

2

$$\begin{aligned} x'_i \hat{\beta} - x'_i \hat{\beta}_{(i)} &= \frac{x'_i \left(X'_{(i)} X_{(i)} \right)^{-1} x_i}{1 + x'_i \left(X'_{(i)} X_{(i)} \right)^{-1} x_i} \left(y_i - x'_i \hat{\beta}_{(i)} \right) \\ &= \frac{x_{1i}^2 + \frac{1}{5} x_{2i}^2}{1 + x_{1i}^2 + \frac{1}{5} x_{2i}^2} \left(y_i - x'_i \hat{\beta}_{(i)} \right). \end{aligned}$$

If the size of x_i is fixed, which means $x_{1i}^2 + x_{2i}^2 = C$ for some constant C , then $x_{1i}^2 + \frac{1}{5} x_{2i}^2$ is maximized when $x_{1i}^2 = C$ and $x_{2i}^2 = 0$, and is minimized if $x_{1i}^2 = 0$ and $x_{2i}^2 = C$. Therefore, the impact is larger if x_{1i} is relatively more deviated from the origin, and is smaller if x_{2i} is relatively more deviated from the origin.

4

Let $X_n(w) : [0, 1] \rightarrow \mathbb{R}$ be as $X_n(w) = \begin{cases} k+1 & \text{if } \frac{n-K+1}{k+1} > w > \frac{n-K}{k+1} \\ 0 & \text{otherwise} \end{cases}$ where $k = \max \left\{ k \in \mathbb{N}^+ : \frac{k(k+1)}{2} \leq n \right\}$ and $K = \frac{k(k+1)}{2}$.

- $\{X_n\}$ converges to 0 in probability because given any small positive ϵ ,

$$\Pr \{|X_n(w) - 0| < \epsilon\} = 1 - \frac{1}{k+1} \rightarrow 0 \quad \text{as } n \rightarrow \infty.$$

- $\{X_n\}$ does not converges to 0 in r th moment for any $r \geq 1$, as

$$E \{|X_n(w) - 0|^r\}^{\frac{1}{r}} = \left(k+1 \frac{1}{k+1} \right)^{\frac{1}{r}} = 1 \not\rightarrow 0.$$

- $\{X_n\}$ does not converges to 0 almost surely, because if we fix some $w \in (0, 1)$, given any positive ϵ for any N , there exists some \tilde{k} and $\tilde{n} > N$ such that

$$\frac{\tilde{n} - \tilde{K} + 1}{\tilde{k} + 1} > w > \frac{\tilde{n} - \tilde{K}}{\tilde{k} + 1}$$

which implies $|X_{\tilde{n}}(w) - 0| = |\tilde{k} + 1| > \epsilon$, which means $X_n(w) \not\rightarrow 0$. Therefore, $\Pr \{\lim_{n \rightarrow \infty} X_n(w) = 0\} = 0 \neq 1$.

5

(a) The CDF for a constant c is

$$F_c(t) = \Pr(c \leq t) = \begin{cases} 0 & \text{if } c > t \\ 1 & \text{if } c \leq t \end{cases}.$$

(b) The condition: for each $t \in \mathbb{R}$,

$$F_n(t) \rightarrow F_c(t) \quad \text{as } n \rightarrow \infty.$$

(c) For any $\epsilon > 0$,

1. $F_n(c + \epsilon) \rightarrow F_c(c + \epsilon) = 1$,
2. $F_n(c - \epsilon) \rightarrow F_c(c - \epsilon) = 0$.

(d)

$$\begin{aligned}
\Pr(|X_n - c| \leq \epsilon) &= \Pr(-\epsilon \leq X_n - c \leq \epsilon) \\
&= \Pr(c - \epsilon \leq X_n \leq c + \epsilon) \\
&= \Pr(X_n \leq c + \epsilon) - \Pr(X_n < c - \epsilon) \\
&= \Pr(X_n \leq c + \epsilon) - \Pr(X_n < c - \epsilon) + \Pr(X_n = c + \epsilon) \\
&= F_n(c + \epsilon) - F_n(c - \epsilon) + \Pr(X_n = c + \epsilon).
\end{aligned}$$

(e) Given any ϵ ,

$$\begin{aligned}
\Pr(|X_n(w) - c|) &= F_n(c + \epsilon) - F_n(c - \epsilon) + \Pr(X_n = c + \epsilon) \\
&\rightarrow 1 + \Pr(X_n = c + \epsilon).
\end{aligned}$$

Notice that

$$\begin{aligned}
\Pr(X_n = c + \epsilon) &= \Pr(X_n \leq c + \epsilon) - \Pr(X_n < c + \epsilon) \\
&\leq \Pr(X_n \leq c + \epsilon) - \Pr\left(X_n < c + \frac{\epsilon}{2}\right) \\
&\rightarrow 1 - 1 \\
&= 0.
\end{aligned}$$

Therefore,

$$\Pr(|X_n(w) - c|) \rightarrow 1$$

which implies X_n converges to c in probability.

6

(a)

$$\begin{pmatrix} A & C \\ B' & D \end{pmatrix} \begin{pmatrix} W & Y \\ X' & Z \end{pmatrix} = \begin{pmatrix} AW + CX' & AY + CZ \\ B'W + DX' & B'Y + DZ \end{pmatrix}$$

If this equals to $\begin{pmatrix} I & 0 \\ 0' & I \end{pmatrix}$, then we have

$$AW + CX' = I \quad (1)$$

$$B'W + DX' = 0' \quad (2)$$

$$AY + CZ = 0 \quad (3)$$

$$B'Y + DZ = I \quad (4)$$

By (2) and (3), we have

$$X' = -D^{-1}(B'W)$$

$$Y = -A^{-1}(CZ)$$

Substitue this into (1)

$$AW - CD^{-1}(B'W) = I$$

$$(A - CD^{-1}B')W = I$$

$$(A - CD^{-1}B')^{-1} = W.$$

and

$$X' = -D^{-1}B'(A - CD^{-1}B')^{-1}.$$

Similarly, in (4) we have

$$\begin{aligned} -B'A^{-1}(CZ) + DZ &= I \\ (D - B'A^{-1}C)Z &= I \\ (D - B'A^{-1}C)^{-1} &= Z. \end{aligned}$$

and

$$Y = -A^{-1}C(D - B'A^{-1}C)^{-1}.$$

(b) By the definition of inverse matrix, if

$$AA^{-1} = I,$$

then

$$A^{-1}A = I.$$

(c) Using the relationship implied in (b), we have

$$\begin{aligned} WA + YB' &= I \\ X'C + ZD &= I, \end{aligned}$$

which implies

$$\begin{aligned} W &= A^{-1} - YB'A^{-1} \\ Z &= D^{-1} - X'CD^{-1}. \end{aligned}$$

Using the results in (1), we have

$$Y = -A^{-1}C(D - B'A^{-1}C)^{-1},$$

therefore,

$$W = A^{-1} + A^{-1}C(D - B'A^{-1}C)^{-1}B'A^{-1}.$$

Similarly, the results in (1) implies

$$X' = -D^{-1}B'(A - CD^{-1}B')^{-1},$$

and therefore

$$Z = D^{-1} + D^{-1}B'(A - CD^{-1}B')^{-1}CD^{-1}.$$