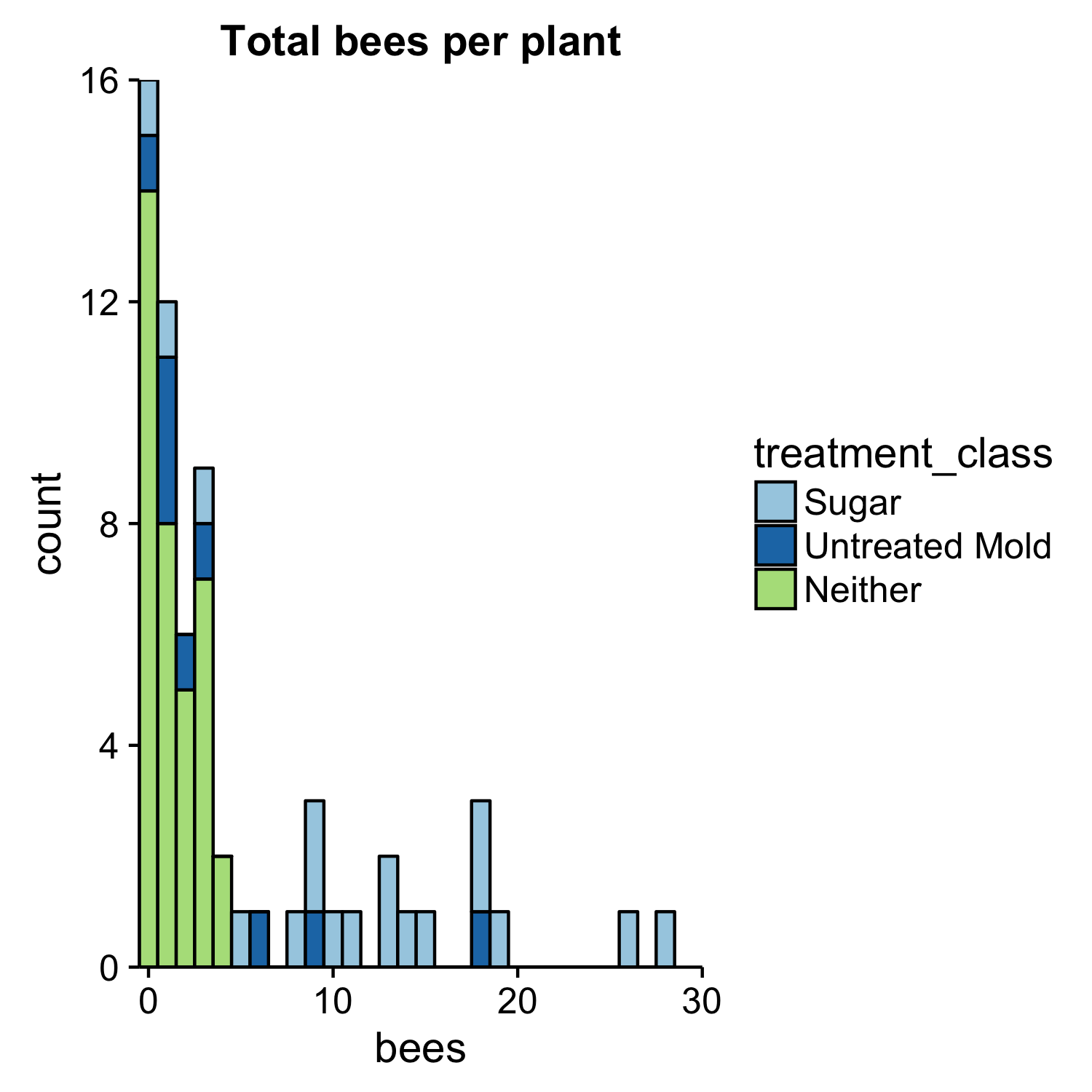
Appendix for 'Bees Without Flowers'

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# Preliminaries

set.seed(1)  
library(lme4) # For model fitting  
library(tidyverse) # For data processing  
library(mvtnorm) # for multivariate Gaussians  
  
# hours\_noon represents the amount of time since noon (or until noon, if  
# the value is negative)  
d = read.csv("Meiners\_BeeHoneydew\_data.csv") %>%   
 mutate(hours\_noon = min\_day / 60 - 12) %>%   
 mutate(treatment\_class = 1 + Sugar + 2 \* Mold -2 \* (Mold \* Insecticide)) %>%   
 mutate(treatment\_class = factor(treatment\_class,   
 labels = c("Neither", "Sugar", "Untreated Mold"))) %>%   
 mutate(treatment\_class = forcats::fct\_relevel(treatment\_class,  
 c("Sugar", "Untreated Mold")))  
  
