

MAST30027: Modern Applied Statistics

Assignment 1 Solution 2021

1. Fit a binomial regression model to the O-rings data from the Challenger disaster, using a *probit* link. You must use R (but without using the `glm` function); I want you to work from first principles.

Your report should include the following:

- (a) (3 marks) Compute MLEs (maximum likelihood estimates) of the parameters in the model.
- (b) (7 marks) Compute 95% CIs for the estimates of the parameters. You should show how you derived the Fisher information.
- (c) (3 marks) Perform a likelihood ratio test for the significance of the temperature coefficient.
- (d) (3 marks) Compute an estimate of the probability of damage when the temperature equals 31 Fahrenheit (your estimate should come with a 95% CI, as all good estimates do).
- (e) (2 marks) Make a plot comparing the fitted probit model to the fitted logit model. To obtain the fitted logit model, you are allowed to use the `glm` function.

Solution

For a binomial regression with a probit link we have $y_i \sim \text{bin}(m_i, \Phi(\eta_i))$, where $\eta_i = \mathbf{x}_i^T \boldsymbol{\beta}$, so

$$\begin{aligned} l(\boldsymbol{\beta}) &= \sum_i [y_i \log \Phi(\eta_i) + (m_i - y_i) \log(1 - \Phi(\eta_i))] \\ \frac{\partial l(\boldsymbol{\beta})}{\partial \beta_0} &= \sum_i \left[\frac{y_i \phi(\eta_i)}{\Phi(\eta_i)} - \frac{(m_i - y_i) \phi(\eta_i)}{1 - \Phi(\eta_i)} \right] \\ \frac{\partial l(\boldsymbol{\beta})}{\partial \beta_1} &= \sum_i \left[\frac{y_i \phi(\eta_i) x_{i1}}{\Phi(\eta_i)} - \frac{(m_i - y_i) \phi(\eta_i) x_{i1}}{1 - \Phi(\eta_i)} \right] \\ \frac{\partial^2 l(\boldsymbol{\beta})}{\partial \beta_0^2} &= \sum_i \left[\frac{-y_i \phi(\eta_i)^2}{\Phi(\eta_i)^2} + \frac{-y_i \phi(\eta_i) \eta_i}{\Phi(\eta_i)} \right. \\ &\quad \left. - \frac{(m_i - y_i) \phi(\eta_i)^2}{(1 - \Phi(\eta_i))^2} - \frac{-(m_i - y_i) \phi(\eta_i) \eta_i}{1 - \Phi(\eta_i)} \right] \\ \frac{\partial^2 l(\boldsymbol{\beta})}{\partial \beta_1^2} &= \sum_i x_{i1}^2 \left[\frac{-y_i \phi(\eta_i)^2}{\Phi(\eta_i)^2} + \frac{-y_i \phi(\eta_i) \eta_i}{\Phi(\eta_i)} \right. \\ &\quad \left. - \frac{(m_i - y_i) \phi(\eta_i)^2}{(1 - \Phi(\eta_i))^2} - \frac{-(m_i - y_i) \phi(\eta_i) \eta_i}{1 - \Phi(\eta_i)} \right] \\ \frac{\partial^2 l(\boldsymbol{\beta})}{\partial \beta_0 \partial \beta_1} &= \sum_i x_{i1} \left[\frac{-y_i \phi(\eta_i)^2}{\Phi(\eta_i)^2} + \frac{-y_i \phi(\eta_i) \eta_i}{\Phi(\eta_i)} \right. \\ &\quad \left. - \frac{(m_i - y_i) \phi(\eta_i)^2}{(1 - \Phi(\eta_i))^2} - \frac{-(m_i - y_i) \phi(\eta_i) \eta_i}{1 - \Phi(\eta_i)} \right] \\ -\mathbb{E} \frac{\partial^2 l(\boldsymbol{\beta})}{\partial \boldsymbol{\beta} \partial \boldsymbol{\beta}^T} &= \sum_i m_i \phi(\eta_i)^2 \left[\frac{1}{\Phi(\eta_i)} + \frac{1}{1 - \Phi(\eta_i)} \right] \begin{bmatrix} 1 & x_{i1} \\ x_{i1} & x_{i1}^2 \end{bmatrix} \end{aligned}$$

