance: $\sigma_A^2 = 0.5925$

Posterior for θ_B :

$$\theta_B|y_B \sim \text{Gamma}(116, 14)$$

• Mean: $\mu_B = 8.29$

• Variance: $\sigma_B^2 = 0.5918$

The 95% quantile-based Bayesian confidence intervals can be derived using R. I applied qgamma(c(0.025, 0.975), shape = alphaA, rate = betaA) and qgamma(c(0.025, 0.975), shape = alphaB, rate = betaB) to get the 95% Bayesian confidence intervals.

[1] "95% Confidence Interval for theta_A: [10.39, 13.41]"

[1] "95% Confidence Interval for theta_B: [6.85, 9.86]"





We can plot the posterior expectation of θ_B as a function of n_0 in the prior Gamma $(12n_0, n_0)$ using R with my favorite package ggplot2,



The posterior expectation helps us understand how much our previous knowledge (or assumptions) affects the results we get after seeing the actual data. By plotting and analyzing this, we can see how sensitive our results are to changes in what we previously thought. This is especially useful when our initial guesses or information might be a bit uncertain or not based on a lot of data. This way, we can check if our conclusions are reliable or if they're too dependent on shaky assumptions.