  
  
d %>%   
 group\_by(Plant\_Code, treatment\_class) %>%   
 summarize(bees = sum(Bee\_Count)) %>%   
 ggplot(aes(x = bees, fill = treatment\_class)) +   
 geom\_histogram(binwidth = 1, color = "black") +   
 cowplot::theme\_cowplot() +   
 scale\_fill\_brewer(type = "qual", palette = 3) +   
 coord\_cartesian(expand = FALSE, xlim = c(-.5, 30)) +   
 ggtitle("Total bees per plant")



d %>%   
 group\_by(Plant\_Code, treatment\_class) %>%   
 summarize(bees = sum(Bee\_Count)) %>%   
 group\_by(treatment\_class) %>%   
 summarize(mean(bees))

## # A tibble: 3 × 2  
## treatment\_class `mean(bees)`  
## <fctr> <dbl>  
## 1 Sugar 12.222222  
## 2 Untreated Mold 4.555556  
## 3 Neither 1.305556

# Core model formulas

We will be focusing on models that include fixed effect for the experimental manipulations and for site, as well as a continuous time-of-day variable to capture variation in bee activity associated with diurnal patterns. We modeled variation among days (e.g. due to differences in recent weather events) and among individual plants using random effects. Because the number of sites (3) was too small to estimate site-to-site variance, we treated Site as a fixed effect.

While additional variables were recorded during sampling, their primary purpose was to keep the sampling effort focused on a narrow range of environmental conditions (because environmental effects such as humidity were not related to our primary hypotheses). We thus did not expect these variables to vary enough in our samples to substantially affect the results, and did not include them in most of our analyses. In the final section of this Appendix, we show that the exclusion of these variables (and of a continuous measure of seasonality) do not affect the statistical significance or point estimates associated with any treatment effects, and that they do not significantly improve model fit in terms of or AIC.

raw\_formula = "Bee\_Count ~ Mold \* Insecticide +   
 Sugar \* Paint +   
 hours\_noon +   
 Site +  
 (1|Plant\_Code) +   
 (1|julDate)"  
  
formula = as.formula(raw\_formula)

We also fit a model that did not include sugar as a predictor variable, to assess whether its inclusion substantially improves our ability to predict bee density.

# Drop "Sugar" and the asterisk from the formula  
no\_sugar\_formula = as.formula(  
 gsub("Sugar \\\* ", "", raw\_formula)  
)  
  
print(no\_sugar\_formula)

## Bee\_Count ~ Mold \* Insecticide + Paint + hours\_noon + Site +   
## (1 | Plant\_Code) + (1 | julDate)

# Model fitting

We modeled the bee counts with negative binomial and Poisson mixed models with the default log link.

# Some versions of the model only reach the maximum-likelihood  
# estimate without warnings when we use this optimizer  
control = glmerControl(optimizer = "bobyqa")  
  
Honeydew <- glmer.nb(  
 formula,  
 data=d,  
 control = control  
)  
  
  
Honeydew\_poisson <- glmer(  
 formula,  
 data=d,  
 family = poisson,  
 control = control  
)  
  
Honeydew\_no\_sugar <- glmer.nb(  
 no\_sugar\_formula,  
 data=d,  
 control = control  
)

# Model comparison

Which elements of the models fit above are essential? See what happens to model performance (AIC) when various degrees of freedom are removed from the full Honeydew negative binomial model.

# Drop predictors from the model & reformat the output for  
# subsequent work (e.g. removing headings, renaming columns).  
# When `drop` says the number of degrees of freedom is NA, it actually   
# means zero, so replace the NAs.  
# If the model in `x` is already simplified, then report a larger  
# reduction in degrees-of-freedom.  
make\_dropped\_df = function(x, distribution, n\_fewer\_df = 0){  
 drop1(x) %>%   
 structure(heading = NULL) %>%   
 rownames\_to\_column(var = "dropped") %>%   
 cbind(distribution = distribution) %>%   
 mutate(Df = ifelse(is.na(Df), 0, Df)) %>%   
 mutate(Df = Df + n\_fewer\_df) %>%   
 rename\_(`df reduction` = "Df")  
}  
  
