

Statistical Modeling

CH.4 - Qualitative Variables

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Wir geben Impulse

1 Organizational Information

2 Qualitative Variables as Predictors

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1 Organizational Information

2 Qualitative Variables as Predictors

- Qualitative or categorical variables (such as gender, marital status, etc) are useful predictors and are usually called **indicator** or **dummy variables**.
- Those variables usually only take two values, 0 and 1, which signify that the observation belongs to one of two possible categories.
- The numerical values of indicator variables **do not reflect quantitative ordering**.
- **Example Variable:** Gender, coded as 1 for *female* and 0 for *male*.
- Indicator variables can also be used in a regression equation to distinguish between three or more groups.
- The response variable is still a quantitative continuous in all discussed cases.

Example: Salary Survey Data

P130

##		S	X	E	M
## 1	13876	1	1	1	
## 2	11608	1	3	0	
## 3	18701	1	3	1	
## 4	11283	1	2	0	
## 5	11767	1	3	0	
## 6	20872	2	2	1	
## 7	11772	2	2	0	
## 8	10535	2	1	0	
## 9	12195	2	3	0	
## 10	12313	3	2	0	
## 11	14975	3	1	1	
## 12	21371	3	2	1	
## 13	19800	3	3	1	
## 14	11417	4	1	0	
## 15	20263	4	3	1	
## 16	13231	4	3	0	
## 17	12884	4	2	0	
## 18	13245	5	2	0	
## 19	13677	5	3	0	
## 20	15965	5	1	1	
## 21	12336	6	1	0	
## 22	21352	6	3	1	
## 23	13839	6	2	0	
## 24	22884	6	2	1	
## 25	16978	7	1	1	
## 26	14803	8	2	0	

Your turn

Salary survey of computer professionals with objective to identify and quantify variables that determine salary differentials.

S Salary (Response)

X Experience, measured in years

E Education, 1 (High School/HS), 2 (Bachelor/BS), 3 (Advanced Degree/AD)

M Management 1 (is Manager), 0 (no Management Responsibility)

Example: Salary Survey Data

- **Experience:** We assume linearity, which means that each additional year is worth a fixed salary increment.
- **Education:** Can be used in a linear or categorical form.
 - Using the the variable in its raw form would assume that each step up in education is worth a fixed increment in salary. This may be too restrictive.
 - Using education as categorical variable can be done by defining **two indicator variables**. This allows to pick up the effect of education wether it is linear or not.
- **Management:** Is also an indicator variable, that allows to distinguish between management (1) an regular staff positions (0).

Indicator Variables

When using indicator variables to represent a set of categories, the number of these variables required is **one less than the number of categories**. For *education* we can create two indicators variables:

$$E_{i1} = \begin{cases} 1, & \text{if the } i\text{-th person is in the HS category} \\ 0, & \text{otherwise.} \end{cases}$$

$$E_{i2} = \begin{cases} 1, & \text{if the } i\text{-th person is in the BS category} \\ 0, & \text{otherwise.} \end{cases}$$

These two variables allow representing the three groups (HS, BS, AD).

HS: $E_1 = 1, E_2 = 0$, BS: $E_1 = 0, E_2 = 1$, AD: $E_1 = 0, E_2 = 0$

- The regression equation from the Salary Survey Data is:

$$S = \beta_0 + \beta_1 X + \gamma_1 E_1 + \gamma_2 E_2 + \delta_1 M + \epsilon$$

- The regression equation from the Salary Survey Data is:

$$S = \beta_0 + \beta_1 X + \gamma_1 E_1 + \gamma_2 E_2 + \delta_1 M + \epsilon$$

- There is a different valid regression equation for each of the six (three education and two management) categories.

