

Assessing Variable Importance Nonparametrically using Machine Learning Techniques

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Variable importance

- Data O_1, O_2, \dots, O_n from unknown distribution $P_0 \in \mathcal{M}$
 - $O_i := (X_i, Y_i)$
 - Covariate vector $X_i := (X_{i1}, X_{i2}, \dots, X_{ip}) \in \mathbb{R}^p$
 - Outcome $Y_i \in \mathbb{R}$
- Estimate $\mu_{P_0}(x) := E_{P_0}(Y \mid X = x)$
- Which features contribute most to variation in $\mu_{P_0}(x)$?
 - Consider $\mu_{P_0,s}(x) := E_{P_0}(Y \mid X_{(-s)} = x_{(-s)})$
 - $X_{(-s)}$ is the vector with the element(s) in $s \subseteq \{1, 2, \dots, p\}$ removed

Variable importance (continued)

- Fundamental questions:
 - How do we estimate μ_{P_0} and $\mu_{P_{0,S}}$?
 - How do we quantify variable importance?
- Approaches:
 - Parametric, e.g., ANOVA; must be **correctly specified**
 - Model-agnostic:
 - **Technique-specific measures**, e.g., random forests [Breiman (2001)]
 - **Technique-agnostic measures** [Doksum and Samarov (1995)], [van der Laan (2006), Chambaz et al. (2012), Sapp et al. (2014)]

Our goals

Flexible, Interpretable

- Estimate μ_{P_0} and $\mu_{P_{0,s}}$ using state-of-the-art methods
- Estimate a scientifically meaningful parameter consistently and efficiently
- Properly quantify the uncertainty in our estimates

The parameter of interest

- Additional proportion of variability in Y explained by including X_s in the regression:

$$\psi_{0,s} \equiv \Psi_s(P_0) := \frac{\int \{E_{P_0}(Y | X = x) - E_{P_0}(Y | X_{(-s)} = x_{(-s)})\}^2 dP_0(x)}{\text{var}_{P_0}(Y)}$$

- $\Psi_s(P_0)$ is a property of the data generating mechanism
- Interpretation **does not change** with estimating procedure
- Equivalent to difference in R^2 between the two regressions:

$$\Psi_s(P_0) = \frac{E_{P_0}[\{Y - \mu_{P_0}(X)\}^2]}{\text{var}_{P_0}(Y)} - \frac{E_{P_0}[\{Y - \mu_{P_{0,s}}(X)\}^2]}{\text{var}_{P_0}(Y)}$$

Efficient influence function?

- MLE $\hat{\theta}_n$ of θ_0 ; information $I(\theta_0)$, score $\dot{\ell}(\theta_0 | X)$
- Let $\tilde{\ell}(\theta_0 | X) = I^{-1}(\theta_0)\dot{\ell}(\theta_0 | X)$:
 - This is the **efficient influence function (EIF)** for θ_0
 - $\sqrt{n}(\hat{\theta}_n - \theta_0) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \tilde{\ell}(\theta_0 | X_i) + o_p(1)$
 - $\sqrt{n}(\hat{\theta}_n - \theta_0) \rightarrow_d N \left[0, E_{P_0} \left\{ \tilde{\ell}(\theta_0 | X)^2 \right\} \right] = N\{0, I^{-1}(\theta_0)\}$
- Given an EIF for a nonparametric parameter:
 - Estimator with influence function = EIF **is efficient**
 - Can use similar distribution theory to parametric case

The EIF for $\Psi_s(P)$ relative to \mathcal{M}

$$\begin{aligned}\mu_P(x) &= E_P(Y \mid X = x) \\ \mu_{P,s}(x) &= E_P(Y \mid X_{(-s)} = x_{(-s)}) \\ \phi_s(P) &= \int \{\mu_P(x) - \mu_{P,s}(x)\}^2 dP(x)\end{aligned}$$

Then

$$\begin{aligned}o \mapsto D_{P,s}^*(o) &:= \frac{2\{y - \mu_P(x)\}\{\mu_P(x) - \mu_{P,s}(x)\} + \{\mu_P(x) - \mu_{P,s}(x)\}^2}{\text{var}_P(Y)} \\ &\quad - \phi_s(P) \left\{ \frac{y - E_P(Y)}{\text{var}_P(Y)} \right\}^2\end{aligned}$$

Asymptotic expansion

- Estimate the relevant components of P_0 using \hat{P}_n
- Linearize Ψ using the EIF $D_{\hat{P}_n,s}^*$ and use the empirical \mathbb{P}_n :

$$\begin{aligned}\Psi_s(\hat{P}_n) - \Psi_s(P_0) &= \int D_{\hat{P}_n,s}^*(o) d(\hat{P}_n - P_0)(o) + R_s(\hat{P}_n, P_0) \\ &= \frac{1}{n} \sum_{i=1}^n D_{P_0,s}^*(O_i) \\ &\quad + \int \{D_{\hat{P}_n,s}^*(o) - D_{P_0,s}^*(o)\} d(\mathbb{P}_n - P_0)(o) \\ &\quad + R_s(\hat{P}_n, P_0) - \frac{1}{n} \sum_{i=1}^n D_{\hat{P}_n,s}^*(O_i)\end{aligned}$$

