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## Today

- Time series: definition
- Structural/non-structural modeling
- Stationarity, weak and strong
- Autocovariance, autocorrelation, partial autocorrelation
- Lag operator
- ARMA models (beginning)
- Tsay "Analysis of Financial Time Series." (1.2, 2.1-2.6) Hamilton "Time Series Analysis" (2.1, 3.1-3.5)

  Stock and Watson "Introduction to Econometrics" (14.1, 14.2)
- ✓ Diebold "Forecasting" (online version: http://www.ssc.upenn.edu/ fdiebold/Teaching221/Forecasting.pdf (6.5, 7.1, 7.2)

#### **Cross-sectional data:**

- The sample is i.i.d. (or at least independent)
- Useful for answering questions about
   causal effects of one variable on another

#### **Time Series:**

- The sample is not i.id., observe variable(s) over time
- Useful for answering questions about dynamic causal effects
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We will study quantitative (non-structural) models of time series to use them later for forecasting:

#### **Structural**

- Has some **economic theory** behind it
- Parameters have meaning and causal interpretation
- Will briefly touch in VAR and ADL topic

#### **Non-structural**

- Models based on fitting data
- Coefficients do not have causal interpretation
- Will be the main topic of the course

### Time Series: definition

- **Informaly:** a set of realizations of a random variable ordered according to time
- Formally:

#### Definition 1

Collection of random variables defined on the sample space  $\{Y_t, t \in T\}$  is called a *stochastic process* 

We will consider  $T = \{..., -1, 0, 1, 2,...\} = \mathbb{Z}$ 

#### Definition 2

A *time series* is a realization of a stochastic process:  $\{y_t, t \in \mathbb{Z}\}$ 

#### **Definition 3**

A time series sample is  $\{y_t, t = 1, ..., T\}$  for some  $T < \infty$ .

But 'time series' can be used as a synonym of 'stochastic process'

### Important concepts

- Goal: forecast values of a random variable using the time series sample
- So, we need the future to be like the past
- Reflected in the concept of *stationarity*

### Definition 4

A process  $\{Y_t, t \in Z\}$  is *strictly stationary* if, for any k, s and any  $t_1, \ldots, t_k$ , the *distributions* of  $(Y_{t_1}, Y_{t_2}, \ldots, Y_{t_k})$  and  $(Y_{t_1+s}, Y_{t_2+s}, \ldots, Y_{t_k+s})$  are *the same* .

In other words, the following distributions are the same:

- of  $Y_1$  and  $Y_{100}$
- of  $(Y_1, Y_2)$  and  $(Y_5, Y_6)$
- of  $(Y_3, Y_{10}, Y_{22})$  and  $(Y_{13}, Y_{20}, Y_{32})$
- and so on ...

Strict stationarity is a complicated concept

Very often people consider *weak stationarity* 

#### Definition 5

A process  $\{Y_t, t \in Z\}$  is weakly, or covariance-, stationary if, for any  $t_1, t_2, s \in Z$ 

$$E[Y_{t_1}] = E[Y_{t_2}],$$

$$2 \quad Cov\left(Y_{t_1}, Y_{t_1+s}\right) = Cov\left(Y_{t_2}, Y_{t_2+s}\right) \neq f\left(t_1\right) \quad \forall s$$

So, only the following has to be the same:

• mean of all 
$$Y_t$$
  $\mathbb{E}[Y_t]$ 

• variance of all 
$$Y_t$$
  $\bigvee_{a} (y_t)$ 

• mean of all 
$$Y_t$$
  $E[Y_t]$   $V_{as}(Y_t)$   $V_{as}(Y_t)$   $V_{as}(Y_t)$   $V_{as}(Y_t)$ 

• covariances between all of the possible pairs of  $Y_t$  that are fixed number of periods away from each Cov (5+, 5++5) / 45 other

### Question

If  $\{Y_t, t \in \mathbb{Z}\}$  is weakly stationary, is it also strictly stationary?

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## Stationarity: extra remarks

**Not all** *weakly stationary* process *strictly stationary*.

But if  $\{Y_t\}$  is gaussian, then it is *weakly stationary*, it is also *strictly stationary*.

### Autocovariance and autocorrelation function

- Want to forecast future by exploring the relation between r.v.
   corresponding to consecutive periods of time
- Autocovariance is a way to quantify this relation

#### Definition 6

- Autocovariance of order k is  $\gamma(k) = Cov(Y_t, Y_{t+k})$
- Autocorrelation of order k is  $\rho(k) = corr\left(Y_t, Y_{t+k}\right) = \frac{\gamma(k)}{Var(Y_t)}$   $\Rightarrow$   $3^{(k)} = \frac{\gamma(k)}{\sqrt{(k)}}$ .
- $\gamma(\bullet)$  is called *autocovariance function* (ACF)
- $\rho(\bullet)$  is called *autocorrelation function* (also ACF)