Category	E	M	Regression Equation
1	1	0	$S = (\beta_0 + \gamma_1) + \beta_1 X + \epsilon$
2	1	1	$S = (\beta_0 + \gamma_1 + \delta_1) + \beta_1 X + \epsilon$
3	2	0	$S = (\beta_0 + \gamma_2) + \beta_1 X + \epsilon$
4	2	1	$S = (\beta_0 + \gamma_2 + \delta_1) + \beta_1 X + \epsilon$
5	3	0	$S = \beta_0 + \beta_1 X + \epsilon$
6	3	1	$S = (\beta_0 + \delta_1) + \beta_1 X + \epsilon$

Indicator Variables

```
d <- P130
d$E1 <- as.numeric(d$E == 1)
d$E2 <- as.numeric(d$E == 2)
mod <- lm(S ~ 1 + X + E1 + E2 + M, data=d)
summary(mod)

##
## Call:
## lm(formula = S ~ 1 + X + E1 + E2 + M, data = d)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -1884.60  -653.60   22.23   844.85  1716.47
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  11031.81    383.22   28.787 < 2e-16 ***
## X              546.18     30.52   17.896 < 2e-16 ***
## E1            -2996.21    411.75   -7.277 6.72e-09 ***
## E2              147.82     387.66    0.381  0.705
## M              6883.53    313.92   21.928 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 1027 on 41 degrees of freedom
## Multiple R-squared:  0.9568, Adjusted R-squared:  0.9525
## F-statistic: 226.8 on 4 and 41 DF,  p-value: < 2.2e-16
```

Your turn

Interpret the regression coefficients. Assume that the residual patterns are satisfactory.

Table 3

	<i>Dependent variable:</i>	
	S	
	(1)	(2)
X	546.184*** (30.519)	570.087*** (38.559)
E1	−2,996.210*** (411.753)	
E2	147.825 (387.659)	
E		1,578.750*** (262.322)
M	6,883.531*** (313.919)	6,688.130*** (398.276)
Constant	11,031.810*** (383.217)	6,963.478*** (665.695)
Observations	46	46
R ²	0.957	0.928
Adjusted R ²	0.953	0.923
Residual Std. Error	1,027.437 (df = 41)	1,312.789 (df = 42)
F Statistic	226.836*** (df = 4; 41)	179.627*** (df = 3; 42)

Note:

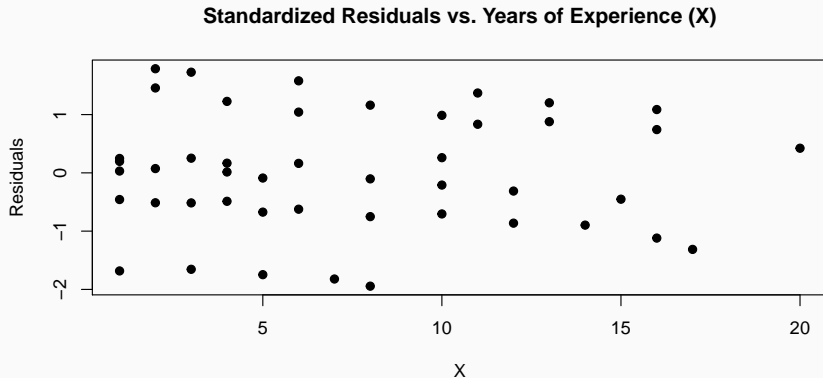
*p<0.1; **p<0.05; ***p<0.01

Before we continue we check the residuals

- 1) Residuals vs. Years of Experience
- 2) Residuals vs. Categories from Dummies

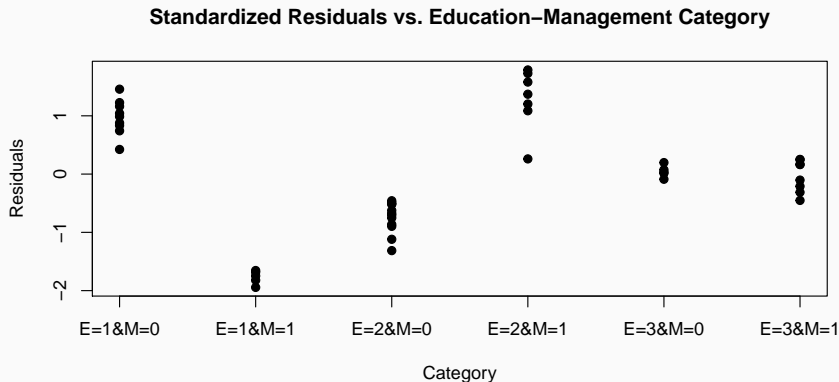
Regression Diagnostics

```
plot(x = d$X, y = rstandard(mod), pch=19,  
     ylab="Residuals", xlab = "X",  
     main = "Standardized Residuals vs. Years of Experience (X)")
```



