

Lecture #2

$$\text{Var}(R|F) = E(R^2|F) - (E(R|F))^2$$

↑
RV ↓
 σ -algebra

Properties:

- ① If R is measurable wrt F then $\text{Var}(R|F) = 0$
- ② If R and F are independent then $\text{Var}(R|F) = \text{Var}(R)$
- ③ "Take out known RVs"

If R, S, T - random variables

S, T are known RV (S, T are measurable wrt F) then

$$\text{Var}(SR+T|F) = S^2 \cdot \text{Var}(R|F)$$

- ④ If S is measurable wrt F then $\text{Cov}(R, S|F) = 0$
- ⑤ $\text{Var}(R+S|F) = \text{Var}(R|F) + \text{Var}(S|F) + 2 \text{Cov}(R, S|F)$

$$\begin{array}{ccc} E(\cdot|F) & & \\ \swarrow \text{Corr}(\cdot, \cdot|F) & \downarrow & \searrow \text{Cov}(\cdot, \cdot|F) \\ \text{Var}(\cdot|F) & & \text{P}(\cdot|F) \end{array}$$

def $\text{Cov}(R, S|F) = E(RS|F) - E(R|F) \cdot E(S|F)$

Wow!

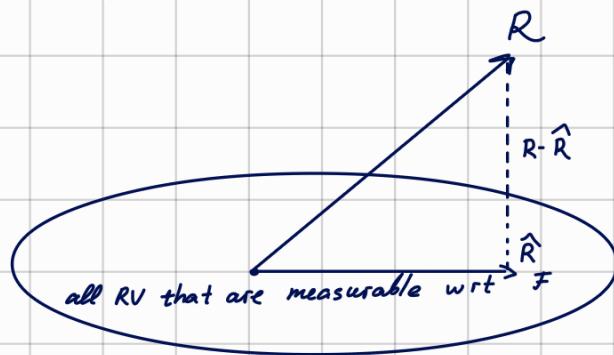
$$\text{Var}(R) = E(\text{Var}(R|F)) + \text{Var}(E(R|F))$$

/ t is Pythagorean theorem
 $c^2 = a^2 + b^2$

If S is measurable wrt F

$$\text{Var}(S|F) = E(S^2|F) - (E(S|F))^2 = S^2 - (S)^2 = 0$$

↓
by def



def $E(R|F) = \hat{R}$

- ① \hat{R} is meas. wrt. F
- ② $E(\hat{R}) = E(R)$
- ③ $\text{Cov}(\hat{R}-R, I_A) = 0$
 if $A \in F$

$$\begin{aligned} \text{Var}(R) &= \text{Var}(\hat{R}) + \text{Var}(R-\hat{R}) + 2 \text{Cov}(R-\hat{R}, \hat{R}) \\ &\quad \text{Var}(E(R|F)) \quad \text{E}(\text{Var}(R|F)) \quad 0 \end{aligned}$$

$$\text{Var}(R - \hat{R}) = E((R - \hat{R})^2) - (E(R - \hat{R}))^2 \quad \text{by def}$$

$$= E((R - \hat{R})^2) = E(E(R - \hat{R})^2 | F))$$

$\nearrow w$

Tower property of $E(\cdot | F)$

$$E(E(w | F)) = E(w)$$

$$\text{Cov}(R - \hat{R}, \hat{R}) = 0$$

omitted

$$= E(E(R^2 - 2R\hat{R} + \hat{R}^2 | F))$$

$$= E(E(R^2 | F) - 2\hat{R}E(R | F) + \hat{R}^2) = E(E(R^2 | F) - 2\hat{R}\hat{R} + \hat{R}^2) = E(E(R^2 | F) - E(R | F))^2$$

\hat{R}

$$= E(E(R^2 | F) - \hat{R}^2)$$

def the process (X_t) is a martingale.

