

y_{up}

нен-и метод моментов

+ геометрический

где построение алгоритма - 20 СИ

$$n = 10000$$

$y_1, \dots, y_n \sim \text{клав.expo}(\lambda)$

a) $\hat{\lambda}$ с помощью метода моментов.

b) построить 95% СИ для клав-го λ
[асимптотический]

$$E(y_i) = \frac{1}{\lambda}$$

$$= \int_0^{\infty} \lambda \cdot \exp(-\lambda y) dy$$

a)

$$\frac{y_1 + \dots + y_n}{n} = \hat{\lambda}$$

$$\hat{\lambda} = \frac{1}{\bar{y}}$$

$$n = 10000$$

$$y_1 + \dots + y_n = 20000$$

$$\hat{\lambda} = \frac{1}{\frac{20000}{10000}} = 0.5$$

мас 1

вих. [геометрический]:

но УПТ. $\bar{y} \approx N(\cdot, \cdot)$

$$\hat{\lambda} = g(\bar{y}) = \frac{1}{\bar{y}} \approx N(\cdot, \cdot)$$

мас 2

задача: приложение к рег. Геодезии

СИ

мас 3

Mom 1. no UMT. $\bar{y} \stackrel{\text{approx}}{\sim} N(?, ?)$

$$E(\bar{y}) = E\left(\frac{y_1 + \dots + y_n}{n}\right) = \frac{1}{n}(E(y_1) + \dots + E(y_n)) = \\ = \frac{1}{n}\left(\frac{1}{\lambda} + \frac{1}{\lambda} + \dots + \frac{1}{\lambda}\right) = \frac{1}{\lambda}$$

$$\text{Var}(\bar{y}) = \text{Var}\left(\frac{y_1 + \dots + y_n}{n}\right) = \frac{1}{n^2} \text{Var}(y_1 + \dots + y_n) = \\ = \frac{1}{n^2} \cdot n \cdot \frac{1}{\lambda^2} = \frac{1}{n\lambda^2}$$

$$\bar{y} \stackrel{\text{approx}}{\sim} N\left(\frac{1}{\lambda}; \frac{1}{n\lambda^2}\right)$$

$$\lambda \xrightarrow{\text{prob}} 1 \quad \lambda = 0,5 \quad \lambda \approx 0,5$$

$$\bar{y} \stackrel{\text{approx}}{\sim} N\left(\frac{1}{0,5}, \frac{1}{10000 \cdot 0,5^2}\right) = N(2; \frac{1}{2500})$$

Mom 2

$$\hat{\lambda} = \frac{1}{\bar{y}}$$

$$\approx g(u) + g'(u) \cdot (\bar{y} - u) =$$

mein Testwerte
z.B. $u = 1$ dann $\bar{y} = \mu = \frac{1}{\lambda}$

$$g(u) = \frac{1}{u}$$

$$g'(u) = -\frac{1}{u^2}$$

$$= \frac{1}{\mu} + \left(-\frac{1}{\mu^2}\right) \cdot \left(\bar{y} - \mu\right) =$$

$$= \lambda - \lambda^2 \cdot \left(\bar{y} - \frac{1}{\lambda}\right)$$

Konkav
aus
aus
no!
 \bar{y} - CB & Boges
aus
 λ - Konkav

no global - energy

$$\hat{\lambda} \stackrel{\text{approx}}{\sim} N(?, ?)$$

$$E(\hat{\lambda}) \approx E\left(\lambda - \lambda^2 \cdot \left(\bar{y} - \frac{1}{\lambda}\right)\right) = \\ = \lambda - \lambda^2 \cdot \left(E(\bar{y}) - \frac{1}{\lambda}\right) = \lambda - \lambda^2 \cdot 0 = \lambda$$

$$\text{Var}(\hat{\lambda}) \approx \text{Var}\left(\lambda - \lambda^2 \cdot \left(\bar{y} - \frac{1}{\lambda}\right)\right) =$$

$$= \text{Var}\left(-\lambda^2 \cdot \bar{y}\right) = \lambda^4 \cdot \text{Var}(\bar{y}) =$$

