Design considerations for subgroup analyses in cluster-randomized trials

Brian D. Williamson, PhD

Kaiser Permanente Washington Health Research Institute

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https://bdwilliamson.github.io/#talks

Acknowledgments

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Yates Coley

KAISER PERMANENTE

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Health Research Institute



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Kaiser Permanente Washington
Health Research Institute



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CRE Center for
Research and Evaluation



Andrea Cook

KAISER PERMANENTE

Kaiser Permanente Washington
Health Research Institute

Heterogeneity of treatment effects: what and why

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If HTE is present:

- resources can be prioritized to maximize benefit and reduce risks [Knol and VanderWeele (2012)]
- can determine whether intervention remediates or exacerbates health disparities [Petkovic et al. (2020)]

ENSPIRE: Engaging Staff to Improve COVID-19 Vaccination Response at Long-term Care Facilities

Background:

- Rates of COVID-19 disease and death higher in long-term care than general population [Chidambaram (2022)]
- Lag in COVID-19 vaccinations among long-term care workers
- Long-term care (LTC) workers often low-wage, represent diverse social, cultural, racial, and ethnic backgrounds

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Design:

- Enrolled 40 LTC centers, encompassing approximately 4000 staff members
- Recruited from 2 US states (Georgia and Washington)
- Recruited from urban (Atlanta or Seattle) and suburban/rural locations

Primary outcomes:

- 1. LTC center COVID-19 booster rate (cluster-level)
 - o chosen due to concerns about data capture
- 2. LTC care staff net promoter score (individual-level)

Intervention: booster promotion materials co-designed with LTC staff

VS.

"Usual care" (Standard materials)

Why You Should Stay Up to Date with Your COVID-19 Vaccines

Getting vaccinated and boosted greatly reduces your risk of severe illness, hospitalization, and death.

COVID-19 vaccines, including boosters, are effective against known variants, including Omicron.

All COVID-19 vaccines currently approved or authorized in the U.S. are proven to be safe.

"Co-design" (Materials designed by teams of staff)





Co-design: tailoring messages to specific audiences

- In ENSPIRE, focused on language, cultural, and ethnic affinity groups
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HTE analysis natural for interventions, particularly co-designed:

- Planned analyses based on state, urban vs suburban/rural, baseline booster uptake
- Self-reported race/ethnicity both interesting and important, however:
 - neither sites nor other entity had comprehensive and accurate race and ethnicity data for staff
 - o staff survey would only capture a portion of staff members
 - also, didn't know what this would look like with a cluster-level outcome!

Individual-level data structure

Data on n_c staff members in each cluster $c=1,\ldots,K$:

- A_{ci} : indicator of intervention assignment (cluster-level)
- $Z_{ci} \in \mathbb{R}^q$: cluster-level subgroup variables
- $W_{ci} \in \mathbb{R}^p$: baseline covariates
- Y_{ci} : binary outcome

HTE with individual-level outcomes

Suppose we are interested in HTE by a single binary variable W_{ci} .

Target estimand: the difference in the risk difference based on levels of W_{ci} .

Estimate using linear regression:

$$E(Y_{ci} \mid A_{ci} = a, Z_{ci} = z, W_{ci} = w) = \beta_0 + \beta_1 a + \beta_2 z + \beta_3 w + \beta_4 aw.$$

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Under standard assumptions, β_4 quantifies HTE.

