# Testing for Network Effects in Field Experiments: Examples from Legislative Studies

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#### **Abstract**

Most social processes involve complex interaction and dependence among units in a network. The stable unit treatment value assumption (SUTVA)—the assumption that a unit's outcome is unaffected by other units' treatment statuses—is required in conventional approaches to causal inference. When SUTVA is violated, as in networked social interaction, treatment effects spread to control units through the network structure. We evaluate the evidence for spillover effects in data from three field experiments on US state legislatures. Randomized field experiments represent the gold standard in causal inference when studying political elites. It is not possible to bring political elites into the lab, and causal identification with observational data is fraught with problems. We propose new specifications for treatment spillover models, and construct networks through geographical or ideological proximity and cosponsorship. Considering different combinations of spillover models and networks, we evaluate the robustness of recently developed non-parametric tests for interference. The approaches we illustrate can be applied to any experimental setting in which interference is suspected.

### 1 For discussion on 04/06

- Reviewing the poster
- Reviewing the Coppock extension used in the poster
- I would look forward to discussing the poster and the other idea of modeling nearest neighbors with you. I am happy to meet you on Skype. I really like the idea you have suggested for taking care of tied neighbors. I will try to code it in by tomorrow. After finishing the code that we now have, I was thinking of other ways in which we can further explore the idea of nearest neighbors. Maybe we could look at the

number of nodes for which j is the nearest neighbor and in some way account for the j's influence on that basis? This does not fit into any of the 3 dimensions we have defined for extension. But it does touch upon an idea we had briefly discussed; that of looking at this model from the point of view of treated units instead of only control units. We can also turn it around in the following steps:

1. Look at control unit i 2. Calculate a composite score that indicates the ranking of i for each of its k neighbors (similar to the ranking idea you have suggested) 3. Estimate a parameter to see if change in composite score leads to a change in propensity for changing the outcome 4. It seems intuitive to not model this parameter as linear, but will need to reason through it

For example, say we are under k=3 condition. If I am a control unit and I am the lowest in the proximity of all of my 3 neighbors, I am less likely to receive a spillover; as compared to if I was in the highest proximity of all three of them (the two extremes). We might have to condition this somehow on the treatment status of my neighbors though.

• JSM timeline

#### 2 Introduction

Networks are integral parts of human interaction and hence social science research. If one unit in a network gets treated, the effect may trickle down throughout network. The currently established framework for causal inference relies on SUTVA (Stable Unit Treatment Value Assumption). It assumes that whether or not one person/unit/node is treated, does not affect any other unit. However, SUTVA breaks down in a network setting. It is therefore imperative to take the interference structure into account. Rather, in policy planning or designing marketing campaigns, a researcher may be interested in studying the propogation of treatment effect itself.

In field experiments on social groups, interference may be substantial. In this project we intend to study intererence models for randomized experiments coducted on social networks and causal inference basis this.

To understand, explain and predict social phenomena, social scientists typically look to individual actors' attributes to explain their behavior (e.g., an increase in an individual's wealth will result in a decrease in that individual's support for government spending on social welfare), or to attributes of the macro context (e.g., an increase in the unemployment rate will lead an individual's support for the party of the president to decrease). However, these two conceptual approaches to explaining individual behavior leave out a potentially powerful class of social dynamics – interpersonal influence. That is, the behavior of one individual may depend upon the behavior of one or more others (e.g., a person may decide to vote due to their friends claiming to have voted (?)). Inferences regarding influence involve the analysis of individual behaviors and the behaviors of those adjacent or nearby in some contact network. However, as in most settings, it is generally not possible to identify the causal effects that map onto the process of social influence in observational data (?). As such, we need experimental methods to identify causal influence effects.

#### 2.1 tasks

- Points about why it is interesting to study propagation. (BD)
- Outline of the paper (SP)

## 3 Background

• Paragraph on each category of papers that serve as relevant background (SP)

- Interference models (diffusion, propagation) (SP–Review)
- Experiments on networks (applications) (SP–Review)
- Approaches to inference or estimation with propagation (SP–Review)
- Potential outcomes framework (SP find papers & Review)
- Review of political networks (SP–Review)
- Review of field experiments (SP–Review)

**-** (??????)

## 4 Research Design

We plan to re-analyze data from past field experimental studies to understand how conclusions regarding direct effects and interference effects depend upon the network structure.