$\lambda = \text{const}$

$$= \lambda^q \cdot \frac{1}{n\lambda^2} = \frac{\lambda^2}{n}$$

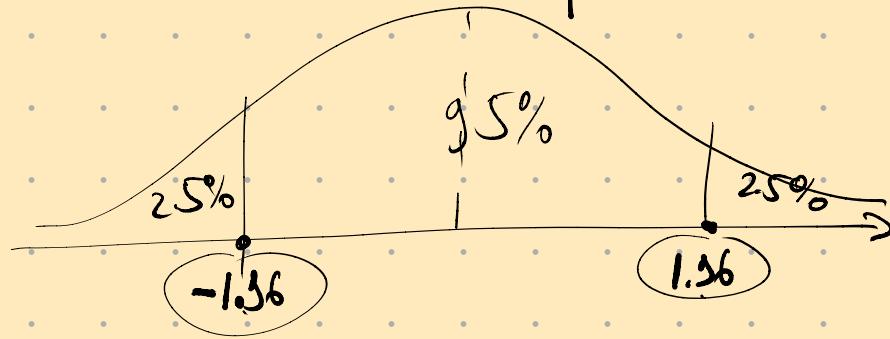
$$\hat{\lambda} \stackrel{\text{approx}}{\sim} N\left(\lambda; \frac{\lambda^2}{n}\right)$$

Mar 3

$$\text{CI} \quad N(\mu; \sigma^2) \rightarrow N(0; 1)$$

$$\frac{\hat{\lambda} - \lambda}{\sqrt{\frac{\lambda^2}{n}}} \stackrel{\text{approx}}{\sim} N(0; 1)$$

$$\frac{\hat{\lambda} - \lambda}{\sqrt{\frac{\lambda^2}{n}}} \stackrel{\text{approx}}{\sim} N(0; 1) \xrightarrow[\text{prob dist}]{} \begin{matrix} \uparrow \\ \text{p. norm. gdw } N(0; 1) \end{matrix}$$



$$P\left(-1.96 \leq \frac{\hat{\lambda} - \lambda}{\sqrt{\frac{\lambda^2}{n}}} \leq 1.96\right) \xrightarrow{n \rightarrow \infty} 0.95$$

$$-1.96 \leq \frac{\hat{\lambda} - \lambda}{\sqrt{\frac{\lambda^2}{n}}} \leq 1.96 \quad \text{x-1}$$

$$-1.96 \leq \frac{\lambda - \hat{\lambda}}{\sqrt{\frac{\lambda^2}{n}}} \leq 1.96$$

$$-1.96 \sqrt{\frac{\lambda^2}{n}} \leq \lambda - \hat{\lambda} \leq 1.96 \sqrt{\frac{\lambda^2}{n}}$$

$$\hat{\lambda} - 1.96 \sqrt{\hat{\lambda}^2/n} \leq \lambda \leq \hat{\lambda} + 1.96 \sqrt{\hat{\lambda}^2/n}$$

верно

$$P(\lambda \in [\hat{\lambda} - 1.96 \sqrt{\hat{\lambda}^2/n}, \hat{\lambda} + 1.96 \sqrt{\hat{\lambda}^2/n}]) \xrightarrow{n \rightarrow \infty} 0.95$$

asy CT give $\lambda_{(N)}$

$$[\hat{\lambda} - 1.96 \sqrt{\hat{\lambda}^2/n}, \hat{\lambda} + 1.96 \sqrt{\hat{\lambda}^2/n}]$$

$$\hat{\lambda} = 0.5$$

$$\hat{\lambda} = \frac{1}{g}$$

нужн. коеф
недостаточн.

$$[0.5 - 2 \cdot \sqrt{0.5^2/10000}, 0.5 + 2 \cdot \sqrt{0.5^2/10000}]$$

$$[0.5 - 0.01, 0.5 + 0.01]$$

$$[0.49, 0.51]$$

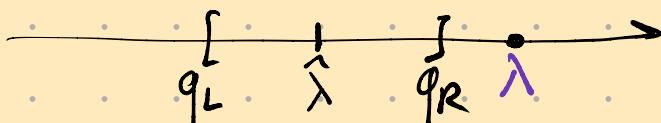
реально:

$$P(\lambda \in [0.49; 0.51]) = 0.95$$

λ -коеф. коеф.

