

# **Statistical Modeling**

CH.9 - Working with Collinear Data

SS 2021 || Prof. Dr. Buchwitz

Wirgeben Impulse

#### **Outline**

- 1 Organizational Information
- 2 Multicollinearity
- **3** Principal Components
- 4 Principal Component Regression
- 5 Ridge Regression

## **Course Contents**

Session	Торіс
1	Simple Linear Regression
2	Multiple Linear Regression
3	Regression Diagnostics
4	Qualitative Variables as Predictors
5	Transformation of Variables
6	Weighted Least Squares
7	Correlated Errors
8	Analysis of Collinear Data
9	Working with Collinear Data
10	Variable Selection Procedures
11	Logistic Regression
12	Further Topics

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#### Introduction

- When multicollinearity is present, the least squares estimates of the individual regression coefficients ten to be unstable and can lead to erroneous inferences.
- In the last session we discussed the problem of multicollinearity and ways to diagnose this problem. We found that eliminating predictors from the analysis does not always work and in most analytical settings is not a feasible option.
- We consider two alternative approaches for dealing with multicollinearity:
  - Imposing or searching for constraints on the regression parameters.
  - Using alternative estimation techniques (e.g. principal components regression and ridge regression).

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- The principal components method is based on the fact that any set of p predictors  $X_1, X_2, \ldots, X_p$  can be **transformed** to a set of p **orthogonal** variables.
- The new orthogonal variables are known as the **principal** components and are denoted by  $C_1, C_2, \ldots, C_p$ .
- Each variable  $C_j$  is a linear function of the standardized variables  $\tilde{X}_1, \tilde{X}_2, \dots, \tilde{X}_p$ .

$$C_j = v_{1j}\tilde{X}_1 + v_{2j}\tilde{X}_2 + \ldots + v_{pj}\tilde{X}_p$$
 for  $j = 1, 2, \ldots, p$ 

- The coefficients of the linear functions are chosen so that the variables  $C_1, \ldots, C_p$  are orthogonal.
- The coefficients for the *j*-th principal components  $C_j$  are the elements of the *j*-th eigenvector that corresponds to the eigenvalue  $\lambda_j$ , the *j*-th largest eigenvalue of the correlation matrix of the *p* variables.

$$V = \begin{pmatrix} V_1 & V_2 & \cdots & V_p \end{pmatrix} = \begin{pmatrix} v_{11} & v_{12} & \cdots & v_{1p} \\ v_{21} & v_{22} & \cdots & v_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ v_{p1} & v_{p2} & \cdots & v_{pp} \end{pmatrix}$$

## **Example: French Econmony Data**

P241 YEAR IMPORT DOPROD STOCK CONSUM ## 15.9 149.3 4.2 108.1 ## 1 ## 2 50 16.4 161.2 4.1 114.8 19.0 171.5 3.1 123.2 ## 3 51 ## 4 19.1 175.5 3.1 126.9 ## 5 53 18.8 180.8 1.1 132.1 ## 6 54 20.4 190.7 2.2 137.7 ## 7 55 22.7 202.1 2.1 146.0 26.5 212.4 ## 8 56 5.6 154.1 5.0 162.3 ## 9 57 28.1 226.1 ## 10 27.6 231.9 5.1 164.3 58 ## 11 26.3 239.0 0.7 167.6 59 ## 12 60 31.1 258.0 5.6 176.8 ## 13 61 33.3 269.8 3.9 186.6 ## 14 37.0 288.4 3.1 199.7 62 ## 15 63 43.3 304.5 4.6 213.9 ## 16 64 49.0 323.4 7.0 223.8 ## 17 50.3 336.8 1.2 232.0 65

56.6 353.9

4.5 242.9

## 18

66

## **Data Description**

YEAR Year of Observation.

IMPORT Import Volume.

DOPROD Domestic Production.

STOCK Stock Formation.

**CONSUM Domestic Consumption.** 

Variables are measured in billion French francs.