# Use the above function on both of the full models, then manually add  
# a row for the no\_sugar model.  
# Finally, eliminate "<" and ">" to prevent formatting errors  
initial\_dropped\_df = rbind(make\_dropped\_df(Honeydew, "Negative Binomial", 0),   
 make\_dropped\_df(Honeydew\_poisson, "Poisson", 1)) %>%   
 rbind(data\_frame(dropped = "Sugar", `df reduction` = 2,   
 AIC = AIC(Honeydew\_no\_sugar),   
 distribution = "Negative Binomial")) %>%   
 mutate(dropped = gsub("[\\<\\>]", "", dropped))

Omitting site or either of the interaction terms (lines 1, 2, 4 and 5) produces a relatively small change in AIC, compared with the full model (line 3). However, none of the models without overdispersion (i.e. the Poisson-distributed models) had any appreciable AIC weight, nor did the model that removed all sugar effects (line 10).

# Sort, calculate DeltaAIC & AIC weights, format for printing with   
# reasonable precision using knitr's `kable` function for tables.  
initial\_dropped\_df %>%   
 arrange(AIC) %>%   
 mutate(`$\\Delta$AIC` = AIC - AIC[1]) %>%   
 select(-AIC) %>%   
 mutate(`AIC weight (%)` = 100 \* exp(-`$\\Delta$AIC` / 2) /   
 sum(exp(-`$\\Delta$AIC` / 2))) %>%   
 cbind(` ` = 1:nrow(.), .) %>%   
 knitr::kable(digits = c(rep(1, 4), 2, 1), align = c("llclrr"))

|  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- |
|  | dropped | df reduction | distribution | AIC | AIC weight (%) |
| 1 | Site | 2 | Negative Binomial | 0.00 | 59.4 |
| 2 | Sugar:Paint | 1 | Negative Binomial | 1.58 | 26.9 |
| 3 | none | 0 | Negative Binomial | 3.57 | 10.0 |
| 4 | Mold:Insecticide | 1 | Negative Binomial | 5.68 | 3.5 |
| 5 | hours\_noon | 1 | Negative Binomial | 11.40 | 0.2 |
| 6 | Site | 3 | Poisson | 28.16 | 0.0 |
| 7 | Sugar:Paint | 2 | Poisson | 29.97 | 0.0 |
| 8 | none | 1 | Poisson | 31.94 | 0.0 |
| 9 | Mold:Insecticide | 2 | Poisson | 34.06 | 0.0 |
| 10 | Sugar | 2 | Negative Binomial | 38.69 | 0.0 |
| 11 | hours\_noon | 2 | Poisson | 71.34 | 0.0 |

tests show the same result: omitting sugar effects or overdispersion significantly reduces model performance (P < .000001).

anova(Honeydew, Honeydew\_no\_sugar)

## Data: d  
## Models:  
## Honeydew\_no\_sugar: Bee\_Count ~ Mold \* Insecticide + Paint + hours\_noon + Site +   
## Honeydew\_no\_sugar: (1 | Plant\_Code) + (1 | julDate)  
## Honeydew: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + hours\_noon +   
## Honeydew: Site + (1 | Plant\_Code) + (1 | julDate)  
## Df AIC BIC logLik deviance Chisq Chi Df  
## Honeydew\_no\_sugar 11 820.74 864.02 -399.37 798.74   
## Honeydew 13 785.62 836.77 -379.81 759.62 39.12 2  
## Pr(>Chisq)   
## Honeydew\_no\_sugar   
## Honeydew 3.201e-09 \*\*\*  
## ---  
## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

anova(Honeydew, Honeydew\_poisson)

## Data: d  
## Models:  
## Honeydew\_poisson: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + hours\_noon +   
## Honeydew\_poisson: Site + (1 | Plant\_Code) + (1 | julDate)  
## Honeydew: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + hours\_noon +   
## Honeydew: Site + (1 | Plant\_Code) + (1 | julDate)  
## Df AIC BIC logLik deviance Chisq Chi Df  
## Honeydew\_poisson 12 813.99 861.21 -395.00 789.99   
## Honeydew 13 785.62 836.77 -379.81 759.62 30.374 1  
## Pr(>Chisq)   
## Honeydew\_poisson   
## Honeydew 3.562e-08 \*\*\*  
## ---  
## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

# Description of the full Honeydew model

summary(Honeydew, correlation = FALSE)