- | | |
|---------------------------|-------------------------|
| ■ linear term; | (1 st order) |
| ■ empirical process term; | (2 nd order) |
| ■ remainder term; | (2 nd order) |
| ■ problem term! | (irregular) |

A naive estimator of $\Psi_s(P_0)$

$$\psi_{0,s} \equiv \Psi_s(P_0) = \frac{\int \{\mu_{P_0}(x) - \mu_{P_{0,s}}(x)\}^2 dP_0(x)}{\text{var}_{P_0}(Y)}$$

- Given estimators $\hat{\mu}(x)$ and $\hat{\mu}_s(x)$
- Plug in:

$$\hat{\psi}_{\text{naive},s} = \frac{n^{-1} \sum_{i=1}^n \{\hat{\mu}(X_i) - \hat{\mu}_s(X_i)\}^2}{n^{-1} \sum_{i=1}^n (Y_i - \bar{Y}_n)^2}$$

Problems with the naive estimator

$$\begin{aligned}\psi_s(\hat{P}_n) - \psi_s(P_0) &= \frac{1}{n} \sum_{i=1}^n D_{P_0,s}^*(O_i) + R_s(\hat{P}_n, P_0) - \frac{1}{n} \sum_{i=1}^n D_{\hat{P}_n,s}^*(O_i) \\ &\quad + \int \{D_{\hat{P}_n,s}^*(o) - D_{P_0,s}^*(o)\} d(\mathbb{P}_n - P_0)(o)\end{aligned}$$

- “Bias” incurred from estimating components of P_0
- Generally neither efficient nor regular and asymptotically linear

The one-step estimator

- Remove bias and get regularity, asymptotic linearity, and efficiency by adding on $\frac{1}{n} \sum_{i=1}^n D_{\hat{P}_{n,s}}^*(O_i)$:

$$\hat{\psi}_{n,s} = \hat{\psi}_{\text{naive}, s} + \frac{1}{n} \sum_{i=1}^n D_{\hat{P}_{n,s}}^*(O_i),$$

or equivalently

$$\hat{\psi}_{n,s} = \hat{\psi}_{\text{naive}, s} + \frac{n^{-1} \sum_{i=1}^n 2\{Y_i - \hat{\mu}(X_i)\}\{\hat{\mu}(X_i) - \hat{\mu}_s(X_i)\}}{n^{-1} \sum_{i=1}^n (Y_i - \bar{Y}_n)^2}$$

Asymptotic behavior of the one-step estimator

Under some regularity conditions,

$$\sqrt{n}(\hat{\psi}_{n,s} - \psi_{0,s}) = n^{-1/2} \sum_{i=1}^n D_{P_{0,s}}^*(O_i) + o_P(1)$$

and

$$\sqrt{n}(\hat{\psi}_{n,s} - \psi_{0,s}) \rightarrow_d N \left[0, E_{P_0} \{ D_{P_{0,s}}^*(O)^2 \} \right].$$

- Consistent, regular, efficient
- Regularity conditions:
 - $\psi_{0,s} \neq 0$
 - $\hat{\mu}, \hat{\mu}_s$ converge quickly enough
 - $D_{\hat{P}_{n,s}}^*$ falls in a P_0 -Donsker class with probability tending to one
- Estimate variance of $\hat{\psi}_{n,s}$ empirically

Simulations with a low-dimensional vector of covariates

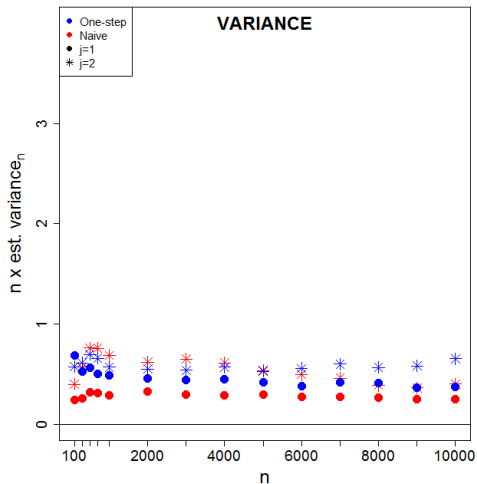
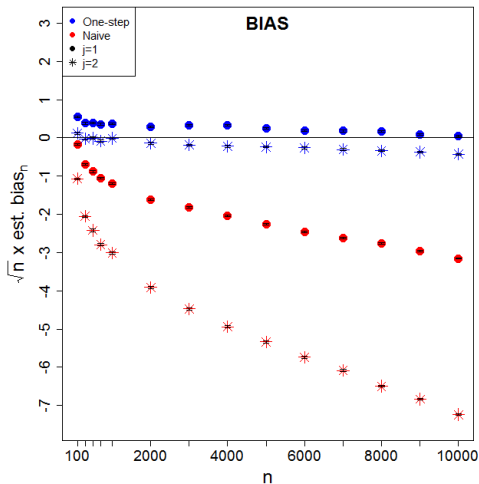
- Data:

$X_1, X_2 \stackrel{iid}{\sim} \text{Unif}(-1, 1)$ and $\epsilon \sim N(0, 1)$ independent of (X_1, X_2)

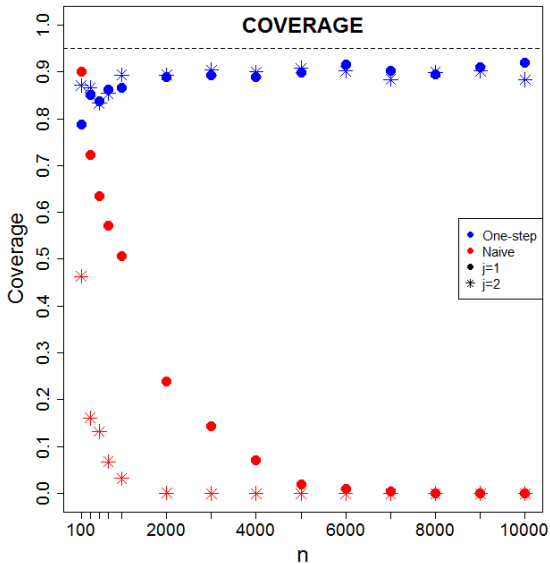
$$Y = X_1^2 \left(X_1 + \frac{7}{5} \right) + \frac{25}{9} X_2^2 + \epsilon$$

- Truths: $\psi_{0,1} \approx 0.158$, $\psi_{0,2} \approx 0.342$
- Locally-constant loess, five-fold CV to obtain optimal bandwidths
- Percentile bootstrap for naive confidence intervals

Results



Results



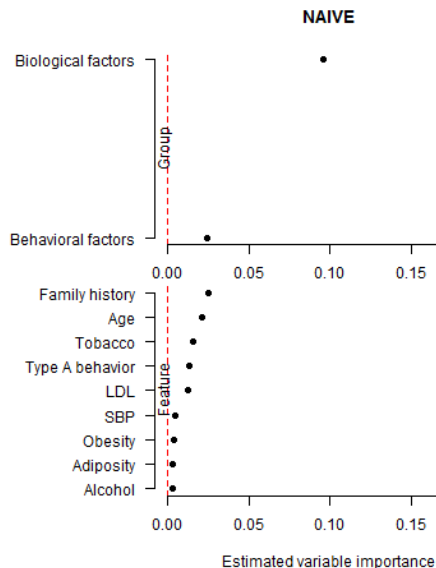
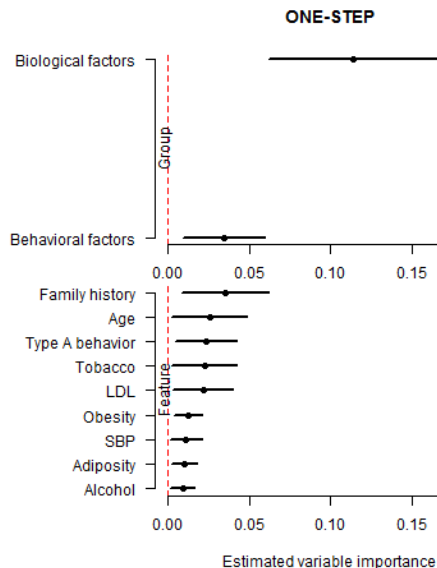
The CORIS data [Rousseaw et al. (1983)]

$n = 462$, outcome = presence of MI

- Behavioral:
 - tobacco consumption,
 - alcohol consumption,
 - type-A behavior
- Biological:
 - systolic blood pressure,
 - LDL cholesterol,
 - adiposity,
 - obesity,
 - family history,
 - age

Super learner [van der Laan et al. (2007)] with boosted trees, elastic net, GAMs, random forests, and five-fold CV

Results from the CORIS data



Conclusions

- **Interpretable**: Additional proportion of variability in Y explained by including X_s in the estimation technique
- **Flexible**: Valid CIs with state-of-the-art methods!
- Consistently and efficiently estimate a property of the data generating mechanism
- Reasonable performance in simulation and in data analysis
- Implemented in R package `vimp` and Python package `vimpy`
- Future work:
 - dealing with a boundary null hypothesis,
 - working in a structured model (e.g., additive models),
 - nested case-control study data,
 - censoring

References

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