Regression Diagnostics

```
d$cat <- factor((paste0("E=", d$E, "&M=", d$M)))
plot(x = as.numeric(d$cat), y = rstandard(mod), pch=19, xaxt="n",
     ylab="Residuals", xlab = "Category",
     main = "Standardized Residuals vs. Education-Management Category")
axis(1, at=1:6, labels=levels(d$cat))
```



What is wrong with the residuals:

- Depending on the category the residuals are almost entirely positive or negative.
- The **pattern of the residuals is highly moderated by the associated group** (education-management category). This makes it clear that the combinations of education and management have not been treated sufficiently in the model.
- The residual plots provide evidence that the effects of education and management status on salary determination are **not additive**.

The multiplicative pattern needs to be embedded in the model!

- Interaction effects are *multiplicative* effects that allow capturing nonadditive effects in variables.
- Interaction variables are products of existing indicator variables.
- Using the Salary Survey Data this can be achieved by creating the two interaction effects $(E_1 \cdot M)$ and $(E_2 \cdot M)$ and **adding** them to the model.
- The interaction effects **do not replace** the indicator variables.

Interaction Effects

```
mod <- lm(S ~ 1 + X + E1 + E2 + M + E1*M + E2*M, data=d)
summary(mod)
```

Your turn

Is that model sufficient?

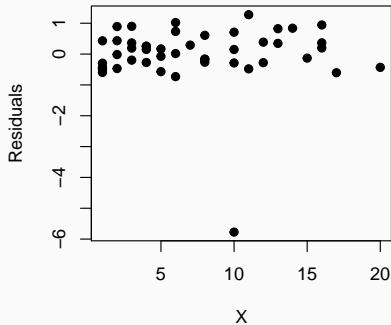
```
##
## Call:
## lm(formula = S ~ 1 + X + E1 + E2 + M + E1 * M + E2 * M, data = d)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -928.13  -46.21   24.33   65.88  204.89
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11203.434      79.065  141.698 < 2e-16 ***
## X              496.987       5.566   89.283 < 2e-16 ***
## E1            -1730.748     105.334  -16.431 < 2e-16 ***
## E2             -349.078      97.568   -3.578 0.000945 ***
## M              7047.412     102.589   68.695 < 2e-16 ***
## E1:M           -3066.035     149.330  -20.532 < 2e-16 ***
## E2:M           1836.488      131.167   14.001 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 173.8 on 39 degrees of freedom
## Multiple R-squared:  0.9988, Adjusted R-squared:  0.9986
## F-statistic: 5517 on 6 and 39 DF, p-value: < 2.2e-16
```

Regression Diagnostics

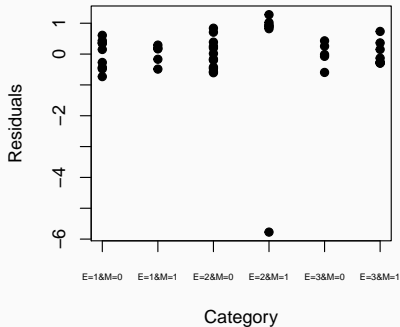
```
summary(rstandard(mod))
```

##	Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
##	-5.773499	-0.285974	0.149530	0.000711	0.418380	1.277256

**Standardized Residuals vs.
Years of Experience (X)**



**Standardized Residuals vs.
Education-Management Category**



Regression Diagnostics

```
d$res      <- residuals(mod)
d$res_std  <- rstandard(mod)
tail(d, n=15)
```

