

STAT 8700 Final Question 2

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2. (a) Consider a random sample y_1, y_2, \dots, y_n taken from a Normal population with mean $= \mu$ and known variance $= \sigma^2$. Show that the likelihood is equivalent to the likelihood of a single observation of \bar{y} taken from a Normal population with a mean of μ' and $\sigma^{2'}$ where \bar{y} is the mean of the y 's. Find the appropriate expressions for μ' and $\sigma^{2'}$.

Given multiple observations of a Normal distribution with mean μ and known variance σ^2 , the likelihood for observations y_1, y_2, \dots, y_n is given as

$$\begin{aligned} p(y_1, y_2, \dots, y_i | \sigma^2) &= \prod_{i=1}^n p(y_i | \sigma^2) \\ &= \prod_{i=1}^n \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2\sigma^2} (y_i - \mu)^2} \\ &= \left[\frac{1}{\sigma\sqrt{2\pi}} \right]^n e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2} \end{aligned}$$

where the data only appear in the exponential, $e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2}$. Using a trick from [1],

$$\begin{aligned} p(y_1, y_2, \dots, y_i | \sigma^2) &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2} \\ &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n [(y_i - \bar{y}) - (\mu - \bar{y})]^2} \\ &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \bar{y})^2 + \sum_{i=1}^n (\bar{y} - \mu)^2 - 2 \sum_{i=1}^n (y_i - \bar{y})(\mu - \bar{y})} \\ &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \bar{y})^2 + \sum_{i=1}^n (\bar{y} - \mu)^2 - (\mu - \bar{y}) \left(\sum_{i=1}^n y_i - n\bar{y} \right)} \\ &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \bar{y})^2 + \sum_{i=1}^n (\bar{y} - \mu)^2 - (\mu - \bar{y})(n\bar{y} - n\bar{y})} \\ &\propto e^{-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \bar{y})^2 + \sum_{i=1}^n (\bar{y} - \mu)^2} \\ &\propto e^{-\frac{1}{2\sigma^2} ns^2 + n(\bar{y} - \mu)^2} \\ &\propto e^{-\frac{ns^2}{2\sigma^2}} e^{-\frac{n(\bar{y} - \mu)^2}{2\sigma^2}} \end{aligned}$$

And since σ^2 is constant, we can ignore it,

$$p(y_1, y_2, \dots, y_i | \sigma^2) \propto e^{-\frac{n(\bar{y} - \mu)^2}{2\sigma^2}}$$

Now, we recognize this as a member of the exponential family of distributions, specifically a Gaussian,

$$p(y_1, y_2, \dots, y_i | \sigma^2) \sim \text{Normal}\left(\bar{y} | \mu, \frac{\sigma^2}{n}\right)$$

This shows the equivalence to the likelihood for a single observation of \bar{y} , and hence, $\mu' = \mu$ and $\sigma^{2'} = \frac{\sigma^2}{n}$. ■

(b) Suppose that y_i is Normally distributed with mean $= \mu$ and known variance $= \sigma_i^2$, for $i = 1, \dots, n$. Show that if a uniform prior for μ is used then the posterior distribution of μ is Normal with mean $= \frac{\sum_{i=1}^n y_i / \sigma_i^2}{\sum_{i=1}^n 1 / \sigma_i^2}$ and variance $= (\sum_{i=1}^n 1 / \sigma_i^2)^{-1}$

Show all working.

Using a uniform prior on μ we have the improper prior,

$$p(\theta) \propto 1$$

Then the conditional distribution given in **3.2** is

$$\mu | \sigma^2, y \sim N(\bar{y}, \sigma^2 / n)$$

which can be rewritten as

$$\mu | \sigma^2, y \sim N\left(\frac{1}{n} \sum_{i=1}^n y_i, \frac{1}{n} \sum_{i=1}^n \sigma_i^2\right)$$

Note that

$$\theta | \sigma_i^2, y_i \sim N\left(\frac{\sum_{i=1}^n y_i / \sigma_i^2}{\sum_{i=1}^n 1 / \sigma_i^2}, \left(\sum_{i=1}^n 1 / \sigma_i^2\right)^{-1}\right)$$

can be written as

$$\theta | \sigma_i^2, y_i \sim N\left(\sum_{i=1}^n y_i, \sum_{i=1}^n \sigma_i^2\right)$$

for $n = 1, 2, \dots, n$.

Thus, the distribution given in **3.2** is a single observation of y of all y_i from $i = 1, 2, \dots, n$. ■

References

- [1] Kevin P. Murphy
Conjugate Bayesian analysis of the Gaussian distribution
 University of British Columbia, 2007
<https://www.cs.ubc.ca/~murphyk/Papers/bayesGauss.pdf>