- (a) (3 marks) Compute MLEs (maximum likelihood estimates) of the parameters in the model.

```
> library(faraway)
> data(orings)
> logL <- function(beta, orings) {
+   y <- orings$damage
+   X <- cbind(1, orings$temp)
+   zeta <- X %>% beta
+   p <- pnorm(zeta)
+   return(sum(y*log(p) + (6 - y)*log(1 - p)))
+ }
> (betahat <- optim(c(10, -.1), logL, orings=orings, control=list(fnscale=-1))$par)
[1] 5.5917242 -0.1058008
```

- (b) (7 marks) Compute 95% CIs for the estimates of the parameters. You should show how you derived the Fisher information.

```
> X <- cbind(1, orings$temp)
> zetahat <- X %>% betahat
> a <- dnorm(zetahat)^2*(1/pnorm(zetahat) + 1/(1-pnorm(zetahat)))
> I11 <- sum(6*X[,1]^2*a)
> I12 <- sum(6*X[,1]*X[,2]*a)
> I22 <- sum(6*X[,2]^2*a)
> Iinv <- solve(matrix(c(I11, I12, I12, I22), 2, 2))
> c(betahat[1] - 1.96*sqrt(Iinv[1,1]), betahat[1] + 1.96*sqrt(Iinv[1,1]))
[1] 2.239700 8.943748
> c(betahat[2] - 1.96*sqrt(Iinv[2,2]), betahat[2] + 1.96*sqrt(Iinv[2,2]))
[1] -0.15784765 -0.05375385
```

Comparing with glm output, we see that the estimates and standard errors agree with ours to four significant figures.

```
> probitmod <- glm(cbind(damage, 6-damage) ~ temp, family=binomial(link=probit), orings)
> summary(probitmod)
```

Call:

```
glm(formula = cbind(damage, 6 - damage) ~ temp, family = binomial(link = probit),
    data = orings)
```

Deviance Residuals:

	Min	1Q	Median	3Q	Max
	-1.0134	-0.7761	-0.4467	-0.1581	1.9983

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	5.59145	1.71055	3.269	0.00108 **
temp	-0.10580	0.02656	-3.984	6.79e-05 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

Null deviance: 38.898 on 22 degrees of freedom
 Residual deviance: 18.131 on 21 degrees of freedom
 AIC: 34.893

Number of Fisher Scoring iterations: 6

- (c) (3 marks) Perform a likelihood ratio test for the significance of the temperature coefficient.

First we calculate the deviance for the model including temperature.

```
> y <- orings$damage
> n <- rep(6, length(y))
> ylogxy <- function(x, y) ifelse(y == 0, 0, y*log(x/y))
> phat <- pnorm(zetahat)
> (D <- -2*sum(ylogxy(n*phat, y) + ylogxy(n*(1-phat), n - y)))

[1] 18.13058

> (df <- length(y) - length(betahat))

[1] 21
```

Next we fit the null model and use a likelihood ratio test.

```
> (phatN <- sum(y)/sum(n))

[1] 0.07971014

> (DN <- -2*sum(ylogxy(n*phatN, y) + ylogxy(n*(1-phatN), n - y)))

[1] 38.89766

> (dfN <- length(y) - 1)

[1] 22

> pchisq(DN - D, dfN - df, lower=FALSE) # p-value

[1] 5.186684e-06
```

We have very strong evidence that $\beta_1 \neq 0$.

Note that our deviance calculations agree with the output from `glm`.