Intuit. The best forecast of the next value is the current value

$$E(X_{t+1} | X_t, X_{t-1}, X_{t-2}, X_{t-3}, \dots, X_1) = X_t$$

ex. s. X_1, X_2, X_3, \dots iid $P(X_t = 1) = P(X_t = -1) = \frac{1}{2}$

$$S_t = X_1 + X_2 + \dots + X_t$$

a) $E(X_t)$ $E(S_t)$ $E(M_t)$

$$M_t = S_t^2 - t$$

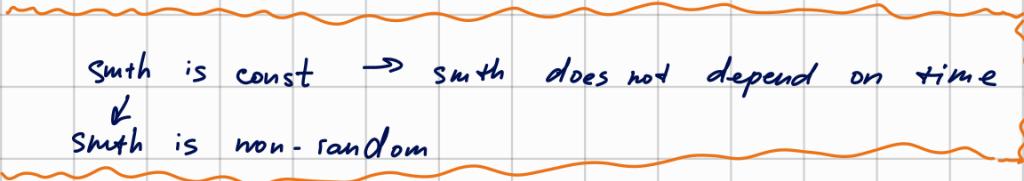
b) which one is a martingale?

$$E(X_t) = 0$$

$$E(S_t) = E(X_1 + X_2 + \dots + X_t) = 0 + 0 + \dots + 0 = 0$$

$$E(M_t) = E(S_t^2) - t = \text{Var}(S_t) - t = t \cdot \text{Var}(X_1) - t = t(E(X_1^2) - 0) - t = t - t = 0$$

S_t is const $\rightarrow S_t$ does not depend on time
 \downarrow
 S_t is non-random



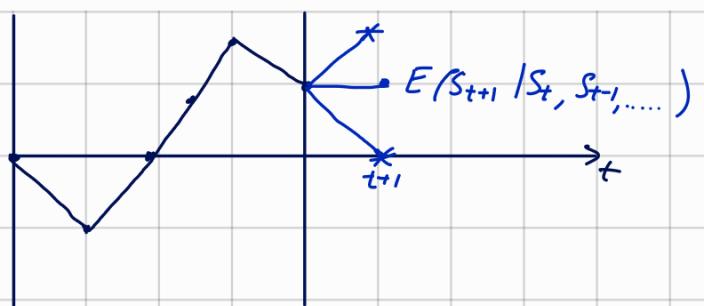
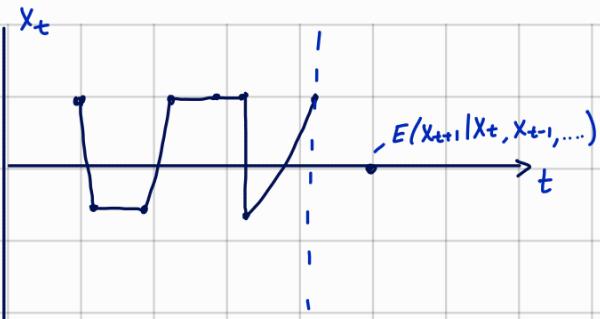
$$E(X_{t+1} | \underbrace{X_t, X_{t-1}, \dots, X_1}_{\text{indep}}) = E(X_{t+1}) = 1 \cdot \frac{1}{2} + (-1) \cdot \frac{1}{2} = 0$$

is $E(X_{t+1} | X_t, X_{t-1}, \dots, X_1) = ?$ X_t

No, X_{t+1} is not a martingale

$$E(S_{t+1} | S_t, S_{t-1}, S_{t-2}, \dots, S_1) = E(S_t + X_{t+1} | \underbrace{S_t, S_{t-1}, \dots, S_1}_{\text{known}}) =$$

$$= S_t + E(X_{t+1} | X_t + X_{t-1} + X_{t-2}, \dots, X_1, \dots, X_2 + X_1, \dots, X_1) = S_t + E(X_{t+1}) = S_t - \text{is a martingale}$$