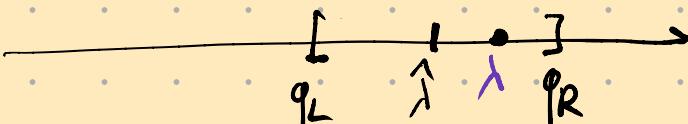
эксп. N1

$$n=10000$$



эксп. N2

$$n=10000$$



Случаи нахождения
всех экспериментов

эксп. N Nepp

Nepp, вероятно в 95%
всех экспериментов
коэф. коеф λ попадет в 1-е

Варианте решения задачи №1.

$$\bar{y} \underset{\text{approx}}{\sim} N\left(\frac{1}{\lambda}; \frac{1}{n\lambda^2}\right)$$

$$\frac{\bar{y} - \frac{1}{\lambda}}{\sqrt{\frac{1}{n\lambda^2}}} \underset{\text{approx}}{\sim} N(0; 1)$$

$$\underset{\text{approx}}{\sim} N(0; 1)$$

$$\bar{y} = \frac{1}{\lambda} \quad \text{approx } N(0; 1) \quad \hat{\lambda} \approx \lambda$$

$$\sqrt{\frac{1}{n\lambda^2}}$$

$$P(-1,96 \leq \frac{\bar{y} - 1/\lambda}{\sqrt{1/n\lambda^2}} \leq 1,96) \xrightarrow{n \rightarrow \infty} 0,95$$

↓ римские крк-бо

$$P\left(\frac{1}{\lambda} \in [\bar{y} - 1,96 \cdot \sqrt{1/n\lambda^2}, \bar{y} + 1,96 \cdot \sqrt{1/n\lambda^2}]\right) \xrightarrow{0,95}$$

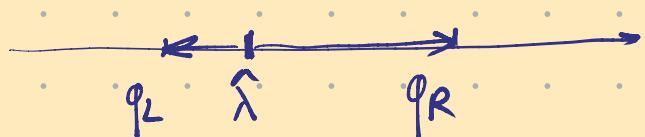
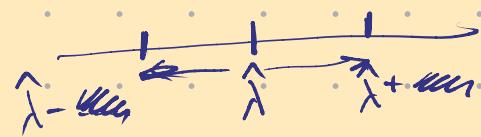
$$P\left(\lambda \in \left[\frac{1}{\bar{y} + 1,96 \cdot \sqrt{1/n\lambda^2}}; \frac{1}{\bar{y} - 1,96 \cdot \sqrt{1/n\lambda^2}}\right]\right) \xrightarrow{n \rightarrow \infty} 0,95$$

95%asy CI

$\text{CI}_{N1} \left[\frac{1}{\bar{y} + 1,96 \cdot \sqrt{1/n\lambda^2}}; \frac{1}{\bar{y} - 1,96 \cdot \sqrt{1/n\lambda^2}} \right]$

(CI N1)

симметр.↑
отк-ко ↑



$y_{1:n}$

$y_1, \dots, y_n \sim \text{некая } U[0; a]$

$[E(y_i)]$

a)

а) среднее момента

δ)

асы 95% CI симметр. без отк-ко а.

$$E(y_i) = \frac{a+0}{2}$$

$$\frac{y_1 + \dots + y_n}{n} = \frac{\hat{a} + 0}{2}$$

$$\hat{a} = 2\bar{y}$$

a

$$\bar{y} \sim N(?, ?)$$

$$E(\bar{y}) = E(y_1) = \frac{\alpha}{2}$$

$$\text{Var}(\bar{y}) = \frac{\text{Var}(y_1)}{n} = \frac{\alpha^2/12}{n}$$

$$\text{Var}(y_1) = \underbrace{E(y_1^2)}_{\int_0^a y^2 f(y) dy} - (E(y_1))^2$$

$$= \int_0^a y^2 \cdot \frac{1}{\alpha} dy$$

$$\bar{y} \underset{n \rightarrow \infty}{\sim} N\left(\frac{\alpha}{2}; \frac{\alpha^2}{12n}\right)$$

$$\hat{\alpha} = 2 \cdot \bar{y} \quad \leftarrow \text{име интересное значение}$$

$$\hat{\alpha} \underset{n \rightarrow \infty}{\sim} N\left(\alpha; \frac{\alpha^2}{3n}\right) \quad \begin{array}{l} \text{одновероятное} \\ \text{с помощью Гаусса} \\ \text{не нужно} \end{array}$$

$$\frac{\hat{\alpha} - \alpha}{\sqrt{\alpha^2/3n}} \underset{n \rightarrow \infty}{\sim} N(0; 1)$$

$$\frac{\hat{\alpha} - \alpha}{\sqrt{\hat{\alpha}^2/3n}} \underset{n \rightarrow \infty}{\sim} N(0; 1)$$

