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Problem 1: Golden State Warriors

a)

ttest wage, by (guard)

t-statistic: 2.0530

p-value: .0410

Since p-values are below .05, we reject that the population mean of salaries are equal. Another way of saying the same thing: t=2.053>1.96. We could not plausibly reject this p-value at .01.

$$t = 2.05 > 1.96$$

 $p = .041 < .05$

Reject that guards in the NBA are paid more than different positions (lower average salary). This is the one-sided hypothesis.

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. ttest wage, by(guard)

Two-sample t test with equal variances

0.5	917					
Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	156 113	1529.706 1277.658	87.72705 78.31673	1095.71 832.5182	1356.412 1122.484	1703.001 1432.833
combined	269	1423.828	60.9573	999.7741	1303.811	1543.844
diff		252.048	122.7685		10.33043	493.7656
diff :	= mean(0) - = 0	- mean(1)		degrees	t of freedom	
Harris d				•	Hara di	14.E.E. 0

Ha: diff < 0 Pr(T < t) = **0.9795** Ha: diff != 0 Pr(|T| > |t|) = **0.0410** Ha: diff > 0 Pr(T > t) = **0.0205**



gen degree = coll >= 4 ttest wage, by(degree)

Greater than 4 years NBA experience More education is correlated with a smaller salary

t-statistic: 2.3705 p-value: .0185

Reject the null hypothesis that NBA players who obtain a college degree earn the same amount as players with no college education at 10% and 5% significance; however, this does not hold up at the 1% significance level. This may happen because players who leave college early are usually better players who have the opportunity to play in the NBA and make a higher salary at a younger age.

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- . gen degree = coll >= 4
- . ttest wage, by(degree)

Two-sample t test with equal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	49 220	1727.408 1356.212	156.2117 65.22361	1093.482 967.4225	1413.323 1227.666	2041.493 1484.758
combined	269	1423.828	60.9573	999.7741	1303.811	1543.844
diff		371.1963	156.5877		62.89262	679.5001

$$\label{eq:diff} \mbox{diff} = \mbox{mean}(\mathbf{0}) - \mbox{mean}(\mathbf{1}) \\ \mbox{Ho: diff} = \mbox{0} \\ \mbox{degrees of freedom} = \\ \mbox{2.3705} \\ \mbox{degrees of freedom} = \\ \mbox{267}$$

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0
$$Pr(T < t) = \textbf{0.9908} \qquad Pr(|T| > |t|) = \textbf{0.0185} \qquad Pr(T > t) = \textbf{0.0092}$$

c)

Index

ppg = "points" total minutes = "min" total games played = "games"

gen productivity = points/(minutes/games)
ttest productivity, by (guard)

t-statistic: -.9746 p-value: .3307

Fail to reject the null that guards are as productive as other positions.

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There is no statistical difference that there is a difference in the overall productivity of guards in comparison to other positions.

- . gen productivity = points/(minutes/games)
- . ttest productivity, by(guard)

Two-sample t test with equal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	156 113	.4038649 .4177012	.0098823 .009596	.1234298 .102007	.3843435 .3986879	.4233862 .4367145
combined	269	.4096771	.0070068	. 11492	.3958818	. 4234725
diff		0138363	.0141975		0417896	.0141169
diff Ho: diff	= mean(0) = 0	- mean(1)		degrees	t of freedom	= -0.9746 = 267
Ha: d	iff < 0		Ha: diff !=	0	Ha: d	iff > 0

Pr(T < t) = 0.1653 Pr(|T| > |t|) = 0.3307 Pr(T > t) = 0.8347



correlate points assists rebounds

Points, rebounds, and assists are positively correlated with each other.

The correlation between assists and rebounds is weak at .06.

Reasonably strong correlation between points and assists at .5393.

Reasonably strong between points and rebounds at .5633.

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. correlate points assists rebounds (obs=269)

	points	assists	rebounds
points	1.0000		
assists	0.5393	1.0000	
rebounds	0.5633	0.0600	1.0000

. correlate points rebounds assists rebounds, cov (obs=269)

	points	rebounds	assists	rebounds
points	34.6037			
rebounds	9.58661	8.36933		
assists	6.81438	.372593	4.61444	
rebounds	9.58661	8.36933	.372593	8.36933



gen index = points + rebounds + 2*assists ttest index, by(guard)

Judging by the ttest, guards are usually more productive during games. As a result, the null hypothesis should be rejected. This means to reject the null that the performance of guards is the same as other positions' performance during games at the 10% and 5% significance level. We

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cannot reject at the 1% significance level.

