

Condition for an unbiased estimate

- Simple regression:

$$y_i = \beta_0 + \beta_1 x_i + u_i$$

We can estimate

in any case!

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$

But will estimates be good?

- On lecture: $E(\hat{\beta}_1 | x) = \beta_1$

$$E(E(\hat{\beta}_1 | x)) = E\hat{\beta}_1 = \beta_1$$

unbiased estimate

if $E(u_i | x) = 0$

↑
our best guess about u_i (based on information about x) is zero.

① weaker
 u and x are independent \leftarrow $E(u_i | x_i) = 0$ \rightarrow ② weaker $E(u) = 0$

③ $\text{corr}(u, x) = 0$

Week 2

we are here

①

	y_3	y_3	y_3
weather u	-1	1	0
mood x	0	0	1

$$E(u|x) = \begin{cases} 0 & x=0 \\ 0 & x=1 \end{cases}$$

But

$$E(x|u) = \begin{cases} 0 & u=\{-1\} \\ 1 & u=0 \end{cases}$$

→ not independent

②

	$1/3$	$1/3$	$1/3$
weather u	-1	0	1
mood x	0	1	1

$$E(u|x) = \begin{cases} -1, & x=0 \\ \frac{1}{2}, & x=1 \end{cases}$$

but

$$E(u)=0$$

③

$$X \sim U(-1, 1) \Rightarrow \bullet \text{cov}(x, u) = E(xu) - E_x E_u$$

$$u = x^2$$

$$\bullet E(xu) = E(x^3) \Leftrightarrow$$

$$\Leftrightarrow \int_{-1}^1 x^3 \cdot \frac{1}{2} dx = \frac{1}{2} \cdot \frac{1}{4} x^4 \Big|_{-1}^1 = 0$$

$$\bullet E x = 0 \Rightarrow \text{cov}(x, u) = 0$$

But $E(u|x) = E(x^2|x) = x^2 \neq 0$
(usually)

so $\text{cov}(x,u) = 0 \not\Rightarrow E(u|x) = 0$

$E(u) = 0 \not\Rightarrow E(u|x) = 0$

$E(u|x) = 0 \not\Rightarrow \text{independ.}$

$$\begin{cases} \text{cov}(u,x) = E(ux) - E u E x = E(ux) = 0 \\ E(u|x) = 0 \Rightarrow E(E(u|x)) = E u = 0 \end{cases}$$

$$E(ux) = E(E(ux|x)) = E(X E(u|x)) = 0$$

$$\Rightarrow E(u|x) = 0 \Rightarrow \text{cov}(u,x) = 0$$

usefull

$$\text{cov}(x,u) \neq 0 \Rightarrow E(u|x) \neq 0$$

$$\text{OK, } \forall i=1, \dots, n \quad E(u_i|x_i) = 0$$

Why it is so important?

Sometimes we can build a model where

the assumption does not hold

this assumption goes wrong.

Task 1: Missing Var.

salary ↴
educ ↴
profession ↴

real life: $y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + u_i$

Our model $y_i = \beta_0 + \beta_1 X_{i1} + v_i$

$E(u_i | X_i) = 0$ Is $E(\hat{\beta}_1) = \beta_1$?

$\text{cov}(X_{i1}, X_{i2}) \neq 0$

Our estimation $\hat{\beta}_1 = \frac{\sum (X_{i1} - \bar{X}_1)(y_i - \bar{y})}{\sum (X_{i1} - \bar{X}_1)^2} =$
or the slope

$$= \frac{\sum (X_{i1} - \bar{X}_1)y_i}{\sum X_{i1}(X_{i1} - \bar{X}_1)} =$$

$$= \frac{\sum (X_{i1} - \bar{X}_1)(\beta_0 + \beta_1 X_{i1} + v_i)}{\sum X_{i1}(X_{i1} - \bar{X}_1)} =$$