It can be shown that the variance of the *j*-th principal component is  $Var(C_j) = \lambda_j$  for j = 1, 2, ..., p. Therefore the variance-covariance matrix of the principal components is

$$\begin{pmatrix} \lambda_1 & 0 & \cdots & 0 \\ 0 & \lambda_2 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \lambda_p \end{pmatrix}$$

- All the off-diagonal elements are zero because the principal components are orthogonal. The value of the *j*-th diagonal element  $\lambda_i$  is the variance of  $C_i$ , the *j*-th principal component.
- The principal components are arranged so that  $\lambda_1 \geq \lambda_2 \geq \dots \lambda_p$ , which means that the first component has the largest variance.

```
d <- head(P241[ ,c("DOPROD", "STOCK", "CONSUM")], 11)
d.pca <- prcomp(d, center=TRUE, scale=TRUE)
C <- d.pca$x
round(C, 4)</pre>
```

```
##
         PC1
                 PC2
                        PC3
## 1 -2.1259 0.6387 -0.0207
  2 -1.6189 0.5555 -0.0711
## 3 -1.1152 -0.0730 -0.0217
## 4 -0.8943 -0.0824 0.0108
  5 -0.6442 -1.3067 0.0726
  6 -0.1904 -0.6591 0.0266
##
  7 0.3596 -0.7437 0.0428
##
  8 0.9718 1.3541 0.0629
## 9
     1.5593 0.9640 0.0236
## 10
     1.7670 1.0152 -0.0450
## 11
     1.9311 -1.6627 -0.0806
```

```
cormat <- cor(d)
eigen(cormat) # Eigen Decomposition of Correlation Matrix

## eigen() decomposition
## $values
## [1] 1.999154934 0.998154176 0.002690889
##
## $vectors
## [,1] [,2] [,3]
## [1,] 0.70633041 0.03568867 0.706982083
## [2,] 0.04350059 -0.99902908 0.006970795
## [3,] 0.70654444 0.02583046 -0.707197102</pre>
```

```
round(var(C),4) # Variance-Covariance Matrix of PCs
```

```
## PC1 PC2 PC3
## PC1 1.9992 0.0000 0.0000
## PC2 0.0000 0.9982 0.0000
## PC3 0.0000 0.0000 0.0027
```

#### Remember

Multicollinearity leads to heterogeneous sizes of eigenvalues so that one eigenvalue is much smaller than the others. When one eigenvalue is exactly zero a perfect linear relationship (special case of extreme multicollinearity) among the original variables exists.

The variance-covariance matrix of the new variables only has entries on the main diagonal (which correspond to the eigenvalues) and zeros in all other places (as the variables are orthogonal).

- The principal components lack simple interpretation as they are a mixture of the (standardized) original variables.
- Since  $\lambda_j$  is the variance of the j-th principal component, a value of  $lambda_j \approx 0$  shows that the respective principal component  $C_j$  is equal to a constant. That constant is the mean value of  $C_j$  (which is zero as the variables have been standardized).
- Inspecting the eigenvectors of the previous example shows that *only* the variables CONSUM and DOPROD play a relevant role when determining  $C_3$ .

```
## X1_tilde X2_tilde X3_tilde
## 0.706982083 0.006970795 -0.707197102
```

$$\tilde{X}_1 pprox \tilde{X}_3$$
 as  $v_{23} pprox 0.007 pprox 0$ 

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■ We consider the model for the French Economony Dataset

$$\mathsf{IMPORT} = \beta_0 + \beta_1 \mathsf{DOPROD} + \beta_2 \mathsf{STOCK} + \beta_3 \mathsf{CONSUM} + \epsilon$$

■ This model expressed using the standardized variables  $\tilde{Y} = (y_i - \bar{y})/s_y$  and  $\tilde{X}_j = (x_{ij} - \bar{x}_j)/s_{x_j}$  yields

$$\tilde{Y} = \theta_1 \tilde{X}_1 + \theta_2 \tilde{X}_2 + \theta_3 \tilde{X}_3 + \epsilon'$$

 Utilizing the principal components of the standardized predictors the model can be written as

