Q. IV condition for time-series.

but E(T+1 U+) +0 -> Endogeneity phoblem.

⇒ To handle the endogeneity, use IV.

## IV candidate: Xt-1

- 1 E(Xt-1 Ut) = O (This is likely given that E(UtiXt)=0)
- ② E(Xt-1 Tt-1) is correlated.

  "B3 ≠ 0" means "Tt-1 and Xt-1 GWelated."

  i.e., Tt-1 = B1 + B2 Tt-2 + B3 Xt-1 + Ut-1.
- \* Lagged Dependent Variables + Serial arrelation. (continued:)
  - @ Coming up with IV in Data (Natural Experiments).

Examples: Twins, Natural phenomenon, weather, earthquakes e.t.c. Some form of discontinuity arising from "participation" constraint.

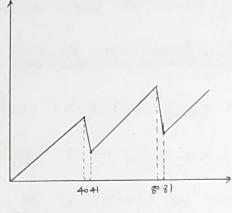
"Class-size effect on test-score". Angrist -Lavy. QJE.

Tis =  $\beta_1 + \beta_2 Dis + \beta_3 X_{25} + U_{25}$ , Dis = 1 if class size is small o if " is not small.

Let's think of only one school case.

 $T_i = \beta_1 + \beta_2 D_i + \beta_3 X_i + U_i$ ,  $D_i = \begin{pmatrix} 1 & i + class size is small \\ 0 & i + \end{pmatrix}$ , not small  $\begin{pmatrix} c_{40} \end{pmatrix}$ 

class size



When we see 20 or 30 or even 40,

it is just "I" class so that it is small.

But if the total student is 41, there should exist "2" class

so that it is not small.

total student in a grade.

In this rase, if we compare 35-40 vs 41-45. In a bandwidth selection publican.

 $\rightarrow$  Thus. adjusting its bandwidth, 35-40 vs 20.5~22.5 38-40 vs 20.5~21.5 We can compare them.

(We will talk about it again at Non-parametric part).

## \* How to test the validity of IV?

## Test of over-identifying restrictions.

From IV model, g:(B) = Z:(y:-x:/B).

Then, O E(U, Zi) =0 Q E(Zi(Xi) has full rank.

=> This should be checked.

Test of over-identifying restrictions: Check O E(U; Zi) = 0.

Note / Hansen's J-test>

Under E(XiVi) +0, (Endogeneity),

 $J = J(\hat{\beta}_{gmm}) \xrightarrow{d} \mathcal{X}'_{(J-K)}$ 

 $H_0: E(U_i, Z_i) = 0 \Rightarrow Pr(J > C | H_0) \rightarrow C$ (Significant level)

(rejection)

· It we have K IVs, (i.e., J=K), \(\sigma \sum\_{\infty} \in \text{Z} \text{(T-X'\(\beta\_{\text{LV}}\))=0}\)
\(\begin{array}{c} \begin{array}{c} \begin{array}{c

biv solves & (T-\*b)=0

When J=k. if we use IV residual ûs to test E(u; Zi)=0.

\[ \int\_{\infty}^{\infty} Z\_{\infty} \infty = \int\_{\infty}^{\infty} Z\_{\infty} \left( \infty - \times \infty \infty \right) = 0 \quad \text{50} \text{ that this is always zero.} \]

( When J=k, E(UiZi)=0 always holds).

· However, if we have more than k IVs, (J>k), then NOT all equations can be set to zero,

So that we need to test "extra equalites J-K": over-identifying restriction test.

· Which K equations are set to zero to GIMM?

At first, Let's think of GIMM estimator, when J>K

min g(b) Wg(b) where g(b) = /Z'(T-xb).

Now. Z is NXJ matrix.

min / (1-x6) ZW / Z'(1-x6)