#### 4.1 tasks

The key factors that we will consider in building propagation models are:

- 1. Distance from the nearest treated node  $(d_i)$
- 2. Number/proportion of treated nodes neighboring each  $d_i$
- 3. Form of spread (linear or non-linear)

Table 1: The Results

	Distance from treated < 5		Distance from treated > 5	
	Linear	Non-linear	Linear	Non-linear
Number of treated neighbors	1	2	5	6
Proportion of treated neighbors	3	4	7	8

Tabulates the basic idea of the models we will look at. The linear/non-linear refers to the form taken by the spread of effect. Distance from the nearest treated node is vaguely separated as <5 or >5 for now. Finally, we want to look at the difference in spillovers effects caused due to number of treated neighbors as against proportion of treated neighbors. Potentially, the number/proportion at hops of different lengths can also be considered.

The types of networks we will consider in the analysis are:

- 1. Grographical proximity
- 2. Ideological similarity
- 3. Co-sponsorship

Each of these factors is such, that a treatment such as the message sent through emails in New Hampshire, would possibly spread to untreated units as well. Legislators from adjoining districts may affect each other's opinions through geographic proximity as well as potentially via common issues faced by citizens in their constituencies.

Ideological similarity is a tricky variable because it could be hard to distinguish that from party affiliation. In the New Hampshire paper, there is a need to include this effect in

the model since matched pairs are created based on party affiliation. If a Republican candidate receives the treatment, the chances that through various communication channels, he/she will convey the message to the control group candidate from the same party and district, are very high.

Finally, serving on the same committee can also contribute to spreading the effect of a treatment. We must test for any dependence across these three fators before incorporating them into our model. Therefore it is important that we propose and test propogation models that consider the spread of treatment through our network.

\*\*Notes:

- 1. Ideological similarity: This should get highest priority in modeling, since I believe, it would be easiest to affect an undecided legislator's vote through similarity in ideas and belief about citizen's issues and how to resolve them. I would propose that we model immediate neighbours to have a 50
- 2. Grographical proximity: An untreated legislator from adjoining district would be
- 3. Co-sponsorship: Serving on the same committee increases the chances of

## 5 Analysis

We begin this section by reviewing prior methodological work. We will look at two spillover specifications and testing frameworks given in Bowers et. al. and Coppock papers. We will extend their analysis by considering various other specifications for spillover effects. Additionally, we will look into tests other than the Kolmogorov-Smirnov (KS) test.

### 5.1 Review of existing methods

- Bowers et. al. method: This paper introduces a Fisherian inference algorithm to test for spillover of treatment effect. The model  $\mathscr{H}$  is compared against the observed data. The hypothesized model of interference is specified by the researcher. The steps involved in conducting this test are as follows:
  - 1. We assume the "sharp null hypothesis of no effects" i.e. we assume that the treatment assignment has no effect on any unit
  - 2. We begin by specifying the causal model which describes the change in potential outcomes when treatment assignment changes from  $\mathbf{u}$  to  $\mathbf{w}$ ;  $\mathcal{H}(y_{i,\mathbf{u}},\mathbf{w},\theta)$ . If spillover effects are theoretically motivated, we must also specify treatment assignment for  $u_j$  and  $w_j$  where  $i \neq j$
  - 3. The potential outcomes from the causal model must be mapped to the observed outcomes  $y_z$ . The treatment assignment in the experiment (z) must be mapped to the uniformity trial which is based on a no-treatment assignment i.e. every unit is a control unit. In this condition, all  $z_i$ s are zero and we refer to this as the baseline condition. Uniformity trial is specified as  $\mathcal{H}(y_z, 0, \theta) = y_0$ . These mappings should give us the hypothesized value of  $\theta$  and a data adjusted to the model
  - 4. The test statistic we consider is  $\mathscr{T}$  and the key characteristic is that it should be a small value when distribution of treated and control outcomes in the adjusted data (mentiond in point 3) are similar.  $\mathscr{T}$  should be larger when distributions are dissimilar. Since we want the similarity to depend on not just the center but

also higher-order moments of a distribution (spread, skewness etc.), we need a sensitive measure. Bowers et al recommend using the Kolmogorov-Smirnov (KS) test statistic.