$$P(-1.96 \leq \frac{\hat{\alpha} - \alpha}{\sqrt{\hat{\alpha}^2/3n}} \leq 1.96) \xrightarrow{n \rightarrow \infty} 0.95$$

$$P(a \in [\hat{\alpha} - 1.96 \sqrt{\hat{\alpha}^2/3n}; \hat{\alpha} + 1.96 \sqrt{\hat{\alpha}^2/3n}]) \xrightarrow{n \rightarrow \infty} 0.95$$

$$\hat{\alpha} = \bar{y}$$

y_{np}

y_1, y_2, \dots

$y_n \sim \text{Geom}(p)$

y_i - число попыток до 1го
успеха ($p = \text{вероятность}$
успеха в однократной
попытке)

$n = 10000$

a) \hat{p} вероятность успехов

b) аsey 95% CI для p (см. слайд оценка \hat{p})

$$E(y_i) = \frac{1}{p}$$

$$\underbrace{y_1 + \dots + y_n}_n = \frac{1}{\hat{p}}$$

$$\hat{p} = \frac{1}{\bar{y}}$$

но YMT $\bar{y} \approx N(?, ?)$

$$E(\bar{y}) = E(y_i) = \frac{1}{p} = \mu$$

$$\text{Var}(\bar{y}) = \text{Var}\left(\frac{y_1 + \dots + y_n}{n}\right) = \frac{1}{n^2} n \cdot \text{Var}(y_i) = \frac{\text{Var}(y_i)}{n}$$

$$\text{Mean } \bar{y} = E(y_i)$$

Медиана на
б.р. $\mu = E(y_i)$

$$= \frac{1-p}{p^2} \cdot \frac{1}{n}$$

функция - мерид.

$$g(u) = \frac{1}{u}$$

$$g'(u) = -\frac{1}{u^2}$$

$$\hat{p} = g(\bar{y}) \approx$$

$$\approx g(u) + g'(u) \cdot (\bar{y} - u)$$

$$\hat{p} \approx g\left(\frac{1}{p}\right) + g'\left(\frac{1}{p}\right) \cdot \left(\bar{y} - \frac{1}{p}\right) = p - p^2 \cdot \left(\bar{y} - \frac{1}{p}\right)$$

$$E(\hat{p}) \approx E\left(p - p^2 \cdot \left(\bar{y} - \frac{1}{p}\right)\right) = p - p^2 \cdot \underbrace{\left(E(\bar{y}) - \frac{1}{p}\right)}_{=0} = p$$

$$\text{Var}(\hat{p}) \approx \text{Var}\left(p - p^2 \cdot \left(\bar{y} - \frac{1}{p}\right)\right) =$$

$$= \text{Var}\left(-p^2 \cdot \bar{y}\right) = p^4 \cdot \text{Var}(\bar{y}) = p^4 \cdot \frac{(1-p)}{n^2} =$$

$$= p^2 \cdot (1-p) \frac{1}{n}$$

$$\hat{p} \text{ approx } N(p; \frac{p^2 \cdot (1-p)}{n})$$

$$\frac{\hat{p} - p}{\sqrt{\frac{p^2 \cdot (1-p)}{n}}} \text{ approx } N(0; 1)$$

$$\frac{\hat{p} - p}{\sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}} \text{ approx } N(0; 1)$$

$$P(-1.96 \leq \frac{\hat{p} - p}{\sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}} \leq 1.96) \xrightarrow{n \rightarrow \infty} 0.95$$

$$P(p \in [\hat{p} - 1.96 \sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}, \hat{p} + 1.96 \sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}]) \xrightarrow{n \rightarrow \infty} 0.95$$

asy 95% CI für p : $[\hat{p} - 1.96 \sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}, \hat{p} + 1.96 \sqrt{\frac{\hat{p}^2 \cdot (1-\hat{p})}{n}}]$

erste Berechnung: $\sum_{i=1}^n y_i = 20000 \quad n=10000 \Rightarrow \bar{y}=2 \quad \hat{p} = \frac{1}{\bar{y}} = 0.5$

$$[0.5 - 2 \sqrt{\frac{0.5 \cdot 0.5}{10000}}; 0.5 + 2 \cdot \sqrt{\frac{0.5 \cdot 0.5}{10000}}]$$

$$[0.5 - 0.007; 0.5 + 0.007]$$

realis. $\rightarrow [0.493; 0.507]$
gab 20 verschiedene.