- . gen index = points + rebounds + 2*assists
- . ttest index, by(guard)

Two-sample t test with equal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0	156	18.45513	.8462608	10.56979	16.78344	20.12682
1	113	21.26549	.9520769	10.12072	19.37907	23.1519
combined	269	19.63569	. 6375833	10.45714	18.38038	20.891
diff		-2.810359	1.282715		-5.335882	2848355
diff :	= mean(0) -	- mean(1)			t	= -2.1909
Ho: diff	= 0			degrees	of freedom	= 267
Ha: d	iff < 0		Ha: diff !=	0	Ha: d	iff > 0
Pr(T < t) = 0.0147	Pr(T > t) =	0.0293	Pr(T > t)) = 0.9853



gen indexper1k = index/wage ttest indexper1k, by(guard)

Performance index per \$1000 shows that guards outperform their salary in comparison to other positions, which indicates good value for what they are paid relative to their production.

Reject the null hypothesis that claims that guards are paid as much as other positions on the court relative to their performance at the 10% and 5% significance level; however, not at the 1% level, but we cannot reject at the 1% level.

p-value = .0115

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T-statistic = -2.5462

gen indexper1k = index/wage

ttest indexper1k, by(guard)

「wo-sample t test with equal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	156 113	.0179347 .0241632	.001336 .0022052	.0166866 .0234414	.0152956 .0197939	.0205738 .0285325
combined	269	.0205512	.0012197	.0200042	.0181498	.0229525
diff		0062285	.0024462		0110449	0014121

 $\label{eq:diff} \mbox{diff} = \mbox{mean(0)} - \mbox{mean(1)} \\ \mbox{lo: diff} = \mbox{0} \\ \mbox{degrees of freedom} = \mbox{267}$

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0 $Pr(T < t) = \textbf{0.0057} \qquad Pr(|T| > |t|) = \textbf{0.0115} \qquad Pr(T > t) = \textbf{0.9943}$

Problem 2: Crime on Campus

a)

	Enrollment	Police	Crime
Number of observations	97	97	97
Sample mean	16076.35	20.49	394.45
Sample median	11990	16	187

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Sample standard deviation	12298.99	15.63	460.78
Sample skewness	1.22	1.31	1.85
Mean, if public	17473.01	22.15	432.8
Mean, if private	6183.33	8.75	122.83

Summarize enrollment police crime, detail

Number of observations, sample mean, sample median (50%), sample standard deviation, sample skewness

by private, sort: summ enrollment police crime

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. summarize enrollment police crime, detail

enrollment

	Percentiles	Smallest		
1%	1799	1799		
5%	2446	1859		
10%	3712	2123	0bs	97
25%	6485	2167	Sum of Wgt.	97
50%	11990		Mean	16076.35
		Largest	Std. Dev.	12298.99
75%	21836	43030		
90%	33376	49961	Variance	1.51e+08
95%	42790	54311	Skewness	1.224383
99%	56350	56350	Kurtosis	4.192367
		police		
	Percentiles	Smallest		
1%	1	1		
5%	5	2		
10%	6	3	0bs	97
25%	9	4	Sum of Wgt.	97
50%	16		Mean	20.49485
		Largest	Std. Dev.	15.63058
75%	27	60		
90%	48	63	Variance	244.3151
95%	51	65	Skewness	1.310145
99%	74	74	Kurtosis	4.283173

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crime

	Percentiles	Smallest		
1%	1	1		
5%	25	1		
10%	44	15	0bs	97
25%	85	17	Sum of Wgt.	97
50%	187		Mean	394.4536
		Largest	Std. Dev.	460.7839
75%	491	1712		
90%	1145	1839	Variance	212321.8
95%	1384	1962	Skewness	1.852137
99%	2052	2052	Kurtosis	6.005492

. by private, sort: $\ensuremath{\mathsf{summ}}$ enrollment police crime

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Variable	0bs	Mean	Std. Dev.	Min	Max
enrollment police	85 85	17473.01 22.15294	12468.03 15.84712	1859 1	56350 74
crime	85	432.8	477.9452	1	2052

