

# Chapter 7: Sampling Distributions

Juhhyung Lee

Department of Statistics  
University of Florida

# Sampling Distributions

Recall that we defined a statistic to be a numerical summary of a sample taken from the population. Here is a more rigorous definition.

## Definition (Statistic)

A function of one or more random variables that does not depend upon any unknown parameter is called a **statistic**. A statistic is a random variable.

## Example

- The random variable  $Y = \sum_{i=1}^n X_i$  is a statistic.
- The random variable  $Z = (X_1 - \mu)/\sigma$  is not a statistic unless  $\mu$  and  $\sigma$  are known numbers.

## Definition (Sampling Distribution)

The probability distribution of a statistic is called a **sampling distribution**.

# Sampling Distribution of Means and the Central Limit Theorem

## Theorem

Suppose  $X_1, X_2, \dots, X_n$  are independent and identically distributed (iid) with  $E(X) = \mu$  and  $\text{Var}(X) = \sigma^2 < \infty$ . Let  $\bar{X} = n^{-1} \sum_{i=1}^n X_i$  denote the sample mean. Then

$$E(\bar{X}) = \mu \text{ and } \text{Var}(\bar{X}) = \frac{\sigma^2}{n}.$$

## Theorem

Suppose  $X_1, X_2, \dots, X_n$  are iid  $N(\mu, \sigma^2)$ . Then

$$\bar{X} \sim N\left(\mu, \frac{\sigma^2}{n}\right).$$

# Sampling Distribution of Means and the Central Limit Theorem

## Theorem (Theorem 8.2 of WMMY: Central Limit Theorem (CLT))

If  $\bar{X}$  is the mean of a random sample of size  $n$  taken from a population with mean  $\mu$  and finite variance  $\sigma^2$ , then the limiting form of the distribution of

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}},$$

as  $n \rightarrow \infty$ , is the standard normal distribution  $N(0, 1)$ .

The sample size  $n = 30$  is a guideline to use for the CLT.

# Sampling Distribution of Means and the Central Limit Theorem

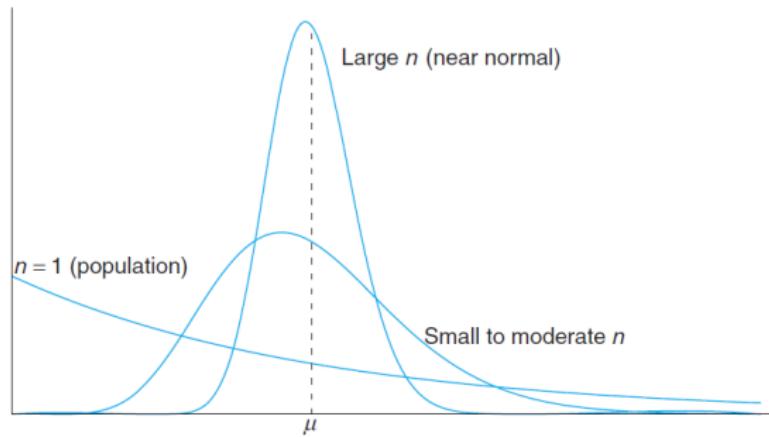


Figure 8.1: Illustration of the Central Limit Theorem (distribution of  $\bar{X}$  for  $n = 1$ , moderate  $n$ , and large  $n$ ).

## Example

An electrical firm manufactures light bulbs that have a length of life that is distributed with mean equal to 800 hours and a standard deviation of 40 hours. Find the probability that a random sample of 64 bulbs will have an average life of less than 790 hours.

# Sampling Distribution of Means and the Central Limit Theorem

The CLT can be easily extended to the two-sample, two-population case.

## Theorem (Theorem 8.3 of WMMY)

If independent samples of size  $n_1$  and  $n_2$  are drawn at random from two populations, discrete or continuous, with means  $\mu_1$  and  $\mu_2$  and variances  $\sigma_1^2$  and  $\sigma_2^2$ , respectively, then the sampling distribution of the differences of means,  $\bar{X}_1 - \bar{X}_2$ , is approximately normally distributed with mean and variance given by

$$E(\bar{X}_1 - \bar{X}_2) = \mu_1 - \mu_2 \text{ and } \text{Var}(\bar{X}_1 - \bar{X}_2) = \frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}.$$

Hence,

$$Z = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\sigma_1^2/n_1 + \sigma_2^2/n_2}} \sim N(0, 1).$$

The normal approximation is usually good if  $n_1 \geq 30$  and  $n_2 \geq 30$ .