##	S	X	E	M	E1	E2	cat	res	res_std
## 32	23174	10	3	1	0	0	E=3&M=1	-46.71590	-0.2885031
## 33	23780	10	2	1	0	1	E=2&M=1	-928.12616	-5.7734989
## 34	25410	11	2	1	0	1	E=2&M=1	204.88683	1.2772560
## 35	14861	11	1	0	1	0	E=1&M=0	-78.54257	-0.4795797
## 36	16882	12	2	0	0	1	E=2&M=0	63.79979	0.3865825
## 37	24170	12	3	1	0	0	E=3&M=1	-44.68992	-0.2783867
## 38	15990	13	1	0	1	0	E=1&M=0	56.48341	0.3464916
## 39	26330	13	2	1	0	1	E=2&M=1	130.91281	0.8226405
## 40	17949	14	2	0	0	1	E=2&M=0	136.82577	0.8383399
## 41	25685	15	3	1	0	0	E=3&M=1	-20.65096	-0.1315808
## 42	27837	16	2	1	0	1	E=2&M=1	146.95178	0.9436856
## 43	18838	16	2	0	0	1	E=2&M=0	31.85175	0.1983361
## 44	17483	16	1	0	1	0	E=1&M=0	58.52238	0.3647875
## 45	19207	17	2	0	0	1	E=2&M=0	-96.13526	-0.6047045
## 46	19346	20	1	0	1	0	E=1&M=0	-66.42566	-0.4309873

```
d <- d[-33, ] # Remove problematic observation
```

Interaction Effects

Model Summary

## R	1.000	RMSE	6
## R-Squared	1.000	Coef. Var	0.392
## Adj. R-Squared	1.000	MSE	4504.951
## Pred R-Squared	1.000	MAE	51.794

Note: The level accuracy with which the model explains the data is very rare! Usually Goodness of fit indicators are worse.

ANOVA

##	Sum of				
##	Squares	DF	Mean Square	F	Sig.
## Regression	957607113.080	6	159601185.513	35427.955	0.0000
## Residual	171188.120	38	4504.951		
## Total	957778301.200	44			

Parameter Estimates

##	model	Beta	Std. Error	Std. Beta	t	Sig	lower	upper
##	(Intercept)	11199.714	30.533		366.802	0.000	11137.902	11261.525
##	X	498.418	2.152	0.557	231.640	0.000	494.062	502.774
##	E1	-1741.336	40.683	-0.304	-42.803	0.000	-1823.693	-1658.979
##	E2	-357.042	37.681	0.052	-9.475	0.000	-433.324	-280.761
##	M	7040.580	39.619	0.738	177.707	0.000	6960.376	7120.785
##	E1:M	-3051.763	57.674	-0.149	-52.914	0.000	-3168.519	-2935.008
##	E2:M	1997.531	51.785	0.103	38.574	0.000	1892.697	2102.364

Interaction Effects

Note: The notation is slightly different here as the equations are automatically generated. However, it does not really matter whether you use a β , δ or any other greek letter for the (interaction) effects.

```
mod <- lm(S ~ 1 + X + E1 + E2 + M + E1*M + E2*M, data=d)
equatiomatic::extract_eq(mod, use_coefs=F, intercept="beta", wrap=T)
```

$$S = \beta_0 + \beta_1(X) + \beta_2(E1) + \beta_3(E2) + \\ \beta_4(M) + \beta_5(E1 \times M) + \beta_6(E2 \times M) + \epsilon$$

```
equatiomatic::extract_eq(mod, use_coefs=T, coef_digits=4, wrap=T)
```

$$\hat{S} = 11199.7138 + 498.4178(X) - 1741.3359(E1) - 357.0423(E2) + \\ 7040.5801(M) - 3051.7633(E1 \times M) + 1997.5306(E2 \times M)$$