## Generalized linear mixed model fit by maximum likelihood (Laplace  
## Approximation) [glmerMod]  
## Family: Negative Binomial(1.8463) ( log )  
## Formula: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + hours\_noon +   
## Site + (1 | Plant\_Code) + (1 | julDate)  
## Data: d  
## Control: control  
##   
## AIC BIC logLik deviance df.resid   
## 785.6 836.8 -379.8 759.6 365   
##   
## Scaled residuals:   
## Min 1Q Median 3Q Max   
## -1.0060 -0.5083 -0.3610 0.1850 4.6619   
##   
## Random effects:  
## Groups Name Variance Std.Dev.  
## Plant\_Code (Intercept) 0.3982 0.6311   
## julDate (Intercept) 0.1712 0.4138   
## Number of obs: 378, groups: Plant\_Code, 63; julDate, 9  
##   
## Fixed effects:  
## Estimate Std. Error z value Pr(>|z|)   
## (Intercept) -2.09367 0.50337 -4.159 3.19e-05 \*\*\*  
## Mold 1.15201 0.49708 2.318 0.02047 \*   
## Insecticide 0.45650 0.52686 0.866 0.38624   
## Sugar 2.41458 0.47326 5.102 3.36e-07 \*\*\*  
## Paint -0.46287 0.60338 -0.767 0.44300   
## hours\_noon 0.16880 0.05426 3.111 0.00186 \*\*   
## SiteB 0.28098 0.45612 0.616 0.53788   
## SiteC 0.04974 0.46442 0.107 0.91471   
## Mold:Insecticide -1.48408 0.71781 -2.068 0.03869 \*   
## Sugar:Paint 0.08231 0.70727 0.116 0.90735   
## ---  
## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

anova(Honeydew)

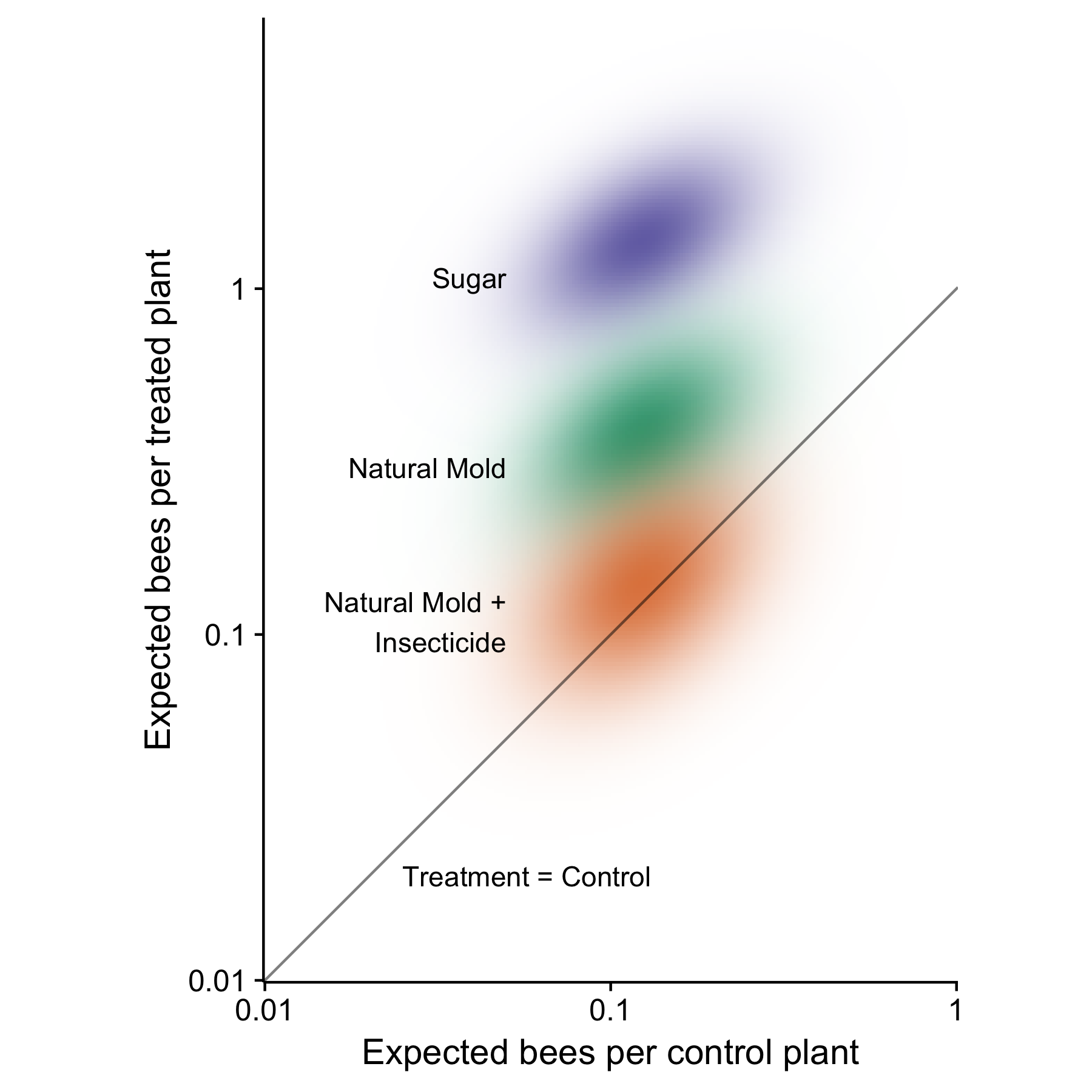
## Analysis of Variance Table  
## Df Sum Sq Mean Sq F value  
## Mold 1 5.002 5.002 5.0019  
## Insecticide 1 12.643 12.643 12.6429  
## Sugar 1 39.648 39.648 39.6479  
## Paint 1 2.841 2.841 2.8408  
## hours\_noon 1 10.158 10.158 10.1582  
## Site 2 0.414 0.207 0.2069  
## Mold:Insecticide 1 4.880 4.880 4.8797  
## Sugar:Paint 1 0.013 0.013 0.0133