As noted in footnote 12 of the paper, KS statistic is the maximum difference between the empirical cumulative distribution functions (ECDFs) of treated  $(F_1)$  and control  $(F_0)$  units. So under the baseline condition,

$$\mathcal{T}_{\mathbf{y}_0,\mathbf{z}} = \max_{1 \le i \le n} [F_1(y_i,0) - [F_0(y_i,0)]$$

where  $F(x) = \frac{1}{n} \sum_{i=1}^{n} I(x_i \le x)$  is the proportion of x below  $x_i$ 

- 5. We must form hypothesis for interference. Here we assume that treatment only spreads through edges and the spillover effect only depends on the number of neighbours treated. The model for interference is explained in the immediately next section. However, here we note that the spillover effect is modeled using a growth curve  $\beta + (1 \beta)e^{-\tau^2 \mathbf{z}^T \mathbf{S}}$
- 6. Now we generate the distribution of test statistic under our hypothesis. The exact distribution is specified by computing  $t_k = \mathcal{T}(\mathbf{y}_0, \mathbf{Z}_k)$  for each  $\mathbf{Z}_k \in \Omega$ . Essentially, we are evaluating this for every possible treatment assignment. Alternatively, we can use sampling methods and limit theorems to estimate the distribution from data.
- 7. Finally, the p-value for our test can be calculates using the following formula:

$$\frac{\sum_{k=1}^{abs(\Omega)} I(x_i > t_k)}{abs(\Omega)}$$

 Coppock method: Coppock builds upon the New Mexico Legislator experiment conducted by Butler and Nickerson (2011). This paper also works with the idea of sharp null of no effects and uniformity trial. Using ideological similarity scores in the adjacency matrix, Coppock separates out the direct treatment effect and spillover occurring through the network.

In the first extension, we consider k nearest neighbors in terms of the ideology score. As of now, I have not taken care of ties. In 10 out of the 70 cases, there are in fact more than 3 neighbors within the same set of 3 distances.

\*Could we use the idea of communities to model spread of treatment across the network? \*Explore how alternative assumptions regarding interference change results

#### 5.2 tasks

- Replicate Bergan. (SP)
- Find other network data for the New Mexico legislature. (BD)
- Geography and ideology data for New Hampshire (BD)
- Produce cosponsorship, ideology and geography estimates for both Bergan and Nickerson (SP & BD)
- For at least two spreading models (SP & BD)
- Replicate Nyhan (SP)

Records of standing committee membership in the 16 standing committees in place during the 2008 regular session was obtained from the New Mexico Legislative Council Service Librarian.

## 6 Summary

So far, we have replicated the two papers mentioned earlier; (?) and (?). As of now, we will go ahead with the Coppock method alone. Bowers method is suitable if the outcome (y) is continuous. For binary outcomes, the uniformity trial conversion leads to ties in the sub-vectors we will use to perform KS test. In our datasets, outcome variable is voting decision; essentially binary. The codes are available in the appendix. So far we have worked on the New Mexico legislators dataset from (?) and considered network based on ideological similarity. The plan ahead for this paper is:

- Consider different diffusion models by varying the distance from treated node, number/proportion of treated neighbors and form of spread of treatment
- Consider other legislator networks depending on geographical proximity and cosponsorship
- Consider additional test statistic such as the Anderson-Darling test and other tests mentioned in (?) (Mann-Whitney U test, Control Median test etc.)

### 6.1 Notes from lab meetings

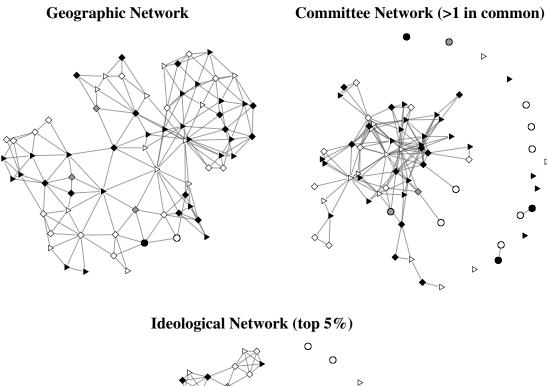
- Would we like to give a motivating example for why people should care about spillover effects? Or should this paper itself become a motivation for people to understand the importance of it?
- In what format do we want to present the results? Potentially give visuals of p-values under 8 different models using each of the networks?
- Consider committee network
- Estimate a  $\gamma$  for weighted combination of networks under consideration
- What would we like to include in the paper about the possible attenuation bias? (Reference: email dated 11/11/15 with subject line 'attenuation')