Cluster-level data structure

In many settings, we observe cluster-level data:

- ullet $\overline{Y}_c:=rac{1}{n_c}\sum_{i=1}^{n_c}Y_{ci}$: cluster-level outcome proportion
- A_c : indicator of intervention assignment
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- ullet $X_c \in \mathbb{R}^p$: (possibly aggregated) baseline covariates

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Possible options for aggregating an individual-level covariate W_{cj} :

$$ullet$$
 $X_{cj}=V_{cj}-rac{1}{K}\sum_{c=1}^K V_{cj}$, where $V_{cj}=rac{1}{n_c}\sum_{i=1}^{n_c} W_{cji}$

$$ullet$$
 $X_{cj}=I\left(V_{cj}>t
ight)$ for threshold t

Aggregated data: complications

Individual-level effects may not be identifiable based on aggregate data

- Often referred to as ecological bias [Wakefield (2008)]
- Exposure-outcome association can be biased: [Greenland & Morgenstern (1989)]
 - by aggregated confounders
 - by aggregated HTE variables

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Compositional data cause challenges in interpretation:

- Arise when aggregating mutually-exclusive binary variables
- Changing the value of one proportion necessarily changes at least one other

Now, outcome \overline{Y}_c is a proportion; interest in HTE by single variable X_c .

Target estimand: difference in difference of proportions based on values of X_c .

Estimate using linear regression:

$$E(\overline{Y}_c \mid A_c = a, Z_c = z, X_c = x) = lpha_0 + lpha_1 a + lpha_2 z + lpha_3 x + lpha_4 a x.$$

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If W represents a single binary variable, then aggregation determines interpretation:

- Using mean-centered proportion: difference in difference for 1-unit increase from the across-cluster mean proportion
- Using threshold: difference in difference comparing "high" to "low"

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If W is compositional, this is further complicated:

- how are the categories transformed to enable model fitting?
- what does a one-unit change imply?

Simulation setup

Data mimic FNSPIRF trial:

- Sample $K \in \{30, 40, 80\}$ clusters within two regions
- Clusters are urban or suburban/rural
- Randomization constrained to balance on region, urban vs suburban/rural
- Each center has 100 staff members, who each:
 - have record of receiving COVID-19 booster vaccine
 - have self-reported race (recorded as single binary variable)
- Define cluster-level booster rate \overline{Y}_c
- Define two aggregations of self-reported race:
 - center-level proportion
 - o indicator of whether proportion > 0.5

Simulation setup

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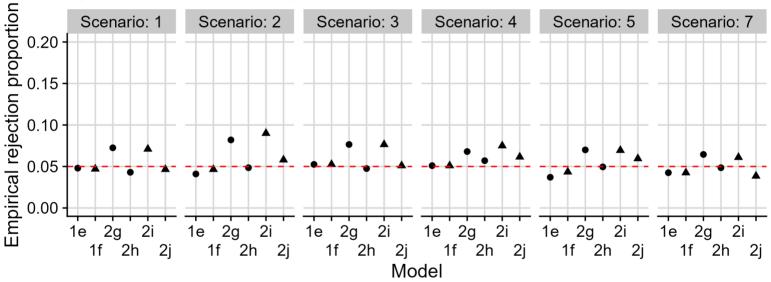
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Fit both individual-level and cluster-level regression models

Tested for HTE by self-reported race, investigated type I error and power

Results: type I error



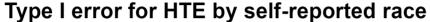


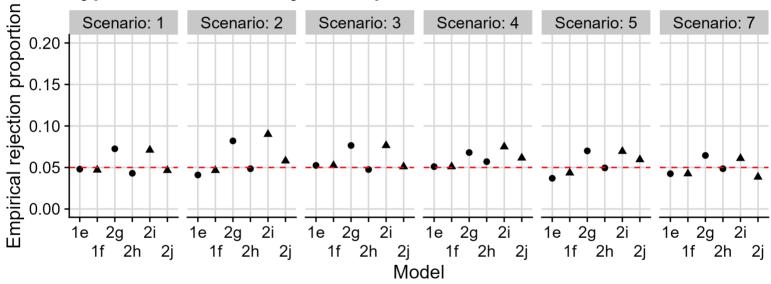
- Interaction variable(s)
- Self-reported race
- Self-reported race and urban status

Scenarios encode presence of true differences in booster rate or HTE:

- 1, 2, 5: no difference in booster rate by self-reported race, no HTE
- 3, 4, 7: difference in booster rate by self-reported race, no HTE

Results: type I error





Interaction variable(s)

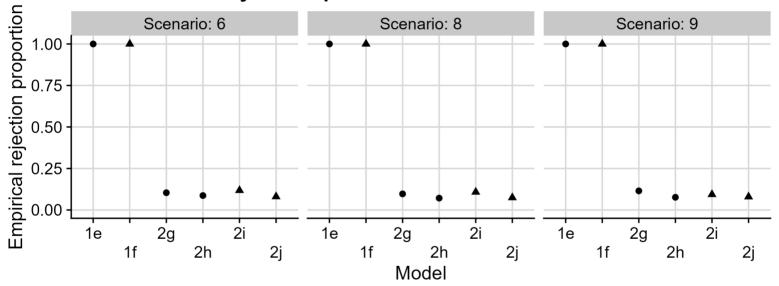
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Models encode individual- (1) vs cluster-level (2), adjustment and interaction variables:

- 1e, 1f: adjust for self-reported race, allow HTE by self-reported race (1f also by urban)
- 2g, 2i: adjust for center-level race proportion, allow HTE by self-reported race (2i also by urban)
- 2h, 2j: adjust for proportion > 0.5, allow HTE by self-reported race (2j also by urban)

Results: power

Power for HTE by self-reported race



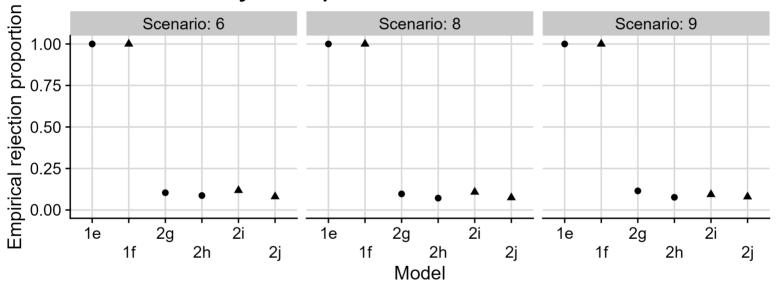
- Interaction variable(s)
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- Self-reported race and urban status

Scenarios encode presence of true differences in booster rate or HTE:

- 6: difference in booster rate by self-reported race, HTE by self-reported race
- 8: difference in booster rate by urban and self-reported race, HTE by self-reported race
- 9: difference in booster rate by urban and self-reported race, HTE by urban and self-reported race

Results: power

Power for HTE by self-reported race



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Closing thoughts

High power, controlled type I error to detect HTE in the individual-level analysis

where possible, this approach is key for identifying an remedying disparities

Low power to detect HTE by an aggregated individual-level predictor

- even in simple setting with only two possible categories!
- likely exacerbated in more realistic settings
- elevated type I error when using proportion in some settings

However, meaningful estimated differences across groups can still be acted upon

Power calculations should aggregate simulated individual-level data, rather than simulating cluster-level data directly.

For more details, including more realistic simulation settings:

Williamson BD, Coley RY, Hsu C, McCracken CE, Cook AJ (2023). Considerations for subgroup analyses in cluster-randomized trials based on individual-level predictors. *Prevention Science*.

References

- Knol MJ and VanderWeele TJ (2012). Recommendations for presenting analyses of effect modification and interaction. *International Journal of Epidemiology*.
- Petkovic et al. (2020). Reporting of health equity considerations in cluster and individually randomized trials. *Trials*.
- Chidambaram P (2022). Over 200,000 residents and staff in long-term care facilities have died from COVID-19. Kaiser Family Foundation.
- Wakefield J (2008). Ecologic studies revisited. Annual Review of Public Health.
- Greenland S & Morgenstern H (1989). Ecological bias, confounding, and effect modification.
 International Journal of Epidemiology.