Let's shorten the notation :

$$\mathcal{F}_t = \sigma(X_1, X_2, \dots, X_t)$$

$$E(M_{t+1} | \mathcal{F}_t) = E(S_{t+1}^2 - (t+1) | \mathcal{F}_t) = E((S_t + X_{t+1})^2 | \mathcal{F}_t) - t - 1 =$$

$$= E(S_t^2 + \underbrace{X_{t+1}^2}_{\text{known}} + 2S_t X_{t+1} | \mathcal{F}_t) - t - 1 = E(S_t^2 + 2S_t X_{t+1} | \mathcal{F}_t) + 1 - t - 1 =$$

$$= S_t^2 + 2S_t \cdot E(X_{t+1} | \mathcal{F}_t) - t = S_t^2 + 2S_t \cdot E(X_{t+1}) - t = S_t^2 + 0 - t = M_t \quad \text{is a martingale}$$

indep

def (\mathcal{F}_t) is a filtration

$$\mathcal{F}_1 \subseteq \mathcal{F}_2 \subseteq \mathcal{F}_3 \subseteq \mathcal{F}_4 \subseteq \dots$$

every \mathcal{F}_t is a σ -algebra

def the process (M_t) is a martingale wrt filtration (\mathcal{F}_t) if

$$E(M_{t+1} | \mathcal{F}_t) = M_t$$

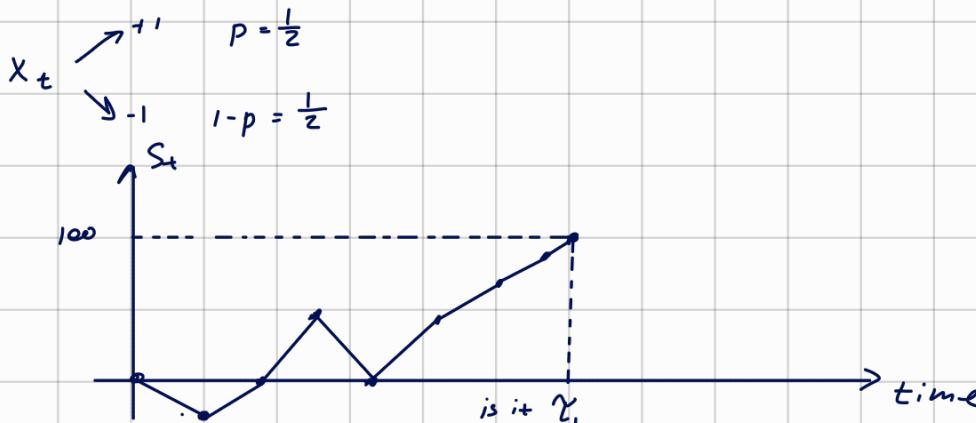
Intuition Behind the stopping time τ - a stopping time

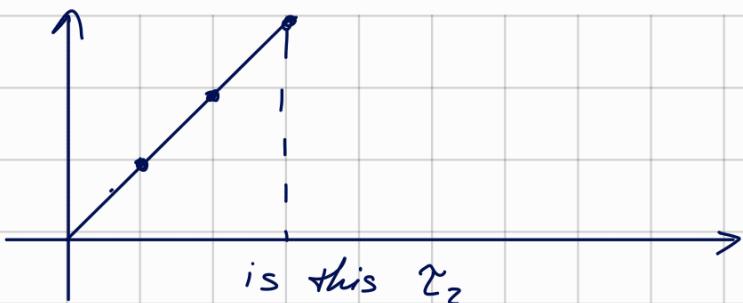
example $(X_t) \sim \text{iid}$ $S_t = X_1 + X_2 + \dots + X_t$

$$\mathcal{F}_t = \sigma(X_1, X_2, X_3, \dots, X_t) \quad \tau_1 = \min_t \{t + 1 \mid S_t = 100\}$$

