

EM

与Kmeans不同, 高斯混合聚类采用概率模型表达聚类。先介绍EM:

考虑 $p(X, Z | \theta) = \prod_{n=1}^N p(x_n, z_n)$, 其中:

observed: $X = \{x_n\}_{n=1}^N$; latent: $Z = \{z_n\}_{n=1}^N$

Goal: Estimate θ via MLE (or MAP).

z :

$$\hat{\theta} = \arg\max_{\theta} \log p(X | \theta) = \arg\max_{\theta} \log \sum_z p(X, z | \theta) \quad (\text{discrete})$$

$$= \arg\max_{\theta} \log \int_z p(X, z | \theta) dz. \quad (\text{continuous})$$

$$\log p(X | \theta) = \log \sum_z p(X, z | \theta) = \log \sum_z q_n(z) \frac{p(X, z | \theta)}{q_n(z)}$$

$$\geq \sum_z q_n(z) \log \frac{p(X, z | \theta)}{q_n(z)} \quad (\text{since } \log(\cdot) \text{ is concave; Jensen})$$

$$= \sum_z q_n(z) \log p(X, z | \theta) - \sum_z q_n(z) \log q_n(z) \quad \text{constant! since irrelevant with } \theta$$

(if $q_n(z) = p(z | X, \theta)$, 则为等号)

故令 $q_n(z) = p(z | X, \theta)$ (独立): $\log p(X | \theta) = [\mathbb{E}[\log p(X, z | \theta)]] + \text{const}^*$

* lower-bound factor! EM就是最大化它!

EM流程: E: 以当前 θ^t 推后验 $P(z | X, \theta^t)$, 计算: 后验 先验

$$Q(\theta | \theta^t) = \mathbb{E}_{z | X, \theta^t} [\log p(z, X | \theta)] = \sum_z p(z | X, \theta^t) \log p(X, z | \theta)$$

M: $\theta^{t+1} = \arg\max_{\theta} \{Q(\theta | \theta^t) + [\log p(\theta)]\}$ 若加: MAP; 不加: MLE

latent variable, cluster中, $|z|=k$ $\sum_{k=1}^K \pi_k = 1$ k 类

GMM: 假设生成 N 个点: $x_n, n = 1, 2, \dots, N$

首先从 K 个高斯中选一个。此选择基于混合成分先验概率 π 的:

$z_n \sim \text{Multinomial}(z_n | \pi)$ (多项式分布)

然后生成点服从 $N(x_n | \mu_k, \Sigma_k)$, 其中 μ_k 是第 k 个高斯的均值,

Σ_k 是第 k 个高斯的协方差 (suppose $z_n = k$)

△: 说白了就是在 $[\pi_i]_{i=1}^K$ 的概率下, 抽一个 z_i



上述描述了在有 π_i, μ_i, Σ_i 下, 如何抽 n 个点出来。

那么学 GMM 呢? 考虑:

observed: $X = \{x_1, \dots, x_N\}$ latent: $Z = \{z_1, \dots, z_N\}$

$z_i \in \{1, \dots, K\} \Rightarrow N$ 点分 K 类

则 complete-data log-likelihood:

$$\log p(X, Z | \theta) = \sum_{i=1}^N \log \pi_{z_i} + \log N(x_i | \mu_{z_i}, \Sigma_{z_i})$$

则 E-step: posterior: $\gamma_{ik} = p(z_i = k | x_i, \theta^{\text{old}})$

$$\gamma_{ik} = \frac{p(x_i | z_i = k, \theta^{\text{old}}) p(z_i = k | \theta^{\text{old}})}{p(x_i | \theta^{\text{old}})}$$

Bayes

$$= \frac{\pi_k^{\text{old}} N(x_i | \mu_k^{\text{old}}, \Sigma_k^{\text{old}})}{\sum_{j=1}^K \pi_j^{\text{old}} N(x_i | \mu_j^{\text{old}}, \Sigma_j^{\text{old}})}$$

Bring in Gaussian

M-step: $Q(\theta, \theta^{\text{old}}) = \mathbb{E}_{Z|X, \theta^{\text{old}}} [\log p(X, Z | \theta)]$

$$= \sum_{i=1}^N \sum_{k=1}^K \gamma_{ik} [\log \pi_k + \log N(x_i | \mu_k, \Sigma_k)]$$

先验

后验

, constraint: $\sum_k \pi_k = 1$

$$\mathcal{L} = Q(\theta, \theta^{\text{old}}) + \lambda (1 - \sum_{k=1}^K \pi_k)$$

$$\frac{\partial \mathcal{L}}{\partial \pi_k} = \sum_{i=1}^N \frac{\gamma_{ik}}{\pi_k} - \lambda = 0 \Rightarrow \pi_k \propto \sum_{i=1}^N \gamma_{ik}$$

$$\text{又 } \sum_k \pi_k = 1, \text{ 故 } \sum_{k=1}^K \sum_{i=1}^N \gamma_{ik} = 1$$

$$\begin{cases} \pi_k^{\text{new}} = \frac{1}{N} \sum_{i=1}^N \gamma_{ik} \\ \mu_k^{\text{new}} = \frac{\sum_{i=1}^N \gamma_{ik} x_i}{\sum_{i=1}^N \gamma_{ik}} \\ \Sigma_k^{\text{new}} = \frac{\sum_{i=1}^N \gamma_{ik} (x_i - \mu_k^{\text{new}})(x_i - \mu_k^{\text{new}})^T}{\sum_{i=1}^N \gamma_{ik}} \end{cases}$$

$$\text{同理: } \frac{\partial Q}{\partial \mu_k} = \sum_{i=1}^N \gamma_{ik} \sum_k^{-1} (x_i - \mu_k) = 0$$

$$\frac{\partial Q}{\partial \Sigma_k^{-1}} = \frac{1}{2} \sum_{i=1}^N \gamma_{ik} [\Sigma_k - (x_i - \mu_k)(x_i - \mu_k)^T] = 0$$



附: PCA: 主成分分析 $\rightarrow \in \mathbb{R}^d$

Input: $D = \{x_1, x_2, \dots, x_m\}$, 低维空间维数 d'

Algorithm: 中心化: $x_i \leftarrow x_i - \frac{1}{m} \sum_{i=1}^m x_i$

计算协方差矩阵: XX^T

特征值分解

取最大 d' 个特征值对应的 eigenvector $w_1, w_2, \dots, w_{d'}$

\Rightarrow Output $W^* = (w_1, w_2, \dots, w_{d'}) \in \mathbb{R}^{d' \times d}$

$x_i \cdot W^{*T}$ 后便进入低维空间

PCA 是最常用的降维方法!