## 7 Appendix

### 7.1 Appendix 1A

In this section, we will look at user-defined R-functions that replicate the Bowers et. al. methodology ((?)). There will be four steps in this:

- A function to transform the observed outcomes into potential outcomes for any treatment assignment w
- A function to separate the hypothesized treatment effect



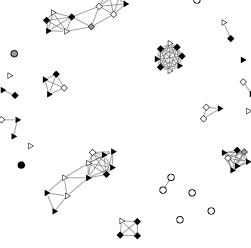


Figure 1: Different networks among New Mexico legislators. Colors denote outcome: black means voted with district, gray means abstained, white means voted against. Shape denotes treatment status. Triangles are treated. Squares are adjacent to treated. Circles are isolated from treatment

- A function to calculate test statistic
- A function to calculate the p-value.

The results from the ks.test function in R for calculating Kolmogorov-Smirnoff test statistic are verified with that in Footnote 12 of the paper.

**Function 1: calculating potential outcomes** 

```
set. seed (132)
2 library (doParallel)
3 library (foreach)
4 library (kSamples)
5 library (network)
6 library (permute)
8 #### Potential outcomes ####
10 #### Transform uniformity trial outcome into observed outcome
unif.to.z <- function(z, S, y.0, beta, tau){
    # z: observed treatment assignment
    # S: adjacency matrix
13
    # y.0: outcome vector for uniformity trial
    # beta: growth curve parameter
15
    # tau: rate of growth parameter
17
    scalar \leftarrow as.vector(t(z)\%*\%S)
18
    spillover \leftarrow rep(NA, n)
19
20
    spillover \leftarrow beta + ((1-z) * (1-beta) * exp(-tau^2 * scalar))
21
```

```
22
    # This is equation 4
23
    h.y0.z <- spillover*y.0
25 }
27 #### Transform observed outcome into uniformity trial outcome
z.to.unif \leftarrow function(z, S, y.z, beta, tau)
    # z: initial treatment assignment
    # S: adjacency matrix
30
    # y.z: observed outcome vector
    # beta: growth curve parameter
32
    # tau: rate of growth parameter
34
    scalar \leftarrow as.vector(t(z)\%*\%S)
35
    spillover \leftarrow rep(NA, n)
36
37
    # Equation (3)
38
    spillover \leftarrow beta + ((1-z) * (1-beta) * exp(-tau^2 * scalar))
40
    # This is equation 5
41
    h.yz.0 \leftarrow (1/spillover)*y.z
42
43 }
45 #### Transform observed outcome into outcome for ANY other assignment w
z.to.w \leftarrow function(z, S, w, y.z, beta, tau)
    # z: initial treatment assignment
    # S: adjacency matrix
    # w: new treatment assignment
```

```
# y.z: vector of outcomes for z
    # beta: growth curve parameter
    # tau: rate of growth parameter
52
53
    scalar.z \leftarrow as.vector(t(z)\%*\%S)
54
    scalar.w \leftarrow as.vector(t(w)\%*\%S)
56
    spillover.z \leftarrow rep(NA, n)
57
    spillover.z \leftarrow beta + ((1-z) * (1-beta) * exp(-tau^2 * scalar.z))
58
60
    spillover.w \leftarrow rep(NA, n)
    spillover.w \leftarrow beta + ((1-w) * (1-beta) * exp(-tau^2 * scalar.w))
62
64
    # Below is the actual function that transforms observed outcomes into
65
       potential outcomes
    # Equation (6)
67
    h.z.to.w <- (spillover.w / spillover.z) * y.z
68
69 }
71 #### Testing and p-value calculation ####
p.val \leftarrow function(z, y.z)
    c1 <- makeCluster(4) #Setup for parallel computing</pre>
75
    registerDoParallel(cl)
```

```
77
     # Calculate the outcome vector after taking away the effect of
78
      treatment
    y.0 \leftarrow z.to.unif(z=z, S=S, y.z=y.z, beta=beta, tau=tau)
80
     # Calculate test statistic
81
     test. stat \leftarrow ks. test (y.0[z==1], y.0[z==0],
82
                             alternative = "g") $ statistic
     sign <- noquote(strsplit(names(test.stat), NULL)[[1]])[3]</pre>
84
     if (sign=="+"){
       test.stat <- test.stat
86
     } else {
87
       test.stat <- test.stat*-1
88
     }
89
90
     # Calculate a vector of test statistic using permutations
91
     results <- foreach (i = 1:perms) %dopar%{
92
       require ( permute )
93
       perm.z \leftarrow z[sample(1:length(z), length(z), rep=F)]
94
       perm. test. stat \leftarrow ks. test(y.0[perm.z==1], y.0[perm.z==0],
                                     alternative = "g") $ statistic
96
       sign <- noquote(strsplit(names(perm.test.stat), NULL)[[1]])[3]</pre>
97
98
       if (sign=="+"){
99
         return (perm.test.stat)
100
       } else {
101
         return (perm. test. stat*-1)
102
       }
103
```

```
stopCluster(cl)

# A vector of test statistics
all.test.stat.vals <- unlist(results)

# Calculating p-value
pval <- sum(all.test.stat.vals > test.stat)/perms
return(pval)

# Calculating p-value
```