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-> private = 1

Variable	0bs	Mean	Std. Dev.	Min	Max
enrollment	12	6183.333	3347.161	1799	13570
police	12	8.75	6.538348	2	26
crime	12	122.8333	129.4441	15	426

. summ enrollment police crime if private == 1

	Variable	0bs	Mean	Std. Dev.	Min	Max
- 12	enrollment	12	6183.333	3347.161	1799	13570
	police	12	8.75	6.538348	2	26
	crime	12	122.8333	129.4441	15	426

. summ enrollment police crime if private == 0

Variable	0bs	Mean	Std. Dev.	Min	Max
enrollment	85	17473.01	12468.03	1859	56350
police	85	22.15294	15.84712	1	74
crime	85	432.8	477.9452	1	2052



correlate enrollment police crime, cov correlate enrollment police crime

Each variable in this set are positively correlated, which may mean that college campuses with a higher number of police officers also have a higher number of crimes committed. Additionally,

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larger campuses have more crime and more police officers present. This positive correlation makes sense because the values are absolute, not relative.

. correlate enrollment police crime, cov (obs=97)

	enroll~t	police	crime
enrollment	1.5e+08		
police	137462	244.315	
crime	4.7e+06	5209.51	212322

. correlate enrollment police crime (obs=97)

	enroll~t	police	crime
enrollment	1.0000		
police	0.7151	1.0000	
crime	0.8360	0.7233	1.0000



ttest crime, by(private) unequal ttest crime, by(private)

If unequal variance is assumed of public and private schools, then the null hypothesis claiming that crimes at the two schools are the same at the 5% and 1% significance level and rejected for both levels. There is a small p-value. Pr(|T|>t)=0.0000

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Assuming equal variance of public and private schools, then the null is rejected at the 5% significance level, but not at the 1% significance level. p-value= .0284 is < .05, but > .01.

Pr(|T|>t) = 0.0284p-value: .0284

t-statistic: 2.2259

. ttest crime, by(private) unequal

Two-sample t test with unequal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	85 12	432.8 122.8333	51.84044 37.36731	477.9452 129.4441	329.7096 40.58844	535.8904 205.0782
combined	97	394.4536	46.78551	460.7839	301.5851	487.3221
diff		309.9667	63.9042		182.2784	437.655

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. ttest crime, by(private)

Two-sample t test with equal variances

79						
Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0	85	432.8	51.84044	477.9452	329.7096	535.8904
1	12	122.8333	37.36731	129.4441	40.58844	205.0782
combined	97	394.4536	46.78551	460.7839	301.5851	487.3221
diff		309.9667	139.2572		33.5062	586.4271
diff =	= mean(0) -	mean(1)			t	= 2.2259
Ho: diff =	= 0			degrees	of freedom	= 95
Ha: d:	iff < 0		Ha: diff !=	0	Ha: d	iff > 0
) = 0.9858	Pr(T > t) =) = 0.0142

d)

gen crimerate = 1000*(crime/enrollment)

ttest crimerate, by(private) unequal

When adjusting for the greater enrollments at public vs. private schools crime is narrowly higher at the public schools than at the private schools.

22.34 public vs. 21.20 private per 1000 students.

This narrow difference is not statistically significant.

p-value: .8491 t-value: .1940

This p-value is significantly higher than .05 (or.1 or .01) significance level.

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- . gen crimerate = 1000*(crime/enrollment)
- . ttest crimerate, by(private) unequal

Two-sample t test with unequal variances

N-						- 6
Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0 1	85 12	22.34158 21.20101	1.861962 5.57801	17.16644 19.32279	18.63887 8.92389	26.0443 33.47812
combined	97	22.20048	1.760902	17.34288	18.70512	25.69584
diff		1.140577	5.880569		-11.50996	13.79112

diff = mean(0) - mean(1) t = 0.1940

Ho: diff = 0 Satterthwaite's degrees of freedom = 13.5659

Ha: diff < 0 Ha: diff != 0 Ha: diff > 0 Pr(T < t) = 0.5755 Pr(|T| > |t|) = 0.8491 Pr(T > t) = 0.4245