$$= \frac{\beta_0 \cdot 0}{\sum X_{i1}(X_{i1} - \bar{X}_1)} + \beta_1 \frac{\sum X_{i1}(X_{i1} - \bar{X}_1)}{\sum X_{i1}(X_{i1} - \bar{X}_1)} + \frac{\sum v_i(X_{i1} - \bar{X}_1)}{\sum (X_{i1} - \bar{X}_1)^2} =$$

$$= \beta_1 + \frac{\sum v_i(X_{i1} - \bar{X}_1)}{\sum (X_{i1} - \bar{X}_1)^2}$$

$$E(\hat{\beta}_1 | X_1) = \beta_1 + \frac{\sum (X_{i1} - \bar{X}_1) E(v_i | X_{i1})}{\sum (X_{i1} - \bar{X}_1)^2} \neq \beta_1$$

in fact $V_i = \beta_2 X_{i2} + u_i$

$$\begin{aligned} E(V_i | X_{i1}) &= \beta_2 E(X_{i2} | X_{i1}) + E(u_i | X_{i1}) = \\ &= \beta_2 E(X_{i2} | X_{i1}) \neq 0 \end{aligned}$$

because $\text{cov}(X_{i1}, X_{i2}) \neq 0 \Rightarrow$

$\hat{\beta}_1$ is biased in this case

you can only miss regressors

which are uncorrelated with yours

Task 2: hands shake

real model: $y_i = \beta_0 + \beta_1 x_i + u_i$

$$E(u_i | x_i) = 0$$

But observed $\tilde{x}_i = x_i + v_i$

v_i is independent
with $x_i, u_i \dots$

so available model $y_i = \beta_0 + \beta_1 \tilde{x}_i + \epsilon_i \Leftrightarrow$

$$\therefore E(\hat{\beta}_1) = \beta_1 + \frac{\sum (\tilde{x}_i - \bar{x}) E(\epsilon_i | \tilde{x}_i)}{\sum (\tilde{x}_i - \bar{x})^2}$$

$$\text{in our task } e_i = \beta_1 X_i - \beta_1 \tilde{X} + u_i = \\ = -\beta_1 v_i + u_i$$

$$E(e_i | \tilde{X}_i) = E(-\beta_1 v_i + u_i | \tilde{X}_i) \neq 0$$

because $\text{cov}(-\beta_1 v_i + u_i, \tilde{X}_i) =$

$$= \text{cov}(-\beta_1 v_i + u_i, X_i + V_i) =$$

$$= -\beta_1 \text{cov}(v_i, X_i) - \beta_1 \text{cov}(v_i, V_i) +$$

$$+ \text{cov}(u_i, X_i) + \text{cov}(u_i, V_i) \quad \begin{matrix} \parallel & \parallel \\ 0 & 0 \end{matrix} \quad \begin{matrix} * \\ 0 \end{matrix} \quad = -\beta_1 \text{Var}[V_i]$$

so shaking hand lead to
biased estimate

We discussed $E(\hat{\beta})$ but
what about $\text{Var}(\hat{\beta}) - ?$

It is also important because we need
to calculate the error of our estimation
(and to test hypotheses)

Task 3

$$y_i = \beta_0 + \beta_1 x_i + u_i$$

$$a) \hat{\beta}_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} = \frac{\sum (x_i - \bar{x})(\beta_0 + \beta_1 x_i + u_i)}{\sum (x_i - \bar{x})^2}$$

$$= 0 + \beta_1 \frac{\sum x_i(x_i - \bar{x})}{\sum x_i(x_i - \bar{x})} + \frac{\sum u_i(x_i - \bar{x})}{\sum (x_i - \bar{x})^2} =$$

$$= \beta_1 + \frac{\sum u_i(x_i - \bar{x})}{\sum (x_i - \bar{x})^2}$$

$$\text{Var}(\hat{\beta}_1 | x) = 0 + \text{Var}\left(\frac{\sum u_i(x_i - \bar{x})}{\sum (x_i - \bar{x})^2} | x\right) =$$