- (d) (3 marks) Compute an estimate of the probability of damage when the temperature equals 31 Fahrenheit (your estimate should come with a 95% CI, as all good estimates do).

```
> si2 <- matrix(c(1, 31), 1, 2) %*% Iinv %*% matrix(c(1, 31), 2, 1)
> (p31 <- pnorm(betahat[1] + betahat[2]*31))

[1] 0.9896084

> pnorm(betahat[1] + betahat[2]*31 - 1.96*sqrt(si2))[1]

[1] 0.7108118

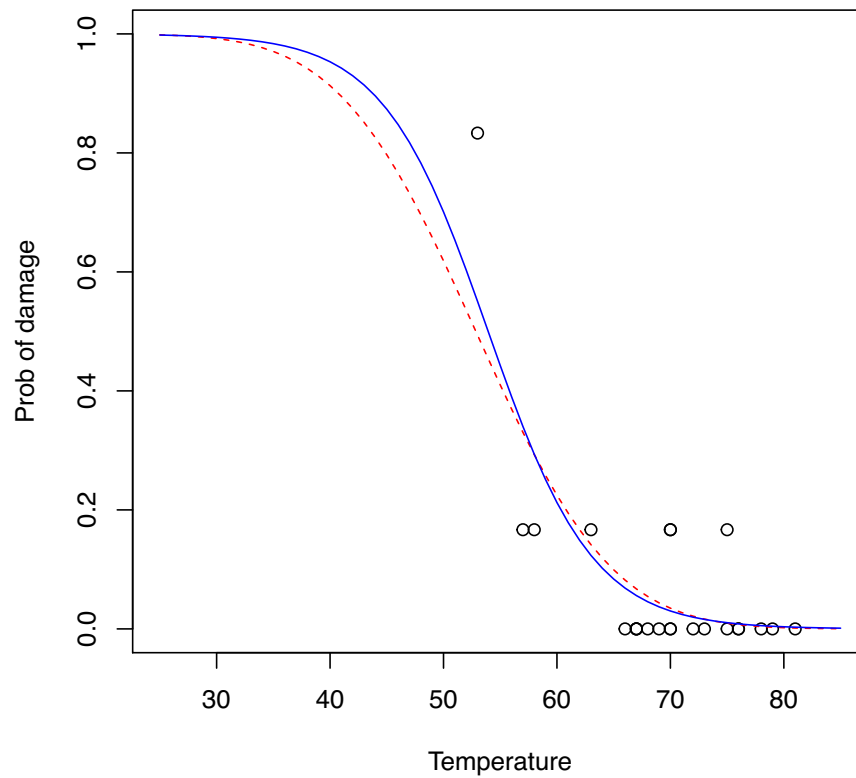
> pnorm(betahat[1] + betahat[2]*31 + 1.96*sqrt(si2))[1]

[1] 0.9999763
```

- (e) (2 marks) Make a plot comparing the fitted probit model to the fitted logit model. To obtain the fitted logit model, you are allowed to use the `glm` function.

They are very close, but the probit model puts a little more weight in the tails.

```
> plot(damage/6 ~ temp, orings, xlim=c(25,85), ylim=c(0,1),
+      xlab="Temperature", ylab="Prob of damage")
> x <- seq(25,85,1)
> lines(x, pnorm(betahat[1] + betahat[2]*x), col="red", lty=2)
> betalgit <- glm(cbind(damage,6-damage) ~ temp, family=binomial, orings)$coefficients
> lines(x, ilogit(betalgit[1] + betalgit[2]*x), col="blue")
```



2. The data frame 'pima_subset' contains a subset of the `pima` data set. For details of the `pima` data set, please see the practical problem 2 for the week 2. You can obtain 'pima_subset' using the commands:

```
> library(faraway)
> missing <- with(pima, missing <- glucose==0 | diastolic==0 | triceps==0 | bmi == 0)
> pima_subset = pima[!missing, c(6,9)]
> str(pima_subset)
'data.frame': 532 obs. of 2 variables:
 $ bmi : num 33.6 26.6 28.1 43.1 31 30.5 30.1 25.8 45.8 43.3 ...
 $ test: int 1 0 0 1 1 1 1 1 1 0 ...
```

Using the 'pima_subset' data set, we will fit a binomial regression with a logit link with `test` as a response and `bmi` as a predictor to see the relationship between the odds of a patient showing signs of diabetes and his/her bmi. The odds o and probability p are related by

$$o = \frac{p}{1-p} \quad p = \frac{o}{1+o}.$$

- (3 marks) Please estimate the amount of increase in the log(odds) when the bmi increases by 7.
- (3 marks) Compute a 95% CI for the estimate.

You are allowed to use the `glm` function.

Solution

- (a) (3 marks) Please estimate the amount of increase in the log(odds) when the bmi increases by 7.