# Monte Carlo comparison of treatment effects

# Indicator for, "was experimental manipulation i applied to treatment j?""  
Mold = c(1, 1, 0, 0, 0, 0, 0)  
Insecticide = c(0, 1, 0, 1, 0, 0, 0)  
Sugar = c(0, 0, 0, 0, 0, 1, 1)  
Paint = c(0, 0, 0, 0, 1, 0, 1)  
  
treat\_names = c("Natural Mold", "Natural Mold + Insecticide", "Control",   
 "Insecticide", "Black Paint", "Sugar", "Sugar + Black Paint")  
  
# Ask the model about expected visitation rates under the  
# following conditions:  
# \* Treatments as specified above  
# \* Time of day is noon  
# \* Site A (i.e. SiteB's effect and SiteC's effect are 0)  
# \* "Typical" plant and "typical" date (random effects set to 0)  
newdata = cbind(  
 `(Intercept)` = 1,  
 Mold = Mold,  
 Insecticide = Insecticide,  
 Sugar = Sugar,  
 Paint = Paint,  
 hours\_noon = 0,  
 SiteB = 0,  
 SiteC = 0,  
 `Mold:Insecticide` = Mold \* Insecticide,  
 `Sugar:Paint` = Sugar \* Paint  
)  
row.names(newdata) = treat\_names  
newdata

## (Intercept) Mold Insecticide Sugar Paint  
## Natural Mold 1 1 0 0 0  
## Natural Mold + Insecticide 1 1 1 0 0  
## Control 1 0 0 0 0  
## Insecticide 1 0 1 0 0  
## Black Paint 1 0 0 0 1  
## Sugar 1 0 0 1 0  
## Sugar + Black Paint 1 0 0 1 1  
## hours\_noon SiteB SiteC Mold:Insecticide  
## Natural Mold 0 0 0 0  
## Natural Mold + Insecticide 0 0 0 1  
## Control 0 0 0 0  
## Insecticide 0 0 0 0  
## Black Paint 0 0 0 0  
## Sugar 0 0 0 0  
## Sugar + Black Paint 0 0 0 0  
## Sugar:Paint  
## Natural Mold 0  
## Natural Mold + Insecticide 0  
## Control 0  
## Insecticide 0  
## Black Paint 0  
## Sugar 0  
## Sugar + Black Paint 1

# mean and variance from lme4's Laplace approximation  
parameter\_mu = fixef(Honeydew)  
parameter\_sigma = as.matrix(vcov(Honeydew))  
  
# Generate Monte Carlo samples from lme4's approximate likelihood surface  
posterior\_samples = rmvnorm(1E6, parameter\_mu, parameter\_sigma) %\*%  
 t(newdata)  
  
mu = colMeans(posterior\_samples)  
  