τ_1 = the time of a first local maximum of S_t

not a stopping time





ex. $\tau_3 = \tau_1 + \tau$ (+) stopping time
 $\tau_4 = \tau_1 - \tau$ (-) not

Standard class #2

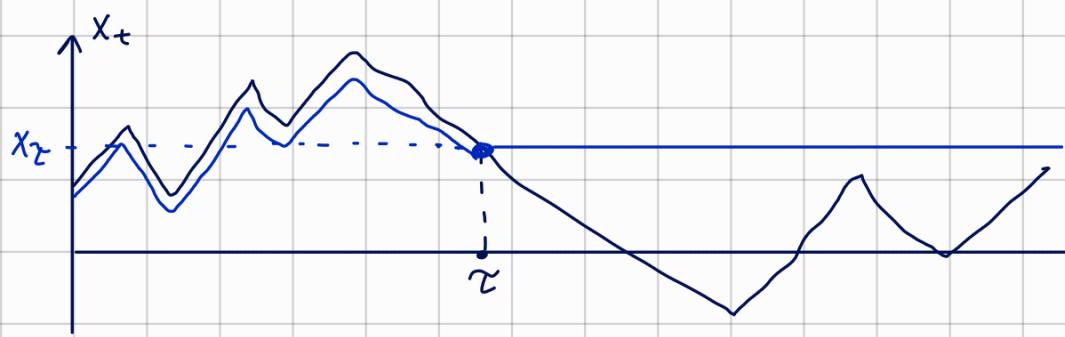
def τ is a stopping time wrt to filtration (\mathcal{F}_t)
if $\forall t \ \{\tau \leq t\} \in \mathcal{F}_t$

notation $x \wedge y = \min(x, y)$

$$x \vee y = \max(x, y)$$

$$x \wedge (y \vee z) = (x \wedge y) \vee (x \wedge z)$$

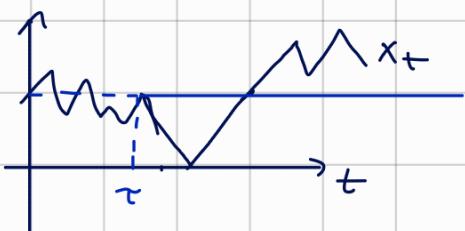
Theorem If (X_t) is a martingale and τ is a stopping time wrt filtration (\mathcal{F}_t) then $Y_t = X_{t \wedge \tau}$ is also a martingale wrt to (\mathcal{F}_t)



$$Y_t \rightarrow$$

$$Y_\tau = X_{\tau \wedge \tau} = X_\tau$$

$$Y_{12} = X_{12 \wedge \tau} = X_\tau$$



Theorem (Optional stopping time / Doob's theorem, Teorema Doobă)

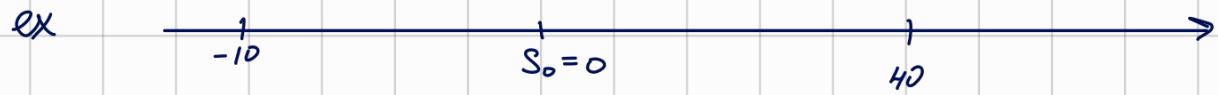
If τ is stopping time and (Y_t) is a martingale wrt filtration (\mathcal{F}_t) and at least one of the following conditions are satisfied: then $E(Y_\tau) = E(Y_0)$

(A) $P(\gamma = +\infty) = 0$, $Y_{t+1} \gamma$ is bounded

[there is a constant c
such that $|Y_{t+1} \gamma| < c$]

(B) $E(\gamma) < \infty$, $d_t = E(Y_{(t+1)\gamma} - Y_{t+1}\gamma | F_t)$ is bounded

(C)



$$S_t = X_1 + X_2 + \dots + X_t$$

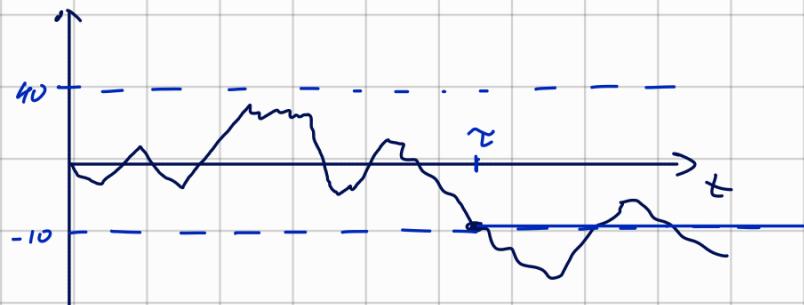
$$\gamma = \min \{t | S_t = -10 \text{ OR } S_t = 40\}$$

$$X_t \sim \text{iid} \begin{cases} \nearrow +1 & (p = \frac{1}{2}) \\ \searrow -1 & (p = \frac{1}{2}) \end{cases}$$