Let o_x , η_x , o_{x+7} , η_{x+7} be the odds and linear response for a woman with bmi at x and $x + 7$ respectively. Then, for binomial regression with logit link,

$$\begin{aligned}\log(o_{x+7}) - \log(o_x) &= \eta_{x+7} - \eta_x \\ &= 7\beta_{bmi}\end{aligned}$$

We fit a binomial regression.

```
> library(faraway)
> missing <- with(pima, missing <- glucose==0 | diastolic==0 | triceps==0 | bmi == 0)
> pima_subset = pima[!missing, c(6,9)]
> str(pima_subset)

'data.frame':      532 obs. of  2 variables:
 $ bmi : num  33.6 26.6 28.1 43.1 31 30.5 30.1 25.8 45.8 43.3 ...
 $ test: int   1 0 0 1 1 1 1 1 0 ...

> model <- glm(cbind(test, 1-test)~., family=binomial, data=pima_subset)
> summary(model)

Call:
glm(formula = cbind(test, 1 - test) ~ ., family = binomial, data = pima_subset)

Deviance Residuals:
    Min       1Q   Median       3Q      Max
-1.9227  -0.8920  -0.6568   1.2559   1.9560

Coefficients:
            Estimate Std. Error z value Pr(>|z|)
(Intercept) -4.03681    0.52783  -7.648 2.04e-14 ***
bmi          0.09972    0.01528   6.524 6.84e-11 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for binomial family taken to be 1)

    Null deviance: 676.79  on 531  degrees of freedom
Residual deviance: 627.46  on 530  degrees of freedom
AIC: 631.46

Number of Fisher Scoring iterations: 4
A point estimate for  $7\beta_{bmi}$  is
```

$$7 \times 0.09972 = 0.69804.$$

- (b) (3 marks) Compute a 95% CI for the estimate.

The standard error of the estimate for $7\beta_{bmi}$ is $7 \times$ standard error of the estimate for β_{bmi} . 95% CI for the estimate is

$$7(0.09972 \pm 1.959964 \times 0.01528) = (0.48840, 0.90768)$$

3. The gamma distribution with shape $\nu > 0$ and rate $\lambda > 0$ has p.d.f.

$$f(x; \nu, \lambda) = \frac{\lambda^\nu}{\Gamma(\nu)} x^{\nu-1} e^{-\lambda x}$$

for $x > 0$.

- (a) (5 marks) Show that the gamma distribution is an exponential family.
 (b) (5 marks) Obtain the canonical link and the variance function.

Solution

- (a) (5 marks) Show that the gamma distribution is an exponential family.
 The gamma distribution with shape $\nu > 0$ and rate $\lambda > 0$ has log density

$$\begin{aligned}\log f(x; \nu, \lambda) &= (\nu - 1) \log(x) - \lambda x + \nu \log(\lambda) - \log(\Gamma(\nu)) \\ &= \frac{x(-\lambda/\nu) + \log(\lambda/\nu)}{1/\nu} - \nu \log(1/\nu) + (\nu - 1) \log(x) - \log(\Gamma(\nu))\end{aligned}$$

Put $\theta = -\lambda/\nu$ and $\phi = 1/\nu$ then we have

$$\log f(x; \nu, \lambda) = \frac{x\theta - \log(-1/\theta)}{\phi} - \frac{\log(\phi)}{\phi} + \left(\frac{1}{\phi} - 1\right) \log(x) - \log(\Gamma(1/\phi))$$

This is in the form of an exponential family, with

$$\begin{aligned}b(\theta) &= \log(-1/\theta) \\ a(\phi) &= \phi \\ c(x, \phi) &= \frac{-\log(\phi) + (1 - \phi) \log(x) - \phi \log(\Gamma(1/\phi))}{\phi}\end{aligned}$$

Note that with this parameterisation we have $\theta < 0$ and $\phi > 0$.

- (b) (5 marks) Obtain the canonical link and the variance function.

For the canonical link g we have $g(\mu) = \theta$. Here $\mu = \nu/\lambda = -1/\theta$, so $g(x) = -1/x$. (Note that in practice people tend to use the inverse link $x \mapsto 1/x$ rather than $x \mapsto -1/x$, because it is convenient to keep things positive.) The variance is $\nu/\lambda^2 = \phi\mu^2 = a(\phi)v(\mu)$. That is, the variance function is $v(\mu) = \mu^2$.