# Density of bivariate normal between Control and a named treatment  
bivariate\_normal\_control\_density = function(x, name){  
 names = c("Control", name)  
 dmvnorm(x,   
 mu[names],   
 cov(posterior\_samples)[names, names])  
}  
  
label\_df = data.frame(  
 x = c(rep(log(.05), 3)),  
 y = c(mu[c("Sugar", "Natural Mold", "Natural Mold + Insecticide")] - 0.25),  
 label = c("Sugar", "Natural Mold", "Natural Mold +\nInsecticide")  
)  
line\_df = data.frame(x = log(.025), y = log(.02),   
 label = "Treatment = Control")  
  
# for each set of x and y values, calculate bivariate densities,  
# then tidy up the results for ggplot (with an optional `theme` for  
# improved visual display)  
plot\_data = expand.grid(x = seq(log(.01), log(1), length = 250),   
 y = seq(log(.01), log(6), length = 250)) %>%   
 mutate(Sugar = bivariate\_normal\_control\_density(., "Sugar"),  
 `Natural Mold` = bivariate\_normal\_control\_density(., "Natural Mold"),  
 `Natural Mold + Insecticide` =   
 bivariate\_normal\_control\_density(., "Natural Mold + Insecticide")) %>%   
 gather(key = treatment, value = likelihood, Sugar,   
 `Natural Mold`, `Natural Mold + Insecticide`) %>%   
 mutate(scaled\_likelihood = likelihood / max(likelihood))  
  
plot\_data %>%   
 ggplot(aes(x = x, y = y, alpha = scaled\_likelihood,   
 fill = treatment)) +   
 geom\_tile() +  
 scale\_alpha\_continuous(range = c(0, 1), guide = FALSE) +   
 geom\_abline(intercept = 0, slope = 1, color = alpha("black", .5)) +  
 cowplot::theme\_cowplot() +  
 scale\_fill\_brewer(palette = "Dark2", guide = FALSE) +   
 xlab("Expected bees per control plant") +  
 ylab("Expected bees per treated plant") +   
 coord\_equal() +  
 scale\_x\_continuous(breaks = log(10^seq(-10, 10)), labels = 10^seq(-10, 10),  
 expand = c(0, 0)) +  
 scale\_y\_continuous(breaks = log(10^seq(-10, 10)), labels = 10^seq(-10, 10),  
 expand = c(0, 0)) +  
 geom\_text(data = label\_df, aes(x = x, y = y, label = label),  
 inherit.aes = FALSE, hjust = "right") +   
 geom\_text(data = line\_df, aes(x = x, y = y, label = label),   
 inherit.aes = FALSE, hjust = "left")



# Post-hoc comparisons among treatment pairs

names = names(sort(colMeans(posterior\_samples), decreasing = TRUE))  
grid = combn(names, 2) %>% t() %>% as.data.frame(stringsAsFactors = FALSE)  
grid$lower = NA  
grid$mean = NA  
grid$upper = NA  
grid$P = NA  
for (i in 1:nrow(grid)) {  
 # One-sided P-values  
 p = mean(posterior\_samples[ , grid[[1]][i]] > posterior\_samples[ , grid[[2]][i]])  
 ratios = exp(posterior\_samples[ , grid[[1]][i]] - posterior\_samples[ , grid[[2]][i]])  
 # Two-sided P-values based on Monte Carlo samples  
 grid$P[i] = 1 - 2 \* abs(0.5 - p)  
 grid$lower[i] = quantile(ratios, .025)  
 grid$mean[i] = mean(ratios)  
 grid$upper[i] = quantile(ratios, .975)  
  
}  
  
my\_format = function(x, d){format(x, digits = d, trim = TRUE)}  
table = grid %>%   
 mutate(`False Discovery Rate` = p.adjust(P, method = "fdr")) %>%   
 cbind(ratio = paste0(my\_format(.$mean, 3),   
 " (",   
 my\_format(.$lower, 2),   
 "-",   
 my\_format(.$upper, 3),   
 ")")) %>%   
 select(V1, V2, ratio, P, `False Discovery Rate`)

Significance and False Discovery Rates for selected post-hoc comparisons between treatments. The false discovery rate is a way to correct for multiple comparisons without sacrificing too much statistical power. See ?p.adjust and references therein. The treatment with the larger expected visitation rate is listed in the left column of each row.