- a) $P(X_\gamma = 40) = \alpha \quad E(X_\gamma) ?$
 b) $E(\gamma) = ?$

S_t is martingale

$$M_t = S_t^2 - t \text{ is martingale}$$



$$F_t = \sigma(X_1, X_2, \dots, X_n)$$

$$E(S_\gamma) = E(S_0) = E(0) = 0$$

$$S_0 = 0$$

$$S_\gamma \xrightarrow{-10} \xrightarrow{40}$$

$$E(S_\gamma) = 40 \cdot P(S_\gamma = 40) + (-10) \cdot P(S_\gamma = -10) = 0$$

$$40\alpha - 10 \cdot (1-\alpha) = 0$$

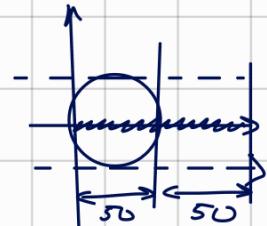
$$40\alpha + 10\alpha = 10$$

$$P(S_\gamma = 40) = \alpha = \frac{10}{10+40} = \frac{1}{5}$$

From A: condition is $Y_{t+1} \gamma$ is satisfied

$P(\gamma = +\infty)$ Probability never quit the game

$$P(\gamma = +\infty) \leq \left(1 - \left(\frac{1}{2}\right)^{50}\right) / \left(1 - \left(\frac{1}{2}\right)^{50}\right) \cdot \dots \cdot () \cdot () \cdot \dots = 0$$



$$E(M_2) = E(M_0) = E(S_0^2 - 0) = 0$$

$$\begin{aligned} S_2^2 &\xrightarrow{-10^2} \text{prob} = \frac{4}{5} \\ S_2 &\xrightarrow{40^2} \text{prob} = \frac{1}{5} \end{aligned}$$

$$E(\Sigma) = E(S_2^2) = -10^2 \cdot \frac{4}{5} + 40^2 \cdot \frac{1}{5} = \frac{400 + 1600}{5} = 400$$

Example when we cannot use Doob's theorem

$$(X_t) \text{ iid } \begin{cases} \xrightarrow{+1} \text{prob} = \frac{1}{2} \\ \xrightarrow{-1} \text{prob} = \frac{1}{2} \end{cases}$$

$$S_t = X_1 + X_2 + \dots + X_t$$

$$J = \min \{t \mid S_t = -1\}$$

$$(\mathcal{F}_t) = \sigma(X_1, X_2, \dots, X_t)$$

$$E(S_2) = -1$$

$$E(S_0) = 0$$

$S_{t \wedge J}$ is bounded?

New rule: I quit the game
when I loose first ruble

$$S_2 = -1$$

- t is not bounded
(no upper-boundary)

$$E(J) ?$$

J - random time necessary to go from 0 to -1

$$P(J=1) = \frac{1}{2}$$

$$P(J=2) = 0$$

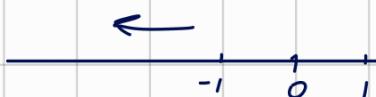
$$P(J=3) = \frac{1}{8} \quad \dots$$

J' is a random time necessary to go from 1 to 0

$$E(J') = E(J)$$

J^n is a random time necessary to go from n to 0

$$E(J) ?$$



$$E(J) = \frac{1}{2} \cdot 1 + \frac{1}{2} E(J+J'+1)$$



$$\begin{aligned} E(J) &= 1 + \frac{1}{2} E(J) + \frac{1}{2} E(J') \\ E(J) &= 2 + E(J') \end{aligned}$$

$$E(V) = E(V')$$

$$E(V) = E(V') = +\infty$$

\Rightarrow Conditions are violated

Check up class #2

(def) (X_t) is a martingale

$$E(X_{t+1} | X_1, X_2, \dots, X_t) = X_t \quad (\forall t)$$

(def) (X_t) is a martingale wrt filtration (\mathcal{F}_t)

$$E(X_{t+1} | \mathcal{F}_t) = X_t \quad (\forall t)$$

$$\mathcal{F}_t = \sigma(X_1, Y_1, Z_1, X_2, Y_2, Z_2, X_3, Y_3, Z_3, \dots, X_t, Y_t, Z_t)$$