Problem 3: Post Katrina Employment in New Orleans



	Е	U	Т
Returned	139	8	147
Not	79	23	102
Total	218	31	249

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Joint Probabilities	Employed (Y=1)	Unemployed (Y=0)	Marginal
Returned to pre-storm address (X=1)	.558	.032	.590
Have not returned (X=0)	.317	.092	.410
Total	.875	.124	

Work

P(X=1, Y=1) = 139/249 = .558
P(X=1, Y=0) = 8/249 = .032
P(X=0, Y=0) = 23/249 = .092
P(X=0, Y=1) = 79/249 = .317
P(X=1) = (139+8)/249 = .590
P(Y=1) = (139+79)/249 = .876
P(X=0) = (79+23)/249 = .410
P(Y=0) = (8+23)/249 = .124

Conditional probability table

Conditional	Employed (Y=1)	Unemployed (Y=0)
Given returned to pre-storm address (X=1)	.946	.054
Given have not returned yet (X=0)	.775	.225

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$$P(Y=1 \mid X=1) = 139/147 = .946$$

$$P(Y=1 \mid X=0) = 79/102 = .775$$

$$P(Y=0 \mid X=1) = 8/147 = .054$$

$$P(Y=0 \mid X=0) = 23/102 = .225$$

b)

1 = returned

0 = not returened

1 = employed

0 = unemployed

E of being employed conditional on returning to pre-storm address

$$E[Y \mid X = 1] = P(Y=1 \mid X=1)(1) + P(Y=0 \mid X=1)(0) = .946$$

$$E[Y \mid X = 0] = P(Y=1 \mid X=0)(1) + P(Y=0 \mid X=0)(0) = .775$$

law of iterated expectations: E[Y] = E[E(Y|X)]

$$E[E(Y|X)] = P(X=1)*E(Y|X=1) + P(X=0)*E[Y|X=0] = (.590*.946)+(.410*.775) = .876$$

.876 matches the unconditional expectation of E[Y] = P(Y=1)(1) + P(Y=0)(0) = .876

c)

Covariance of "returned to pre-storm address" status and "employment" status

$$S_{XY} = 1/(n-1) \sum (Xi-X)(Yi-\bar{Y})$$
 summation from i=1 to n
= $1/248 \sum (Xi-E[X])(Yi-E[Y])$ summation from i=1 to 249
= $1/248 (139*(1-.590)(1-.875) + 8*(1-.590)*(0-.875) + 79(0-.590)(1-.875) + 23(0-.590)(0-.875))$
= $.0415$

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d)

Since $.875 \neq .946 = (\text{or P}(Y=1) \neq P(Y=1 \mid X=1))$ "employment" status is not independent from "return" status, so we are able to conclude that X and Y are positively correlated with each other (relationship).

e)

2 reasons to explain the difference in employment status between returned vs. not returned:

- 1: People who are unemployed after the evacuation may have not have had good work prior to the storm, so they gave up on the town and cut their losses. People without money, job, or housing have no reason or resources to motivate them to return to whatever is left of their homes.
- 2: People who returned may have better luck finding jobs or job opportunities in town vs. people who have not returned to town.

Problem 4: Pollution

a)

by oecd, sort: summ gdppc co2pc

gdppc mean oecd: 31423.91 vs. gdppc mean non-oecd: 5585.894 greater for oecd gdppc sd oecd:17180.88 vs. gdppc sd non-oecd: 9827.863 greater for oecd co2pc mean oecd: 8.984034 vs. co2pc mean non-oecd: 4.074291 greater for oecd co2pc sd oecd: 4.099933 vs. co2pc sd non-oecd: 6.430327 greater for non-oecd

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. by oecd, sort: summ gdppc co2pc

-> oecd = 0					
Variable	0bs	Mean	Std. Dev.	Min	Max
gdppc co2pc	150 164	5585.894 4.074291	9827.863 6.430327	150.7424 .0091268	76393.83 40.31009
-> oecd = 1					
Variable	0bs	Mean	Std. Dev.	Min	Max
gdppc co2pc	34 34	31423.91 8.984034	17180.88 4.099933	7833.529 3.763572	80276.01 21.36027



Since the sample standard deviations of each subsample are extremely different, unequal variance should be assumed.

Reject the null hypothesis at 5% significance level.