# Sampling Distribution of Means and the Central Limit Theorem

## Example (Example 8.6 of WMMY)

The television picture tubes of manufacturer  $A$  have a mean lifetime of 6.5 years and a standard deviation of 0.9 year, while those of manufacturer  $B$  have a mean lifetime of 6.0 years and a standard deviation of 0.8 year. What is the probability that a random sample of 36 tubes from manufacturer  $A$  will have a mean lifetime that is at least 1 year more than the mean lifetime of a sample of 49 tubes from manufacturer  $B$ ?

## Sampling Distribution of $S^2$

- The sampling distribution of  $\bar{X}$  is used to learn about the population mean  $\mu$ .
- Similarly, the sampling distribution of  $S^2$  is used to learn about the population variance  $\sigma^2$ .

### Theorem (Theorem 8.4 of WMMY)

If  $S^2$  is the variance of a random sample of size  $n$  taken from a normal population having the variance  $\sigma^2$ , then the statistic

$$\frac{(n-1)S^2}{\sigma^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sigma^2} \sim \chi_{n-1}^2.$$

## *t*-Distribution

- In practice, a direct use of the CLT is often restricted due to lack of knowledge on the population variance  $\sigma^2$ .
- Then the unknown  $\sigma^2$  is estimated by the sample variance  $S^2$  and the *t*-distribution arises.

### Theorem (Theorem 8.5 of WMMY)

Suppose  $Z \sim N(0, 1)$ ,  $V \sim \chi_{\nu}^2$ , and  $Z \perp\!\!\!\perp V$  (i.e.,  $Z$  and  $V$  are independent).

Then

$$\frac{Z}{\sqrt{V/\nu}} \sim t_{\nu},$$

a ***t*-distribution** with  $\nu$  degrees of freedom.

### Theorem (Corollary 8.1 of WMMY)

Suppose  $X_i \stackrel{iid}{\sim} N(\mu, \sigma^2)$ ,  $i = 1, \dots, n$ . Then

$$T \triangleq \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1}.$$

## *t*-Distribution

- Both the standard normal distribution and the *t*-distribution are bell-shaped, symmetric about zero, but the *t*-distribution is more variable and it has heavier tails (i.e., more likely to have large/small values).
- As  $\nu \rightarrow \infty$ ,  $t_\nu \rightarrow N(0, 1)$ .

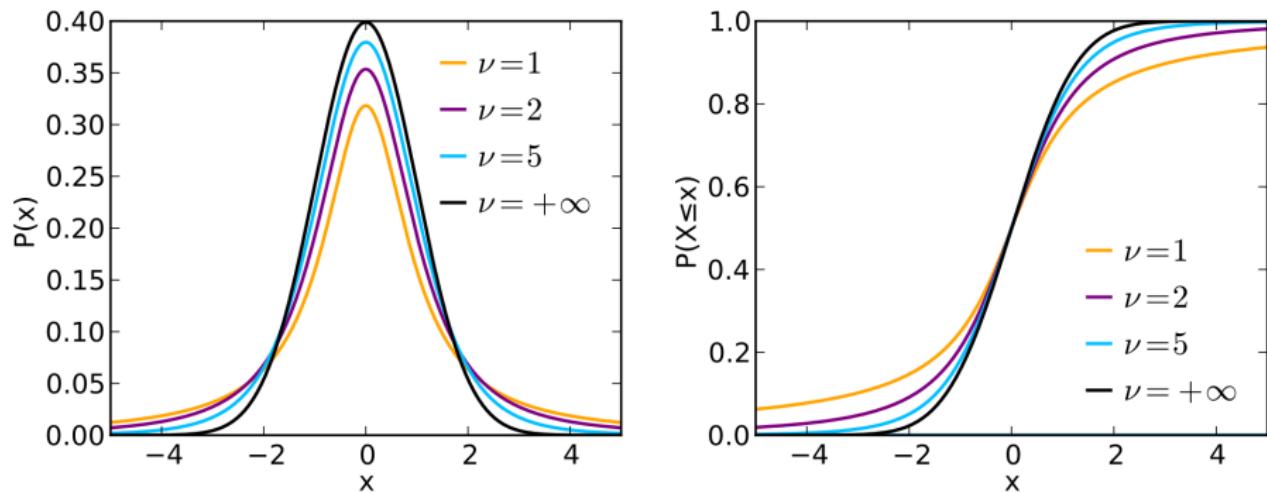


Figure: *t* pdfs and cdfs.

# *t*-Distribution

## *t*-Distribution in R

In R,

- the function `dt()` computes the *t* pdf;
- the function `pt()` computes the *t* cdf;
- the function `qt()` computes the *t* quantiles.