$$\tilde{\mathbf{Y}} = \alpha_1 \mathbf{C}_1 + \alpha \mathbf{C}_2 + \alpha_3 \mathbf{C}_3 + \epsilon'$$

```
# Data Preparation
d <- head(P241,11)
d_scaled <- as.data.frame(scale(d))</pre>
d prcomp <- as.data.frame(cbind(IMPORT=d scaled$IMPORT,</pre>
                                 prcomp(d[,c("DOPROD","STOCK","CONSUM")],
                                        center=TRUE, scale=T)$x))
# Motel Estimation
mod1 <- lm(IMPORT ~ 1 + DOPROD + STOCK + CONSUM, data=d)
(mod2 <- lm(IMPORT ~ -1 + DOPROD + STOCK + CONSUM, data=d scaled))
##
## Call:
## lm(formula = IMPORT ~ -1 + DOPROD + STOCK + CONSUM, data = d_scaled)
##
## Coefficients:
## DOPROD STOCK CONSUM
## -0.3393 0.2130 1.3027
(mod3 <- lm(IMPORT ~ -1 + PC1 + PC2 + PC3, data=d_prcomp))</pre>
##
## Call:
## lm(formula = IMPORT ~ -1 + PC1 + PC2 + PC3, data = d_prcomp)
##
## Coefficients:
      PC1
              PC2
                      PC3
## 0.6900 0.1913 1.1597
```

```
ev <- eigen(cor(d[,c("DOPROD","STOCK","CONSUM")]))$vectors

# Multiply eigenvectors with constant to match output in book
ev[,2] <- ev[,2] * -1; ev[,3] <- ev[,3] * -1

# Eigenvectors
ev</pre>
```

```
## [,1] [,2] [,3]
## [1,] 0.70633041 -0.03568867 -0.706982083
## [2,] 0.04350059 0.99902908 -0.006970795
## [3,] 0.70654444 -0.02583046 0.707197102
```

The coefficients of the principal component regression can be calculated based on the regression coefficients from the model using the standardized values.

$$\alpha_1 = 0.706 \quad \theta_1 + 0.044 \quad \theta_2 + 0.707 \quad \theta_3$$
 $\alpha_2 = -0.036 \quad \theta_1 + 0.999 \quad \theta_2 + -0.026 \quad \theta_3$ 
 $\alpha_3 = -0.707 \quad \theta_1 + -0.007 \quad \theta_2 + 0.707 \quad \theta_3$ 

Conversely this relationship can be turned around to obtain the coefficients from the regression with standardized variables from the principal component regression.

$$\theta_1$$
 = 0.706  $\alpha_1$  + -0.036  $\alpha_2$  + -0.707  $\alpha_3$   
 $\theta_2$  = 0.044  $\alpha_1$  + 0.999  $\alpha_2$  + -0.007  $\alpha_3$   
 $\theta_3$  = 0.707  $\alpha_1$  + -0.026  $\alpha_2$  + 0.707  $\alpha_3$ 

DOPROD

STOCK

## -0.3393426 0.2130484 1.3026815

CONSUM

```
# Calculate alpha (principal components) from theta (standardized variables)
as.vector(coef(mod2) %*% ev)
## [1] 0.6899821 0.1913034 1.1596766
coef(mod3)
         PC1
                   PC2
                             PC3
## 0.6899821 0.1913034 1.1596766
# Calculate theta (standardized variables) from alpha (principal components)
as.vector(ev %*% coef(mod3))
## [1] -0.3393426 0.2130484 1.3026815
coef(mod2)
```

$$\tilde{Y} = \theta_1 \tilde{X}_1 + \theta_2 \tilde{X}_2 + \theta_3 \tilde{X}_3 + \epsilon'$$
$$= \alpha_1 C_1 + \alpha C_2 + \alpha_3 C_3 + \epsilon'$$

- Although the above equations both hold, the C's are **orthogonal**.
- The orthogonality bypasses (but not eliminates) the multicollinearity problem, however, the resulting relationship and therefore the coefficients are not easily interpreted.
- The  $\alpha$ 's unlike the  $\theta$ 's do not have simple interpretations as marginal effects of the original (standardized) predictor variables.

The final estimation results are always restated in terms of the  $\theta$ 's or origininal  $\beta$ 's for interpretation!

Based on the coefficients obtained from regressing the standardized variables the relationship can be expressed in terms of the original  $\beta_j$ 's using the following relationship:

$$\hat{\beta}_j = \frac{s_y}{s_j} \hat{\theta}_j \quad \text{for} \quad j = 1, 2, \dots, p$$

$$\beta_0 = \bar{y} - \hat{\beta}_1 \bar{x}_1 - \hat{\beta}_2 \bar{x}_2 - \dots - \hat{\beta}_p \bar{x}_p$$

This back-transform of the variables to the original scale is crucial for interpretation of the final results!

- Principal component regression can be used to reduce collinearity in the estimation data.
- this can be achieved by using less than the full set of principal components to explain the variation in the response.
- When all principal components are used the OLS solution can be exactly reproduced (as seen before).