F.o.c  $0 = -\frac{1}{N} \frac{1}{N} \frac$ 

Note  $\langle \mathcal{X} \mathcal{Z} \mathcal{W} \left[ \mathcal{Z}' \mathcal{M} - \mathcal{K} \mathcal{B} \right] = 0 \rangle$  $\begin{pmatrix} \mathcal{B}_{1} \\ \mathcal{S}_{N} \end{pmatrix} = \begin{pmatrix} \mathcal{X}_{11} \\ \mathcal{X}_{N1} \\ \mathcal{S}_{N} \end{pmatrix} \begin{pmatrix} \hat{b}_{1} \\ \hat{b}_{K} \end{pmatrix} + \begin{pmatrix} \hat{u}_{1} \\ \hat{u}_{N} \end{pmatrix} \rightarrow \mathcal{B}_{1} = \hat{b}_{1} \mathcal{M}_{11} + \hat{b}_{2} \mathcal{M}_{12} + \dots + \hat{b}_{K} \mathcal{M}_{NK} + \hat{u}_{N}$   $\begin{pmatrix} \mathcal{B}_{1} \\ \mathcal{S}_{N} \end{pmatrix} = \begin{pmatrix} \mathcal{M}_{11} \\ \mathcal{S}_{N1} \\ \mathcal{S}_{N1} \end{pmatrix} \begin{pmatrix} \hat{b}_{1} \\ \hat{b}_{2} \end{pmatrix} + \begin{pmatrix} \hat{u}_{1} \\ \hat{u}_{N} \end{pmatrix} \rightarrow \mathcal{B}_{1} = \hat{b}_{1} \mathcal{M}_{11} + \hat{b}_{2} \mathcal{M}_{12} + \dots + \hat{b}_{K} \mathcal{M}_{NK} + \hat{u}_{N}$ 

\* \* W [ & (1- x6)] = \* \* \* W [ & 'û] = 0

= \( \lambda\_{11} \cdots \lambda\_{NN} \) \( \lambda\_{11} \cdots \lambda\_{11} \cdots \lambda\_{11} \cdots \lambda\_{11} \cdots \lambda\_{12} \cdots \lambda\_{11} \cdots \l

 $= \int_{\mathbb{R}^{N}} X_{N} Z_{N} \qquad \qquad \int_{\mathbb{R}^{N}} X_{N} Z_{N} Z_{N} Z_{N} Z_{N} Z_{N} \qquad \qquad \int_{\mathbb{R}^{N}} X_{N} Z_{N} Z_{N} Z_{N} Z_{N} Z_{N} \qquad \qquad \int_{\mathbb{R}^{N}} X_{N} Z_{N} Z_{N$ 

1000

@ ZII (YI-GIAII- -- GK XIK) + ZZI (Yz-GIAZI - GZAZI - ... - GK XZK) + ... + ZNI(BN-GIANI-...

$$= \sum_{j=1}^{N} Z_{ij} (y_{i} - b_{i} x_{i1} - b_{2} \cdot x_{i2} - \dots - b_{k} x_{ik}) = \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij}) = \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij}) = \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij})$$

$$= \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij}) = \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij})$$

$$= \sum_{j=1}^{N} Z_{ij} (y_{i} - \frac{k}{j} b_{j} x_{ij})$$

Coming back to Main story,

We test how close  $\hat{g}(6)$   $\hat{w}\hat{g}(6)$  is to Zero, using  $W = \hat{\Omega}^{-1}$  (optimal weight)

(\* Homsen's J-test > Use efficient GMM so that W= 121).

LOC: \* & M [ R ( 1 - \* 6] = 0 => BOWN = ( SE - S. ) - X SE - S. L

$$\hat{g}'(\hat{b}) = \frac{1}{N} Z'(\Pi - \chi \hat{b}) = \frac{1}{N} Z'(\chi \beta + u - \chi \hat{b}) = \frac{1}{N} Z'(u - \chi (\hat{b} - \beta))$$

$$\hat{b} = \beta + \chi Z \hat{b}^{-1} Z'(\chi)^{-1} \chi' Z \hat{b}^{-1} Z'(u).$$

$$= A$$

= /Z'(UI - X(A-1) X/ZA-1Z'UI)

= 1/(I- Z\* A-1 \* Z & A-1) Z (U)

Thus.

\$'(6) W \$(6) = \ (11/2 (I-(\hat{O}^{-1})'2/4(A^{-1})'4/2) \hat{O}^{-1} \ \(\hat{I}-2/4A^{-1} \(\hat{Z}-\hat{O}^{-1}) \(\hat{Z}

Note Homsen's J-test:  $J = N \cdot \hat{g}'(\hat{b}) W \hat{g}(\hat{b})$  where  $W = \hat{\Omega}^{-1}$ ,  $\hat{b} = \hat{b}gmm$ 

So, multiply N and adjust them as IN

N. g'(6) W g(6) = UI'\(\infty\) (I-(\hat{A}^{-1})'\(\infty\) (A')\(\infty\) (I-\(\infty\) (A'\(\infty\) (I-\(\infty\) (A'\(\infty\) (I-\(\infty\) (A'\(\infty\) (I-\(\infty\) (I-\(\infty\) (A'\(\infty\) (I-\(\infty\) (I-\(\inft

- 0 / Z'u → N(0, E(u; ZZ')) by CLT.
- By cholesky decomposition,  $\Omega^{-1} = \Omega^{-\frac{1}{2}} \Omega^{-\frac{1}{2}}$  is possible.
- Bgmm-β= (\*/2Ω-12/4) 1/2Ω-12/4 = (-1×2Ω-1-12/4) 1/2Ω-1-1/2/4.