table\_subset = table %>%   
 filter(  
 (V1 == "Sugar" & V2 == "Control") |  
 (V1 == "Sugar" & V2 == "Natural Mold") |  
 (V1 == "Natural Mold" & V2 == "Control") |  
 (V1 == "Natural Mold" & V2 == "Natural Mold + Insecticide") |  
 (V1 == "Natural Mold + Insecticide" & V2 == "Control") |   
 (V1 == "Insecticide" & V2 == "Control") |   
 (V1 == "Sugar" & V2 == "Sugar + Black Paint")  
 )  
table\_subset %>%   
 knitr::kable(digits = 2)

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| V1 | V2 | ratio | P | False Discovery Rate |
| Sugar | Sugar + Black Paint | 1.57 (0.71-3.02) | 0.30 | 0.40 |
| Sugar | Natural Mold | 3.83 (1.61-7.76) | 0.00 | 0.00 |
| Sugar | Control | 12.51 (4.43-28.26) | 0.00 | 0.00 |
| Natural Mold | Natural Mold + Insecticide | 3.15 (1.07-7.26) | 0.04 | 0.06 |
| Natural Mold | Control | 3.58 (1.20-8.39) | 0.02 | 0.04 |
| Insecticide | Control | 1.81 (0.56-4.44) | 0.39 | 0.45 |
| Natural Mold + Insecticide | Control | 1.32 (0.39-3.31) | 0.82 | 0.82 |

Significance and False Discovery Rates for all pairwise comparisons among experimental treatments.

table %>%   
 knitr::kable(digits = 2)

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| V1 | V2 | ratio | P | False Discovery Rate |
| Sugar | Sugar + Black Paint | 1.57 (0.71-3.02) | 0.30 | 0.40 |
| Sugar | Natural Mold | 3.83 (1.61-7.76) | 0.00 | 0.00 |
| Sugar | Insecticide | 7.79 (3.01-16.64) | 0.00 | 0.00 |
| Sugar | Natural Mold + Insecticide | 10.98 (3.99-24.40) | 0.00 | 0.00 |
| Sugar | Control | 12.51 (4.43-28.26) | 0.00 | 0.00 |
| Sugar | Black Paint | 20.44 (6.32-50.04) | 0.00 | 0.00 |
| Sugar + Black Paint | Natural Mold | 2.62 (1.09-5.33) | 0.03 | 0.05 |
| Sugar + Black Paint | Insecticide | 5.34 (2.04-11.49) | 0.00 | 0.00 |
| Sugar + Black Paint | Natural Mold + Insecticide | 7.52 (2.71-16.81) | 0.00 | 0.00 |
| Sugar + Black Paint | Control | 8.56 (3.01-19.50) | 0.00 | 0.00 |
| Sugar + Black Paint | Black Paint | 13.99 (4.31-34.39) | 0.00 | 0.00 |
| Natural Mold | Insecticide | 2.23 (0.81-4.97) | 0.13 | 0.19 |
| Natural Mold | Natural Mold + Insecticide | 3.15 (1.07-7.26) | 0.04 | 0.06 |
| Natural Mold | Control | 3.58 (1.20-8.39) | 0.02 | 0.04 |
| Natural Mold | Black Paint | 5.85 (1.72-14.76) | 0.00 | 0.01 |
| Insecticide | Natural Mold + Insecticide | 1.59 (0.51-3.84) | 0.52 | 0.55 |
| Insecticide | Control | 1.81 (0.56-4.44) | 0.39 | 0.45 |
| Insecticide | Black Paint | 2.96 (0.81-7.75) | 0.11 | 0.17 |
| Natural Mold + Insecticide | Control | 1.32 (0.39-3.31) | 0.82 | 0.82 |
| Natural Mold + Insecticide | Black Paint | 2.15 (0.56-5.79) | 0.32 | 0.40 |
| Control | Black Paint | 1.91 (0.49-5.18) | 0.44 | 0.49 |

# Comparing to a model with environmental predictors & continuous dates

We could have obtained essentially the same results with a much larger model that included a fixed effect for date and environmental conditions (i.e., there would still be a significant sugar effect with point estimate 2.4, a significant mold effect with point estimate 1.2, and a significant mold/insecticide interaction with point estimate -1.5).

However, the eight degrees of freedom associated with environmental conditions and the fixed effect for date do not improve the model fit enough to justify the additional complexity ( > 0.25, higher AIC), so we disregard them outide of this section.