■ The  $C_j$ 's have sample variances  $\lambda_1, \lambda_2, \dots, \lambda_p$  equal to their eigenvalues.

```
eigen(cor(d[,c("DOPROD","STOCK","CONSUM")]))$values
```

```
## [1] 1.999154934 0.998154176 0.002690889
```

Since  $C_3$  has very small variance, the linear function defining  $C_3$  is approximately equal to zero and is the source of collinearity in the data.

■ We exclude  $C_3$  from the analysis and consider the two possible remaining regression models

$$\begin{split} \tilde{\mathsf{Y}} &= \alpha_1 \mathsf{C}_1 + \epsilon \\ \tilde{\mathsf{Y}} &= \alpha_1 \mathsf{C}_1 + \alpha_2 \mathsf{C}_2 + \epsilon \end{split}$$

#### There are two important things to note here:

- In an regression equation where the full set of potential predictor variables under consideration are orthogonal, the estimated values of the regression **coefficients are not altered** when subsets of these variables are either introduced or deleted.
- Both models lead to estimates for **all** three of the original standardized coefficients  $\theta_1$ ,  $\theta_2$  and theta<sub>3</sub>.

	Model 1	Model 2	Model 3				
PC1	0.69***	0.69***	0.69***				
	(0.05)	(0.03)	(0.02)				
PC2		0.19***	0.19***				
		(0.04)	(0.03)				
PC3	1.		1.16				
			(0.61)				
$R^2$	0.95	0.99	0.99				
Adj. R <sup>2</sup>	0.95	0.99	99 0.99				
Num. obs.	11	11	11				
*** < 0.001. ** < 0.01. * < 0.05							

<sup>\*\*\*</sup>p < 0.001; \*\*p < 0.01; \*p < 0.05

Table 2: Statistical models

```
# Coefficients for standardized predictors when using one principal component
coefs1 <- coef(mod3)[1] * ev[,1]
names(coefs1) <- c("DOPROD","STOCK","CONSUM")
coefs1</pre>
```

```
## DOPROD STOCK CONSUM
## 0.48735534 0.03001463 0.48750301
```

```
# Coefficients for standardized predictors when using two principal components
coefs2 <- coef(mod3)[1] * ev[,1] + coef(mod3)[2] * ev[,2]
names(coefs2) <- c("DOPROD", "STOCK", "CONSUM")
coefs2</pre>
```

```
## DOPROD STOCK CONSUM
## 0.4805280 0.2211323 0.4825616
```

DOPROD

## -9.13010782 0.07277981 0.60922012 0.10625939

##

Intercept

```
s <- apply(d[ ,c("IMPORT", "DOPROD", "STOCK", "CONSUM")],2,sd)
m <- apply(d[ ,c("IMPORT", "DOPROD", "STOCK", "CONSUM")],2,mean)</pre>
# Model with one PC for non-standardized data
coefs_org1 \leftarrow s[1]/s[2:4] * coefs1
intercept org1 <- unname(m[1] - sum(m[2:4]*coefs org1))
(beta_org1 <- c(Intercept=intercept_org1, coefs_org1))</pre>
##
     Intercept
                     DOPROD
                                   STOCK
                                               CONSUM
## -7.74582557 0.07381387 0.08269039 0.10734749
# Model with two PCs for non-standardized data
coefs org2 \leftarrow s[1]/s[2:4] * coefs2
intercept org2 <- unname(m[1] - sum(m[2:4]*coefs org2))
(beta_org2 <- c(Intercept=intercept_org2, coefs_org2))</pre>
```

STOCK

CONSUM

- The following table shows that the coefficients are dependent on the number of incorporated principal components.
- As each component explains additional variance the R<sup>2</sup> inceases with the number of considered principal components.

	std_PC1	org_PC1	std_PC2	org_PC2	std_PC3	org_PC3
Intercept	NA	-7.746	NA	-9.130	NA	-10.128
DOPROD	0.487	0.074	0.481	0.073	-0.339	-0.051
STOCK	0.030	0.083	0.221	0.609	0.213	0.587
CONSUM	0.488	0.107	0.483	0.106	1.303	0.287

# **Constraints on the Regression Coefficients**

# **Caution when using Principal Component Regression**

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# **Ridge Regression**