  ⇒ We can use this skill to part ③.

 $\frac{P}{P} = \frac{\mathbb{E}(\mathcal{Z}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1} \mathcal{Z}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{Z}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{Z}_{1} \mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{Z}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{Z}_{1} \mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1} \mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}(\mathcal{X}_{1}^{\prime})}{\mathbb{E}(\mathcal{X}_{1}^{\prime})} \frac{\mathbb{E}$ 

Thus, @ is

 $(I - \mathbb{Z}' / (A^{-1}) / \mathbb{Z} \hat{\Omega}^{-1}) \xrightarrow{P} I - E(\mathbb{Z} / \Omega') [E(\Omega : \mathbb{Z}') \Omega^{-1} E(\mathbb{Z} : \Omega')]^{-1} E(\Omega : \mathbb{Z}') \Omega^{-1}$ Using  $\Omega^{-\frac{1}{2}}$  from @ and combining @ to it,

Considering  $\overline{M}$ ,  $(I-\Omega^{-\frac{1}{2}} \overline{M} \Omega^{-\frac{1}{2}}) \Omega^{\frac{1}{2}} = (I-\Omega^{-\frac{1}{2}} \overline{M} \Omega^{-\frac{1}{2}}) \Omega^{\frac{1}{2}}$   $= (I-\Omega^{-\frac{1}{2}} \overline{M} \Omega^{-\frac{1}{2}}) \Omega^{-\frac{1}{2}} \cdot N(o, E(u, z, z, z, 1))$   $= (I-\Omega^{-\frac{1}{2}} \overline{M} \Omega^{-\frac{1}{2}}) \cdot N(o, I_J)$ 

Therefore.

Therefore.

N. 
$$\hat{g}'(\hat{G}) \otimes \hat{g}(\hat{G}) = \frac{u''''''}{\sqrt{N}} (I - \hat{G}^{-1})'''''''' \otimes \hat{G}^{-1} (I - 2/N A^{-1} \times 2/A^{-1}) \times 2/A^{-1} \times 2/A^{$$

$$\Theta \Omega^{-\frac{1}{2}} = \Omega^{-\frac{1}{2}} E(Z; M;') [E(M; Z;') \Omega^{-1} E(Z; M;')]^{-1} E(M; Z;') \Omega^{-\frac{1}{2}}$$

$$\int_{Jxk} \int_{Kxk} \int_{Kxk} \int_{Kx} \int_{Kx}$$

= 
$$N(0, I_5)(I - B(B'B)^{\dagger}B) (I - B(B'B)^{\dagger}B) \cdot N(0, I_5)$$
  
=  $N(0, I_5)(I_5 - B(B'B)^{\dagger}B) \cdot N(0, I_5) \sim \chi^{\dagger} rank(I_5 - B(B'B)^{\dagger}B)$   
(5) Using the property of trace.  
+ trace  $(I_5 - B(B'B)^{\dagger}B) = trace(I_5) - trace(BCB'B)^{\dagger}B)$   
=  $J - trace(B'B)^{\dagger}B'B) = J - trace(I_k) = J - k$ .

~ NITEK)

Through J-test, we can check Ho: E(u; Z;)=0That is,  $J=N\cdot\hat{g}(b)\cdot\hat{\Omega}^{-1}\hat{g}(b)$   $\longrightarrow \chi_{(J-K)}^{+}$ .

Then. if  $H_0: E(U_1:Z_1)=0$  cannot be rejected, we have  ${}^0E(U_1:Z_1)=0$  in even J>K case cover-identifying restrictions). After that, we need to check  ${}^0F_1II$  rank and then of  $E(Z_1:X_1')$ . It is easy to check  $E(Z_1:X_1')$  has full rank b/c we have data.

Overidentification.

(a)  $E(Z_i X_i')$  is invertible