Reject the sample means co2 emissions per capita are equal to each other.

ttest co2pc, by(oecd) unequal

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. ttest co2pc, by(oecd) unequal

Two-sample t test with unequal variances

Group	0bs	Mean	Std. Err.	Std. Dev.	[95% Conf.	Interval]
0	164	4.074291	.5021242	6.430327	3.082785	5.065798
1	34	8.984034	.7031327	4.099933	7.5535	10.41457
combined	198	4.917379	. 4521244	6.361954	4.025754	5.809004
diff		-4.909743	.8640164		-6.632343	-3.187143
diff = mean(0) - mean(1) $t = -5.6825$						
Ho: diff =	= 0		Satterthwai	te's degrees	of freedom	= 71.4772
Ha: di	iff < 0		Ha: diff !=	0	Ha: d	iff > 0
Pr(T < t)	= 0.0000	Pr(T > t) =	0.0000	Pr(T > t) = 1.0000



generate lgdp = log(gdp)generate lco2 = log(co2)

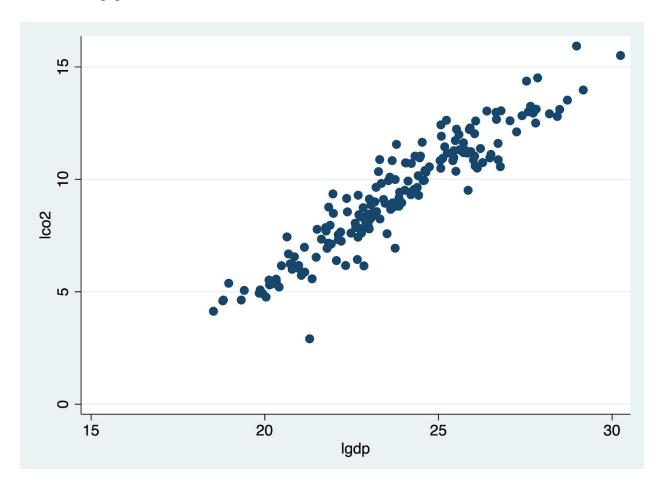
View growth rates, not the absolute levels of emission and GDP since not many countries in the sample have very high levels of emissions and GDP in relation to the other countries in the sample.

When taking the natural log of the two variables, the big picture of the data patterns become easier to visualize. The spread of values shrinks, especially for skewed data, thereby making it easier to visualize.



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scatter lco2 lgdp



GDP growth rates on x-axis CO2 emissions on y-axis

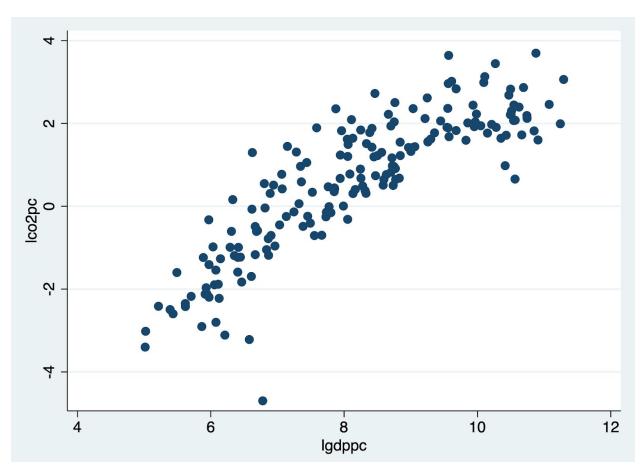


Growth rate per capita GDP on x-axis Growth rate of per capita emissions on y-axis

generate lgdppc = log(gdppc)

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generate lco2pc = log(co2pc) scatter lco2pc lgdppc



GDP and emission rates appear to be positively correlated even after controlling for the population sizes.

The claim of high-GDP countries does not seem convincing.

The scatterplot shows a concave relationship between GDP and emissions growth rates, which begins to tilt downwards for higher GDP growth rates.