### Estimated ACF

$$\bar{Y}_T = \frac{1}{T} \sum_{t=1}^T Y_t$$

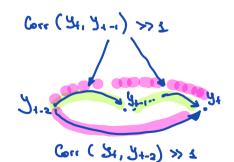
$$\gamma(0) = \frac{1}{T} \sum_{t=1}^{T} (Y_t - \bar{Y}_T)^2$$

$$\gamma(\hat{k}) = \frac{1}{T} \sum_{t=k+1}^{T} (Y_t - \bar{Y}_T)(Y_{t-k} - \bar{Y}_T)$$

Sample autocorrelation function:

$$\hat{\rho}(k) = \frac{\gamma(\hat{k})}{\gamma(\hat{0})}$$

Correlogram: a graph of sample ACF



## Partial Autocorrelation Function (PACF)

- Autocorrelation measures how dependent the data is
- If  $Y_1$  and  $Y_2$  are related, and  $Y_2$  and  $Y_3$  are related, then  $Y_1$  and  $Y_3$  have to be related at least indirectly
- Partial Autocorrelation Function (PACF) measures direct relation between different  $Y_t$ .

#### Definition 6

Partial Autocorrelation Function (PACF) at lag k is

$$\alpha(k) = corr(Y_1 - P(1, Y_2, ..., Y_k)Y_1, Y_{k+1} - P(1, Y_2, ..., Y_k)Y_{k+1}),$$

where  $P(1, Y_2, ..., Y_k)Y_j$  is the linear projection of  $Y_j$  on a constant,  $Y_2, ...,$  and  $Y_k$ .

### Partial Autocorrelation Function (PACF)

Write the linear projection of  $Y_{t+k+1} - \mu$  on  $Y_{t+k} - \mu$ , ...,  $Y_{t+1} - \mu$ :

$$\hat{Y}_{t+k+1} - \mu = \alpha(1) (Y_{t+k} - \mu) + \dots + \alpha(k) Y_{t+1} - \mu$$

PACF of order 1 to k can be found as

$$\begin{pmatrix} \alpha(1) \\ \alpha(2) \\ \dots \\ \alpha(k) \end{pmatrix} = \begin{pmatrix} \gamma(0) & \gamma(1) & \dots & \gamma(k) \\ \gamma(1) & \gamma(0) & \dots & \gamma(k-2) \\ \dots & \dots & \dots & \dots \\ \gamma(k-1) & \gamma(k-2) & \dots & \gamma(0) \end{pmatrix}^{-1} \begin{pmatrix} \gamma(1) \\ \gamma(2) \\ \dots \\ \gamma(k) \end{pmatrix}$$

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Sample PACF: estimate by OLS

$$Y_{t+k+1} - \mu = \alpha(1)(Y_{t+k} - \mu) + \dots + \alpha(k)(Y_{t+1} - \mu) + \varepsilon_{t+k+1}$$

## Lag operator (a.k.a back-shift operator)

#### Definition 7

The lag operator L is a linear operator such that for all t

$$LY_t = Y_{t-1} \qquad \qquad \bigsqcup \left( \mathcal{Y}_t \right) = \mathcal{Y}_{t-1}$$

#### **Properties:**

• 
$$L^2Y_t = L(LY_t) = LY_{t-1} = Y_{t-2}$$

$$L^j L^i Y_t = L^{j+i} Y_t = Y_{t-i-j}$$

• 
$$Lc = c$$

• 
$$(L^{j} + L^{i})Y_{t} = Y_{t-i} + Y_{t-j}$$

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• *L* to a negative power is a *lead operator*:  $L^{-i}Y_t = Y_t$ 

• For 
$$|a| < 1$$
,

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,
$$y_{t} = \begin{bmatrix} 1 + aL + a^{2}L^{2} + \dots + a^{k}L^{k} + \dots \end{bmatrix} = (1 - aL)^{-1}Y_{t} = \underbrace{Y_{t}}_{1 - aL}$$
Time Series and Stochastic Processes

# ARMA models

## Starting with basics

- $\{\varepsilon_t, t \in \mathbf{Z}\}$ :  $\varepsilon_t$  are iid
  - is it stationary?
- MDS (Martingale Difference Sequence):  $\{\varepsilon_t, t \in \mathbf{Z}\}$ :

$$E[\varepsilon_t \mid \varepsilon_{t-1}, \varepsilon_{t-2}, \ldots] = 0$$

- A process like that is closer to economics: many dynamic optimization problems result in a condition of this type.
- White noise: (A more statistical description of innovations)  $\{\varepsilon_t, \ t \in \mathbf{Z}\}$ :

s.t.

$$Cov(\varepsilon_t, \varepsilon_s) = 0$$
, for any  $t \neq s$ 

$$E[\varepsilon_t] = 0$$
, from  $(\varepsilon_t) \Delta^2 \times \Delta_0 \quad \text{for } (\xi_t) = \xi^2 \times \infty$ 



## Moving-Average Models

• Start with  $\{\varepsilon_t\}$ , a white noise.

MA(1) - Moving average of order 1

$$Y_t = \varepsilon_t + \varphi \varepsilon_{t-1}$$

MA(q) - Moving average of order q

$$Y_t = \varepsilon_t + \varphi_1 \varepsilon_{t-1} + \varphi_2 \varepsilon_{t-2} \dots + \varphi_q \varepsilon_{t-q}$$