Page 64 2021 exam $\sqrt{2}$

$$X_0 = 0, 2021$$

$$\mathcal{F}_t = \sigma(X_0, X_1, X_2, \dots, X_t)$$

$$P(Y_t = 1 | \mathcal{F}_{t-1}) = X_t - 1$$

$$X_t = \frac{Y_t + X_{t-1}}{2}$$

a) Is (X_t) a martingale?

b) distribution of $\lim X_t$?

$$Y_t \begin{cases} \rightarrow 1 \text{ with } p = X_0 = 0, 2021 \\ \rightarrow 0 \text{ with } 1-p = 1-X_0 = 0, 7979 \end{cases}$$

$$X = \frac{Y_1 + X_0}{2}$$

$$(Y_2 | \mathcal{F}_1) \begin{cases} \rightarrow 1 \text{ with } p = X_1 \\ \rightarrow 0 \text{ with } 1-p = 1-X_1 \end{cases}$$

$$P(Y_1 = 1, Y_2 = 1, Y_3 = 1) = P(Y_1 = 1) \cdot P(Y_2 = 1 | Y_1 = 1) \cdot P(Y_3 = 1 | Y_1 = 1, Y_2 = 1)$$

$$P(Y_1=1, Y_2=0, Y_3=1) = 0,2021 \cdot P(Y_2=1 | Y_1=1) \cdot P(Y_3=1 | Y_2=1, Y_1=1)$$

$$X_1 = \frac{1 + 0,2021}{2} = 0,60105$$

$$E(x_{t+1} | \mathcal{F}_t) = E\left(\frac{Y_{t+1} + X_t}{2} \mid \mathcal{F}_t\right) =$$

↖ know X_1, X_2, \dots, X_t

$$= \frac{X_t}{2} + \frac{1}{2} \left(1 \cdot X_t + 0 \cdot (1 - X_t) \right) = X_t \quad \text{is a martingale}$$

$$\begin{aligned}
 & x_2 \rightarrow \frac{0,2021}{2} / 2 \\
 & \downarrow \quad \downarrow \quad \downarrow \\
 & \left(1 + \frac{0,2021}{2}\right) / 2 \\
 & \left(1 + \frac{0,2021}{2}\right) / 2 \\
 & \downarrow \\
 & \frac{1 + \frac{0,2021}{2}}{2}
 \end{aligned}$$

$$\text{If } x_t \in [0;1] \Rightarrow x_{t+1} \in [0;1]$$

$$x_0 \in [0; 1] \\ x_1 = \frac{y_1^{L_1} + x_0^{L_1}}{2} \leq 1 \\ \geq 0$$

$$x_2 = \frac{y_2^{\leq 1} - y_1^{\leq 1}}{2} \in [0; 1]$$

Guess $\lim x_t \sim \text{Uniform}[0; 1]$

$$P(X_+ \leq a)$$

for uniform distribution

$$W \sim U[0; 1] \quad P(W \leq 0,7) = 0,7$$

$$P(W \leq 0,3) = 0,3$$



$$\lim P(X_t \leq a) = a$$

$E(E(\cdot | \mathcal{F})) = E(\cdot)$ hard to calculate
 \downarrow easy to calculate

$$P(X_t \leq a) = E(I(X_t \leq a)) = E(E(I(X_t \leq a) | \mathcal{F}_{t-1}))$$

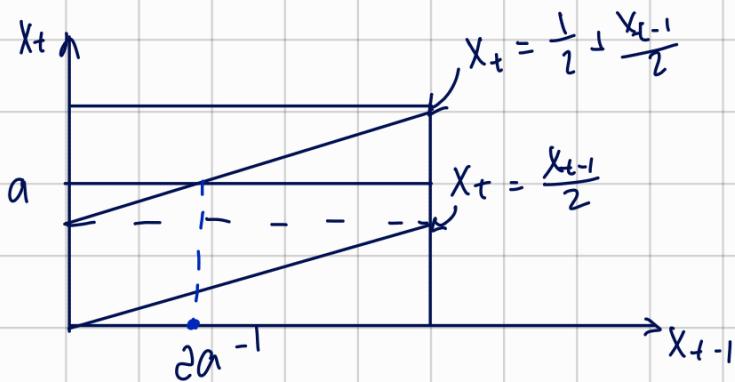
$$E(I(X_t \leq a) | \mathcal{F}_{t-1}) = P(X_t \leq a | X_{t-1}) = P\left(\frac{Y_t + X_{t-1}}{2} \leq a | X_{t-1}\right)$$