Higher lgdppc ≠ higher lco2pc after a certain point

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Problem 5: Wages in Los Angeles



summ wage, detail

Mean: 13.41386

\$13.41386

IQR (interquartile range): 6.5-15.81197

25th percentile-75th percentile

\$6.50-\$15.81

. summ wage, detail

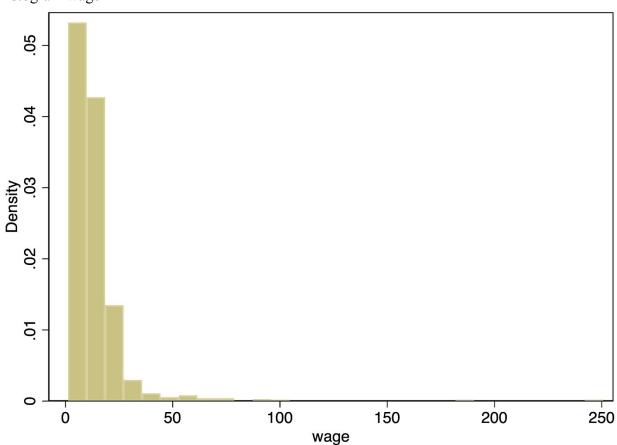
wage

	Percentiles	Smallest		
1%	1.875	1.25		
5%	3.571429	1.25		
10%	4.299754	1.5	0bs	863
25%	6.5	1.593367	Sum of Wgt.	863
50%	10.57692		Mean	13.41386
		Largest	Std. Dev.	14.65928
75%	15.81197	93.99808		
90%	21.92308	97.758	Variance	214.8945
95%	29.16667	187.9962	Skewness	8.17203
99%	63.33333	250.6615	Kurtosis	108.8581

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b)

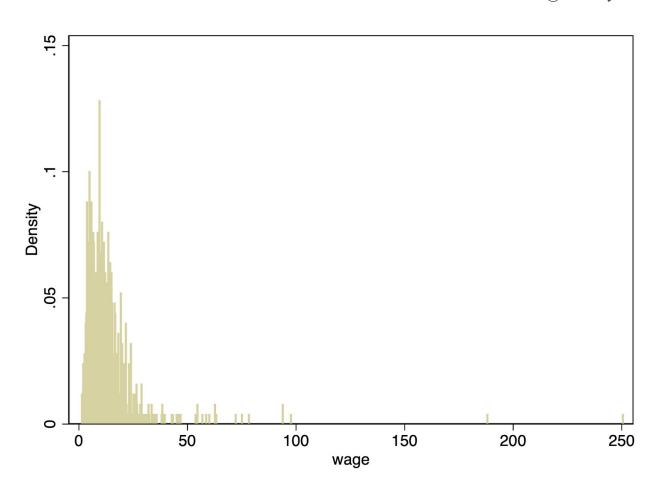
histogram wage



or

histogram wage, bin(863)

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The distribution is skewed

Skewness 8.17203 summ wage, detail Positive, asymmetrically skewed distribution 8.17203 is rightly skewed

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d)

Take log of wage

gen lnwage = ln(wage)

summ lnwage, detail

Mean wage 2.335614 2.34

IQR (interquartile range): 1.871802-2.760767

1.87-2.76 25th-75th

. summ lnwage, detail

lnwage

1				
	Percentiles	Smallest		
1%	.6286086	.2231435		
5%	1.272966	.2231435		
10%	1.458558	.4054651	0bs	863
25%	1.871802	.4658496	Sum of Wgt.	863
50%	2.358675		Mean	2.335614
		Largest	Std. Dev.	.6847209
75%	2.760767	4.543274		
90%	3.08754	4.582495	Variance	.4688427
95%	3.373027	5.236422	Skewness	.2427184
99%	4.148412	5.524104	Kurtosis	4.028374

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How many workers have at least an 8th grade education?

tab education

Education is discrete in this data set, not continuous, proportion ≥ 8 years of education

100-8.69=91.31

91.31% in the same sample have at least an 8th grade education

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. tab education

education	Freq.	Percent	Cum.
0	5	0.58	0.58
3	11	1.27	1.85
7	59	6.84	8.69
9	29	3.36	12.05
10	26	3.01	15.06
11	75	8.69	23.75
12	174	20.16	43.92
13	194	22.48	66.40
14	55	6.37	72.77
16	155	17.96	90.73
18	75	8.69	99.42
20	5	0.58	100.00
Total	863	100.00	χĈ



ln(wages) on y-axis Education on x-axis

twoway (scatter Inwage education)

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