Your Turn

Compare the models `mod1`, `mod2` and `mod3`. Use them to calculate the base salaries (no experience) for each of the six possible education-management categories.

```
# Data Preparation
```

```
d <- P130[-33, ]  
d$cat <- factor((paste0("E=", d$E, "&M=", d$M)))  
d$E.fac <- factor(d$E)
```

```
# Model estimation
```

```
mod1 <- lm(S ~ 1 + X + E.fac + M + E.fac*M, data=d)  
mod2 <- lm(S ~ 1 + X + cat, data=d)  
mod3 <- lm(S ~ 1 + X + E.fac*M, data=d)
```

Interaction Effects

Category	E	M	Estimated Base Salary	95% CI Low	95% CI High
1	1	0	9458.378	9395.539	9521.216
2	2	1	19880.782	19814.090	19947.474
3	3	0	11199.714	11137.902	11261.525
4	1	1	13447.195	13382.933	13511.456
5	2	0	10842.672	10789.719	10895.624
6	3	1	18240.294	18182.503	18298.084

- All models lead to the **same estimates for the base salaries**. This shows that from a technical point using the cat variable (instead of the interaction effects) allows to capture the variation in the data.
- It is still **beneficial to use interaction effects** as we did, because this allows to separate the effects of the three sets of predictor variables education, management and education-management interaction.

A dataset may consists of **two or more distinct subsets**, which may require individual regression equations to avoid bias. Subsets may occur cross-sectional or over time and need to be treated differently:

■ Cross-Sectional Data

- 1 Each group has a separate regression model.
- 2 The models have the same intercept but different slopes.
- 3 the models have the same slope but different intercepts.

■ Time Series Data

- 1 Calendar Effects, e.g. Seasonality
- 2 Stability of regression parameters over time

Example: Preemployment Test

P140

##	TEST	RACE	JPERF
## 1	0.28	1	1.83
## 2	0.97	1	4.59
## 3	1.25	1	2.97
## 4	2.46	1	8.14
## 5	2.51	1	8.00
## 6	1.17	1	3.30
## 7	1.78	1	7.53
## 8	1.21	1	2.03
## 9	1.63	1	5.00
## 10	1.98	1	8.04
## 11	2.36	0	3.25
## 12	2.11	0	5.30
## 13	0.45	0	1.39
## 14	1.76	0	4.69
## 15	2.09	0	6.56
## 16	1.50	0	3.00
## 17	1.25	0	5.85
## 18	0.72	0	1.90
## 19	0.42	0	3.85
## 20	1.53	0	2.95

Your turn

TEST Score on the preemployment test.

RACE Dummy to indicate if individual is part of a minority (1) or not (0).

JPERF Job Performance Ranking after 6 weeks on the job.

Example: Preemployment Test

For simplicity and generality we refer to the job performance as Y and the score on the preemployment test as X . We want to compare the following two models:

Model 1 (Pooled):

$$y_{ij} = \beta_0 + \beta_1 x_{ij} + \epsilon_{ij}$$

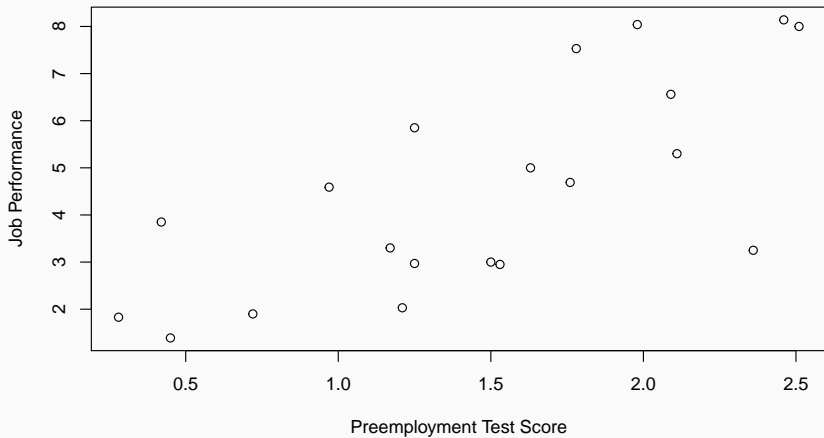
Model 2 (Minority):