$$= \frac{\text{Var}(\sum u_i(x_i - \bar{x}) | x)}{[\sum (x_i - \bar{x})^2]^2} \quad \rightarrow \begin{cases} \text{we need} \\ \text{some assumpt.} \\ \text{about } u_i \end{cases}$$

$$\text{Var}(u_i | x) = \sigma^2$$

$$\Leftrightarrow \frac{\sum \text{Var}(u_i(x_i - \bar{x}) | x)}{[\sum (x_i - \bar{x})^2]^2} = \frac{\sum (x_i - \bar{x})^2 \text{Var}(u_i | x)}{[\sum (x_i - \bar{x})^2]^2} =$$

$$= \frac{\sigma^2 \sum (x_i - \bar{x})^2}{[\sum (x_i - \bar{x})^2]^2} = \frac{\sigma^2}{\sum (x_i - \bar{x})^2} = \text{Var}(\hat{\beta}_1 | x)$$

$$\text{Attention!} \quad \sigma^2 \quad (u_i \sim F(0, \sigma^2))$$

should be known for this formula

8) If we don't know true variance of u_i then we don't know $\text{Var}(\hat{\beta}_1 | x)$: (

What can we do? Estimate it!

$$\sigma^2 \xrightarrow{\text{unknown}} \hat{\sigma}^2 \xrightarrow{\text{estimation}}$$

$$\Rightarrow \text{var}(\hat{\beta}_1 | x_i) \rightarrow \widehat{\text{var}}(\hat{\beta}_1 | x_i) = \frac{\hat{\sigma}^2}{\sum(x_i - \bar{x})^2}$$

$$\hat{\sigma}^2 = \frac{1}{n-2} \sum \hat{u}_i^2 = \frac{\text{RSS}}{n-2}$$

unbiased,

$$E(\hat{\sigma}^2) = \sigma^2$$

$$\Rightarrow \widehat{\text{var}}(\hat{\beta}_1 | x) = \frac{\text{RSS}}{\sum(x_i - \bar{x})^2}$$

$n-k$ for
(k regressors)

Task 4

$$Y_i = \beta_0 + \beta_1 x_i + u_i$$

$$x_i \quad y_i \quad E(u_i | x_i) = 0$$

$$1 \quad 1$$

$$2 \quad 1 \quad \text{Var}(u_i | x_i) = 0,5$$

$$3 \quad 4 \quad \text{cov}(u_i, u_j) = 0$$

a) $\text{Var}(\beta_1), \text{Var}(\hat{\beta}_1 | x), \widehat{\text{Var}}(\hat{\beta}_1 | x)$

• $\text{Var}(\beta_1) = 0 \quad \leftarrow \text{fixed param.}$

• $\text{Var}(\hat{\beta}_1 | x) = \frac{5^2}{\sum(x_i - \bar{x})^2} = \frac{0,5}{(1-2)^2 + 0 + (3-2)^2} = \frac{0,5}{2} =$

= 0,25

• $\widehat{\text{Var}}(\hat{\beta}_1 | x) = \frac{\text{RSS}}{(n-2)} = \frac{\sum \hat{u}_i^2}{(3-2) \cdot 2}$

$$\hat{\beta}_1 = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2} = \frac{3}{2} \quad (\text{from sem 1})$$

$$\hat{\beta}_0 = -1 \quad (\text{from sem 1})$$

$$\hat{y} = \begin{pmatrix} -1 + \frac{3}{2} \cdot 1 \\ -1 + \frac{3}{2} \cdot 2 \\ -1 + \frac{3}{2} \cdot 3 \end{pmatrix} = \begin{pmatrix} 1/2 \\ 2 \\ 7/2 \end{pmatrix} \Rightarrow \hat{u} = \begin{pmatrix} 1/2 \\ -1 \\ 1/2 \end{pmatrix}$$

$$\Rightarrow \sum \hat{u}_i^2 = \frac{1}{4} + 1 + \frac{1}{4} = \frac{3}{2} \Rightarrow$$

$$\Rightarrow \text{Var}(\hat{\beta}_1 | x) = \frac{12}{2} = 0,75$$