### 7.2 Appendix 1B

Below code replicates the (?) results using the framework setup in the Bowers replication code (Version before final corrections by BD)

```
11 library (network)
12 library (permute)
13 library (wnominate)
16 #### Read the original Butler and Nicketson data
17 #### This is the New Mexico dataset
19 data <- read.table("nm.replication.tab", sep="\t", header=TRUE)
z <- data$treatment #observed treatment
22 y.z <- data$sb24 #observed outcome
n \leftarrow length(y.z) #number of observations
t \leftarrow length(z[z==1]) #number of treated units
25 perms <- 10000 #number of permutations to use in generating expected
     exposure
26 perms.test <- 1000 #number of permutations used in testing
29 #### Generate Similarity Scores (this code taken from CoppockJEPS_
     datapreparation .R)
nmhouse2008 <-read.csv("CoppockJEPS_rollcalldata.csv")
bills <- data.frame(nmhouse2008[5:21])</pre>
34 ## Nominate Scores
36 bills_nona <- bills
```

```
bills_nona[bills_nona==99] <- NA
rollcalls <- rollcall(bills_nona)</pre>
39 nominate_scores <- wnominate(rollcalls, polarity=c(1, 2), minvotes=10)
40 dwnom_scores <- nominate_scores$legislators$coord1D
q get. similarity \leftarrow function (x, y) {
    return((2-abs(x-y))/2)
44 }
45
47 ## Create an adjacency/similarity matrix using ideology
48 S. ideo \leftarrow matrix (NA, ncol=70, nrow=70)
49 for (i in 1:70) {
    for (j in 1:70) {
      S.ideo[i,j] <- get.similarity(dwnom_scores[i], dwnom_scores[j])
    }
52
53 }
_{54} diag(S.ideo) \leftarrow 0
55 S.ideo [is.na(S.ideo)==T] \leftarrow 0
58 #### Generate expected exposure
59 perm <- replicate (perms, z[sample(1:length(z),length(z),rep=F)])
expected.exp0 \leftarrow rep(0, n)
expected.exp1 \leftarrow rep(0, n)
64 for(p in 1:ncol(perm)){
```

```
zp <- perm[,p]</pre>
65
    for(i in 1:n){
      if (zp[i] == 1){
67
           expected.exp1[i] \leftarrow expected.exp1[i] + sum(S.ideo[i,-i]*zp[-i])
         }
69
         else {
           expected.exp0[i] <- expected.exp0[i] + sum(S.ideo[i,]*zp)</pre>
71
72
73
74 }
75 num_treat <- apply (perm, 1, sum)
76 num_control <- apply(1-perm, 1, sum)</pre>
expected.exp1 <- expected.exp1/num_treat
78 expected.exp0 <- expected.exp0/num_control</pre>
79
81 #### Generate expected and net exposure
82 #### This is the spillover effect model
84 indirect treatment <- function (permutation, adj.mat) { #any treatment
      assignment vector and adjacency matrix can be used
    # permutation: can be the initial treatment assignment or a
      permutation
    raw.exp \leftarrow rep(NA, n)
86
    for (i in 1:n){
87
      raw.exp[i] <- sum(adj.mat[i,]*permutation)</pre>
88
      }
89
90
```

```
net.exp <- raw.exp - (permutation*expected.exp1 + (1-permutation)*
      expected.exp0)
    standard.exp \leftarrow (net.exp - mean(net.exp))/sd(net.exp) #this is the
      spillover or indirect effect
    return (standard.exp)
93
94 }
95
97 #### We now model the uniformity trial transformation
99 z.to.unif <- function (outcome, beta1, beta2, permutation, adj.mat) {
    # outcome: vector of direct treatment outcomes
    # betal: direct treatment effect parameter
101
    # beta2: indirect treatment effect parameter
    # permutation: vector of a permutation of z (can be z itself)
103
    # adj.mat: adjacency matrix
104
105
    exposure <- indirect.treatment(permutation, adj.mat)
106
    # This is equation 5
107
    h.yz.0 <- outcome - (beta1*permutation) - (beta2*exposure)
108