$$y_{i1} = \beta_{01} + \beta_{11} x_{i1} + \epsilon_{i1}$$

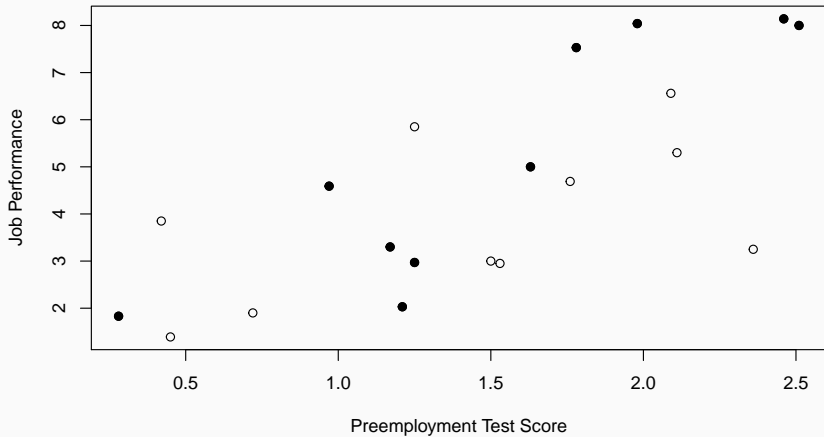
Model 2 (non Minority):

$$y_{i2} = \beta_{02} + \beta_{12} x_{i2} + \epsilon_{i2}$$

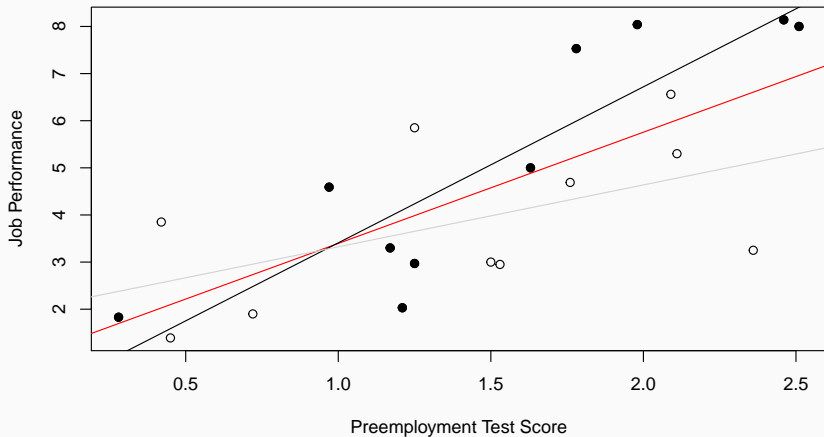
Example: Preemployment Test



Example: Preemployment Test

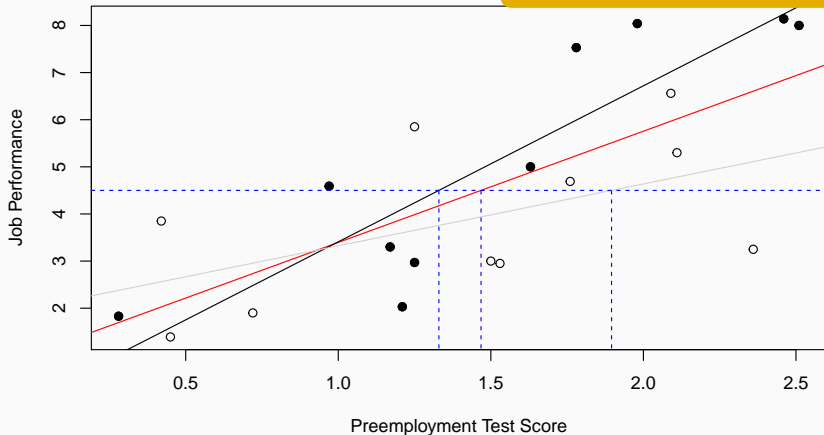


Example: Preemployment Test



Example: Preemployment Test

When different models are required for the groups, this would imply that the required score in the preemployment test that is needed to result in (minimum) job performance (horizontal dashed line) needs to be distinguished by group (vertical dashed lines).