$$a = 0,7$$

$$\frac{Y_t + X_{t-1}}{2} \rightarrow \frac{1+0,5}{2} = 0,75 \text{ prob } 0,5 (X_{t-1})$$

$$X_{t-1} = 0,5$$

$$\frac{0,5}{2} = 0,25 \text{ prob } 1 - 0,5 = 1 - X_{t-1}$$



$$a > \frac{1}{2}$$

$$1 - X_{t-1} + X_{t-1} \cdot I(X_{t-1} \leq 2a-1)$$

$$2a = 1 + X_{t-1}$$

$$P(X_t \leq a) = E(1 - X_{t-1} + X_{t-1} \cdot I(X_{t-1} \leq 2a-1))$$

try at home

$$\lim X_t \stackrel{\text{guess}}{\sim} U[0,1]$$

$$\lim P(Y_t = 1)$$

$$X_{10000000} \approx U[0,1]$$

$$P(Y_t = 1) = E(E(Y_t | X_{t-1})) = E(E(Y_t | X_{t-1})) = E(X_{t-1})$$

exam 2017

√1 page 50

standard dice in times

$w_i \rightarrow 1$ if "five" appears
 $\rightarrow 0$, otherwise

$$Y = w_2 + w_3 + \dots + w_n$$

$$X = w_1 + w_2 + \dots + w_{n-1}$$

$$E(X|Y)$$

$$\text{Var}(Y|X)$$

$$E(Y|X)$$

$$E(X|Y) = E(\underbrace{W_1 + W_2 + \dots + W_{n-1}}_{\text{ind}} | W_2 + W_3 + \dots + W_n) = \underbrace{E(W_1)}_{\frac{1}{6}} + E(W_2 + \dots + W_{n-1} | W_2 + W_3 + \dots + W_n)$$

depends on information

$$E(W_2 + W_3 + \dots + W_n | W_2 + W_3 + W_4 + \dots + W_n) = W_2 + W_3 + \dots + W_n = (n-1)R$$

$$E(W_2 | W_2 + \dots + W_n) = R =$$

$$E(W_3 | W_2 + \dots + W_n) = R =$$

$$E(W_2 | W_2 + W_3) \quad E(W_3 | W_2 + W_3 = 1) = \frac{1}{2}$$

$$W_2 + W_3 = 0 \quad E(W_2 | W_2 + W_3 = 0) = 0 \quad R = \frac{W_2 + W_3 + \dots + W_n}{n-1}$$

$$W_2 + W_3 = 1 \quad E(W_2 | W_2 + W_3 = 1) = \frac{1}{2}$$

$$E(W_2 | W_2 + W_3) = \frac{W_2 + W_3}{2}$$

$$E(W_3 | W_2 + W_3) = \frac{W_2 + W_3}{2}$$

$$E(W_2) = \frac{1}{6}$$

$$E(\underbrace{W_1 + W_2 + \dots + W_{n-1}}_{\text{ind}} | W_2 + \dots + W_n) = \frac{1}{6} + (n-2) \cdot \frac{W_2 + W_3 + \dots + W_n}{n-1}$$

$$E(X|Y) = \frac{1}{6} + \frac{n-2}{n-1} \cdot Y$$

$$E(Y|X) = \text{the same} = \frac{1}{6} + n-2 \cdot \frac{X}{n-1}$$

$$E(W_2 + W_3 + \dots + W_n | \underbrace{W_1 + W_2 + \dots + W_{n-1}}_{\text{unknown}})$$

$$E(\underbrace{W_2 + W_3 + W_4}_{\text{ind}} | W_2 + W_3 + W_4 + W_5 + W_6 = 4) = 3 \cdot E(W_2 | W_2 + \dots + W_6 = 4) =$$

$$= 3 \cdot \frac{4}{5}$$

$$E(W_2 | W_2 + \dots + W_6) = \frac{W_2 + W_3 + \dots + W_6}{5}$$