    return (h.yz.0)
109
110 }
112
#### Testing and p-value calculation
beta 1s < seq (from = -.7, to = 0.2, by = .025)
beta2s < seq (from=-.7, to=0.2, by=.025)
```

```
117
pvals <- matrix (NA, length (betals), length (beta2s))
120 cl <- makeCluster(4) #Setup for parallel computing
registerDoParallel(cl)
pvalues.ideology <- foreach (i = 1:length(betals)) %do% {
    abc \leftarrow foreach (j = 1: length (beta2s)) %do% {
      # Calculate the outcome vector after taking away the effect of
126
      treatment
      y.0 <- z.to.unif(outcome = y.z, beta1 = beta1s[i], beta2 = beta2s[j
      ], permutation = z, adj.mat = S.ideo)
128
      # Calculate observed test statistic
129
      exposure <- indirect.treatment(permutation = z, adj.mat = S.ideo)
130
      test.stat \leftarrow sum((lm(y.0 ~ z + exposure, na.action = na.omit)$resid
      )^2)
      # Calculate a vector of test statistic using permutations
133
134
       results <- foreach (k = 1:perms.test) %dopar% {
         require (permute)
136
         perm. z \leftarrow z[sample(1:length(z), length(z), rep=F)] #Each time we
137
      sample a permutation of z
         perm.y.0 \leftarrow z.to.unif(outcome = y.z, beta1 = beta1s[i], beta2 =
138
      beta2s[j], permutation = perm.z, adj.mat = S.ideo)
```

```
perm.exposure <- indirect.treatment(permutation = perm.z, adj.mat
139
       = S.ideo)
140
         perm.test.stat \leftarrow sum((lm(perm.y.0 ~ perm.z + perm.exposure, na.
141
      action = na.omit)$resid)^2)
143
144
       # A vector of test statistics
145
       all.test.stat.vals <- as.numeric(unlist(results))</pre>
147
       # Calculating p-value
       pval <- sum(all.test.stat.vals < test.stat)/perms.test</pre>
149
150
     as.numeric(unlist(abc))
151
152 }
153
stopCluster(cl)
for (i in 1:length(betals)){
     pvals[i,] <- unlist(pvalues.ideology[i])</pre>
158
159
160 pvals
161
163 ## Creating a plot
image.plot(betals, beta2s, pvals,
```

```
main = "Plot of p-values", xlab = "Direct effects", ylab = "

Indirect effects")

166

167 pdf("pvalues_figure.pdf")

168 image.plot(betals, beta2s, pvals,

169 main = "Plot of p-values",

170 xlab = "Direct effects", ylab = "Indirect effects")

171 dev.off()
```

(Conventional approaches to causal inference rely on the Stable Unit Treatment Value Unit Assumption (SUTVA) i.e. they assume that the outcome of a unit is unaffected by the treatment statuses of other units. They rely on individual-level micro attributes or context-dependent macro attributes to estimate treatment effect. However, most social processes involve complex interaction and dependence among units connected through a network. In this situation, interpersonal interactions influence behavior leading to propagation of treatment effect through network to control units and SUTVA break down. Various factors impact how and how much the treatment effect spreads. It is important to take the interference structure into account in order to estimate the effect of the treatment effect correctly. In policy planning or marketing strategy, understanding the propagation of treatment itself can be of interest. It is generally not possible to identify the causal effects that map onto the process of social influence in observational data (Shalizi and Thomas 2011). Randomized experiments may not always be feasible. Therefore we look at methods suitable for identifying and estimating causal effect using field experiments.)

# References