# *t*-Distribution

## Example

The  $t$ -value with  $\nu = 14$  df that leaves an area of 0.025 to the right, and therefore an area of 0.975 to the left, is

$$t_{0.025, 14} = -t_{0.975, 14} = 2.145.$$

Note that  $t_{0.025, 14} = 2.145 > 1.96 = z_{0.025}$  as the  $t_{14}$  distribution has heavier tails than the standard normal distribution.

```
> qt(0.975, df = 14) # upper 0.025 quantile of t_14
[1] 2.144787
> qnorm(0.975) # upper 0.025 quantile of N(0, 1)
[1] 1.959964
>
> # probabilities to the left and right of 2.145 for t_14
> pt(2.145, df = 14)
[1] 0.9750099
> pt(2.145, 14, lower.tail = FALSE)
[1] 0.02499008
```

## $F$ -Distribution

### Theorem (Theorem 8.6 of WMMY)

Suppose  $U \sim \chi^2_{\nu_1}$ ,  $V \sim \chi^2_{\nu_2}$ , and  $U \perp\!\!\!\perp V$ . Then

$$\frac{U/\nu_1}{V/\nu_2} \sim F_{\nu_1, \nu_2},$$

an  **$F$ -distribution** with  $\nu_1$  and  $\nu_2$  degrees of freedom. Here,  $\nu_1$  is the numerator df and  $\nu_2$  is the denominator df.

### Theorem (Theorem 8.8 of WMMY)

If  $S_1^2$  and  $S_2^2$  are the variances of independent random samples of size  $n_1$  and  $n_2$  taken from normal populations with variances  $\sigma_1^2$  and  $\sigma_2^2$ , respectively, then

$$F \triangleq \frac{S_1^2/\sigma_1^2}{S_2^2/\sigma_2^2} \sim F_{n_1-1, n_2-1}.$$

# $F$ -Distribution

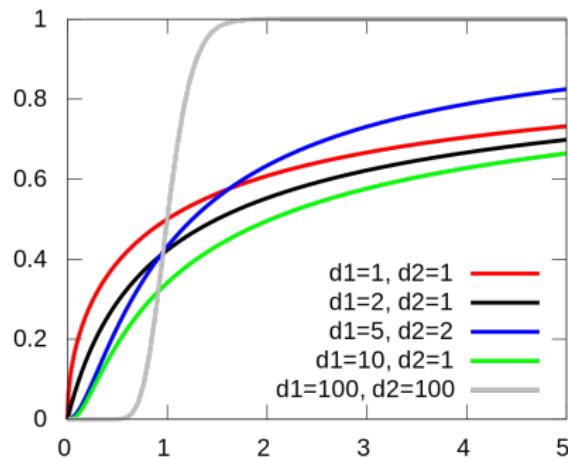
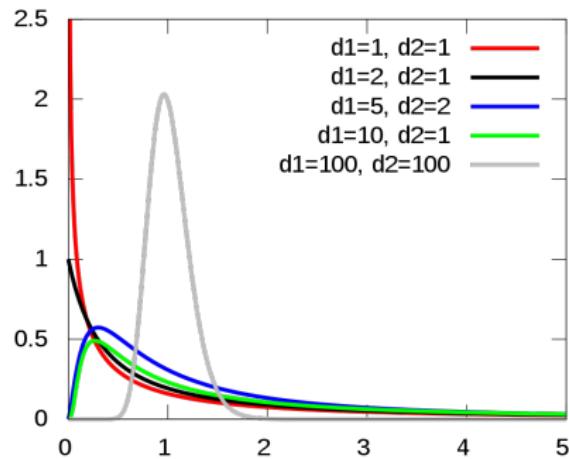


Figure:  $F$  pdfs and cdfs.

## $F$ -Distribution in R

In R,

- the function `df()` computes the  $F$  pdf;
- the function `pf()` computes the  $F$  cdf;
- the function `qf()` computes the  $F$  quantiles.

# $F$ -Distribution

## Example

The  $F$ -value with 6 and 10 df, leaving an area of 0.05 to the right, is

$$F_{0.05,6,10} = 3.217.$$

```
> qf(0.95, df1 = 6, df2 = 10) # upper 0.05 quantile of F_{6,10}
[1] 3.217175
>
> # probabilities to the left and right of 3.217 for F_{6,10}
> pf(3.217, df1 = 6, df2 = 10)
[1] 0.9499925
> pf(3.217, 6, 10, lower.tail = FALSE)
[1] 0.05000754
```