Models with different Slopes and different Intercepts

- What we want to test for the Preemployment Test data is for differences in intercept and slope using the following Null.

$$H_0 : \beta_{11} = \beta_{12}, \beta_{01} = \beta_{02}$$

- This test can be performed using an **interaction term** by using a variable z_{ij} that takes the value 1 if an individual is part of a minority group and 0 otherwise. This leads to two relevant models:

Model 1 (Pooled): $y_{ij} = \beta_0 + \beta_1 x_{ij} + \epsilon_{ij}$

Model 3 (Interaction): $y_{ij} = \beta_0 + \beta_1 x_{ij} + \gamma z_{ij} + \delta(z_{ij} \cdot x_{ij}) + \epsilon_{ij}$

- This model is **equivalent** to the previously discussed Model 2.

Models with different Slopes and different Intercepts

	Model 1	Model 2	Model 2	Model 3
	Pooled	Minority	White	Interaction
(Intercept)	1.03 (0.87)	0.10 (1.04)	2.01 (1.13)	2.01 (1.05)
TEST	2.36*** (0.54)	3.31*** (0.62)	1.31 (0.72)	1.31 (0.67)
RACE				-1.91 (1.54)
TEST:RACE				2.00 (0.95)
R ²	0.52	0.78	0.29	0.66
Adj. R ²	0.49	0.75	0.20	0.60
Num. obs.	20	10	10	20

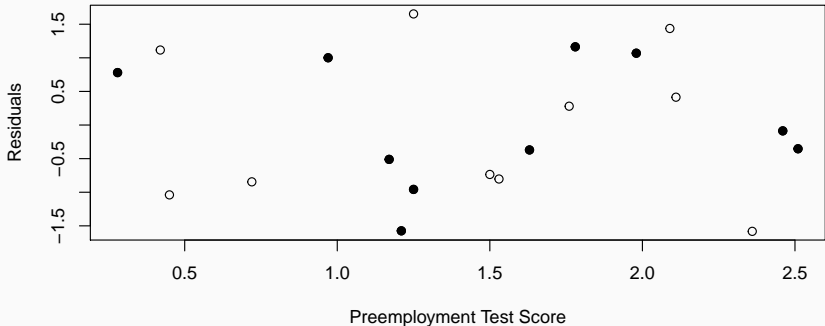
*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 5

- Model 1 can be seen as a restricted version (RM) of model 3, the full model (FM), with $\gamma = \delta = 0$.

Models with different Slopes and different Intercepts

```
df <- cbind(P140, res = rstandard(mod3))  
plot(x=df$TEST, y=df$res,  
     ylab="Residuals", xlab="Preemployment Test Score")  
points(df$TEST[df$RACE == T], df$res[df$RACE == T], pch=19)
```



Models with different Slopes and different Intercepts

- The framework using the models as F-Test for comparison.

$$F = \frac{[SSE(RM) - SSE(FM)]/1}{SSE(FM)/16}$$

```
(SSE_RM <- sum(residuals(mod1)^2))
```

```
## [1] 45.5683
```

```
(SSE_FM <- sum(residuals(mod3)^2))
```

```
## [1] 31.65547
```

```
(F_stat <- ((SSE_RM - SSE_FM)/2)/(SSE_FM/16))
```

```
## [1] 3.516061
```

```
pf(F_stat, df1=2, df2=16, lower.tail=FALSE)
```

```
## [1] 0.05423606
```

Your turn

Interpret the F-Test.

Can you conclude that the relationship is different for the two groups, so that two different equations (intercept + slope) are required?