# rescaling variables that have large values using `scale` improves numerical   
# accuracy, but will not affect AIC.  
Honeydew\_env = glmer.nb(Bee\_Count ~ Mold \* Insecticide +   
 Sugar \* Paint +   
 Site +   
 hours\_noon +   
 scale(Temp\_F) +   
 scale(Wind\_mph) +   
 Conditions +   
 scale(Barometric) +  
 scale(Humidity) +   
 scale(julDate) +  
 (1|Plant\_Code) +   
 (1|julDate),   
 data = d,   
 control = control)  
print(summary(Honeydew\_env), correlation = FALSE)

## Generalized linear mixed model fit by maximum likelihood (Laplace  
## Approximation) [glmerMod]  
## Family: Negative Binomial(2.1074) ( log )  
## Formula:   
## Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + Site + hours\_noon +   
## scale(Temp\_F) + scale(Wind\_mph) + Conditions + scale(Barometric) +   
## scale(Humidity) + scale(julDate) + (1 | Plant\_Code) + (1 |   
## julDate)  
## Data: d  
## Control: control  
##   
## AIC BIC logLik deviance df.resid   
## 791.7 874.3 -374.9 749.7 357   
##   
## Scaled residuals:   
## Min 1Q Median 3Q Max   
## -1.0095 -0.5238 -0.3413 0.1952 4.7630   
##   
## Random effects:  
## Groups Name Variance Std.Dev.  
## Plant\_Code (Intercept) 0.4284 0.6545   
## julDate (Intercept) 0.1009 0.3176   
## Number of obs: 378, groups: Plant\_Code, 63; julDate, 9  
##   
## Fixed effects:  
## Estimate Std. Error z value  
## (Intercept) -1.95538 0.72181 -2.709  
## Mold 1.17430 0.50503 2.325  
## Insecticide 0.42923 0.53530 0.802  
## Sugar 2.39457 0.47918 4.997  
## Paint -0.61127 0.61237 -0.998  
## SiteB 0.73398 0.57844 1.269  
## SiteC 0.50101 0.53278 0.940  
## hours\_noon 0.01093 0.09681 0.113  
## scale(Temp\_F) 0.22330 0.18097 1.234  
## scale(Wind\_mph) 0.19258 0.12820 1.502  
## ConditionsCompletly Cloudy (no Shadow) 0.39652 0.55278 0.717  
## ConditionsFull Sun -0.65723 0.55215 -1.190  
## ConditionsPartly Cloudy (>50% sun) 0.23767 0.44193 0.538  
## scale(Barometric) -0.34483 0.17986 -1.917  
## scale(Humidity) -0.03533 0.19502 -0.181  
## scale(julDate) -0.13315 0.16300 -0.817  
## Mold:Insecticide -1.47040 0.73224 -2.008  
## Sugar:Paint 0.29472 0.71953 0.410  
## Pr(>|z|)   
## (Intercept) 0.00675 \*\*   
## Mold 0.02006 \*   
## Insecticide 0.42265   
## Sugar 5.82e-07 \*\*\*  
## Paint 0.31818   
## SiteB 0.20448   
## SiteC 0.34703   
## hours\_noon 0.91013   
## scale(Temp\_F) 0.21724   
## scale(Wind\_mph) 0.13304   
## ConditionsCompletly Cloudy (no Shadow) 0.47318   
## ConditionsFull Sun 0.23392   
## ConditionsPartly Cloudy (>50% sun) 0.59072   
## scale(Barometric) 0.05521 .   
## scale(Humidity) 0.85623   
## scale(julDate) 0.41398   
## Mold:Insecticide 0.04463 \*   
## Sugar:Paint 0.68210   
## ---  
## Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

anova(Honeydew, Honeydew\_env)

## Data: d  
## Models:  
## Honeydew: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + hours\_noon +   
## Honeydew: Site + (1 | Plant\_Code) + (1 | julDate)  
## Honeydew\_env: Bee\_Count ~ Mold \* Insecticide + Sugar \* Paint + Site + hours\_noon +   
## Honeydew\_env: scale(Temp\_F) + scale(Wind\_mph) + Conditions + scale(Barometric) +   
## Honeydew\_env: scale(Humidity) + scale(julDate) + (1 | Plant\_Code) + (1 |   
## Honeydew\_env: julDate)  
## Df AIC BIC logLik deviance Chisq Chi Df Pr(>Chisq)  
## Honeydew 13 785.62 836.77 -379.81 759.62   
## Honeydew\_env 21 791.70 874.34 -374.85 749.70 9.915 8 0.271