Models with same Slope and different Intercepts

- Assuming we have a reason to believe that only the intercepts for the two groups are different can be achieved using the indicator variable (and omitting the interaction term).

Model 1 (Pooled):

$$y_{ij} = \beta_0 + \beta_1 x_{ij} + \epsilon_{ij}$$

Model 4 (Indicator only):

$$y_{ij} = \beta_0 + \beta_1 x_{ij} + \gamma z_{ij} + \cancel{\delta(z_{ij} \cdot x_{ij})} + \epsilon_{ij}$$

- In the case where $z_{ij} = 1$ (which indicates the non-minority group) the coefficient γ can be added to the intercept β_0 to obtain the effective intercept for that respective group.
- The resulting models represent **two parallel lines** (same slopes) with intercepts β_0 and $\beta_0 + \gamma$.

Models with same Slope and different Intercepts

```
mod4 <- lm(JPERF ~ 1 + TEST + RACE, data=P140)
```

- Significance can be tested using the *F*-Test. As the FM and RM differ by one parameter, results are equivalent to the *t*-Test.

	Model 1	Model 2	Model 2	Model 3	Model 4
	Pooled	Minority	White	Interaction	Indicator
(Intercept)	1.03 (0.87)	0.10 (1.04)	2.01 (1.13)	2.01 (1.05)	0.61 (0.89)
TEST	2.36*** (0.54)	3.31*** (0.62)	1.31 (0.72)	1.31 (0.67)	2.30*** (0.52)
RACE				-1.91 (1.54)	1.03 (0.69)
TEST:RACE				2.00 (0.95)	
R ²	0.52	0.78	0.29	0.66	0.57
Adj. R ²	0.49	0.75	0.20	0.60	0.52
Num. obs.	20	10	10	20	20

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 6

Models with different Slopes and same Intercept

- Finally we can hypothesize that the two groups have the same intercept β_0 but different slopes, which can be done by including only the interaction.

Model 1 (Pooled): $y_{ij} = \beta_0 + \beta_1 x_{ij} + \epsilon_{ij}$

Model 5 (Interaction only): $y_{ij} = \beta_0 + \beta_1 x_{ij} + \gamma z_{ij} + \delta(z_{ij} \cdot x_{ij}) + \epsilon_{ij}$

```
mod5 <- lm(JPERF ~ 1 + TEST + RACE:TEST, data=P140)
```

- Inference for the δ can be carried out using the *F*-Test or the *t*-Test. The FM and RM again only differ by one parameter.

Systems of Regression Equations

- The final results for all discussed cases for the preemployment test data look like follows.

	Model 1	Model 2	Model 2	Model 3	Model 4	Model 5
	Pooled	Minority	White	Full Interaction	Indicator	Interaction
(Intercept)	1.03 (0.87)	0.10 (1.04)	2.01 (1.13)	2.01 (1.05)	0.61 (0.89)	1.12 (0.78)
TEST	2.36*** (0.54)	3.31*** (0.62)	1.31 (0.72)	1.31 (0.67)	2.30*** (0.52)	1.83** (0.54)
RACE				-1.91 (1.54)	1.03 (0.69)	
TEST:RACE				2.00 (0.95)		0.92* (0.40)
R ²	0.52	0.78	0.29	0.66	0.57	0.63
Adj. R ²	0.49	0.75	0.20	0.60	0.52	0.59
Num. obs.	20	10	10	20	20	20

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

Table 7

- Another interesting field of study is temporal structure in the data, which could fill a whole course by itself. Therefore we only briefly look at two ideas.

1) Calendar Effects, e.g. Seasonality

- Can be modeled by including time as regressor, e.g. in the form of (multiple) indicators for e.g. Week/Month/Quarter/Year
- The number of indicator variables is $m - 1$ where m is the frequency of the time effects (e.g. $m = 4$ for Quarters).

2) Stability of Parameters over Time

- By combining indicator and interaction terms one can model intertemporal and interspatial relationships. Insignificance of the interactions with all indicators then provides evidence stability over time.